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# CERTIFICATION TESTING METHODOLOGY FOR COMPOSITE STRUCTURE

## Volume II—Methodology Development

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This research developed a certification testing methodology for composite structures. The existing composite static strength and fatigue life data are analyzed statistically to determine the influence of test parameters on the scatter of composite data. Guidelines to use the composite data scatter in structural certification are recommended.

Various approaches to composite structures certification are analytically evaluated. The approaches evaluated are: scatter factor approach, load enhancement factor approach, ultimate strength approach and change in spectrum approach. The capability, advantages and disadvantages of each approach to determine the minimum life and/or strength are fully discussed. (Keyword)

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19. Abstract

A methodology for certification testing of composite structures is developed and a detailed description of the methodology is presented. Test data interpretation methodology is also developed. The methodology is demonstrated on existing composite structures.

Based on the results of this investigation, composite structure certification testing procedure and requirements are recommended.

Volume I of this report discusses the scatter analysis methods and results of static strength and fatigue life data analysis. Details of the certification approach evaluation, methodology development and demonstrations are given in Volume II. Volume II also contains the recommended certification testing procedure.

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## PREFACE

This report was prepared by the Northrop Corporation, Aircraft Division, Hawthorne, California, covering work performed under U.S. Navy Contract N62269-C-0243 between March 1984 and December 1985. The contract was administerd by the Naval Air Development Center, Warminster, Pennsylvania. Mr. Ed Kautz was the Navy Project Engineer. Partial finding of this effort was provided by the Federal Aviation Administration Technical Center, Atlantic City International Airport, NJ. Mr. L. M. Neri acted as the FAA Technical Manager.

The work was performed in the Northrop's Strength and Life Assurance Research Department under the overall supervision of Dr. R. S. Whitehead. The following Northrop personnel were the major contributors to the program.

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# LIST OF ABBREVIATIONS

| | |
|---|---|
| RTD | Room Temperature Dry |
| RTW | Room Temperature Wet |
| ETW | Elevated Temperature Wet |
| (x/y/z) | ($\%0°$/$\%\pm45°$/$\%90°$) |
| NO L/T | No Load Transfer Specimen |
| INT L/T | Intermediate Load Transfer Specimen |
| HIGH L/T | High Load Transfer Specimen |
| COMPLEX | Complex Specimen |
| C.V. | Coefficient of Variation |
| DLL | Design Limit Load |
| DUL | Design Ultimate Load |
| SRV | Structural Response Variability |
| MSL | Maximum Spectrum Load |
| MLE | Maximum Likelihood Estimates |
| MOL | Material Operating Limit |

## LIST OF SYMBOLS

| | |
|---|---|
| R | Reliability |
| | also used as stress ratio (minimum stress/maximum stress) |
| $\bar{x}$ | Sample mean |
| n | Sample size |
| $\chi^2_\gamma(2n)$ | Chi-square distribution with 2n degrees-of-freedom at $\gamma$ percentage point |
| $\alpha$ | Weibull shape parameter |
| $\hat{\alpha}$ | MLE of Weibull shape parameter |
| $\beta$ | Weibull scale parameter |
| $\hat{\beta}$ | MLE of Weibull scale parameter |
| $\check{\beta}$ | 95% confident Weibull scale parameter |
| p | Probability of survival |
| $N_A$, $N_B$ | A- and B-basis allowables |
| X | Required load or life |
| k | Environmental knockdown factor |
| $\epsilon_E^{150\%}$ | Highest measured strain at 150% DLL |
| $\epsilon_D^E$ | Design allowable strain in worse case environmental |
| $\epsilon_D^{RT}$ | Design allowable strain under room temperature ambient conditions |
| $P_{FP}$ | Predicted room temperature ambient failure load |
| $P_{FR}$ | Required room temperature ambient failure load |
| $\sigma_e^M$ | Mean equivalent static strength from the Sendeckyj analysis |
| $\sigma_e^B$ | B-Basis equivalent static strength from the Sendeckyj analysis |
| N | Test duration |
| F | Load enhancement factor |
| $\sigma$ | Stress level |
| $\Gamma(X)$ | Gamma functions |

Subscripts   S   Static strength

       L   Fatigue life

       R   Residual strength

       F   Failure

       TH   Threshold


Superscripts   c   composite

       m   metal

       B   B-Basis

       M   mean

# SECTION 1

## INTRODUCTION AND BACKGROUND

The application of composite materials to primary aircraft structures requires proven certification procedures to demonstrate their structural integrity. The development of certification procedures for primary composite structures must recognize the inherent differences between metals and composites. These differences are summarized in Figure 1. Composite load-strain response is different from metals in that it is essentially linear to failure. Because of their linear elastic behavior, composites are extremely statically notch sensitive to stress concentrations such as fastener holes. Figure 2 presents typical data which show that composites are statically notch sensitive under both tension and compression loading. For both loading modes, notch sensitivity increases as hole diameter increases. By contrast, the static holed strength of metals is essentially notch insensitive and follows the net section strength reduction line.

Metallic materials are extremely fatigue sensitive to stress concentrations, which are the primary source of fatigue cracking in aircraft structures. In contrast, composites are almost fatigue insensitive to stress concentrations such as fastener holes. Figure 3 presents typical data which show the influence of loaded and unloaded fastener holes on composite fatigue life. It can be seen that compression static strength is very sensitive to the various fastener hole geometries. However, the maximum compression fatigue strain required for a life of $10^7$ cycles at an R-ratio of -1.7 is approximately constant at 4000 $\mu$in/in for all specimen geometries. These data show that, at low cycle lives, fatigue behavior is controlled by the static strength, while at high cycle lives fatigue behavior is controlled by the net section stress level.

| CONDITION | | COMPOSITE BEHAVIOR RELATIVE TO METALS |
|---|---|---|
| LOAD-STRAIN RELATIONSHIP | | MORE LINEAR STRAIN TO FAILURE |
| NOTCH SENSITIVITY | STATIC | GREATER SENSITIVITY |
| | FATIGUE | LESS SENSITIVITY |
| TRANSVERSE PROPERTIES | | WEAKER |
| MECHANICAL PROPERTIES VARIABILITY | | HIGHER |
| SENSITIVITY TO AIRCRAFT HYGROTHERMAL ENVIRONMENT | | GREATER |
| DAMAGE GROWTH MECHANISM | | IN-PLANE DELAMINATION INSTEAD OF THROUGH-THICKNESS CRACKS |

FIGURE 1. DIFFERENCES BETWEEN METALS AND COMPOSITES.
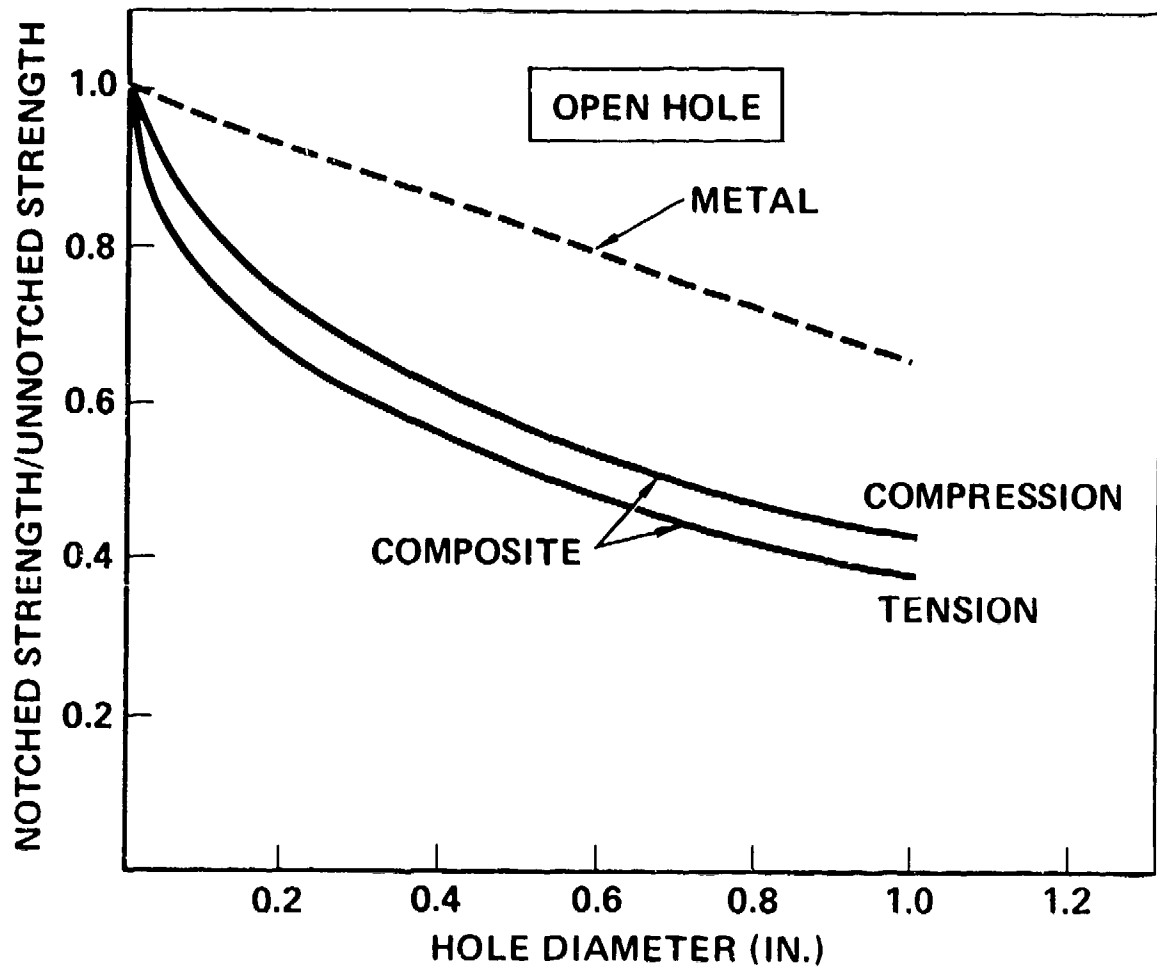
FIGURE 2. COMPARISON OF STATIC NOTCH SENSITIVITY
BETWEEN COMPOSITES AND METALS.
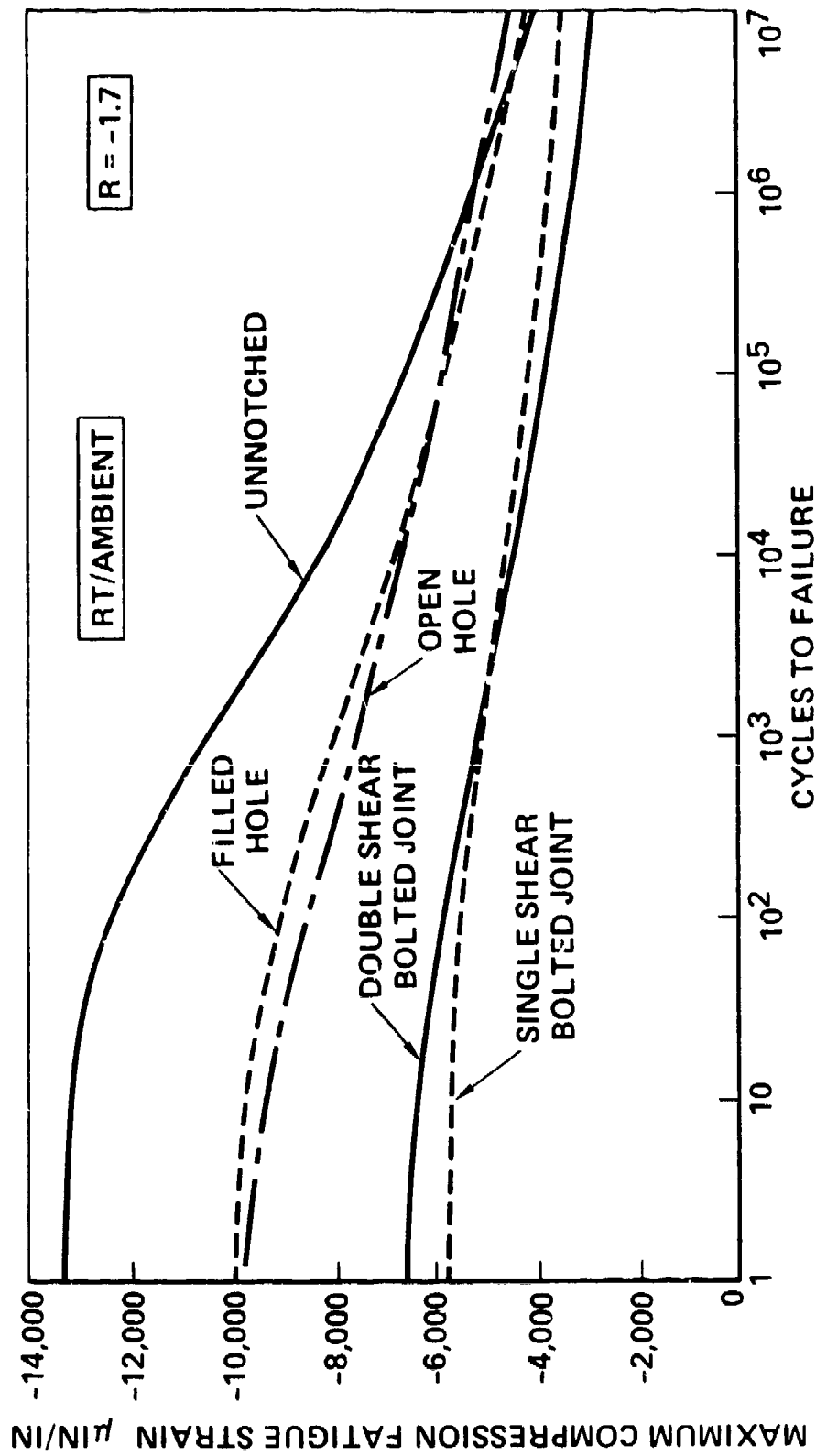
FIGURE 3. TYPICAL FATIGUE BEHAVIOR OF GRAPHITE/EPOXY COMPOSITES.

Figure 4 presents a comparison of metal and composite fatigue behavior under fighter aircraft wing spectrum loading. The comparison is made for an unloaded fastener hole specimen with a Youngs Modulus of approximately 10,000 ksi. The data are plotted for each material's most sensitive fatigue loading mode, which is tension dominated (lower wing skin) for metals and compression dominated (upper wing skin) for composites. Figure 4 shows that composite fatigue properties are markedly superior to metal fatigue properties.

The transverse properties of metallic and composite structures also differ significantly. Because they are isotropic, metallic transverse (thru-thickness) properties are similar to their in-plane properties. However, since composites are laminates manufactured from individual plies, their transverse out-of-plane properties are controlled by interlaminar strength which is considerably weaker than in-plane strength. Typical room temperature/ambient interlaminar strengths for composites are 12 ksi for interlaminar shear and 3 ksi for interlaminar tension. These are low compared to an unnotched in-plane strength of 75 ksi for a typical wing skin laminate. Because of their anisotropic heterogeneous characteristics, composites exhibit significantly higher scatter than metals in both static and fatigue properties. This variability must be accounted for in the design and certification of composites structures.

Composites, which exhibit matrix controlled failure modes (e.g., compression), are sensitive to the aircraft hygrothermal environment. In particular, the effects of temperature and moisture have a synergistic effect. Therefore, the strength degradation of composites in hot/wet environments controls their maximum service temperature application.

Fatigue damage growth mechanisms differ considerably between metals and composites. In metals, fatigue cracks initiate at stress concentrations and generally grow to through-the-thickness cracks under tension dominated loading. In composites, it has been demonstrated that the most common damage growth mechanism is interlaminar separation, known as
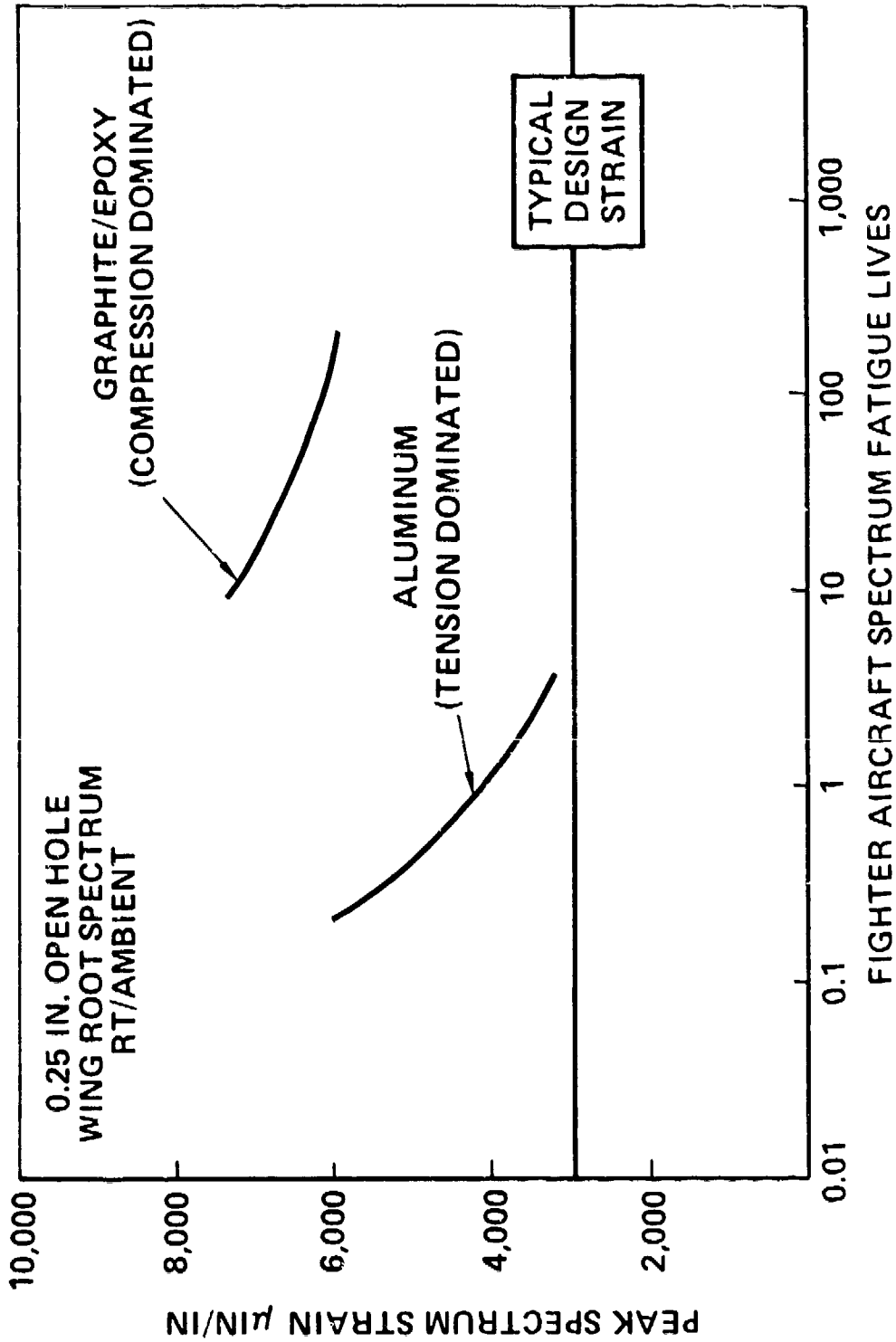
FIGURE 4. COMPARISON OF METAL AND COMPOSITE FATIGUE BEHAVIOR UNDER SPECTRUM LOADING.

delamination. Therefore, damage growth is an in-plane phenome-
non. Prior to delamination growth, matrix cracking and fiber
breakage usually occurs. The analogy with metals is that the
latter would correspond to crack initiation, while the former
would correspond to crack growth. Because delamination growth is
the dominant growth mechanism in composites, they are most sensi-
tive to compression dominated fatigue loading. A second compos-
ite fatigue failure mode has been observed in fastener holes
subject to high bearing stresses. Fatigue failure can occur by
hole wearout rather than delamination induced laminate failure.
In this failure mode, the hole gradually elongates leading to a
bearing failure.

The special properties of composites (fatigue notch
insensitivity, weak transverse properties, matrix dominated
failures, high failure variability, hygrothermal sensitivity and
delamination fatigue growth mechanism) must be addressed in
structural design and certification. It is emphasized that these
properties do not negate the weight efficiency of composite
structures, just that different parameters (from metals) are
important in composite design and certification.

Current practice is to carry out an extensive design
development test effort to:

1. Establish environmental and scatter knockdown for
Strength critical failure modes, and

2. Validate critical design features.

These tests are conducted at the coupon, element and subcomponent
levels. Following these tests, certification culminates in room
temperature ambient full-scale static and fatigue tests.
Usually, only one article is available for each test.

In order to have confidence in the certification com-
pliance of full scale tests, it is imperative to be able to
quantitatively interpret the test data generated. This is a-
chieved by using the design development data not only in their
traditional role in design development but also in the interpre-
tation of full-scale test data.

Current certification practices do not provide an overall testing methodology for the planning and quantitative interpretation of design development and full-scale test data.

The objective of this program is to develop a certification testing methodology for composite aircraft structures. Specifically, the methodology will account for the effects on strength, life, and the scatter in strength and life of variation in structural configuration and complexity, stress or strain level, mixed composite-metal structure and fatigue spectrum shape. Test requirements and procedures for interpreting test results will be defined for the certification of future composite aircraft structure.

The program is composed of four tasks:

- TASK I - SCATTER ANALYSIS

- TASK II - CERTIFICATION APPROACH DEVELOPMENT/EVALUATION

- TASK III - METHODOLOGY DEVELOPMENT

- TASK IV - METHODOLOGY DEMONSTRATION

During Task I, existing composite static strength and fatigue life data were analyzed statistically to determine the influence of different test parameters on the scatter of composite data. The test variables included were: laminate lay-up, specimen type, loading mode, failure mode and test environment for both the static and fatigue data; in addition, stress level, stress ratio, spectrum variation and spectrum shape are investigated for fatigue data. The effects of each variable on static strength and fatigue life data scatter were established by performing statistical tests of significance. As a result of this task, guidelines to use the composite data scatter in structural certification will be recommended and these guidelines will be applied in the subsequent tasks of the program.

In Task II, various approaches to composite structure certification were analytically evaluated. The approaches evaluated were:

1. Scatter factor approach

2. Load enhancement factor approach

3. Ultimate strength approach

4. Change in spectrum approach

The capability, advantages and disadvantages of each approach to determine minimum (B-basis) and mean life and/or strength were fully evaluated. Effects of these approaches on the certification procedure of composite-metal mixed structure were also investigated. The conclusions of this evaluation will then be used in the methodology development.

A methodology for certification testing of composite structures was developed in Task III. The methodology was based on the results of the evaluation in Task II and the scatter analysis in Task I. The number and types of tests required at each level (coupon, element, subcomponent, component, and full-scale) of testing were defined. Test data interpretation methodology was also developed. As part of this task, a detailed description of the development methodology will be presented. This description will include detailed instructions for application and utilization of the methodology within the overall developmental process to satisfy design service life requirements for aircraft utilizing composite structures. The description will also include application of the methodology in an aircraft design/development program and determine the effects on service life resulting from usage change of an aircraft after its introduction into the fleet.

In Task IV, the methodology was demonstrated on an existing composite structure. The full-scale wing and fuselage component from the Composite Wing/Fuselage Program (Reference 1) were selected for this demonstration purpose. The results of the tests that have been performed on these demonstration articles were reevaluated using the methodology developed in Task III.

The scatter analysis methods and results of static strength and fatigue life data analysis are discussed in Volume

I. Details of Task II - Certification Approval Development/Evaluation. Task III - Methodology Development and Task IV - Methodology Demonstration are given in Volume II. Recommendations and certification testing requirements are also documented in Volume II. Computer programs to evaluate structure reliability are appended to Volume II.

# SECTION 2

## CERTIFICATION APPROACH EVALUATION

During Task II of the program detailed evaluation of various certification approaches was conducted. The approaches evaluated were:

1. Scatter factor approach,

2. Load enhancement factor approach,

3. Ultimate strength approach,

4. Change in spectrum approach.

Details of this evaluation are discussed in the following paragraphs. In the evaluation, unless specified otherwise, the modal Weibull shape parameters are used. That is $\alpha = 20.0$ for static strength and $\alpha = 1.25$ for fatigue life.

## 2.1 Scatter Factor Approach

Current Navy certification procedures for composite structures are based on the Scatter Factor Approach. The key elements in the full-scale test requirements are based on experience with metal structures. These are: static tests to a minimum of 150 percent design limit load (DLL) and fatigue tests with a severe load spectrum to a minimum of two lifetimes.

The reliability obtained by using these factors can be determined directly from the Weibull distribution function when the scatter parameter ($\alpha$) is known. Then the reliability at 100 percent DLL or at 1.0 lifetime with $\gamma$ level of confidence is given by

$$R = \exp\left\{ -\frac{\chi^2_\gamma(2n)}{2n} \left[ \frac{\Gamma\left(\frac{\alpha+1}{\alpha}\right)}{\bar{X}} \right]^\alpha \right\} \tag{1}$$

Equation (1) can also be used to interpret static and fatigue test data. In these cases the value of $\bar{x}$ is taken as the average static strength or fatigue life obtained from test.

The ratio of the B-basis static strength (fatigue life) to the mean failure load (fatigue life) is obtained by letting R = 0.9 in equation (1), which can be written as:

$$\frac{\overset{\vee}{N}_B}{\bar{x}} = \frac{1}{\Gamma\left(\frac{\alpha+1}{\alpha}\right)} \left[\frac{-\ln(0.9)}{\chi^2_\gamma(2n)/2n}\right]^{1/\alpha} \tag{2}$$

The A-Basis strength or life factor is obtained simply replacing 0.9 by 0.99 in the above equation.

The influence of changes in load and/or life requirements on the reliability of an aircraft structure can be assessed after the design phase of the structure. This reliability can be computed from the mean test data for static strength or fatigue life. The $\gamma$ level of confidence reliability is given by

$$R = \exp\left\{-\left[\frac{X}{\bar{x}}\Gamma\left(\frac{\alpha+1}{\alpha}\right)\right]^\alpha \cdot \frac{\chi^2_\gamma(2n)}{2n}\right\} \tag{3}$$

An environmental knockdown factor is applied in order to provide increased reliability from an ambient full-scale test on an environmentally critical structure. The environmental knockdown factor, k, is applied based on the assumption that the static strength Weibull shape parameter is not significantly affected by the test environments. This assumption is substantiated by the results of Task I data analysis. From this assumption, the $\gamma$ level confidence reliability can be written as

$$R = \exp\left\{-\left[\left(\frac{kx}{\bar{x}}\Gamma\left(\frac{\alpha+1}{\alpha}\right)\right)^\alpha \cdot \frac{\chi^2_\gamma(2n)}{2n}\right]\right\} \tag{4}$$

## 2.1.1    Static Strength Evaluation

The static factor of 1.5 times limit load is used to safely account for unintentional deviations from service load and scatter in the static strength of a fleet of airframe structures. The degree of safety provided for metallic structures by the 1.5 factor can be calculated accurately, since the scatter in full-scale tests is known.    However, there is insufficient data to establish composite full-scale test data scatter.

To evaluate the reliability of a structure at DLL using a 1.5 static factor approach, the value  of $\bar{x}$ in equation (1) is taken as 1.5  That is assuming the mean static failure load is at 1.5 DLL, or at DUL.  It is seen from equation (1) that the reliability depends on the sample size, n, and the Weibull shape parameter $\alpha$.

Figure 5 shows the 95 percent confidence reliability at DLL as a function of the static strength shape parameter $(\alpha)$. The reliability in this figure is computed assuming structural failure at 150 percent of DLL, or at DUL.  As can be seen from the figure, at the modal value of $\alpha = 20.0$, the 95 percent confidence reliabilities are very high for all sample sizes, and far exceed the B-basis allowables.  This indicates that the static factor of 1.5 times limit load provides a very high degree of safety.  However, the reliability thus obtained does not account for unintentional deviations from service load, non-ambient service environments and structural response variability.  The effects of these factors need to be investigated in order to fully evaluate the degree of safety provided by the 1.5 static factor.

Figure 6 presents the 95 percent reliability at DLL as a function of the mean static failure load (equation (1)).  The figure shows that for a single article static test (n = 1) the B-basis reliability at the DLL can be achieved if the structure failure occurs at 115 percent of DLL.  The same reliability can be achieved if the mean static failure load is 111 percent of Dl for a sample size of 20.  This level of reliability is again determined only by considering the scatter in static strength.
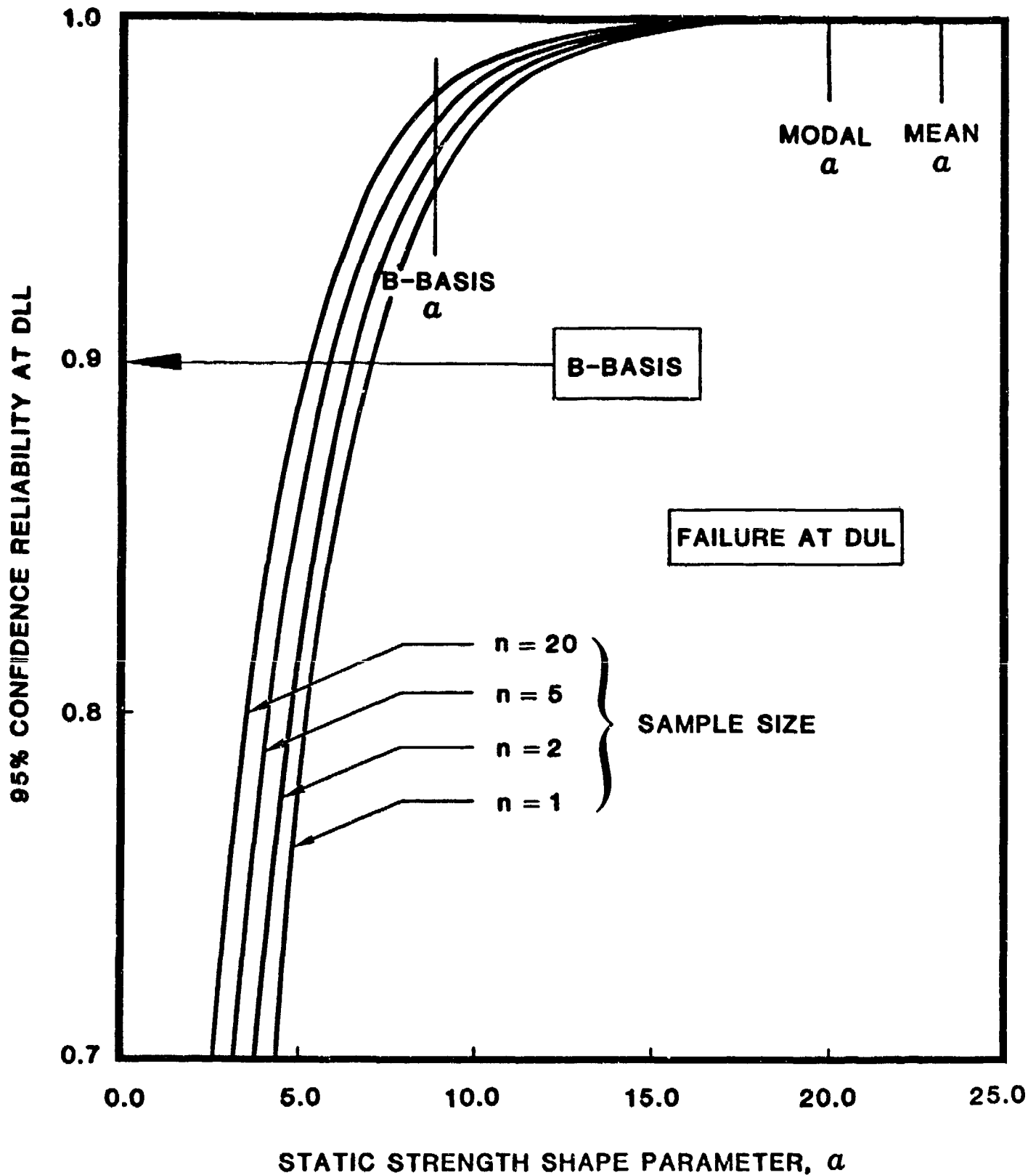
FIGURE 5. INFLUENCE OF WEIBULL SHAPE PARAMETER ON
DEMONSTRATED RELIABILITY OF DESIGN LIMIT LOAD
FOR FAILURE AT DESIGN ULTIMATE LOAD.
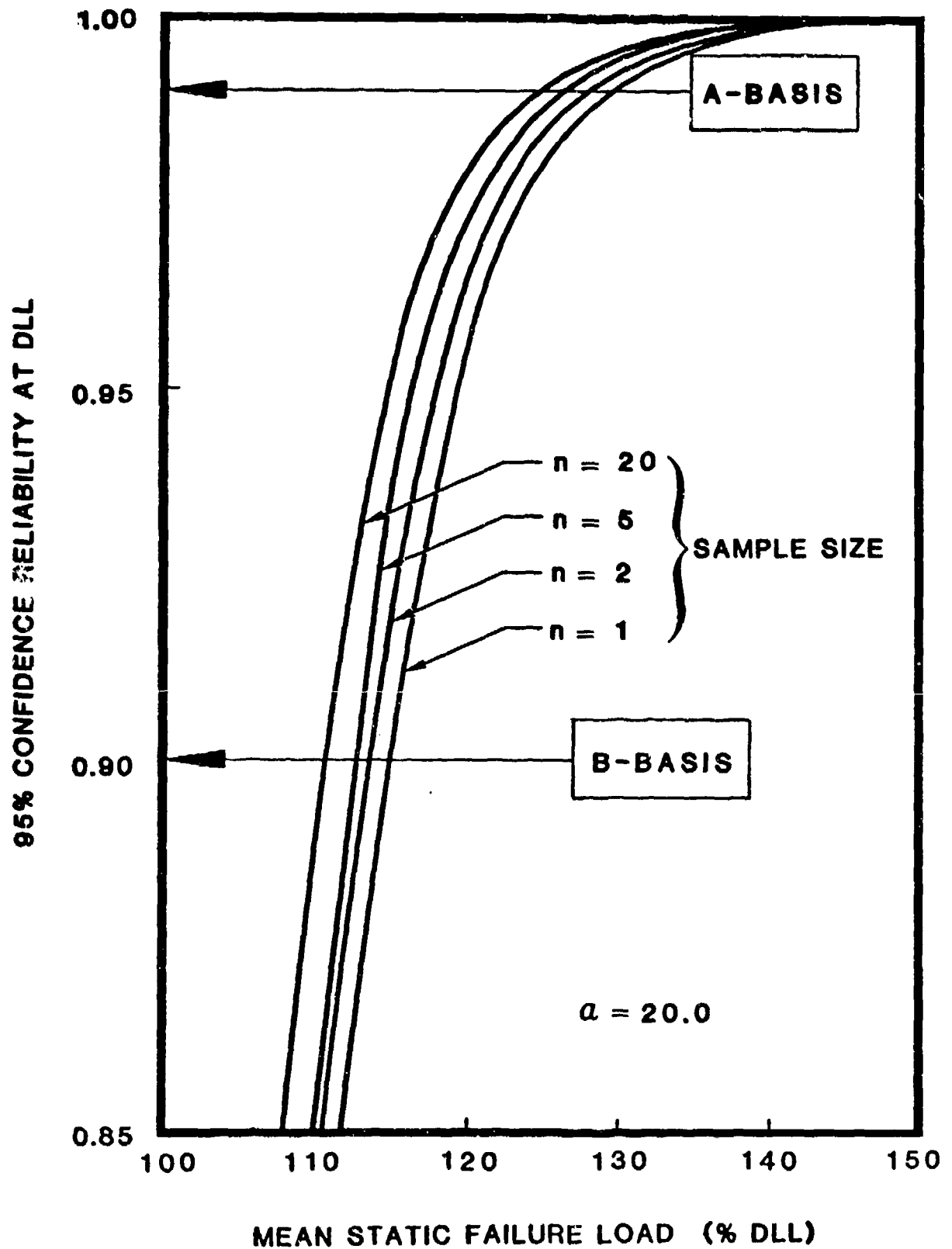
FIGURE 6. INFLUENCE OF MEAN STATIC FAILURE LOAD ON
DEMONSTRATED RELIABILITY AT DESIGN LIMIT LOAD.

The influence of sample size on the B-basis static strength to the mean failure load ratio ($\overset{\vee}{N}_B/\bar{x}$) is determined using Equation (2) and shown in Figure 7. This relation is shown for the mean (23.2), modal (20.0) and B-basis (8.8) values of $\alpha$. As can be seen in the figure, the B-basis to mean failure ratio increases with sample size. For $\alpha=20.0$ (modal value), the value of $\overset{\vee}{N}_B/\bar{x}$ only slightly increase when the sample size is larger than five.

The maximum operating load for a structure can be determined from the actual test results and a desired level of reliability (Equation (3)). This information is plotted in Figure 8. The figure shows the reliability for $\alpha = 20.0$. For B-basis reliability, the maximum load of the structure should not exceed 0.87 of the static test failure load for a sample size of one. Similarly, for A-basis reliability, the maximum load should remain below 0.77 of the failure load.

The effects of environmental knockdown factor on the resulting reliability is shown in Figure 9 (Equation (4)). The figure shows the reliability at different maximum operating load to mean static failure load ratios ($X/\bar{x}$). For a knockdown factor (k) of 1.1 with n = 1, the B-basis reliability maximum operation load is 0.79 of the failure load and the A-basis requires operating below 0.70 of the failure load. These values compare with 0.87 for B-basis and 0.77 for A-basis when no knockdown factor is applied. At k = 1.5 these values become 0.58 and 0.52 for B- and A-basis, respectively.

## 2.1.2    Fatigue Life Evaluation

The use of a fatigue scatter factor of 2-4 for metallic structures has historically been related to a reliability of approximately 699 in 700. Use of the scatter factors for composites cannot be justified because (from the analysis shown in Volume I), the Weibull shape parameter for composite fatigue life is extremely low (large scatter). The reliability of a two-lifetime test is approximately 1 in 4.
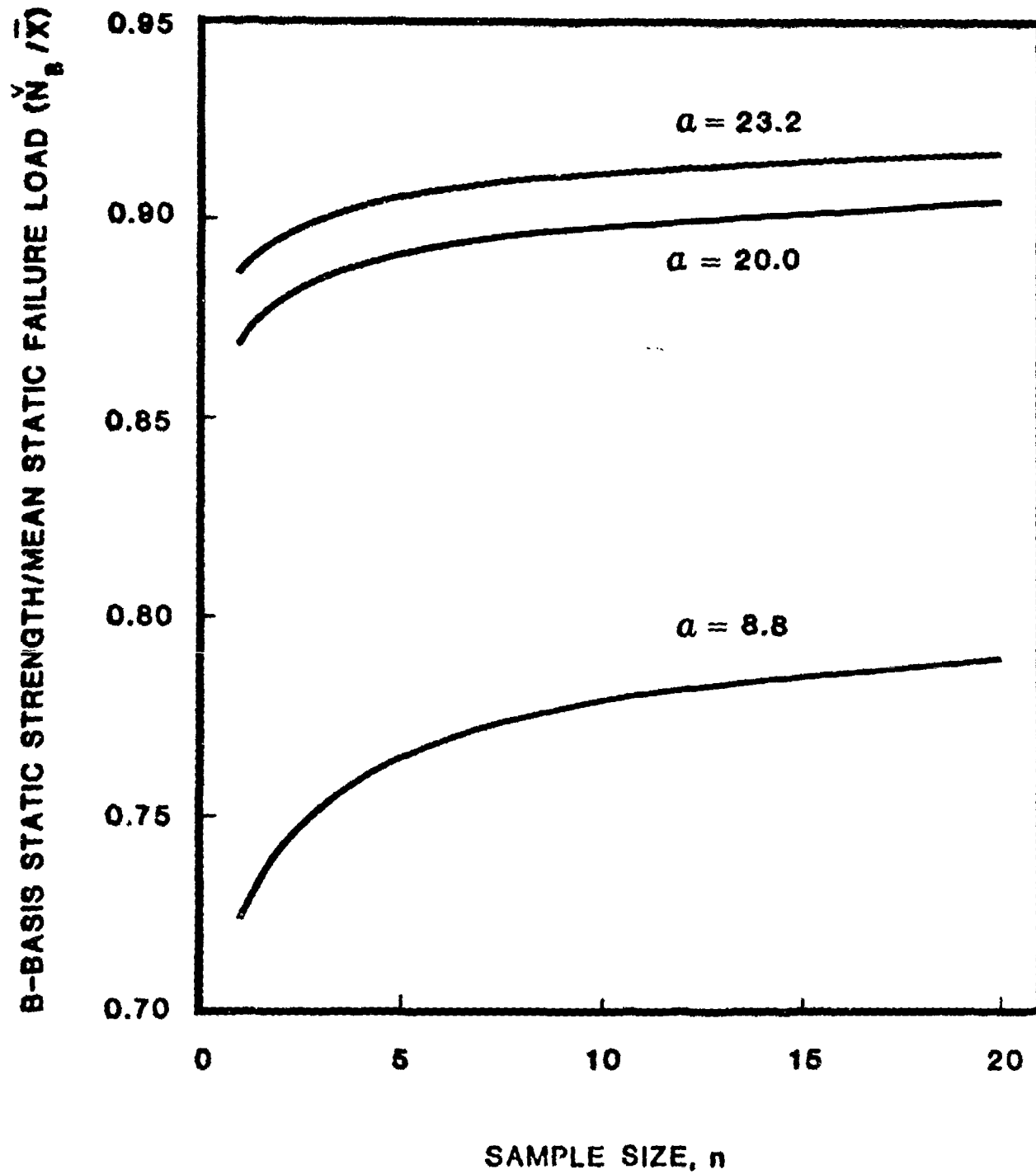
FIGURE 7. INFLUENCE OF SAMPLE SIZE ON B-BASIS
STATIC STRENGTH.

17
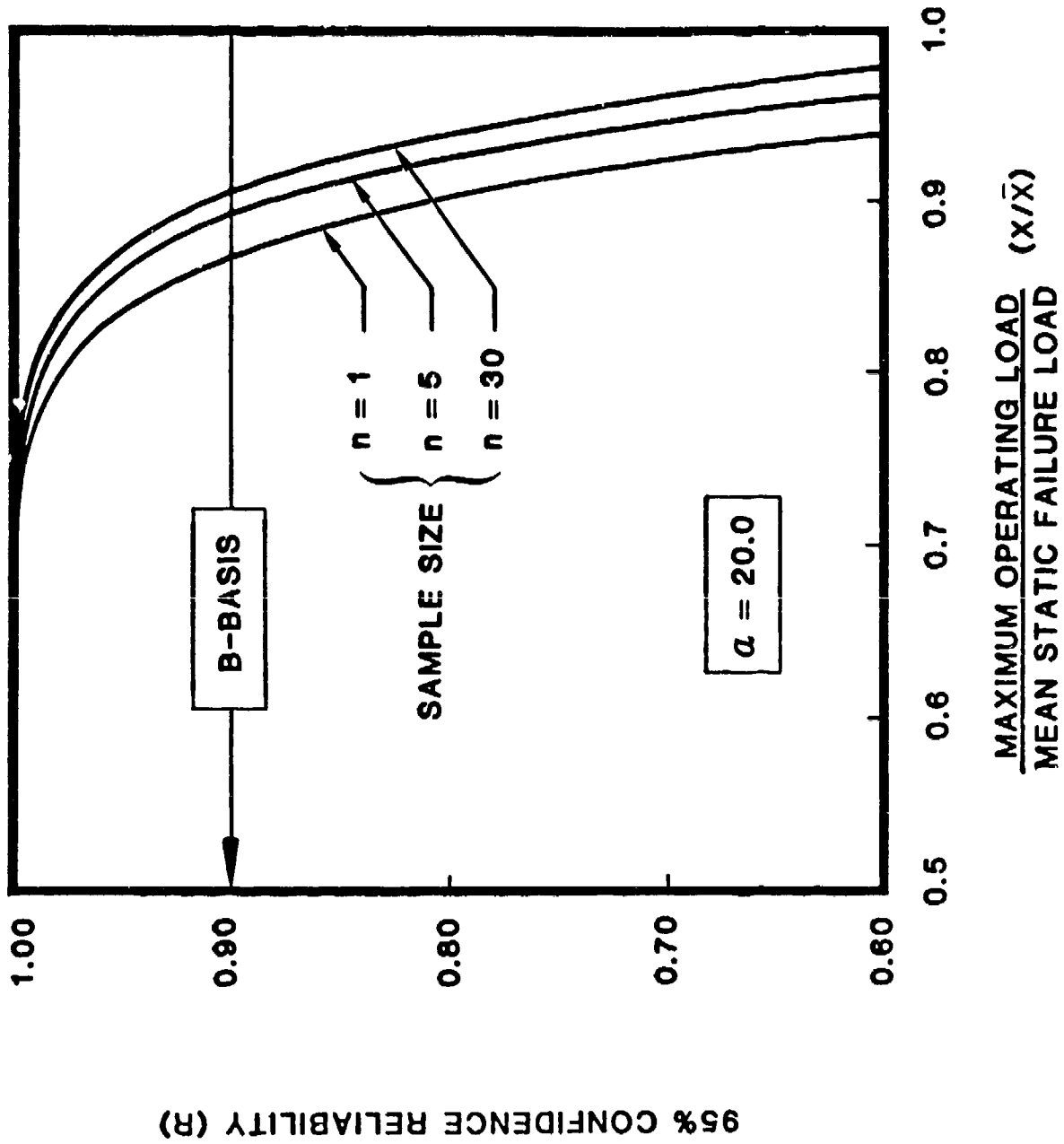
FIGURE 8. INFLUENCE OF MAXIMUM OPERATING LOAD TO MEAN STATIC FAILURE LOAD RATIO ON STATIC STRENGTH RELIABILITY.
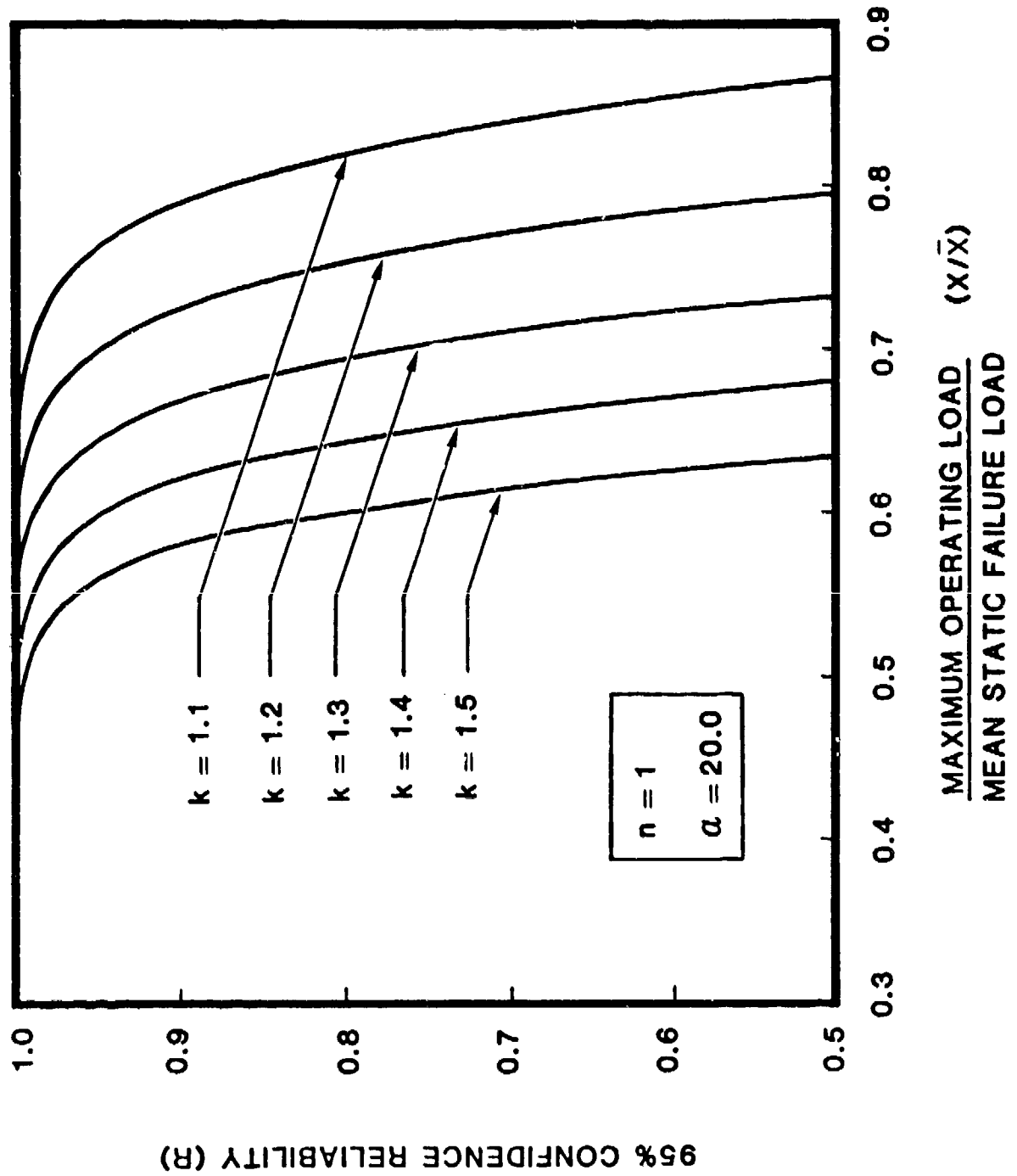
FIGURE 9. INFLUENCE OF ENVIRONMENTAL KNOCKDOWN FACTOR ON STATIC STRENGTH RELIABILITY.

The reliability of a two-lifetime fatigue test can be computed by letting $\bar{x} = 2.0$ in equation (1). The values of reliability are plotted as a function of the fatigue life shape parameter for different sample sizes in Figure 10. As can be seen from this figure, the two-lifetime fatigue test provides very low reliability for $\alpha$'s commonly found for composites ($\alpha\text{mean} = 2.17$ and $\alpha\text{modal} = 1.25$). At the modal value of $\alpha$, the reliability for sample size up to 10 is between 0.32 and 0.55 and is far below the B-basis allowable (R = 0.90). This indicates that a two-lifetime fatigue test does not provide the required level of reliability for composite structures. This conclusion is conservative for a two-lifetime fatigue test with no failure, since there is an unknown additional reliability associated with the assumption of failure at two lifetimes.

The influence of mean test fatigue lifetimes on the 95 percent confidence reliability achieved at one lifetime is shown in Figure 11. The figure shows that for a single article fatigue test, a minimum fatigue life of 13.6 lifetimes is required in order to achieve a B-basis reliability. The required mean life for a ten article fatigue test is 8.1 lifetimes.

The B-basis fatigue life to mean life ratio as a function of sample size at $\alpha = 1.25$ (modal value) and 2.17 (mean value is presented in Figure 12. Figure 12 shows that at the modal value of $\alpha$(1.25), and a sample size of 20, the B-basis to mean life ratio is 0.136. That is, a minimum life factor of 7.35 is required to obtain a B-basis reliability with 20 tests. For a sample size of one the required factor is 13.6. It should be noted that B-basis life reduction factor is very sensitive to the value of $\alpha$. For $\alpha = 2.17$ and a sample size of one, the B-basis reduction factor is reduced to 4.2 Figure 13 presents details of the influence of $\alpha$ on fatigue life ratio.

The 95 percent confidence reliability as a function of required fatigue life to mean failure life ratio $(X/\bar{x})$ is shown in Figure 14 for sample size of 1, 5 and 30. Figure 14 shows that the reliability becomes very low if the required fatigue life is greater than 0.2 of the mean fatigue test life.
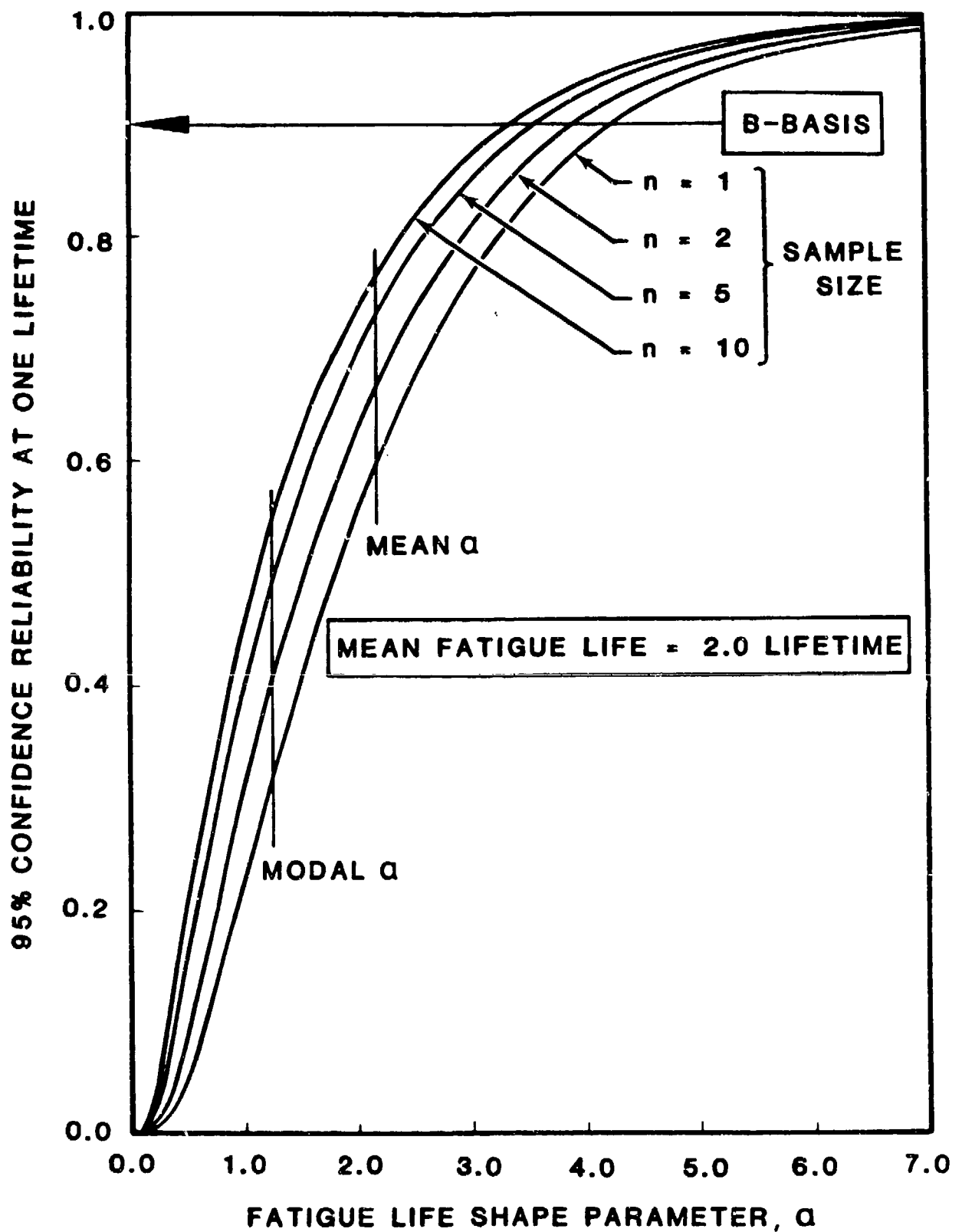
FIGURE 10. INFLUENCE OF FATIGUE LIFE SHAPE PARAMETER ON
RELIABILITY DEMONSTRATED BY A TWO-LIFETIME
FATIGUE TEST.
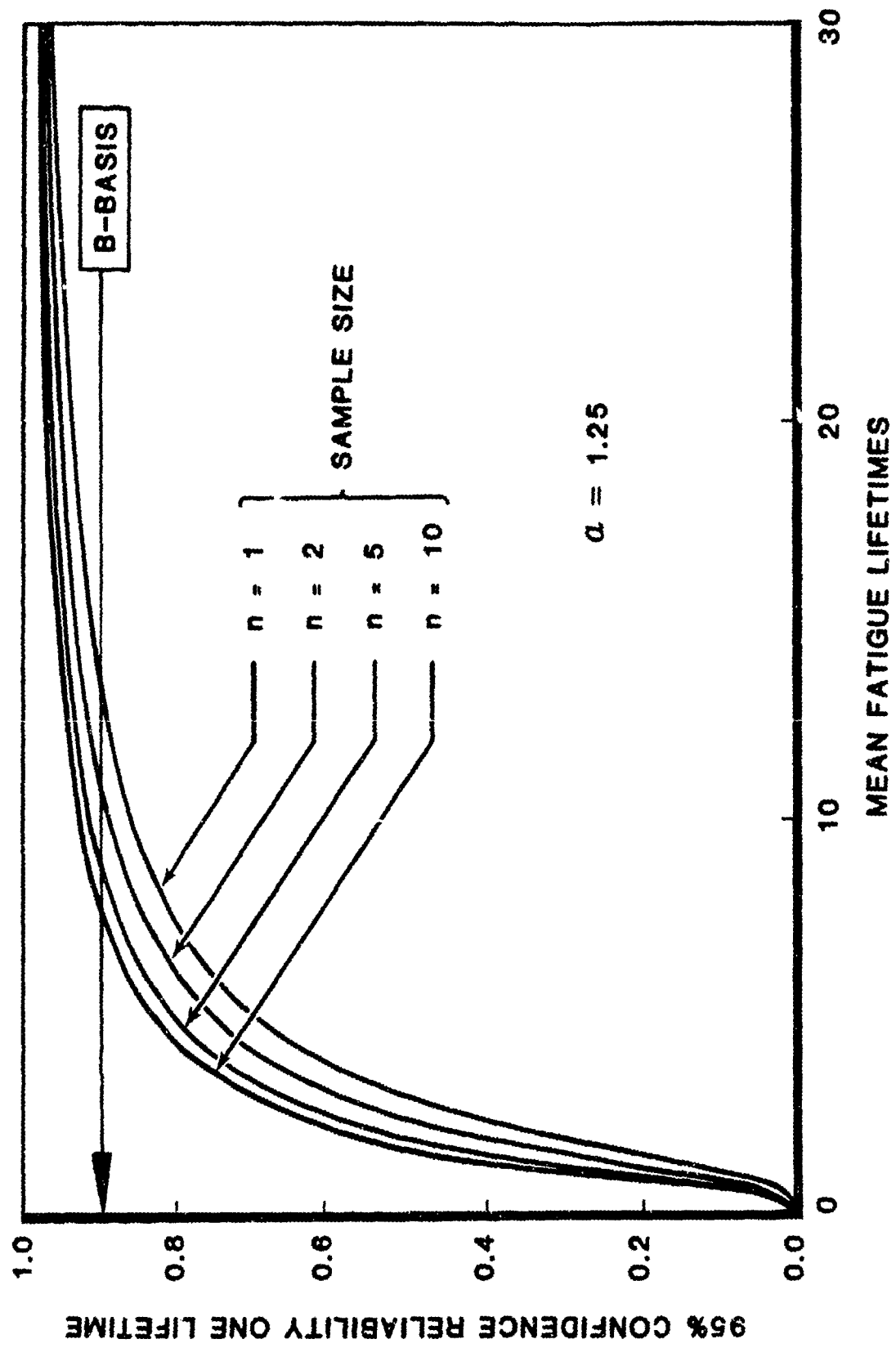
FIGURE 11. INFLUENCE OF MEAN TEST LIFETIME ON THE RELIABILITY ACHIEVED AT ONE LIFETIME.
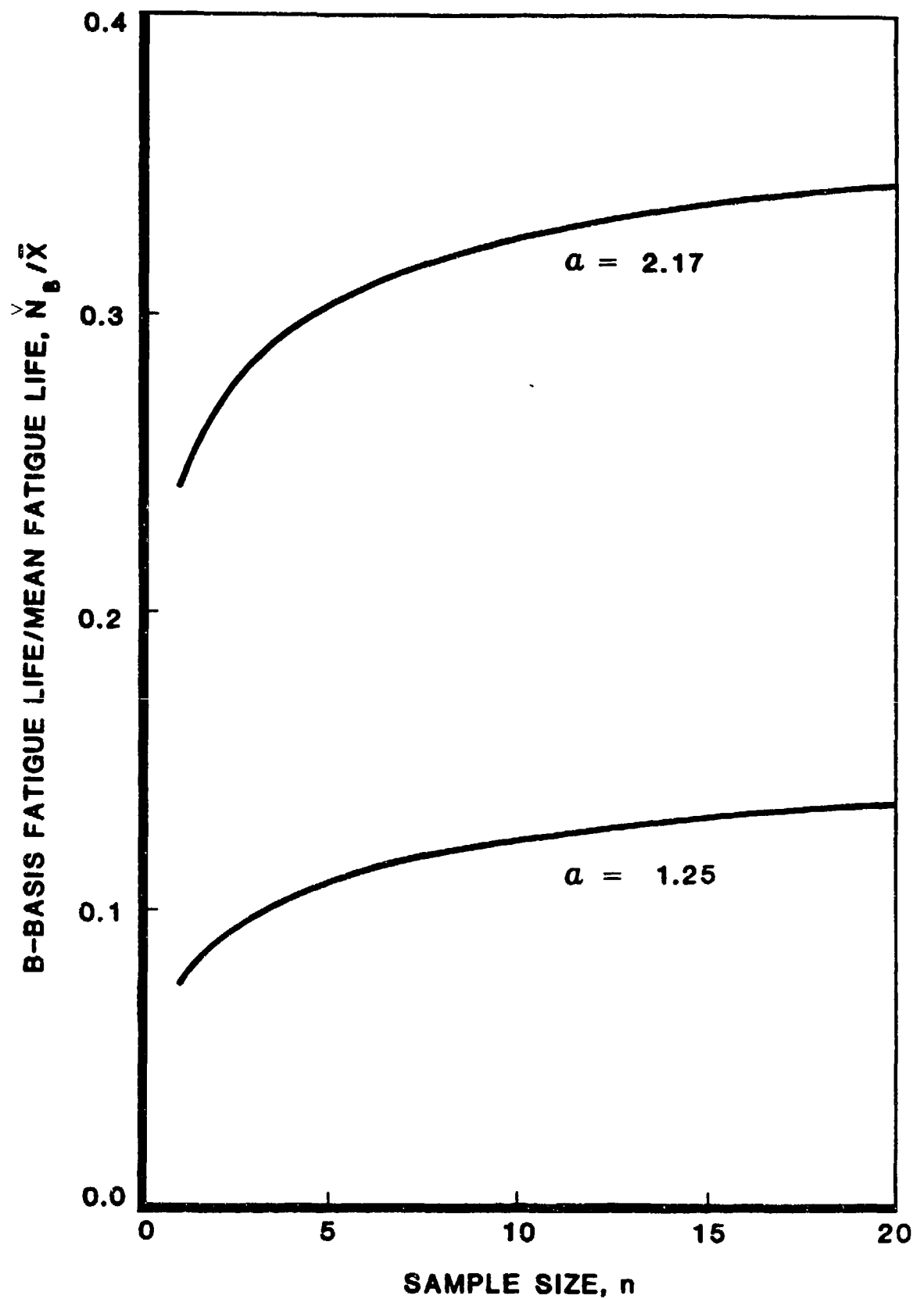
FIGURE 12. INFLUENCE OF SAMPLE SIZE ON B-BASIS FATIGUE LIFE
TO MEAN FATIGUE LIFE RATIO.
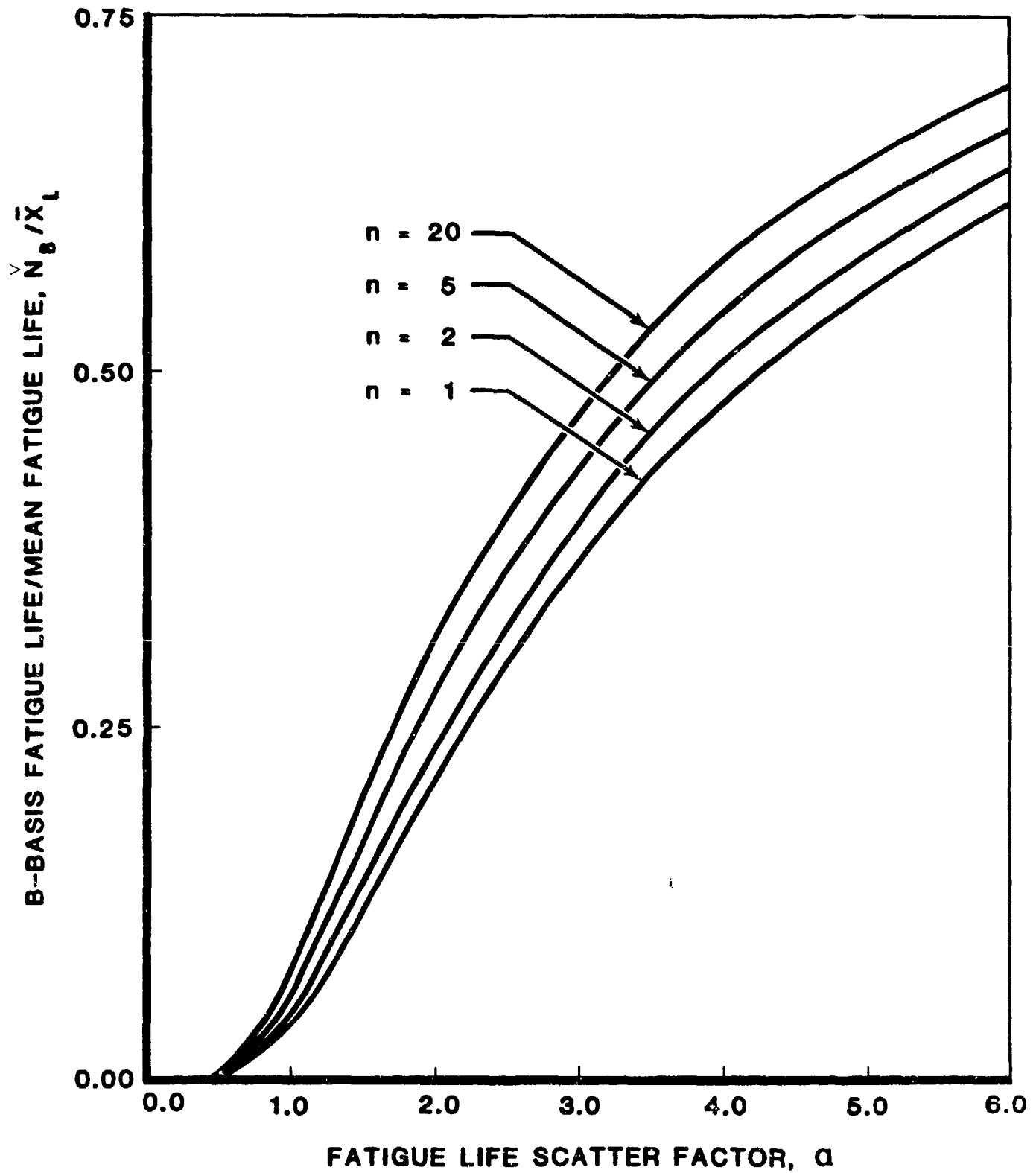
FIGURE 13. INFLUENCE OF FATIGUE LIFE SCATTER ON B-BASIS
FATIGUE LIFE TO MEAN FATIGUE LIFE RATIO.
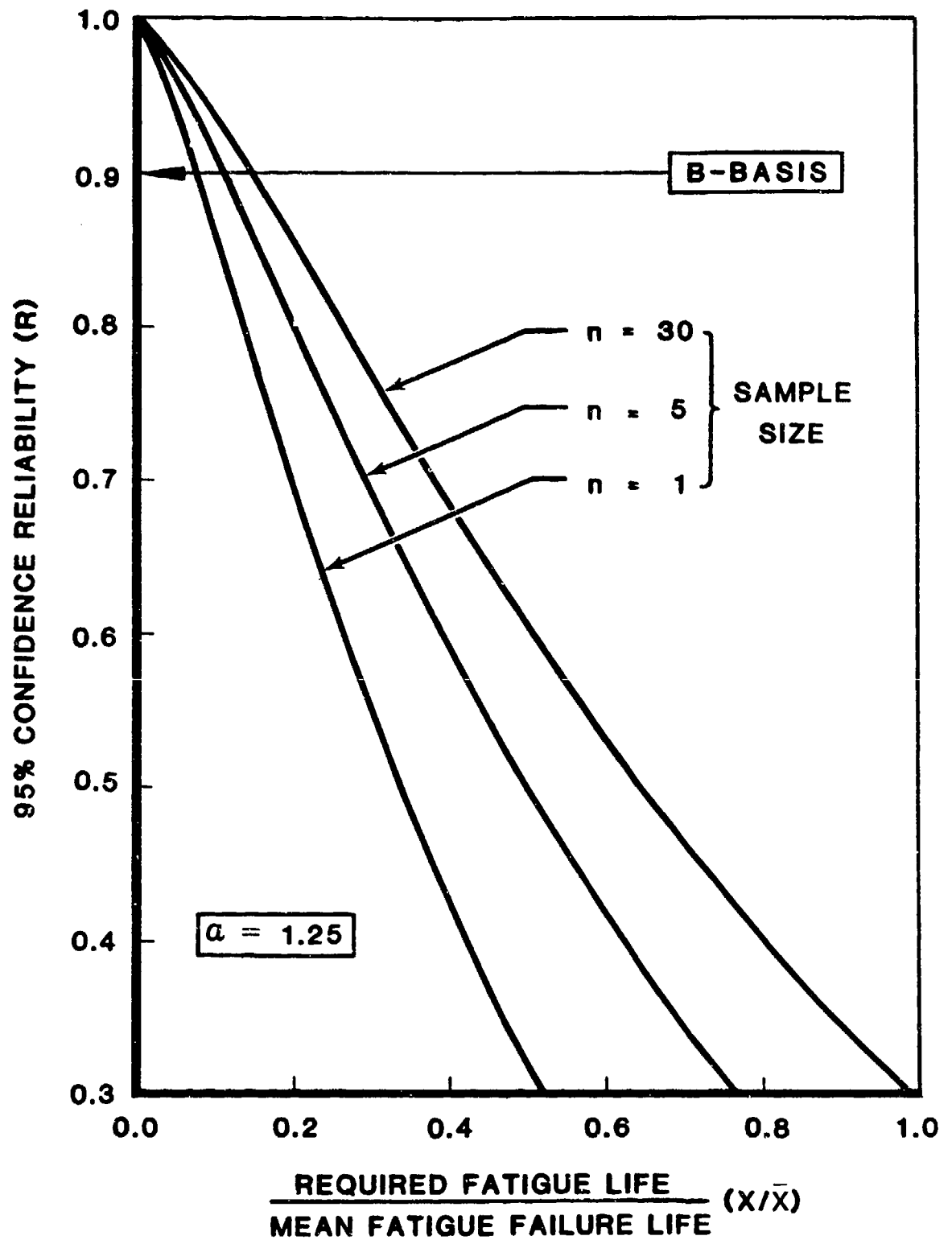
FIGURE 14. INFLUENCE OF REQUIRED FATIGUE LIFE TO MEAN TEST
LIFE RATIO ON FATIGUE RELIABILITY.

## 2.1.3     Assessment of the Navy Certification Approach

The Navy certification approach to primary composite structures was evolved approximately ten years ago.  It has been well documented in References 1 through 4.  The Navy composite certification approach was evolved from their extensive metallic certification experience.  As such, it relies heavily on the scatter factor approach.  An assessment of the Navy certification approach to static strength and fatigue life is discussed below.

### 2.1.3.1    Static Strength

Two requirements are specified for full-scale structural tests:

(1)    The actual room temperature failure load is required to exceed the 150 percent design limit load (DLL) value by a compensation factor dependent on failure location, failure mode, metal or composite structure, environmental test condition and material variability.

(2)    At 150 percent design limit load, all measured and extrapolated strains must not exceed the allowable strain level for the worst environmental condition.

These requirements are shown schematically in Figure 15 for the case of an environmentally sensitive failure in a composite upper skin.  From Figure 15 the requirements above are expressed as

$$\epsilon_T^{150\%} \leq \epsilon_D^E$$

$$P_{FP} \geq P_{FR}$$

$$\frac{P_{FR}}{P_{150\%}} \geq \frac{\epsilon_D^{RT}}{\epsilon_D^E}$$

The use of this approach for full-scale static test requirements is based on three key assumptions:
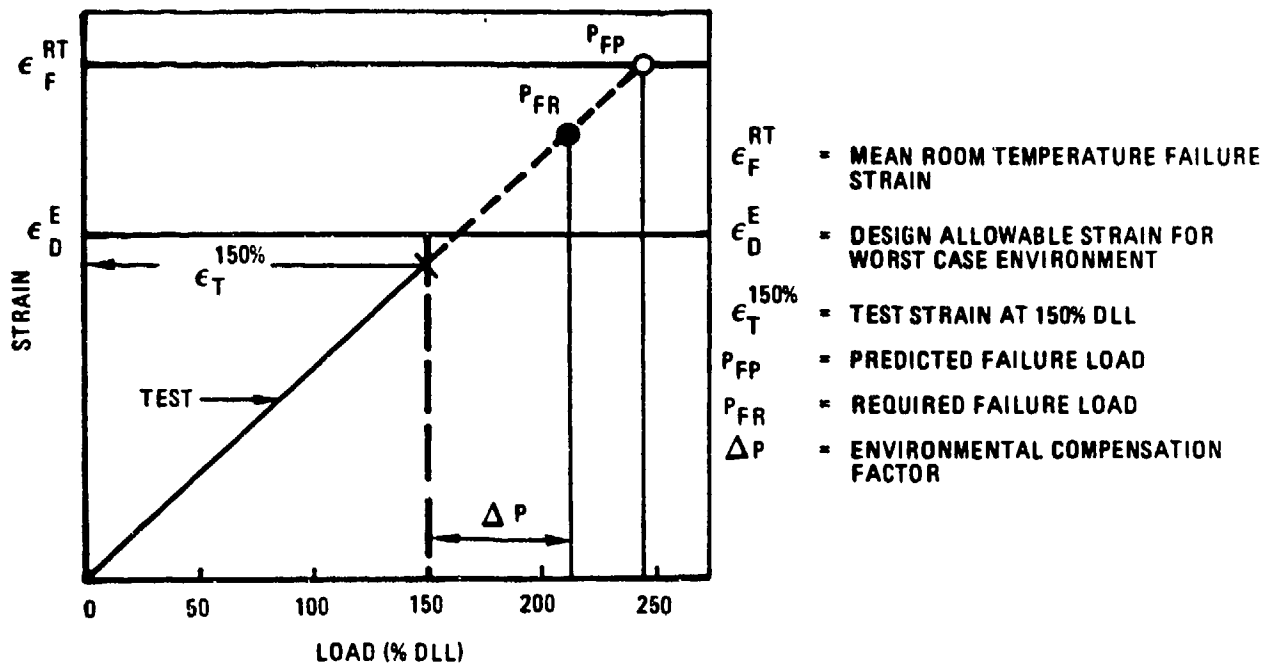
FIGURE 15. F/A-18A AND AV-8B LOAD-STRAIN ASSESSMENT CONCEPT.

Where the diagram on the right lists:

$\epsilon_F^{RT}$ = MEAN ROOM TEMPERATURE FAILURE STRAIN

$\epsilon_D^E$ = DESIGN ALLOWABLE STRAIN FOR WORST CASE ENVIRONMENT

$\epsilon_T^{150\%}$ = TEST STRAIN AT 150% DLL

$P_{FP}$ = PREDICTED FAILURE LOAD

$P_{FR}$ = REQUIRED FAILURE LOAD

$\Delta P$ = ENVIRONMENTAL COMPENSATION FACTOR



$$\epsilon_D = \epsilon_M^{RT} \times K_E \times K_{MV}$$

$$\text{WHERE} \quad K_E = \frac{\epsilon_M^{ETW}}{\epsilon_M^{RT}}, \quad K_{MV} = \frac{\epsilon_D^{RT}}{\epsilon_M^{RT}}$$
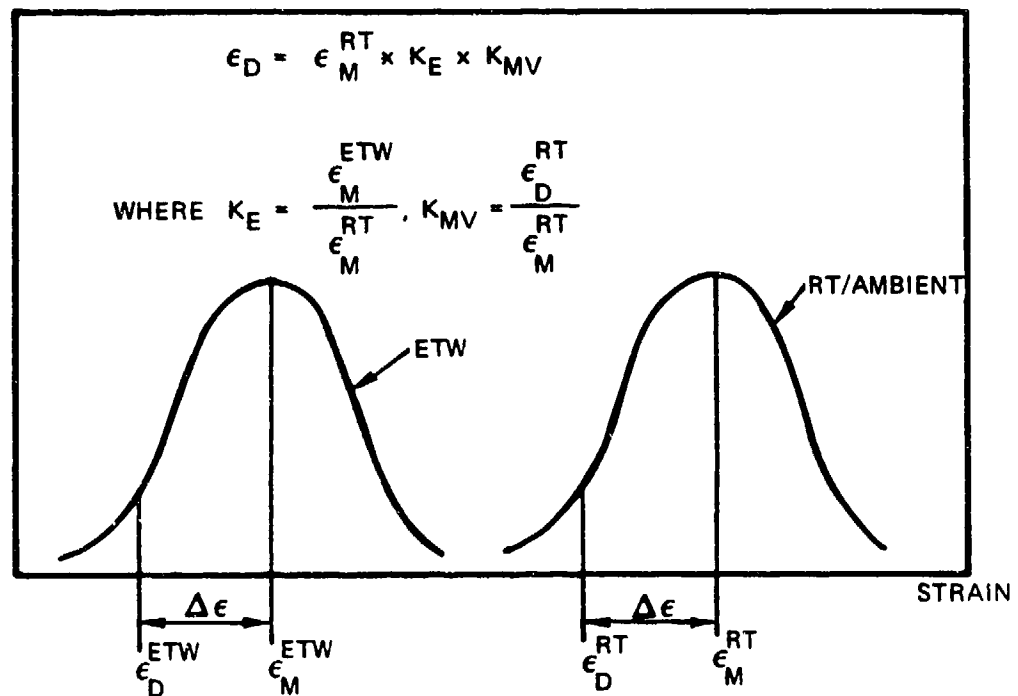
FIGURE 16. SCHEMATIC OF THE CALCULATION OF DESIGN ALLOWABLE STRAIN LEVEL (REFERENCE 1)

(1) The variability of composite strength data is independent of environment,

(2) Load-strain response is linear to structural failure, and,

(3) Failure modes are correctly predicted; i.e., no unexpected hot spot failure occurs.

The validity and implications of these assumptions for structural reliability are discussed below.

## Composite Static Strength Variability

Design strain allowables for each failure mode and its critical environment are determined by assuming composite strength variability are independent of environment. Figure 16 shows an example of this procedure for an environmentally sensitive failure mode (Reference 1). Initially RTA tests are carried out to determine mean strength and material variability. This permits determination of a room temperature design allowable strain, $\epsilon_D^{RT}$. Next, sufficient tests are carried out in the hot/wet environment to establish mean failure strain, $\epsilon_M^{ETW}$. Variability in the hot/wet environment is assumed to be the same as in the room temperature ambient environment. The hot/wet design allowable is then obtained as:

$$\epsilon_D^{ETW} = \frac{\epsilon_D^{RT}}{\epsilon_{Mean}^{RT}} \cdot \epsilon_{Mean}^{ETW} \tag{5}$$

This design allowable, $\epsilon_D^{ETW}$, accounts for environmental sensitivity and inherent material variability, but not structural response variability.

The influence of environment on tension and compression static strength scatter was determined in Task I (see Volume I). A large static strength data base was analyzed. It was shown that for tension failures static strength variability was higher

in ETW environments. In contrast, for the hot/wet sensitive compression failures no influence of environment was observed. However, the work in Task I showed that failure mode more significicantly influenced static strength scatter. The work in Reference 5 also showed that failure mode exercised the greatest effect on static strength scatter.

The test data analysis in Task I, therefore, sugges s that the assumption of environmental independence for static strength scatter may not always be strictly correct. However, the requirement specified in the Navy approach, which mandates the determination of strength variability for each failure mode, is a key requirement because of its dominating influence on static strength variability.

In general, it can be concluded that the Navy approach to static strength variability is soundly based, despite evidence of some influence of environment on strength variability.

The soundness or conservatism of the Navy approach can be improved by specifying conservative strength variability knockdown factors for each failure mode. This would tend to take into account any dependence of variability on environment.

### Linear Load-Strain Response

The second assumption in the Navy certification approach is that of linear load-strain response. This was justified because previous testing of composite structures consistently showed linear load-strain response. However, work in Reference 5 has shown that nonlinear load-strain response can occur under severe hot/wet test conditions. An example, from Reference 5, of nonlinear upper skin load-strain response in a box beam under 250°F/1.3% moisture conditions is shown in Figure 60. It is discussed in more detail in Section 3 of this report.

Figure 17 shows the implications of nonlinear ETW load-strain response for the Navy certification approach. The RTA test load-strain response shown fulfills all certification re-
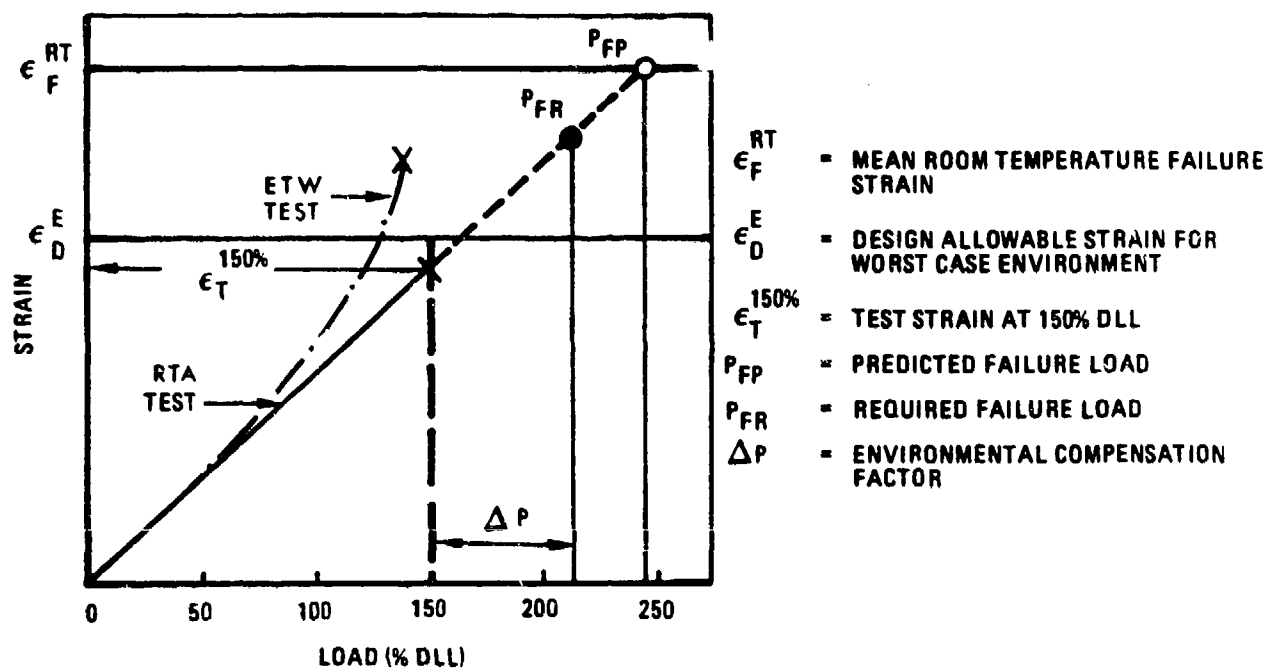
FIGURE 17.   INFLUENCE OF NONLINEAR LOAD-STRAIN RESPONSE
ON THE F/A-18A AND AV-8B LOAD-STRAIN
ASSESSMENT CONCEPT.

quirements and demonstrates adequate margin for the environmental compensation factor. The ETW load-strain response plotted in Figure 17 represents a worst case scenario. The nonlinear ETW response causes the ETW strain allowable $\epsilon_D^E$ to be exceeded before design ultimate load, which leads to failure at less than design ultimate load. This example shows the potential danger in using RTA tests with environmental compensation factors to demonstrate certification compliance. It is important, therefore, in composite structures to conduct environmental tests on realistic subcomponents in order to avoid the scenario described above. The amount and complexity of the environmental testing will be a function of the aircraft service environment. Alternatively, use of composite materials in their nonlinear load-strain response region should be avoided.

Nonlinear load-strain response in composites is induced in ETW environments by loss of resin mechanical properties. Resin-controlled properties, such as compression strength, are the most sensitive to ETW environments. Strength loss and nonlinear load-strain response occur as the glass transition temperature of the material is approached. Good design practice dictates that composites should not be used in this regime. These problems can, therefore, be avoided by setting the Material Operating Limit (MOL) at a safe margin below the glass transition temperature.

The concept of a Material Operating Limit is discussed in detail in Reference 5 and is shown schematically in Figure 18 for an environmentally sensitive property. The decrease in design allowable strain as temperature increases is shown for a constant moisture level. The glass transition temperature (Tg) of the material coincides with a catastrophic rate of strength loss. In order to operate in a safe regime, the MOL should be reduced below the Tg by a safety factor K. This produces the shaded service operational envelope for the material shown in Figure 18. Figure 19 shows how the MOL varies with moisture level, such that a series of MOL's are produced for various
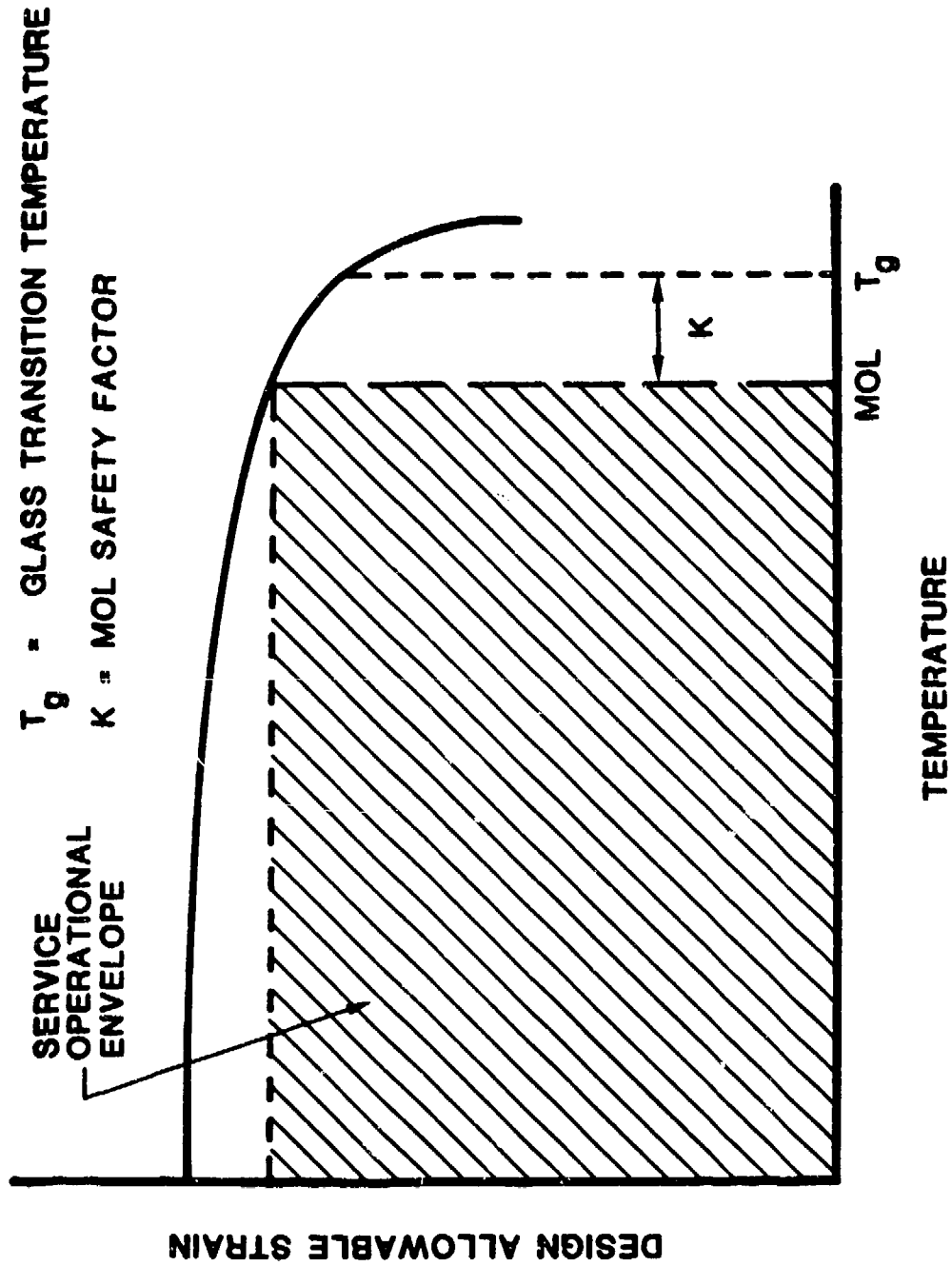
FIGURE 18. MATERIAL SELECTION CRITERION.
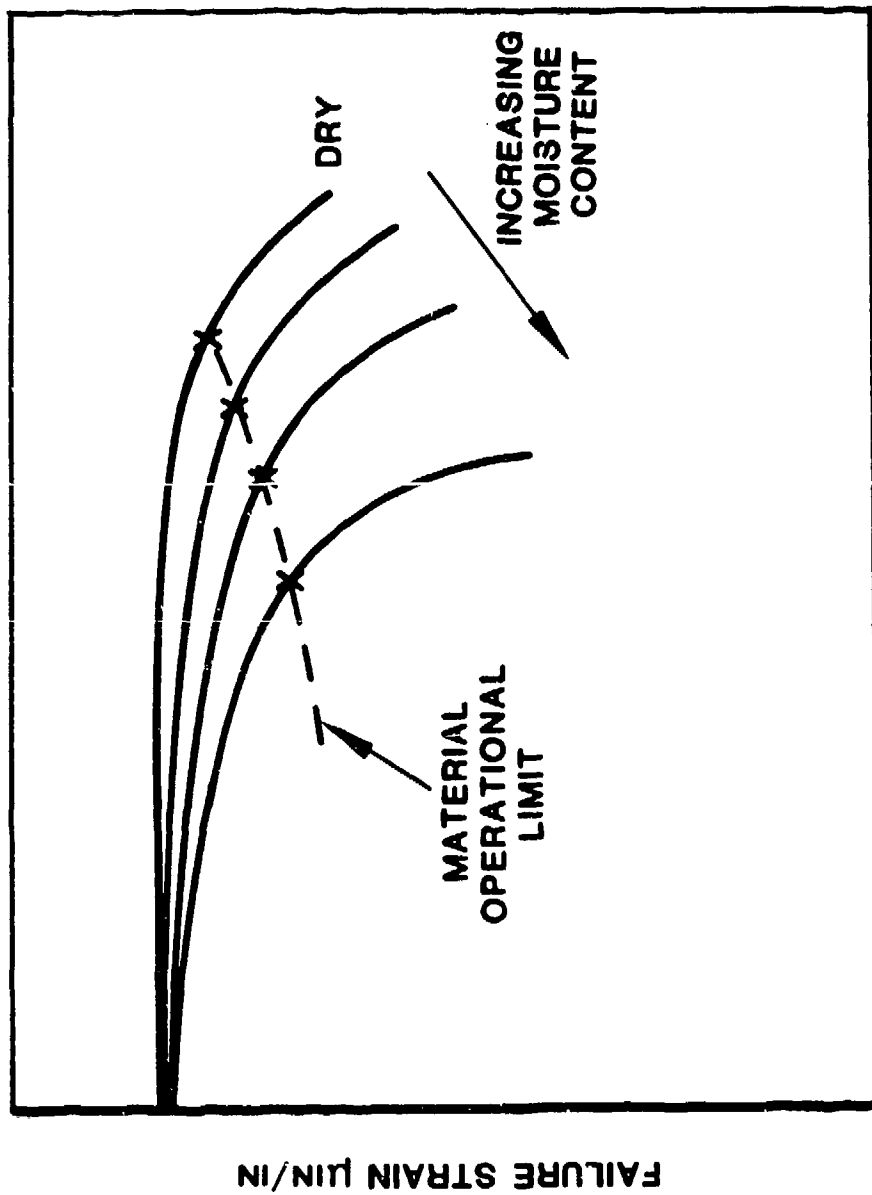
TEMPERATURE °F

FIGURE 19. INFLUENCE OF TEMPERATURE AND MOISTURE ON
MATERIAL OPERATIONAL LIMITS.

moisture levels. If the service environment (temperature/moisture) of the aircraft falls within the MOL, excessive environmental effects can be avoided. If this procedure is adopted, the assumption of linear load-strain response is soundly based.

## Correct Failure Mode Prediction

Use of an environmental compensation factor in an RT/ambient static test, requires the correct prediction of both RTA and ETW failure modes.

A typical application of this approach, would be a RTA test of a wing structure whose critical failure mode is upper skin failure under ETW conditions. In this application, the load-strain response shown in Figure 15 would represent the RTA upper skin load-strain response. The environmental compensation factor $\Delta p$ ($P_{DUL}$ to $P_{FR}$) represents the strength loss at ETW conditions. Extrapolation of the strain data to $P_{FP}$ assumes the correctly predicted upper skin failure mode occurs. If an unanticipated ETW failure mode occurs, it is possible that failure could occur at less than DUL.

This is an important observation because prediction of previous composite full-scale static test failure modes has been inadequate. Examples are given in References 5 through 7. In Reference 5, the predicted RTA failure mode was lower skin failure; however, the observed failure mode was failure of the intermediate spar/lower skin cocured joints. Despite this, the wing box sustained 122 percent of DUL at failure.

Reference 6 reported that the failure mode of a Jaguar aircraft composite wing box was different from that predicted with lower complexity level specimen tests. In Reference 7, the failure modes of the three full-scale structural tests from the ACEE program were reviewed. In all three tests, static failure were induced by unanticipated failure modes. The majority of unexpected failure modes discussed in References 5 through 7 can be attributed to the sensitivity of composites to out-of-plane

loading. This is caused by their relatively low interlaminar tension and shear strength. The major sources of secondary loads are those induced directly such as fuel pressure loading and those induced indirectly by eccentricities, irregular shapes, stiffness changes, discontinuities and loading in the postbuckling regime. The lesson learned from these experiences is that composite structures are much more sensitive to secondary load induced hot spots than metals. Therefore, great care should be exercised in understanding and accounting for these loads in the design process.

The conclusion from the above discussion on the Navy assumption of predicted failure mode is as follows. The failure modes of full-scale composite structures cannot currently be predicted with great confidence. Therefore, a certification process which assumes correct prediction of a full-scale structural failure mode must carry some degree of risk.

2.1.3.2    Fatigue Strength

The Navy approach to fatigue certification of composite structures is similar to that adopted for metals: that is, a two-lifetime RTA fatigue test with a severe design spectrum. This approach has proved successful for identifying fatigue hot spots in metallic structures.

The major problem in certifying composite structures is related to their excellent fatigue resistance. This causes their S-N curves to be relatively flat with significantly higher data scatter. Thus, a two-lifetime test on a composite structure demonstrates a lower reliability than for metal structures.

2.1.4    Example Reliability Calculations

In this subsection, a static strength and a fatigue life problem are selected to demonstrate the reliability calculation procedure using the scatter factor approach. Composite structure as well as metal structure are used in both examples.

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## 2.1.4.1    Static Strength Reliability

Determine the requirements necessary to demonstrate B-basis static strength reliability for an (a) all composite, (b) all metal wing structure of a supersonic fighter aircraft. Assume the following:

(1)    A maximum service temperature of $220^{\circ}F$

(2)    Environmental (ETW) knockdown factor $k^c = 1.2$, $k^m = 1.0$

(3)    Static strength variability $\alpha_s^c = 20$, $\alpha_s^m = 25$

(4)    Maximum operating load, $X = 1.25 \times DLL = 0.833 \times DUL$

(5)    One full-scale test article, $n = 1$

(6)    Ignore structural response variability, SRV = 0

### Composite Wing

From Figure 9, for $k^c = 1.2$ and B-basis reliability (95 percent confidence, 90 percent reliability), the maximum operating load (X) to static failure load ($\bar{x}$) ratio is

$$\frac{X}{\bar{x}} = 0.725$$

Required failure load $\bar{x} = x/0.725$

For $X = 0.833 \times DUL$:

$$\bar{x} = \frac{0.833}{0.725} \ DUL = 1.15 \times DUL$$

Therefore, the static test article must exceed 115 percent DUL to demonstrate B-basis reliability at the maximum operating load of 125 percent DLL.

### Metal Wing

For B-basis reliability, the maximum operating load to static failure load ratio for $\alpha_s^m = 25$ and $k^m = 1.0$ is calculated to be

$$\frac{X}{\bar{x}} = 0.894$$

or the required failure load

$$\bar{x} = \frac{X}{0.894}$$

For  X = 0.833 x DUL

$$\bar{x} = \frac{0.833}{0.894} \text{ x DUL} = 0.932 \text{ x DUL}$$

Therefore, the static strength of the test article must exceed 93 percent of DUL to accommodate B-basis reliability at the maximum operating load of 125 percent DLL.

The calculations above show that, for the example cited, the composite wing would have to achieve a 24 percent higher failure load in order to demonstrate the same structural reliability.

Mixed composite/metal structures pose special certification problems. These have been discussed in detail in Reference 8. In this reference, it was shown that because of their lower variability in mechanical properties and lower environmental sensitivity, the metal portion of a mixed structure would fail first in a RTA test. Historically, metal structures have exhibited full-scale test failure loads of approximately 105 percent DUL. Thus, for a mixed structure where a metal failure occurs at 105 percent DUL, the reliability of the composite structure will be unknown. The proven reliability of the composite structure (based on a minimum strength of 105 percent DUL) will be only 0.54. Thus, certification of both parts of a mixed structure to the same reliability may be difficult to achieve.

2.1.4.2   <u>Fatigue Life Reliability</u>

Determine the requirements necessary to demonstrate B-basis fatigue life reliability for an (a) all composite, (b) all metallic wing structure of a supersonic fighter aircraft. Assume the following:

(1) Fatigue life variability $\alpha_L^c = 1.25$, $\alpha_S^m = 7.5$

(2) One full-scale test article, n = 1

(3) Ignore structural response variability, SRV = 0.

## Composite Wing

The life factor requirement for B-basis reliability for $\alpha_L^c = 1.25$ and n = 1 can be obtained from Figure 11, it is 13.6. Thus, in order to demonstrate B-basis reliability at one lifetime, a successful test to 14 lifetimes must be achieved.

## Metal Wing

The life factor for $\alpha_L^m = 7.5$ and n = 1 is calculated to be 1.5. Thus, a two-lifetime fatigue test is more than adequate to demonstrate B-basis reliability at one lifetime. Actual reliability for a two-lifetime test is 0.99.

The Navy two-lifetime certification test philosophy for composites demonstrates only 0.32 reliability at one lifetime. However, if actual service usage is less severe than the conservative design spectrum, this reliability will be improved. For example, if the actual service loads are 1.13 times lower than the design spectrum a B-basis reliability at one lifetime will be achieved. Thus, the reliability of Navy aircraft will vary and will depend on the conservatism of the severe design spectrum.

The RTA full-scale fatigue test does not account for any environmental fatigue effects. The influence of environment on composite fatigue life has been comprehensively investigated in Reference 5. The issue of environmental test simulation for a composite certification program was addressed. It was shown that the requirements for environmental simulation were closely related to the aircraft temperature spectrum and the relationship between load factor and temperature.

Typical examples, from Reference 5, of these relationships for a fighter aircraft are shown in Figures 20 and 21, respectively. The Mach 2 class aircraft utilized in Reference 5 had a maximum service temperature of 242°F. This design tempera-
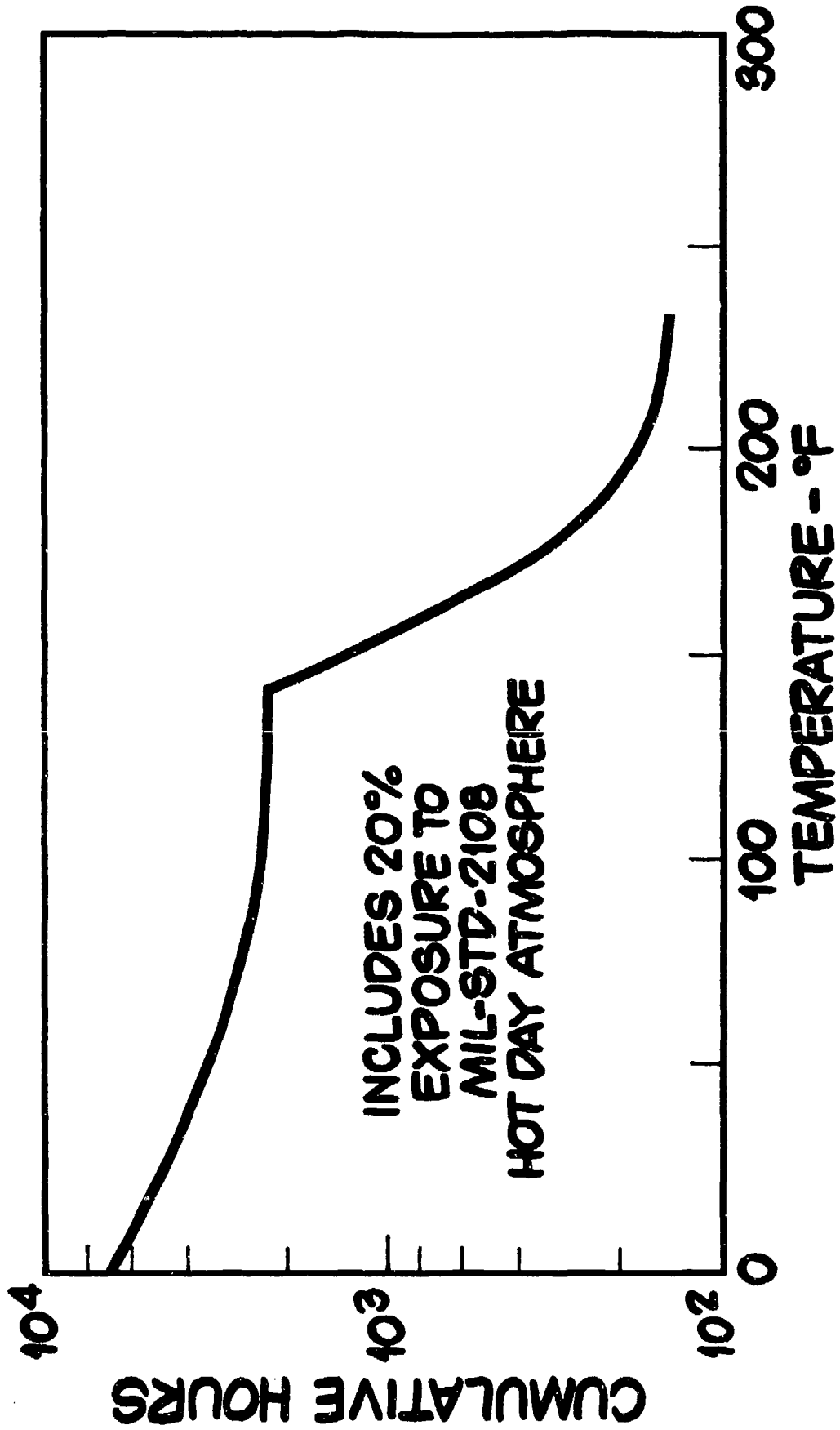
FIGURE 20. TEMPERATURE SPECTRUM FOR MACH 2 FIGHTER AIRCRAFT IN REFERENCE 5.
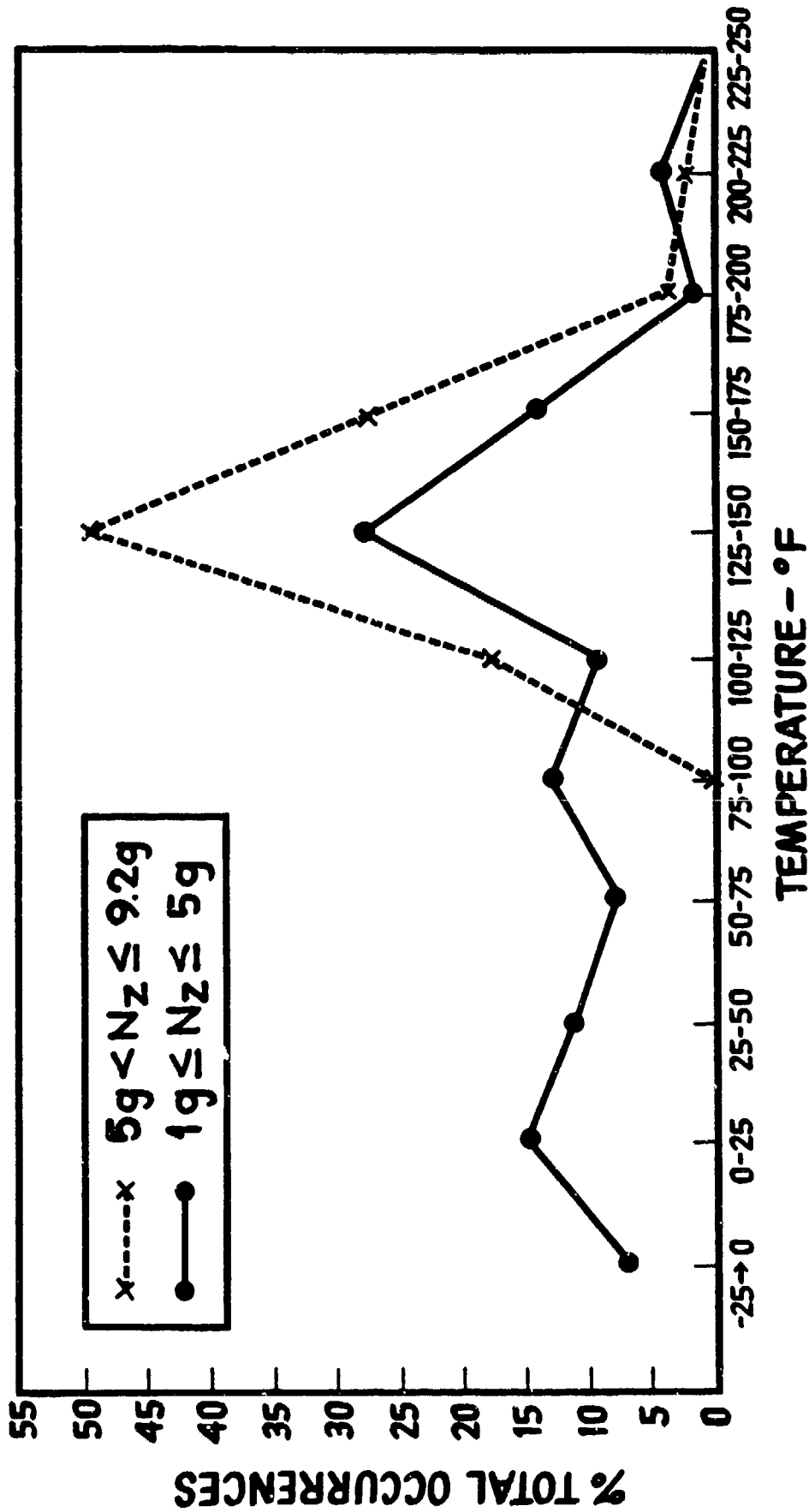
FIGURE 21. LOAD-FACTOR TEMPERATURE RELATIONSHIP FOR MACH 2 FIGHTER AIRCRAFT IN REFERENCE 5.

ture was associated with the most critical static design case.
However, the fatigue temperature spectrum shown in Figure 20
shows two interesting features. First, the maximum temperature
in the fatigue spectrum is 220°F and, second, the aircraft spends
less than 200 hours/lifetime at temperatures above 200°F. This
represents approximately only 2.5 percent of the total life. The
vast majority of the aircraft life is spent at semi-ambient
temperatures.

Figure 21 shows the load factor-temperature relation-
ship of the Reference 5 aircraft. Noticeable features of the
relationship are (1) the low number of load factor occurrences
above 200°F, less than 5 percent of the total, (2) the vast
majority of 5g to 9g loads (the most fatigue damaging) occur at
semi-ambient temperature, 100°F to 175°F.

Because of the low cumulative time at high temperatures
and the semi-ambient temperatures associated with high load
factors, it can be concluded that full-scale RTA testing is
satisfactory for fatigue certification of this type of fighter
aircraft.

The implications of the load-temperature relationships
and the test data generated in Reference 5 for environmental test
complexity have been discussed in more comprehensive detail in
References 9 through 13.

Since the relationship shown in Figures 20 and 21 are
considered to be typical for current fighter aircraft, it can be
concluded that the Navy philosophy is a reasonable approach for
fighter aircraft. It should be noted, however, that, for air-
craft which have significant periods of cumulative time at high
temperatures and/or combinations of high load factors and high
temperature, RTA tests may not be adequate. ETW subcomponent
tests may be required under these circumstances.

2.2     Load Enhancement Factor Approach

The objective of this approach is to increase the
applied loads in the fatigue certification tests so that the same

level of reliability can be achieved with a shorter test dura-
tion. A schematic showing this approach is shown in Figure 22
where the fatigue life scatter represented is typical of that
observed in composites. At one fatigue lifetime a typical resid-
ual strength distribution is shown. If the maximum applied load
in the fatigue test ($P_F$) is increased to the mean residual
strength at one lifetime ($P_T$), then the B-basis residual strength
of the structure would be equivalent to the design maximum
fatigue stress.

Thus, a successful fatigue test to one lifetime at
applied stress $P_T$ or a fatigue test to $N_F$ lifetimes at applied
stress $P_F$ would both demonstrate B-basis reliability. In addi-
tion, combinations of the load enhancement and fatigue life
factors could also be used to demonstrate B-basis life. In order
to use this approach with confidence in a certification methodol-
ogy, a formal relationship between the load enhancement factor
(LEF) and the life factor is required.

The fatigue life factor for a B-Basis reliability at
one fatigue lifetime can be derived from the basic Weibull dis-
tribution, and is given by

$$N_F = \frac{\Gamma\left(\frac{\alpha_L + 1}{\alpha_L}\right)}{\left[\frac{-\ln(0.9)}{\chi_\gamma^2 (2n)/2n}\right]^{1/\alpha_L}} \tag{6}$$

where $N_F$ is the life factor for B-basis reliability at one life-
time.

The residual strength distribution at a certain fatigue
lifetime can be described as a two-parameter Weibull distribu-
tion, as in the static strength distribution. Let $\alpha_R$ and $\beta_R$ be
the shape and scale parameters of the residual strength distribu-
tion and $P_T$ be the mean residual strength. then $P_T$ can be writ-
ten as
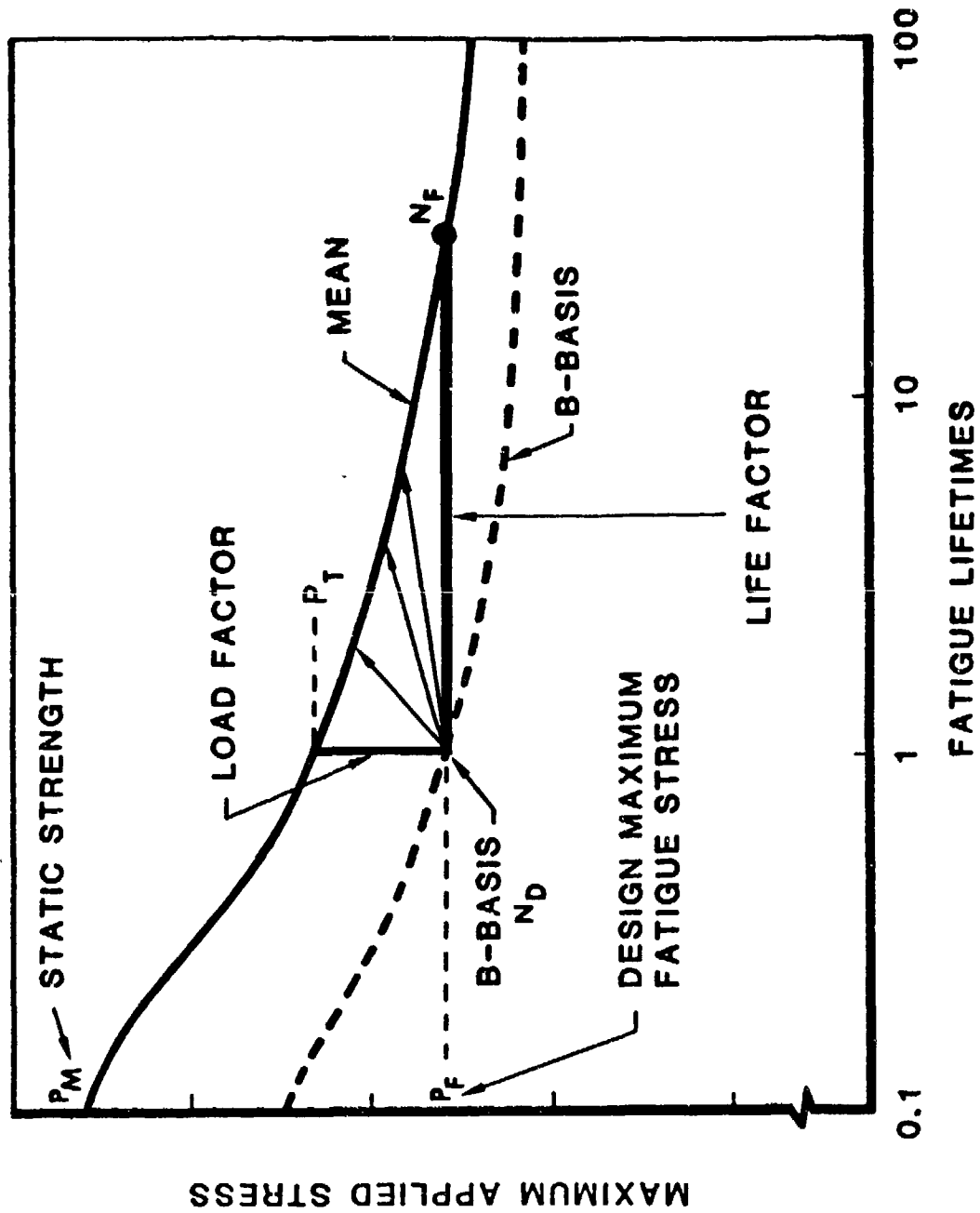
FIGURE 22. SCHEMATIC OF LOAD ENHANCEMENT FACTOR APPROACH FOR COMPOSITES.

$$P_T = \beta_R \Gamma\left(\frac{\alpha_R+1}{\alpha_R}\right) \tag{7}$$

and the B-basis residual strength is

$$\overset{v}{N}_R = \beta_R \left[\frac{-\ln(0.9)}{\chi_\gamma^2(2n)/2n}\right]^{1/\alpha_R} \tag{8}$$

From Figure 22, the load enhancement factor F is defined as

$$F = \frac{P_T}{P_F} \tag{9}$$

where $P_T$ is the maximum fatigue test load

$P_F$ is the maximum design fatigue load.

Since the load enhancement factor approach provides the same reliability as the life factor approach, the factor F can be written as

$$F = \frac{\mu P_T}{\overset{v}{N}_R} = \frac{\mu \Gamma\left(\frac{\alpha_R+1}{\alpha_R}\right)}{\left[\frac{-\ln(0.9)}{\chi_\gamma^2(2n)/2n}\right]^{1/\alpha_R}} \tag{10}$$

where $\mu$ is a coefficient which requires that the load enhancement factor $F = 1.0$ when the test duration is $N_F$.

The probability of survival, at 95 percent confidence level, for a test duration N is given by

$$p = \exp\left[-\left(\frac{N}{\check{\beta}_L}\right)^{\alpha_L}\right] \tag{11}$$

where $\check{\beta}_L$ is the 95 percent confidence scale parameter.

For the requirement that one lifetime is the B-basis life, then the 95 percent confidence $\beta$ is given by

$$\check{\beta}_L = \frac{1}{\left|-\ln(0.9)\right|^{1/\alpha_L}} \tag{12}$$

and equation (11) becomes

$$p = \exp\left|\ln(0.9)N^{\alpha_L}\right| \tag{13}$$

The conditions $N = N_F$, $F = 1.0$ are used to determine the coefficient $\mu$.

$$\mu = \frac{\left[\dfrac{-\ln(0.9)N_F^{\alpha_L}}{X_\gamma^2(2n)/2n}\right]^{1/\alpha_R}}{\Gamma\left(\dfrac{\alpha_R+1}{\alpha_R}\right)} \tag{14}$$

or in a more general form

$$\mu = \frac{\left[\dfrac{-\ln(p)N_F^{\alpha_L+1}}{X_\gamma^2(2n)/2n}\right]^{1/\alpha_R}}{\Gamma\left(\dfrac{\alpha_R+1}{\alpha_R}\right)} \tag{15}$$

In equation (15), $N_F$ is the life factor with p level of reliability at $\gamma$ level of confidence at one-lifetime. By substituting equation (7) into equation (15), it can be shown that

$$\mu = \frac{\left[\Gamma\left(\frac{\alpha_L+1}{\alpha_L}\right)\right]^{\alpha_L/\alpha_R}}{\Gamma\left(\frac{\alpha_R+1}{\alpha_R}\right)} \tag{16}$$

Therefore, the coefficient $\mu$ is a function of $\alpha_L$ and $\alpha_R$ and is independent of the life factor or the sample size.

Finally, the general form of the load enhancement factor is obtained from equation (10) and is written as

$$F = \frac{\mu\,\Gamma\left(\frac{\alpha_R+1}{\alpha_R}\right)}{\left[\frac{-\ln(p)}{x_\gamma^2(2n)/2n}\right]^{1/\alpha_R}} \tag{17}$$

with

$$p = \exp\left[\ln(\ell)N^{\alpha_L}\right] \tag{18}$$

where $\ell$ is the required reliability at $\gamma$ level of confidence ($\gamma = 0.9$ for B-basis and $\gamma = 0.99$ for A-basis).

Equation (17) together with equation (16) can be used to determine the load enhancement factor. The factor F is computed for different values of $\alpha_R$, with a fixed $\alpha_L$, and different values of test duration N.

Figure 23 shows the influence of residual strength scatter, $\alpha_R$, on the load enhancement factor (LEF) required to demonstrate B-basis and A-basis reliability for a one lifetime test. The relationship is shown for various test replications,
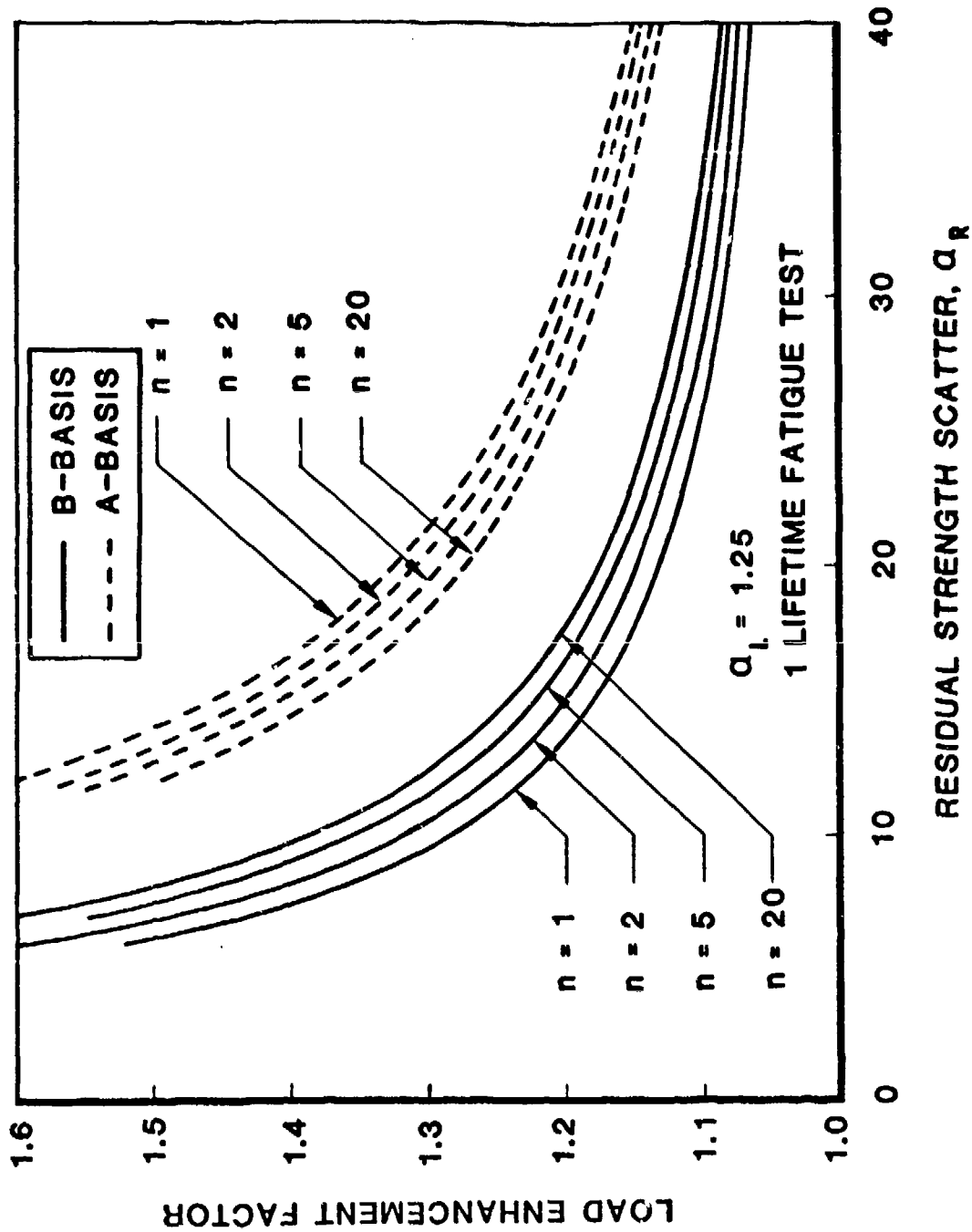
FIGURE 23. INFLUENCE OF RESIDUAL STRENGTH SCATTER $a_R$ ON LOAD ENHANCEMENT FACTOR (ONE LIFETIME FATIGUE TEST)

n.  For typical residual static strength scatter ($\alpha_R$ = 20), the A-basis and B-basis LEF's are 1.33 and 1.18, respectively, for a one replicate test (n = 1).  Figure 24 shows the same relationship as Figure 23 for a two lifetime test.  The extra lifetime of fatigue testing reduces the LEF's for A-basis and B-basis reliability to 1.27 and 1.13, respectively.

It can be seen from Figure 22 that there are three ways to demonstrate B-Basis reliability:

(a)  Load enhancement factor

(b)  Life factor

(c)  Combined load enhancement and life factors

Equations (16) through (18) have been used to develop relationships between load enhancement factor and life factor (test duration).  The relationships  have been determined for various combinations of $\alpha_L$, $\alpha_R$ and n, are shown in Figures 25 and 26.  These plots can be used to specify test or design life requirements for composite structures.

The Sendeckyj fatigue data analysis method can also be used to calculate load enhancement factors from experimental fatigue data.  The Sendeckyj analysis is described in Reference 14 and summarized in Volume I.  This method of analysis was used in Reference 15 to obtain load enhancement factors from experimental data.

The mathematical relationship for LEF's developed in this subsection  is also used to check the accuracy of LEF's calculated by the Sendeckyj analysis.  This is presented in Figure 27, which shows excellent agreement between the two methods.  It can, therefore, be concluded that the Sendeckyj analysis provides good estimates of LEF's from experimental data.  This indicates that the assumptions made in the Sendeckyj fatigue analysis method are valid.
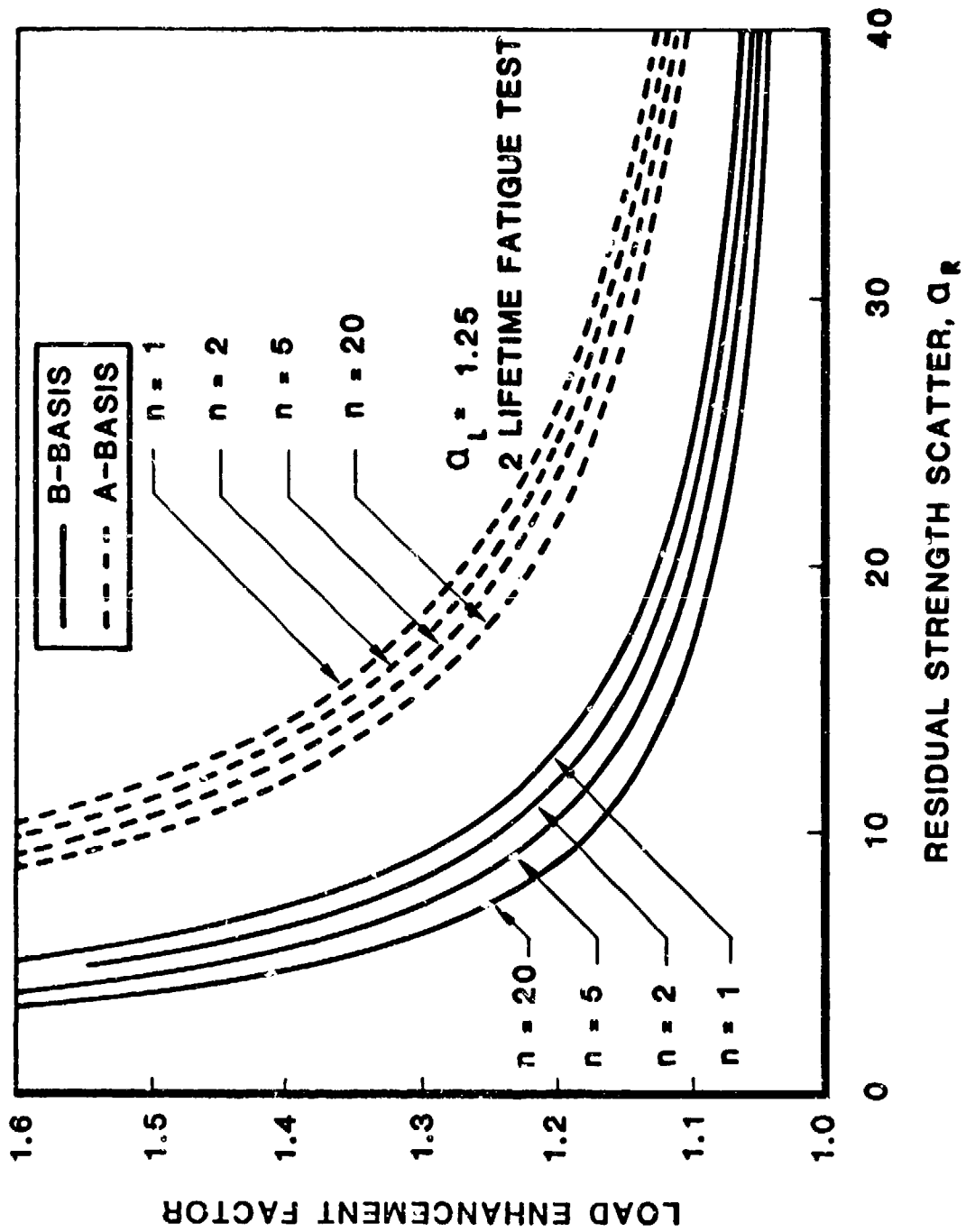
FIGURE 24. INFLUENCE OF RESIDUAL STRENGTH SCATTER, $\alpha_R$, ON LOAD ENHANCEMENT FACTOR (TWO LIFETIME FATIGUE TEST)
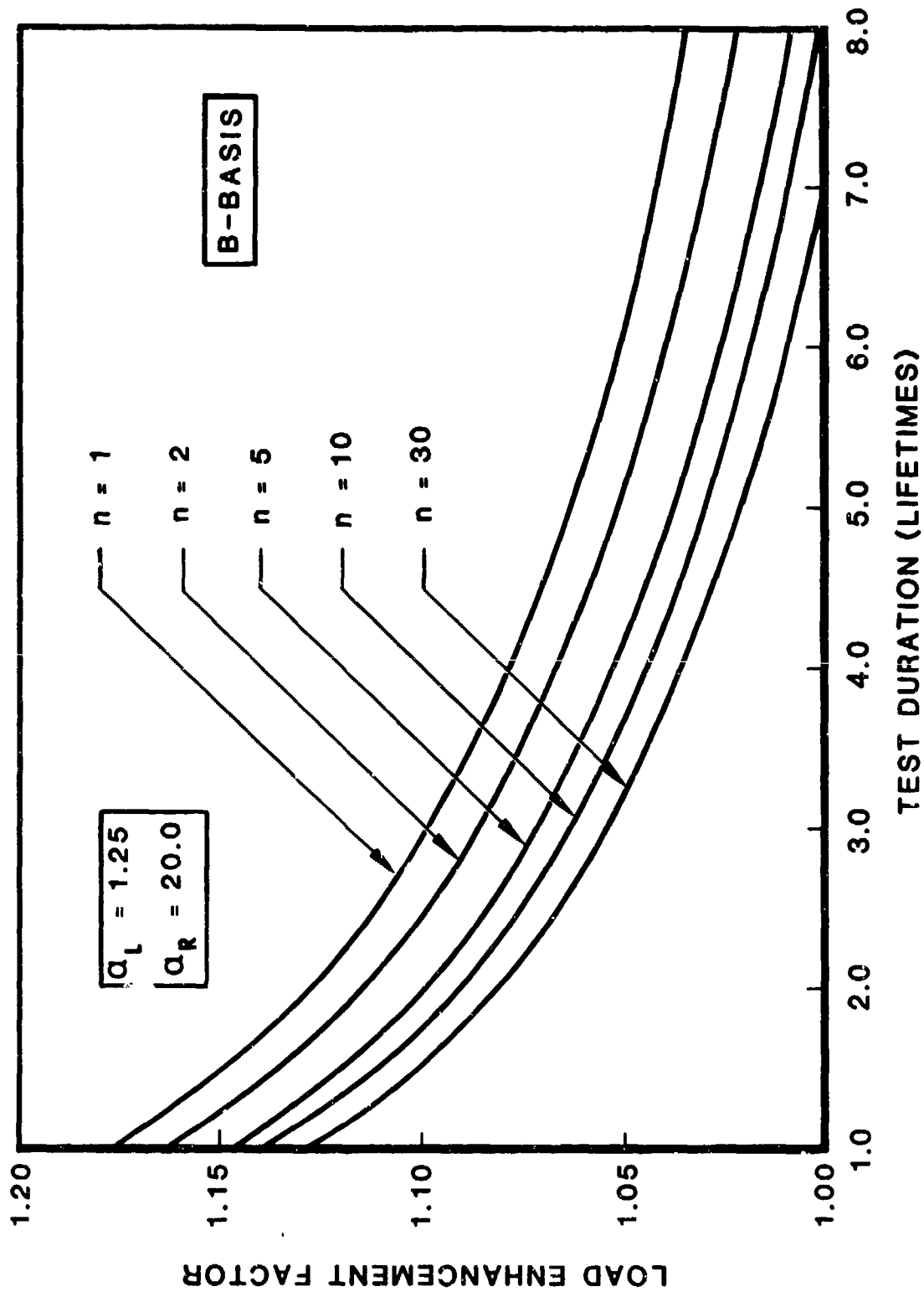
FIGURE 26. INFLUENCE OF TEST DURATION ON B-BASIS LOAD ENHANCEMENT FACTORS.
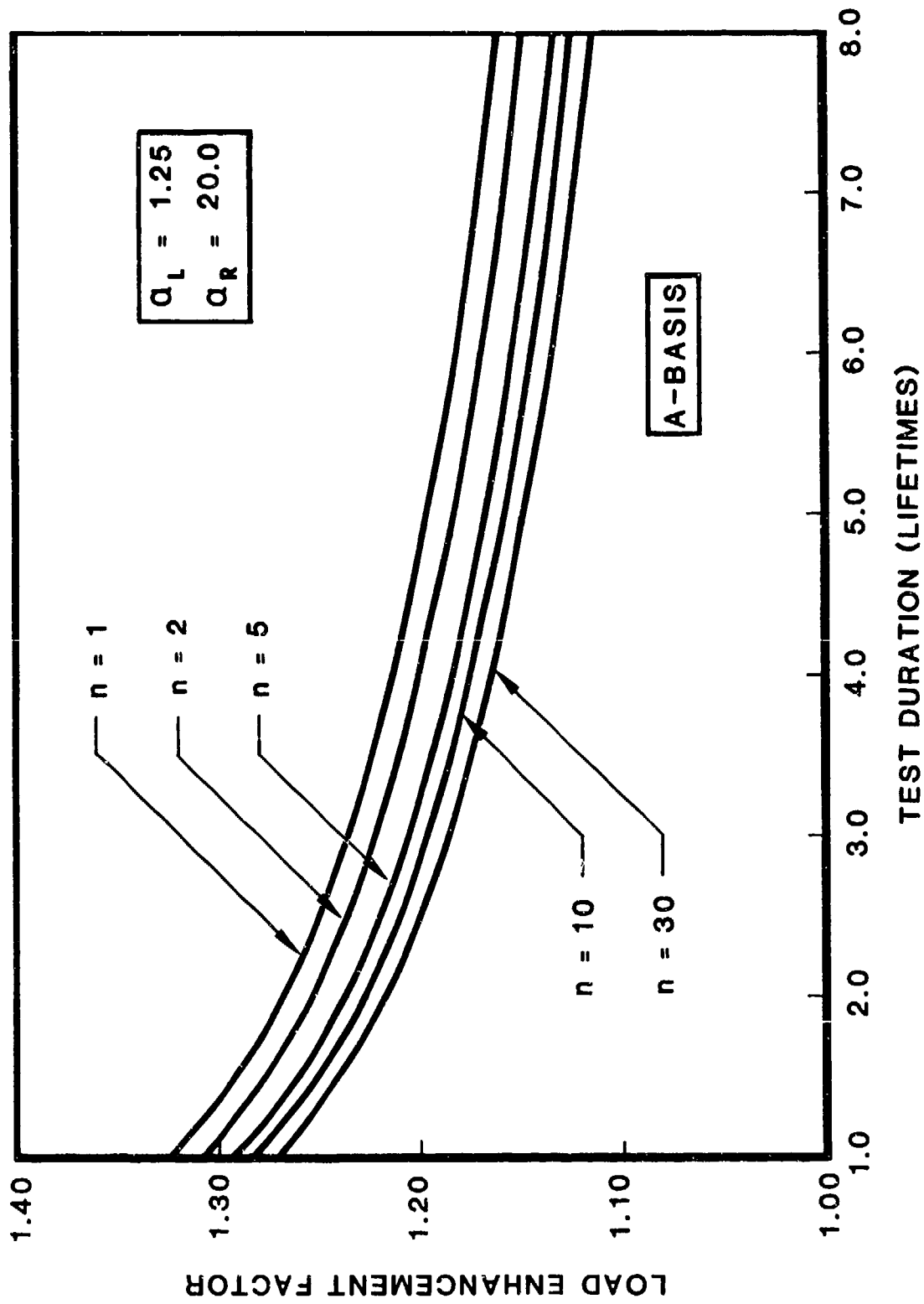
FIGURE 26. INFLUENCE OF TEST DURATION ON A-BASIS LOAD ENHANCEMENT FACTORS.
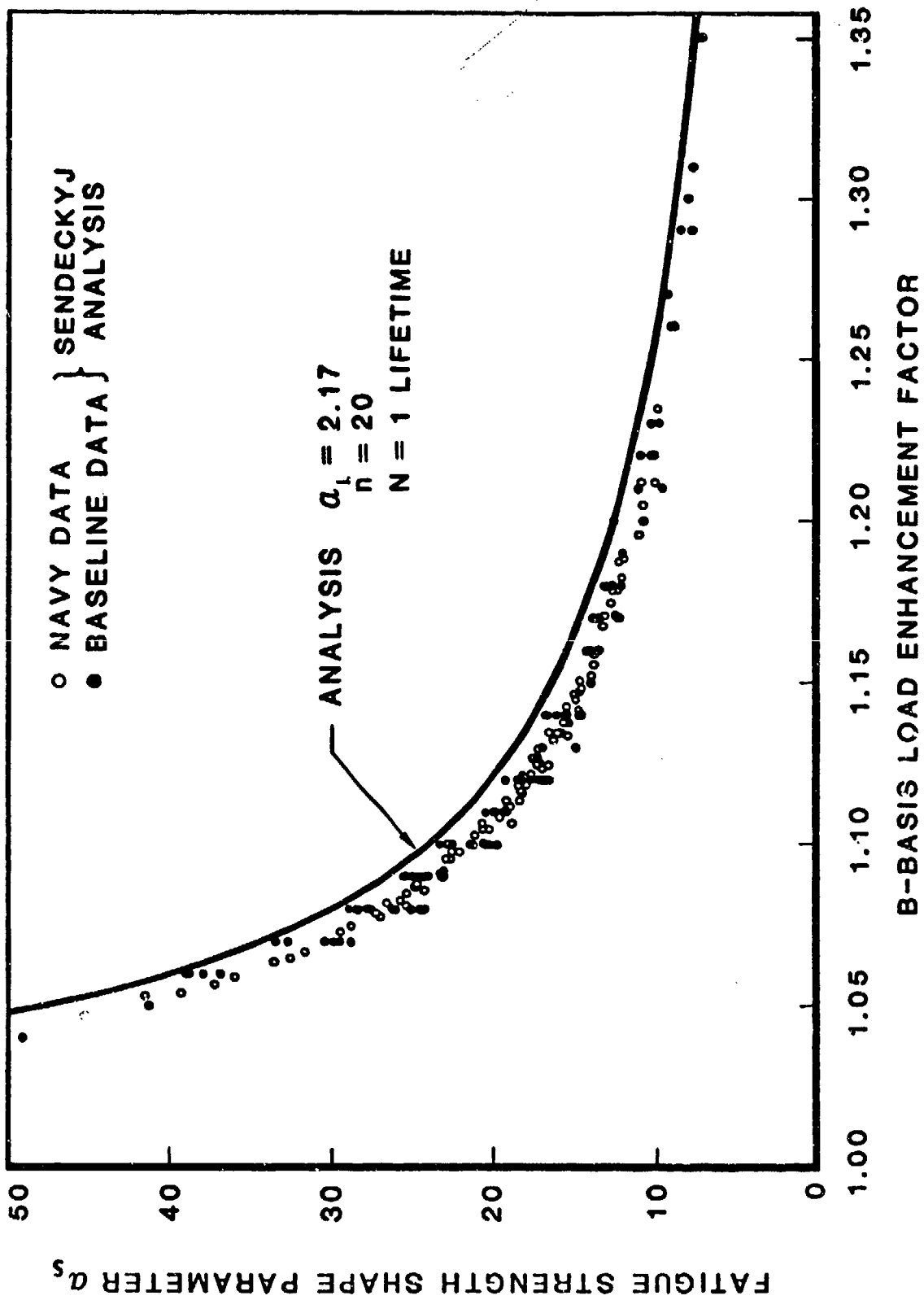
FIGURE 27. COMPARISON OF LOAD ENHANCEMENT FACTORS CALCULATED BY THE SENDECKYJ ANALYSIS WITH THEORETICAL VALUES.

## 2.3    Ultimate Strength Approach

The ultimate strength approach uses an increased static strength margin in conjunction with the fatigue threshold to demonstrate adequate fatigue life. This approach is conservative; however, if it is satisfied no fatigue test is necessary. The concept of a fatigue threshold in composites is shown in Figure 28. This figure shows typical composite fatigue behavior, where a fatigue threshold $\sigma_{TH}$ exists at a relatively high proportion of the static strength. In order to use the ultimate static strength approach it is necessary to design structure such that the maximum spectrum design load ($P_{MSL}$) is no greater than B-basis fatigue threshold stress $\sigma_{TH}$. Thus

$$P_{MSL} \leqslant \sigma_{TH}{}^{B} \tag{19}$$

The relationship between maximum spectrum load ($P_{MSL}$) and design ultimate load ($P_{DUL}$) is a variable which depends on the spectrum type and shape (e.g., wing or tail). Thus we have

$$\frac{P_{MSL}}{P_{DUL}} = X \tag{20}$$

From Figure 28 we can define $\sigma_{TH}$ and as follows:

$$\sigma_{TH}{}^{M}\Big/\sigma_{S}{}^{M} = Y \quad \text{and} \quad \sigma_{TH}{}^{B}\Big/\sigma_{TH}{}^{M} = Z \tag{21}$$

or

$$\sigma_{TH}{}^{B} = YZ\sigma_{S}{}^{M} \tag{22}$$

From equation (19) the requirement becomes

$$P_{MSL} \leqslant YZ\sigma_{S}{}^{M} \tag{23}$$
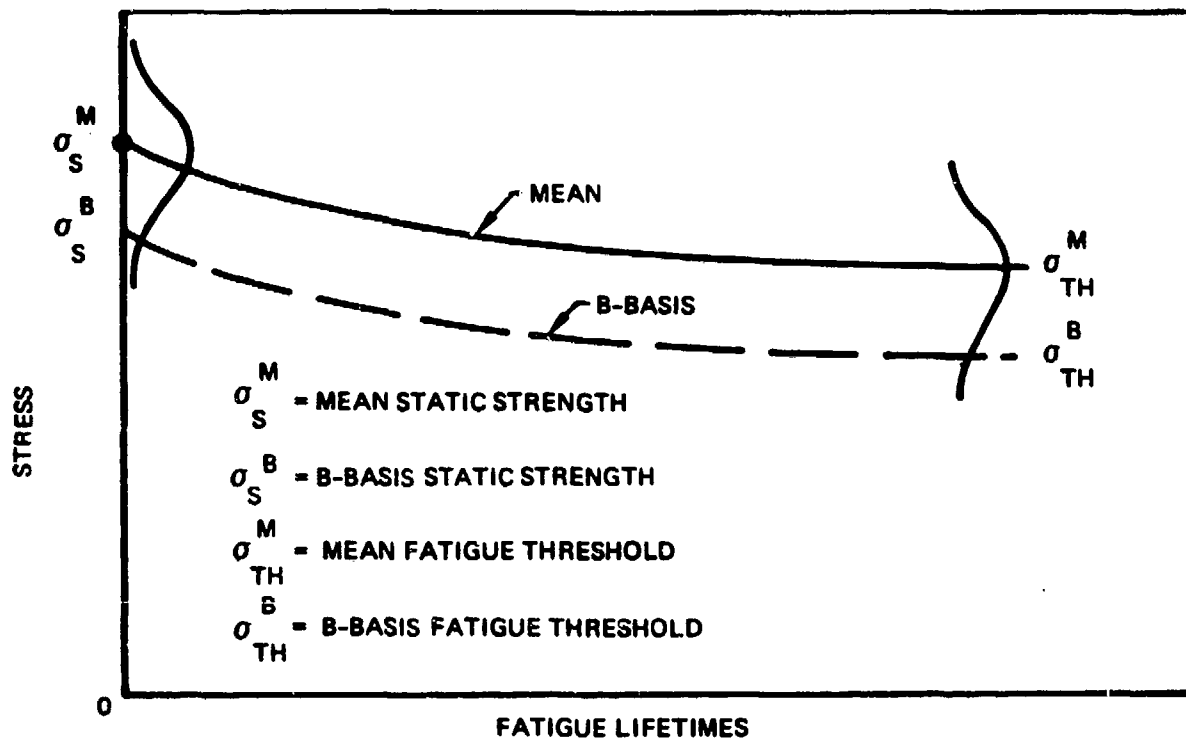
FIGURE 28. COMPOSITE FATIGUE LIFE THRESHOLD APPROACH.

or

$$P_{DUL} \leqslant \frac{YZ}{X} \sigma_S{}^M \qquad\qquad (24)$$

In order to utilize this approach it is necessary to establish the relationship between the B-basis fatigue threshold and mean static strength. This ratio will be a function of the spectrum shape, failure mode and test environment.

The extensive fatigue data base in Reference 16 has been analyzed to determine the relationship between B-Basis fatigue threshold stress $\sigma_{TH}^B$ and mean fatigue threshold stress $\sigma_{TH}^M$. For each S-N data set in Reference 16, the ratio of $\sigma_{TH}^B/\sigma_{TH}^M$ has been determined using the Sendeckyj analysis. The influence of R-ratio and loading mode on this ratio were determined. In addition, the fatigue data scatter analysis conducted in Volume I is used to increase the $\sigma_{TH}^B/\sigma_{TH}^M$ data base. The total data base is then used to establish design knockdown factors for the determination of a B-basis fatigue life threshold from a mean fatigue life threshold. The results of this analysis are discussed below.

The equivalent static strength distribution determined by the Sendeckyj analysis can be used to calculate the ratio of B-basis fatigue threshold $\sigma_{TH}^B$ to mean fatigue life $\sigma_{TH}^M$. The relationship is

$$\frac{\sigma_{TH}^B}{\sigma_{TH}^M} = \frac{\sigma_e^B}{\sigma_e^M} = Z \qquad\qquad (25)$$

The fatigue data in Reference 16 termed Navy data were analyzed using the Sendeckyj analysis to determine values of the $\sigma_e^B/\sigma_e^M$ ratio. Figure 29 shows the influence of R-ratio and loading mode on the B-basis/mean life fatigue threshold ratio $\sigma_e^B/\sigma_e^M$. It can be seen that R-ratio has a small influence on this life ratio. R = -1 loading snows the lowest B-basis/mean life threshold ratio. However, the influence of R-ratio is
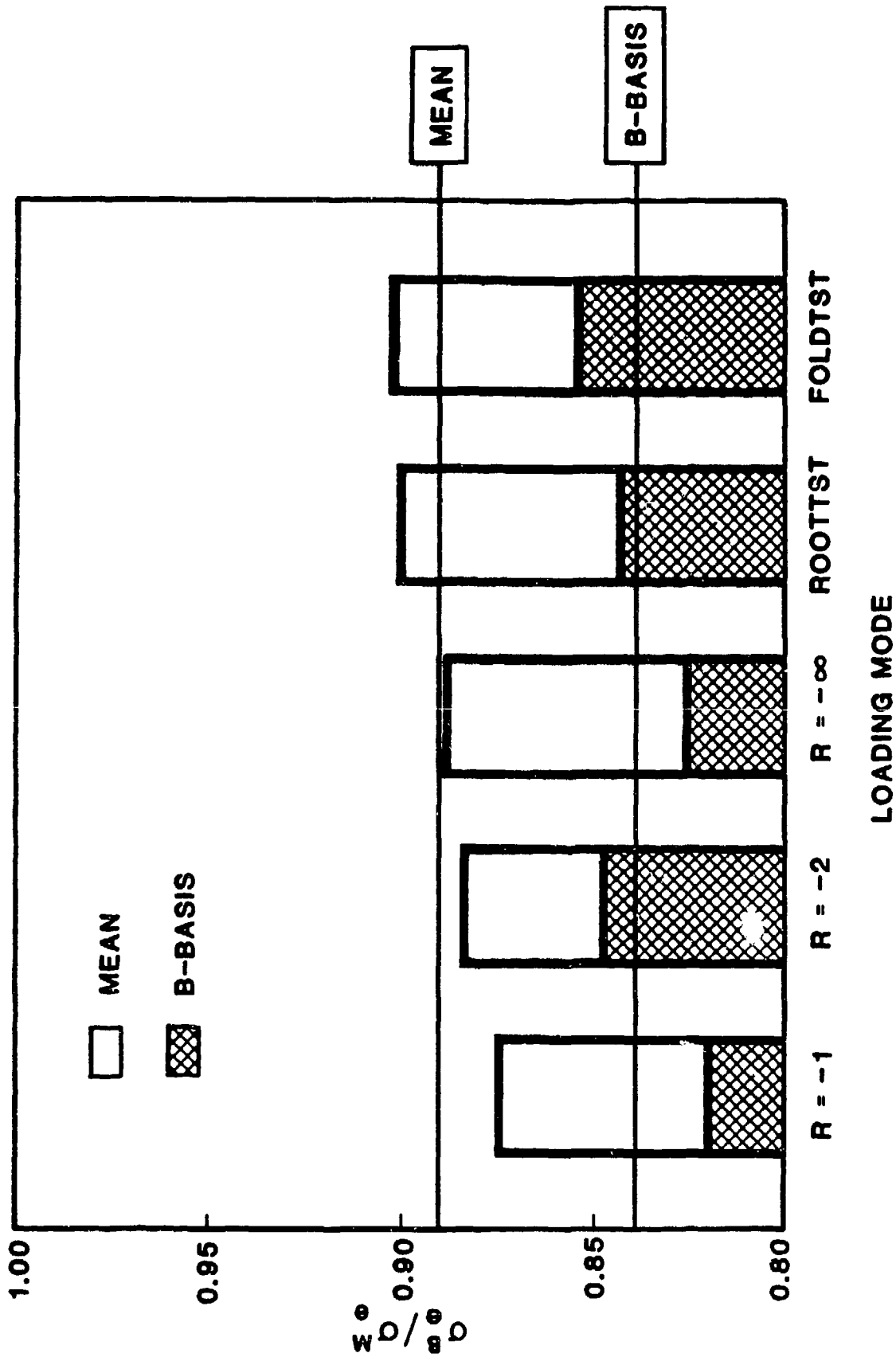
FIGURE 28. INFLUENCE OF LOADING MODE ON THE B-BASIS TO MEAN LIFE FATIGUE THRESHOLD RATIO, $\sigma_e^B / \sigma_e^M$.

different for the B-basis value of the $\sigma_e^B/\sigma_e^M$ ratio. Figures 29 and 30 show that the spectrum loading $\sigma_e^B/\sigma_e^M$ ratios are slightly higher than for constant amplitude loading. The observed differences in the $\sigma_e^B/\sigma_e^M$ ratio in Figures 29 and 30 are not statistically significant. Thus the values of $\sigma_e^B/\sigma_e^M$ from the Navy data set were pooled for statistical analysis. The results are presented in Figure 31, which shows that a modal $\sigma_e^B/\sigma_e^M$ value equal to 0.895 was obtained.

The extensive data base in Reference 15 was also used to determine $\sigma_e^B/\sigma_e^M$ values through the Sendeckyj analysis. The determined $\sigma_e^B/\sigma_e^M$ values are also pooled for statistical analysis. The results are presented in Figure 32, which show that the modal value of $\sigma_e^B/\sigma_e^M$ is equal to 0.905. A comparison of the $\sigma_e^B/\sigma_e^M$ distributions for the Navy and Baseline data sets is shown in Figure 33. It can be seen that the distributions are similar; however, the Baseline data set exhibits a larger spread in $\sigma_e^B/\sigma_e^M$ values. This was anticipated because the Baseline data set contained a wider range of materials, lay-ups and test conditions.

Figure 34 shows analysis results for the combined $\sigma_e^B/\sigma_e^M$ data set (Baseline and Navy data pooled). The modal value of $\sigma_e^B/\sigma_e^M$ is equal to 0.895 for the combined data set. Table 1 summarizes the analysis results for the Navy, Baseline and combined data sets.

Following the philosophy adopted for static strength and fatigue life scatter, the modal value of $\sigma_e^B/\sigma_e^M$ equal to 0.895 will be used for evaluation of the ultimate strength certification approach.

In order to utilize the ultimate strength approach it is also necessary to determine the ratio of the mean fatigue threshold $\sigma_{TH}^M$ to the mean static strength $\sigma_S^M$. The extensive fatigue data in Reference 16 are used to determine the influence of R-ratio, upper wing skin spectrum loading and specimen geometry on the ratio $\sigma_{TH}^M/\sigma_S^M$. Figures 35 through 37 show the influence of R-ratio and spectrum loading on the $\sigma_{TH}^M/\sigma_S^M$ ratio
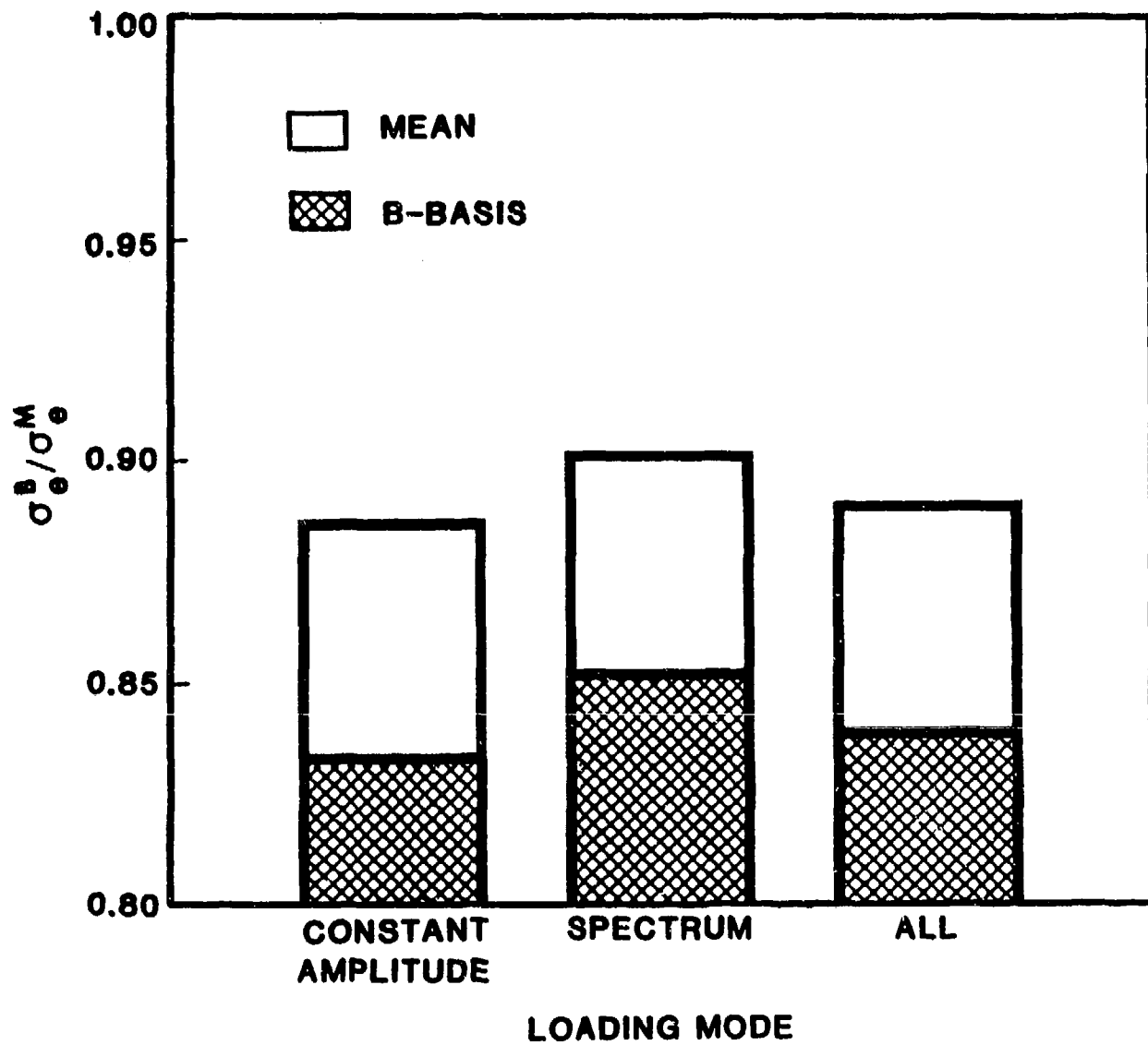
FIGURE 30. INFLUENCE OF LOADING MODE ON THE B-BASIS TO
MEAN LIFE FATIGUE THRESHOLD RATIO, $\sigma_a^B / \sigma_e^M$.
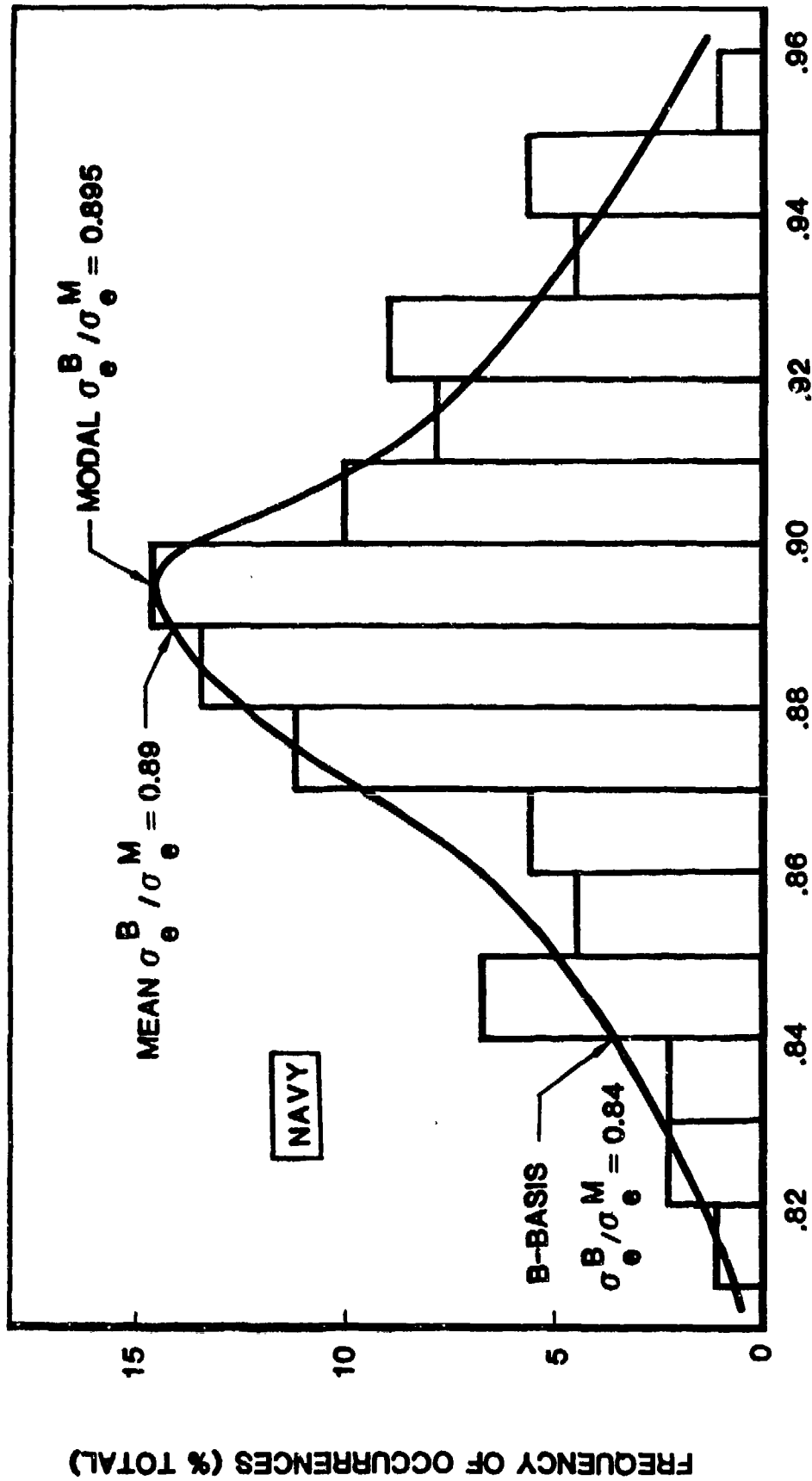
FIGURE 31. DISTRIBUTION OF B-BASIS TO MEAN FATIGUE LIFE THRESHOLDS FOR NAVY DATA.
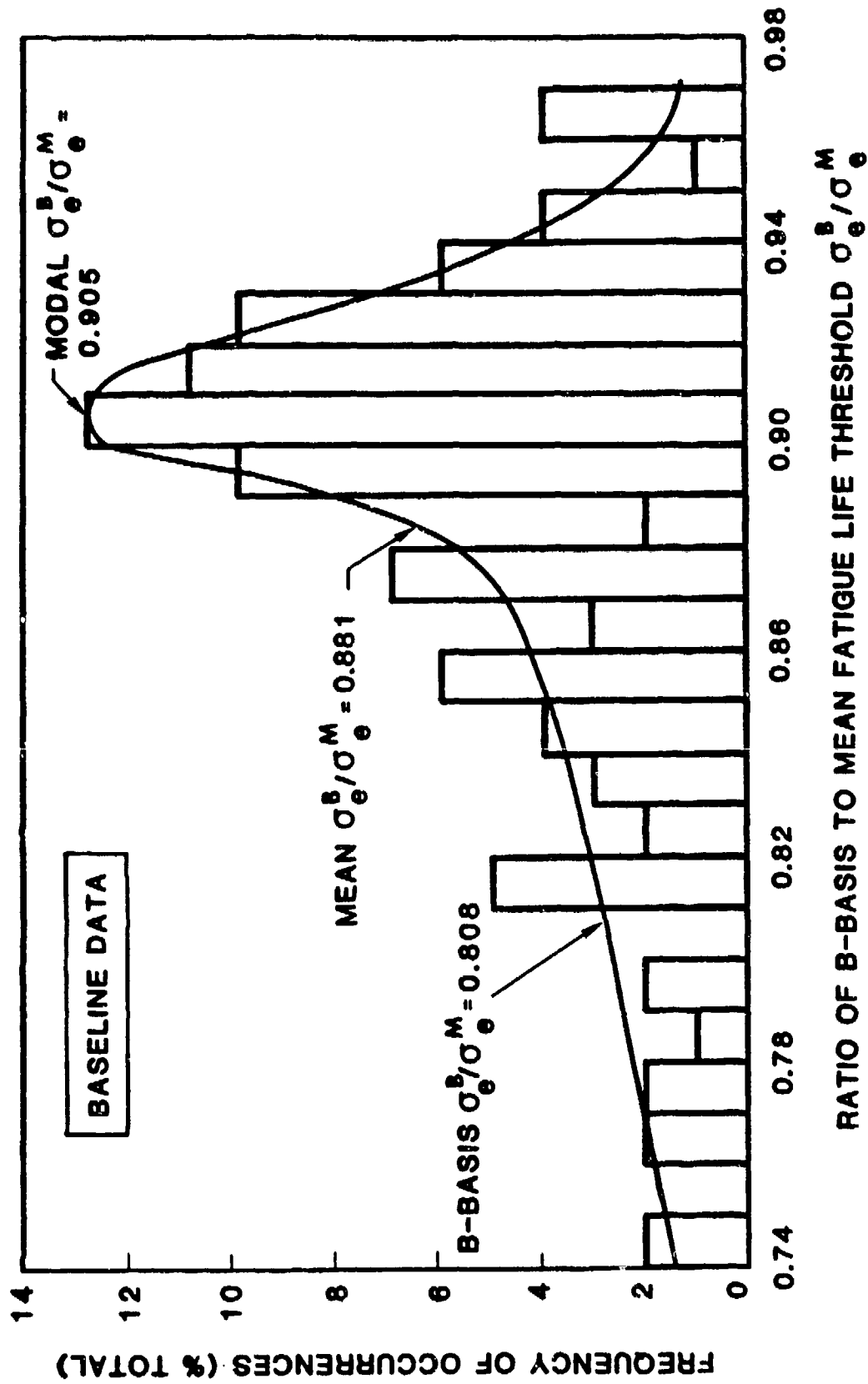
FIGURE 32. DISTRIBUTION OF B-BASIS TO MEAN FATIGUE LIFE THRESHOLDS FOR BASELINE DATA.
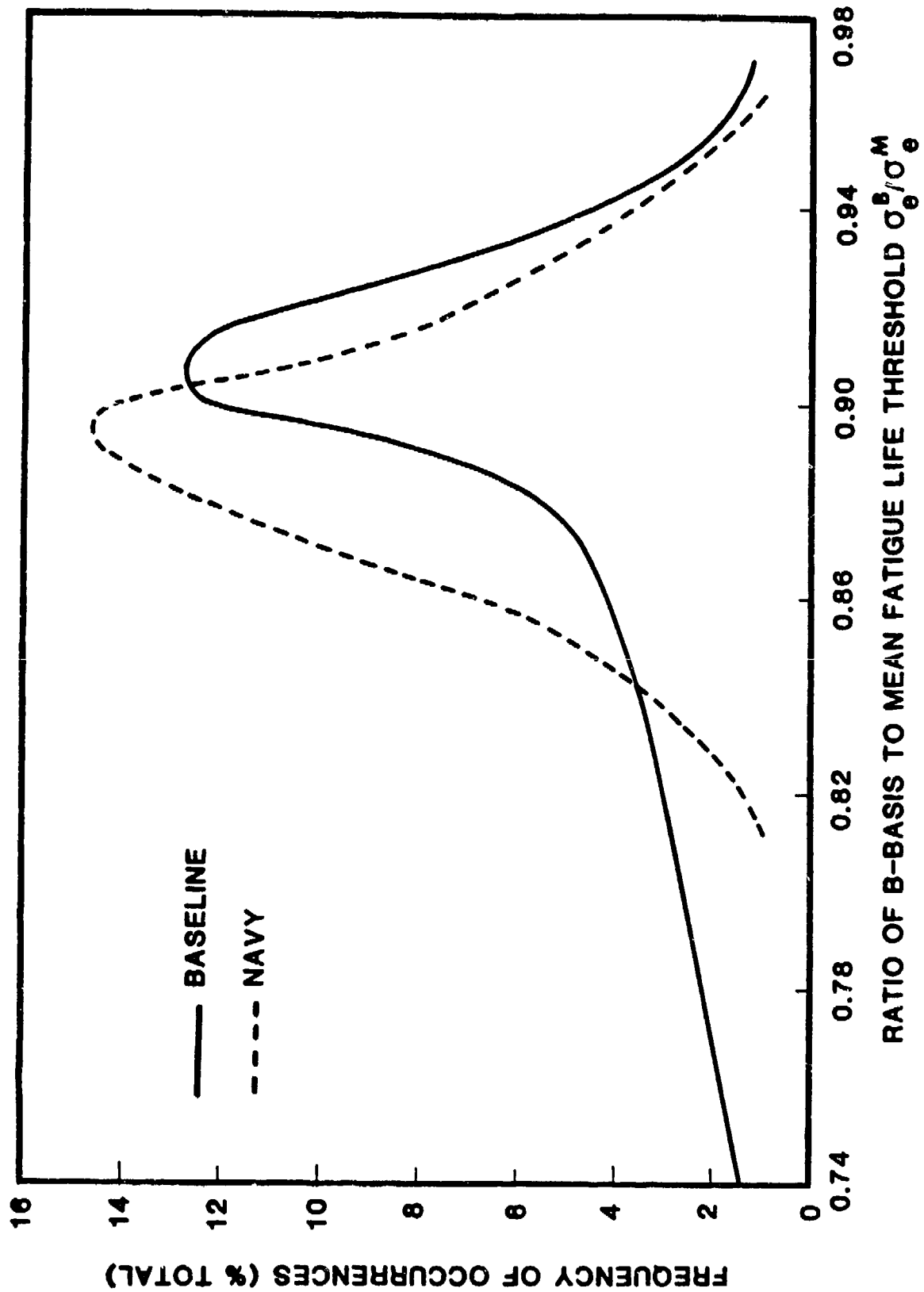
FIGURE 33. COMPARISON OF NAVY AND BASELINE B-BASIS TO MEAN FATIGUE LIFE THRESHOLD DISTRIBUTIONS.
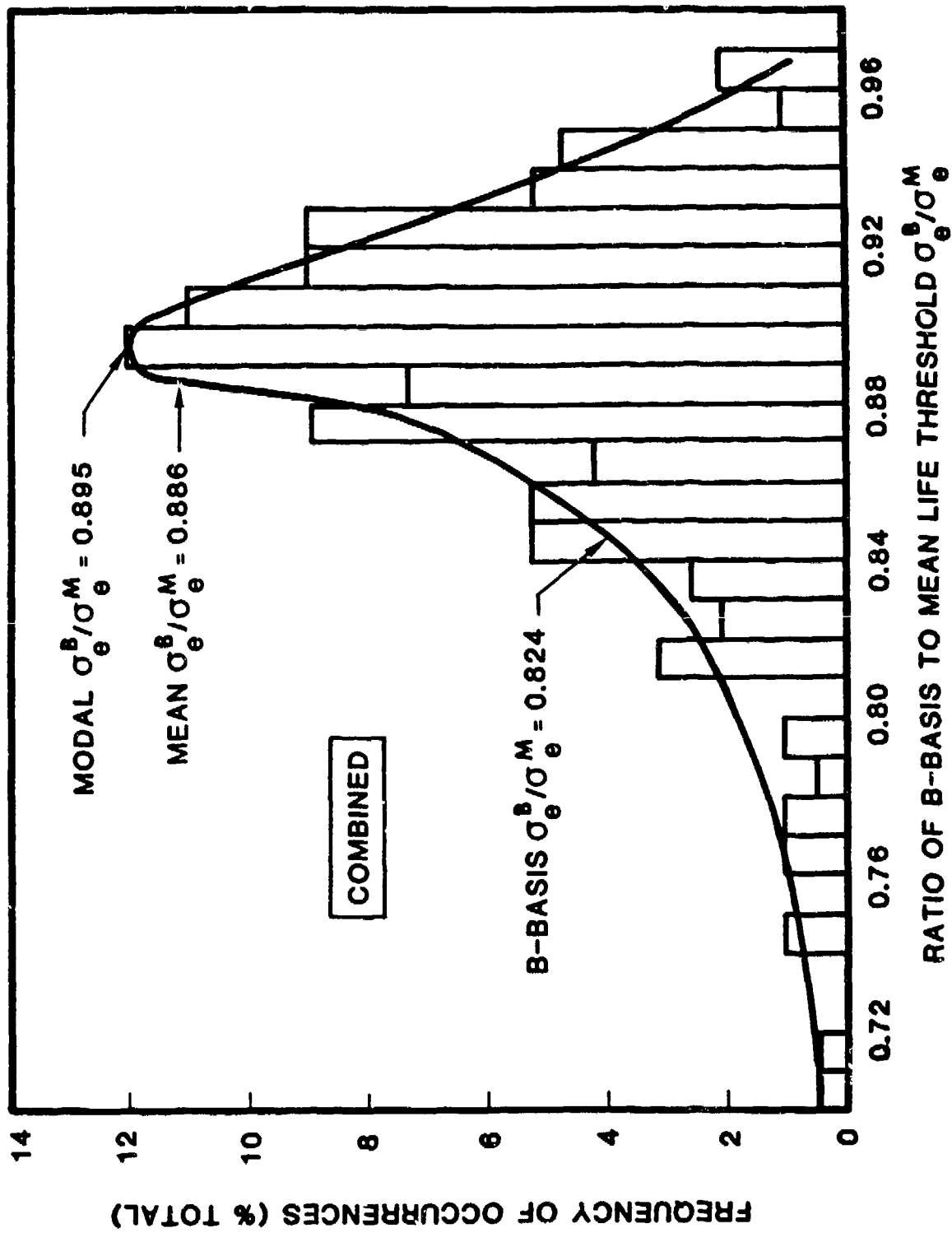
FIGURE 34. DISTRIBUTION OF B-BASIS TO MEAN FATIGUE LIFE THRESHOLDS FOR COMBINED DATA.

TABLE 1. SUMMARY OF THE STATISTICAL ANALYSIS OF THE B-BASIS
TO MEAN FATIGUE LIFE THRESHOLD RATIO VALUES.

| DATA SET | B-BASIS/MEAN FATIGUE LIFE THRESHOLD $\sigma_e^B / \sigma_e^M$ | | |
|----------|------|-------|---------|
| | MEAN | MODAL | B-BASIS |
| NAVY | 0.890 | 0.895 | 0.840 |
| BASELINE | 0.891 | 0.905 | 0.833 |
| COMBINED | 0.886 | 0.895 | 0.824 |
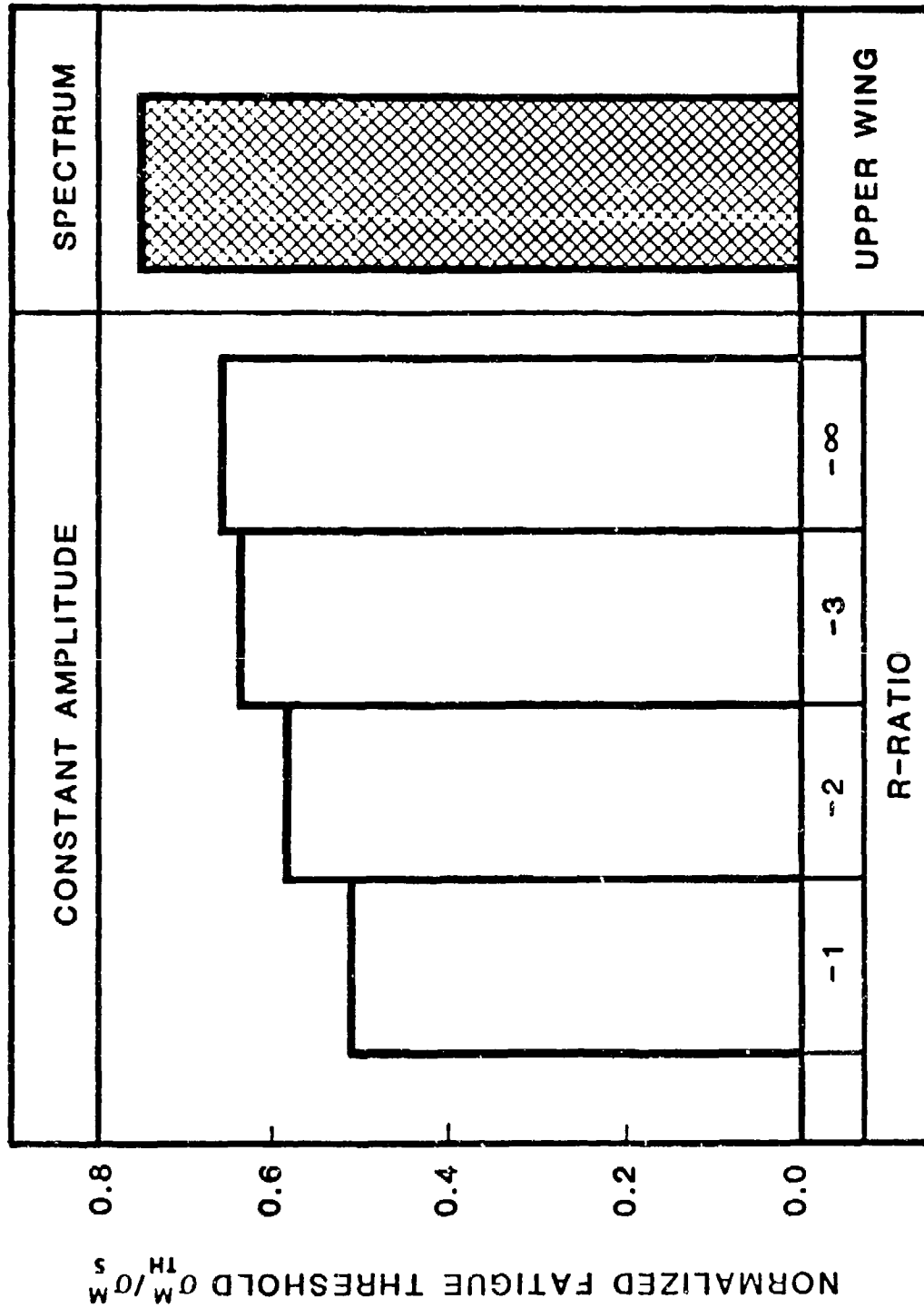
FIGURE 35. INFLUENCE OF FATIGUE LOADING MODE ON NORMALIZED FATIGUE THRESHOLD FOR INTERMEDIATE LOAD TRANSFER SPECIMENS IN REFERENCE 16.
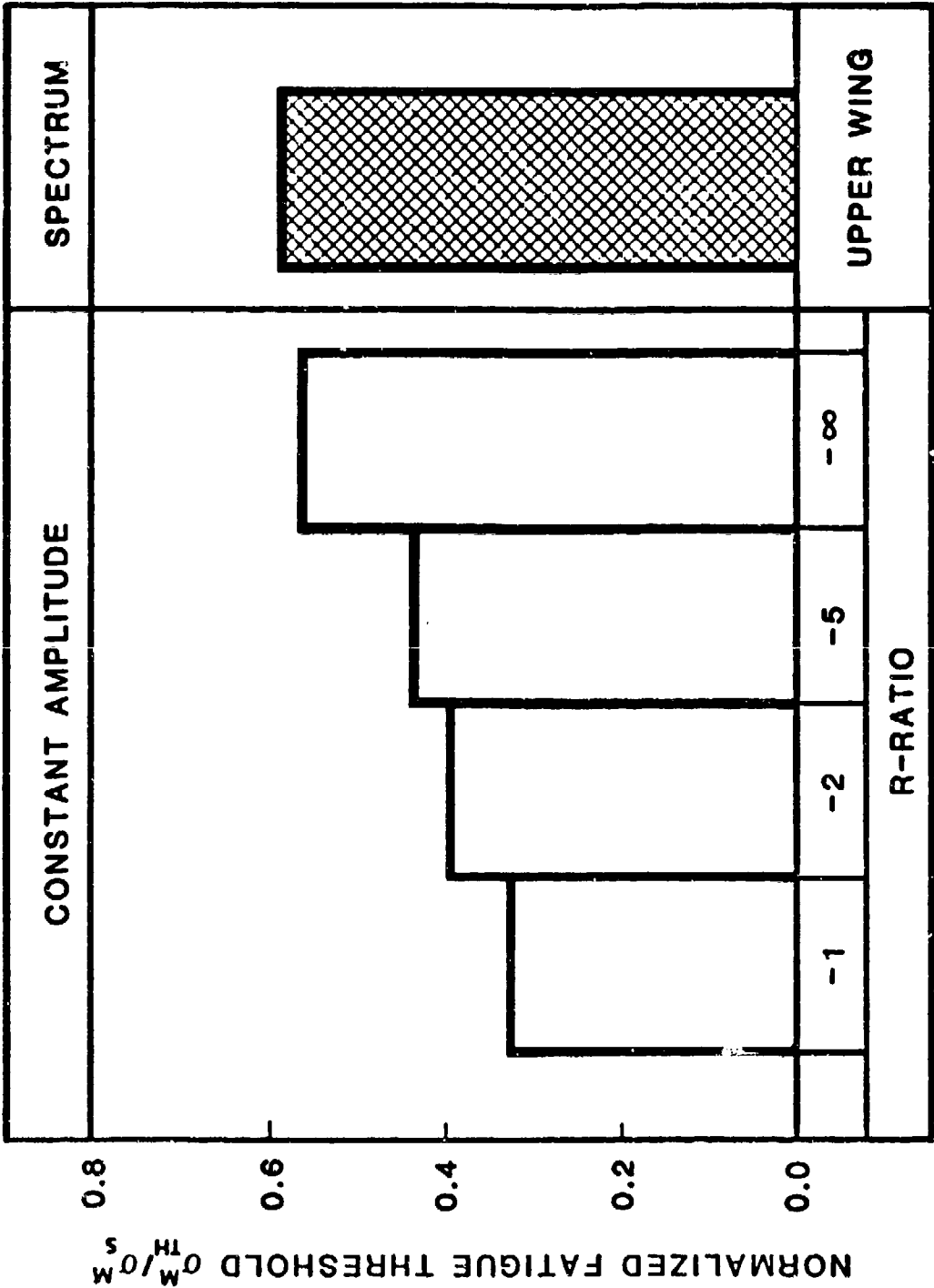
FIGURE 36. INFLUENCE OF FATIGUE LOADING MODE ON NORMALIZED FATIGUE THRESHOLD FOR HIGH LOAD TRANSFER SPECIMENS IN REFERENCE 16.
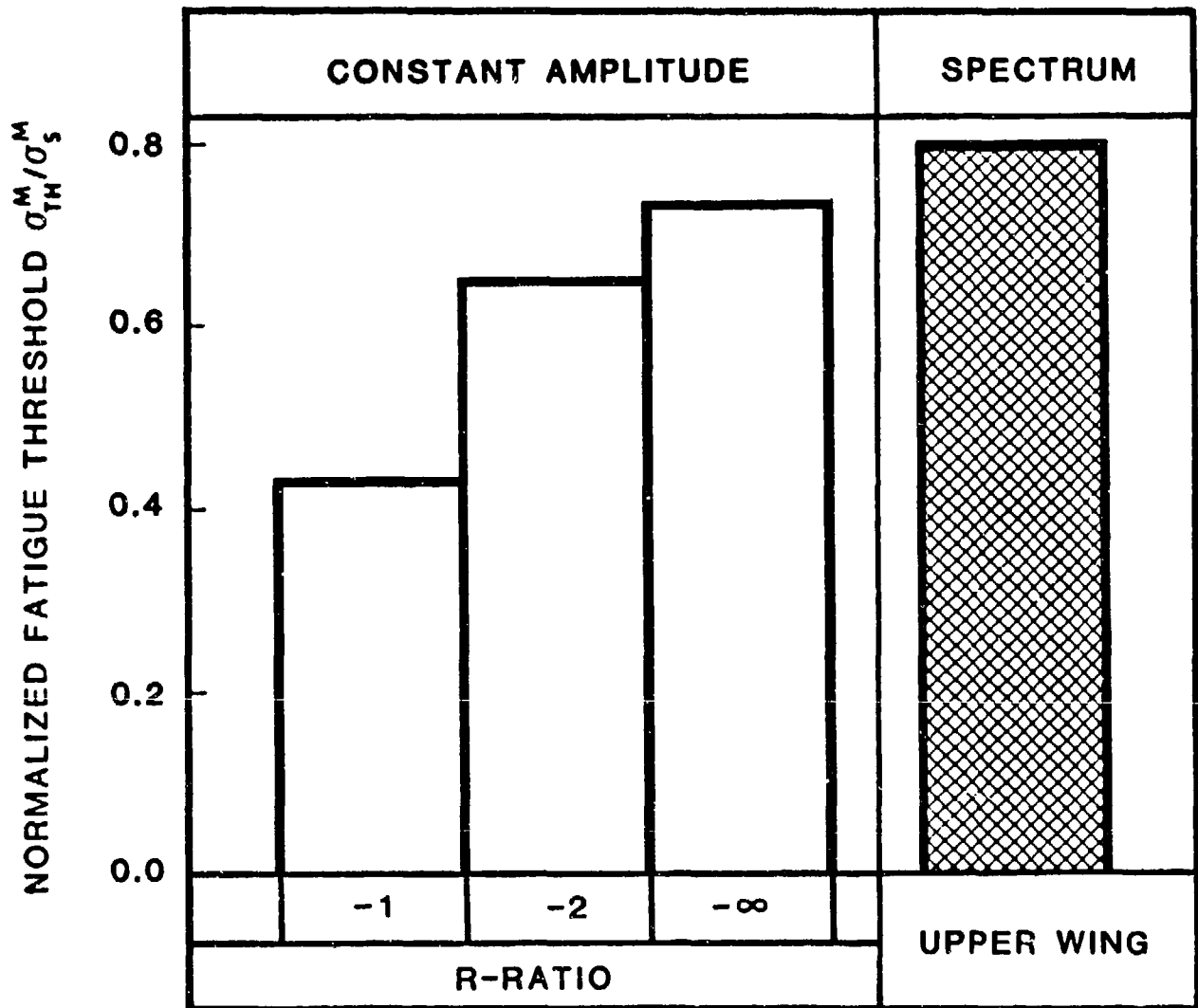
FIGURE 37. INFLUENCE OF FATIGUE LOADING MODE ON NORMALIZED
FATIGUE THRESHOLD FOR COMPLEX SPECIMENS IN
REFERENCE 16.

for three joint types. The data for all three joint types show a similar trend. As R-ratio increases from -1 to -∞ the threshold ratio increases; however, spectrum fatigue threshold ratios are higher than those determined for constant amplitude loading. Thus, it can be concluded that fatigue testing under constant amplitude loading is conservative relative to upper wing skin spectrum fatigue loading. The data for the three joints are pooled and the results are shown in Figure 38. As expected, the pooled data show the same influences of R-ratio and spectrum loading.

The spectrum fatigue loading data in Reference 16 are analyzed in more detail because of its relevance to certification testing. The influence of joint geometry, lay-up, spectrum loading type and test environment are presented in Table 2 and Figure 39. Twenty-three spectrum loading S-N curves are used in the analysis. The results show that high load transfer specimens had lower normalized fatigue threshold than intermediate load transfer and complex specimens. This is probably due to the different failure mode observed in the high load transfer specimens, which was hole wear. The normalized fatigue threshold for these specimens is based on 0.025 inch hole wear, which probably gives a conservative estimate of the fatigue threshold. Figure 39 also shows that the decreasing laminate stiffness lowers the normalized fatigue threshold value. This reflects the higher fatigue sensitivity of the (16/80/4) lay-up. The influence of spectrum loading type on normalized fatigue thresholds is shown to be negligible in Figure 39. This suggests that composite spectrum fatigue life is dominated by the peak compression load and is relatively independent of load reversal severity. Table 2 shows that the normalized fatigue threshold is higher for the ETW environment relative to the RTD environment. This implies that the fatigue degradation rate relative to static strength is lower in the ETW environment. The overall mean value of the normalized fatigue threshold for spectrum loading is determined to be 0.71. This value and the scatter analyses for $\sigma_{TH}^{B}/\sigma_{TH}^{M}$ (described earlier) and for $\sigma_{TH}^{M}/\sigma_{TH}^{S}$ are used to determine the fatigue threshold
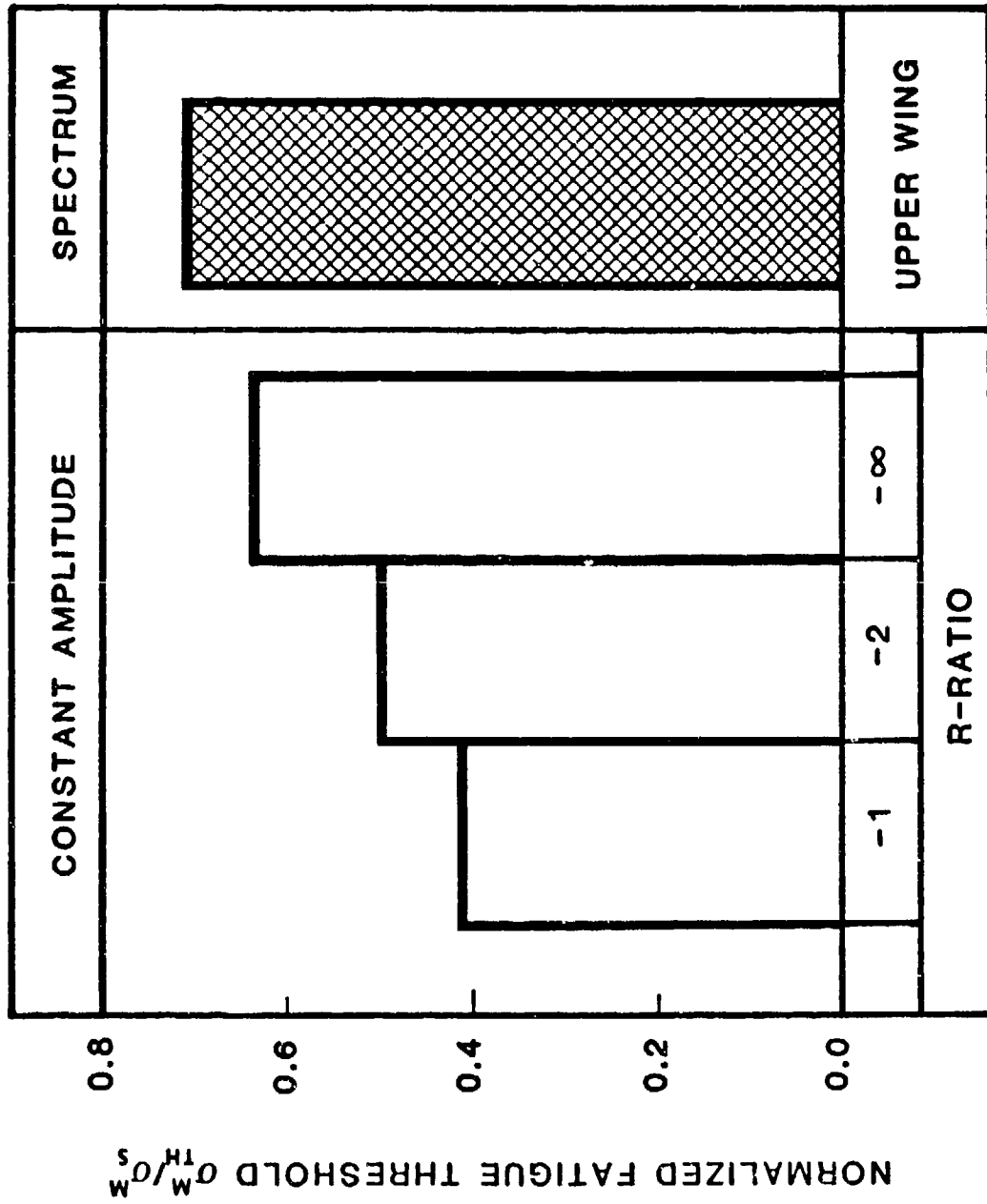
FIGURE 38. INFLUENCE OF FATIGUE LOADING MODE ON NORMALIZED FATIGUE THRESHOLD FOR ALL SPECIMENS IN REFERENCE 16.

## TABLE 2.   INFLUENCE OF TEST VARIABLES ON NORMALIZED FATIGUE THRESHOLDS UNDER SPECTRUM LOADING.

| COMPARISON | NORMALIZED FATIGUE THRESHOLD $\sigma_{TH}^{M}/\sigma_{S}^{M}$ |
|---|---|
| ROOTTST | 0.72 |
| FOLDTST | 0.70 |
| INT. L/T | 0.75 |
| HIGH L/T | 0.59 |
| COMPLEX | 0.80 |
| (48/48/4) | 0.83 |
| (16/80/4) | 0.67 |
| RTD | 0.66 |
| ETW | 0.75 |
| OVERALL MEAN | 0.71 |
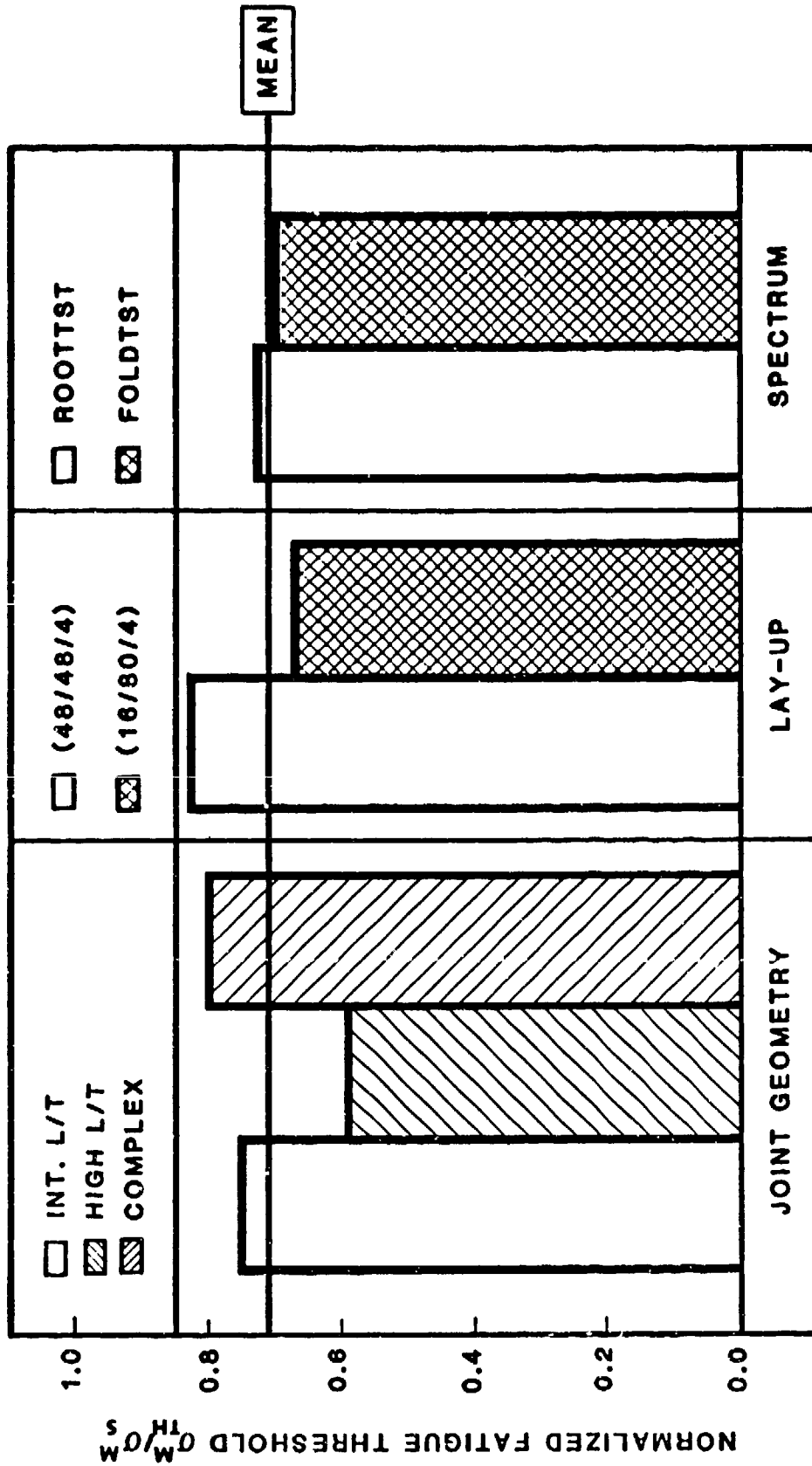
FIGURE 39. INFLUENCE OF JOINT GEOMETRY, LAY-UP AND SPECTRUM TYPE ON NORMALIZED FATIGUE THRESHOLDS FOR SPECTRUM LOADING.

behavior for upper wing spectrum loading. The results are shown in Figure 40. The B-basis fatigue threshold is determined to be 63 percent of the mean static strength. That is, if the maximum spectrum load is set to $\leq 63$ percent of the mean static strength a B-basis fatigue threshold will be established statistically.

Once $\sigma_{TH}^{M}$ and $\sigma_{TH}^{B}$ have been established statistically, as in Figure 40, the static failure load required to guarantee a B-basis fatigue threshold (no fatigue test) is calculated as follows:

No fatigue test is required if:

$$\text{Static Failure Load} = \sigma_{S}^{M} \geq \left( \frac{P_{MSL}}{P_{DUL}} \cdot \frac{\sigma_{S}^{M}}{\sigma_{TH}^{M}} \cdot \frac{\sigma_{TH}^{M}}{\sigma_{TH}^{B}} \right) \cdot P_{DUL}$$

Example calculations for RTD test conditions, a (48/48/4) laminate and F-18 upper wing spectrum, are presented in Table 3. The calculations show that the static failure load requirements for no fatigue test range from 122 percent to 187 percent of design ultimate load.

The influence of spectrum type on the static failure load for B-basis fatigue reliability is shown in Figure 41. These data show that the static failure load requirement ranges from 78 percent to 129 percent DUL. The significant differences in the static overload requirements for the three spectra are caused mainly by the significant differences in their $P_{MSL}/P_{DUL}$ ratios. These are:

| SPECTRUM TYPE | $P_{MSL}/P_{DUL}$ |
|---|---|
| VERTICAL | 0.489 |
| HORIZONTAL | 0.651 |
| WING | 0.814 |

It should be noted that the $\sigma_{TH}^{M}/\sigma_{S}^{M}$ values used for the calculations shown in Figure 41 were typical values. Figure 41 is not, therefore, a design chart.
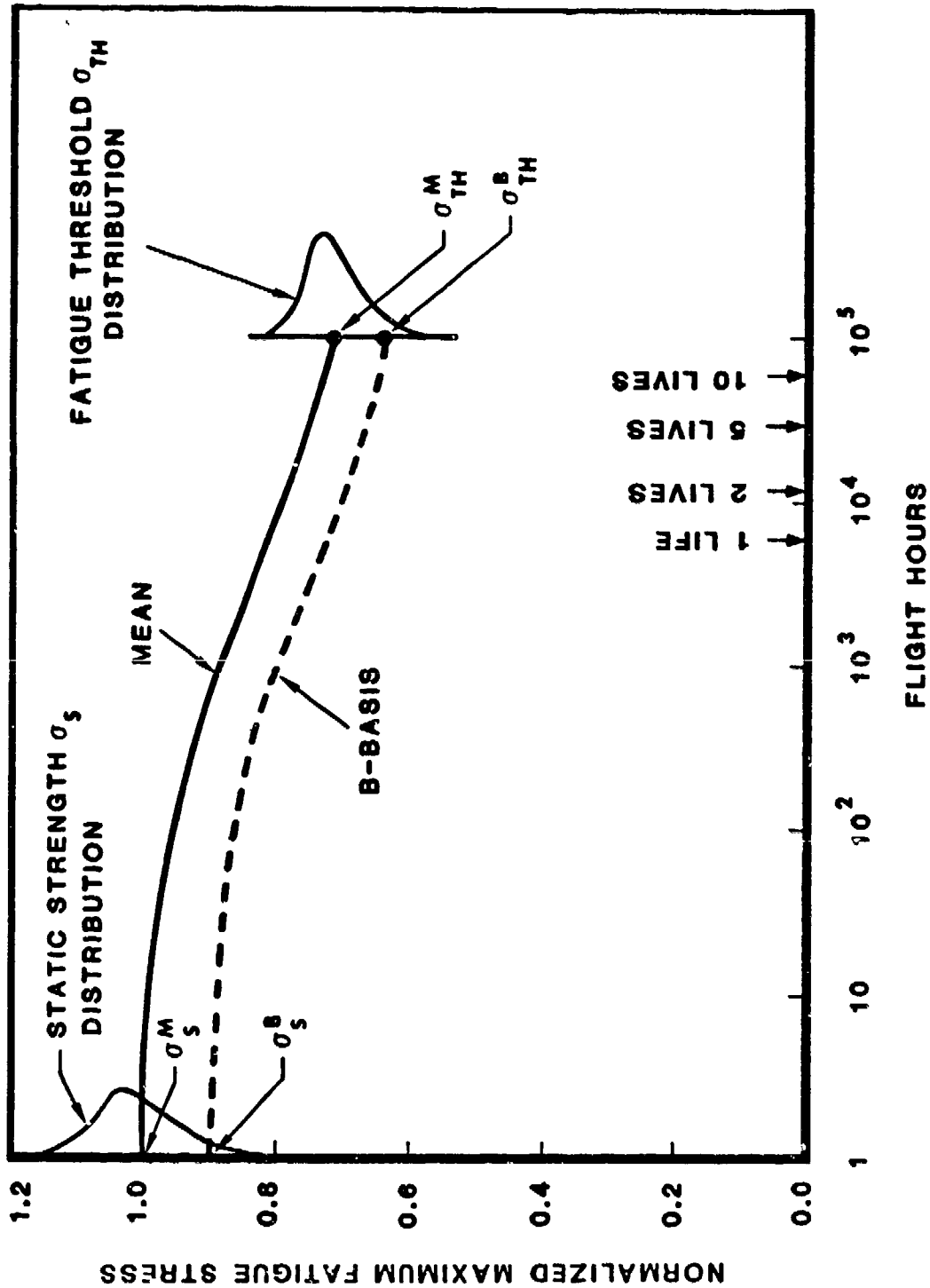
FIGURE 40. STATISTICAL FATIGUE THRESHOLD FOR UPPER WING SKIN SPECTRUM LOADING.

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TABLE 3. STATIC FAILURE LOAD REQUIREMENTS FOR NO FATIGUE TEST OF A (48/48/4) LAMINATE IN A RTD ENVIRONMENT.

| SPECTRUM | SPECIMEN | $\dfrac{P_{MSL}}{P_{DUL}}$ | $\dfrac{\sigma^M_{TH}}{\sigma^M_S}$ | $\dfrac{\sigma^B_{TH}}{\sigma^M_{TH}}$ | STATIC FAILURE LOAD REQUIREMENT % DUL |
|---|---|---|---|---|---|
| ROOTTST | OPEN HOLE | | 0.78 | | 122 |
| | INT. L/T | 0.84 | 0.76 | 0.885 | 125 |
| | HIGH L/T | | 0.67 | | 142 |
| FOLDTST | OPEN HOLE | | 0.82 | | 123 |
| | INT. L/T | 0.894 | 0.54 | 0.885 | 187 |
| | COMPLEX | | 0.84 | | 120 |

FIGURE 41. INFLUENCE OF SPECTRUM TYPE ON STATIC OVERLOAD REQUIREMENT FOR B-BASIS FATIGUE RELIABILITY.

2.4    Change in Spectrum Approach

The problem of mixed composite/metal fatigue testing was summarized in October 1981 by Dr. Someroff of NAVAIR. Figure 42 presents his concern, which was that a two lifetime test did not adequately interrogate composite parts for potential hotspots. This occurred because composites had considerably higher average fatigue lives and exhibit higher life scatter than metals. Someroff suggested changing the test spectrum to reduce composite fatigue life, while maintaining metal fatigue life constant. The ultimate aim would be to make the B-basis stress-life plots overlap for both materials. In principle, this can be accomplished by adding extra high loads to the fatigue spectrum. These loads could significantly reduce composite fatigue life, but maintain metal fatigue life at a constant value because the increased damage accumulation could be cancelled out by increased retardation effects in the metal parts.

The practical situation for mixed structures is more complex than envisaged by Someroff, as shown in Figure 43. Someroff assumed in Figure 42 that the stress-life curves for metals and composites had approximately the same slope. However, Figure 43 shows that this is not the case. Composites have considerably flatter S-N curves. This makes the change in spectrum approach more difficult to apply. The data in Figure 43 are shown for an F-18 wing root spectrum. For each material, sensitivity to the most critical spectrum is shown; that is, compression dominated for composites and tension dominated for metals. Figure 44 shows that this characteristic difference exists for three widely different spectrum types. The curves shown in Figure 44 are analytical. A comparison of the three spectra is shown in Figure 45. All analytical predictions were made using the methods described in References 17 through 19.

Figure 46 shows the influence of overloads on composite fatigue life for an F-18 wing root spectrum. Spectrum overloads in the range 110-120 percent of the maximum spectrum load were selected, with occurrences of these loads ranging 10 to 100 per

FIGURE 42. SCHEMATIC SUMMARY OF MIXED STRUCTURE TESTING PROBLEMS (A. SOMEROFF NAVAIR, OCTOBER 1981).
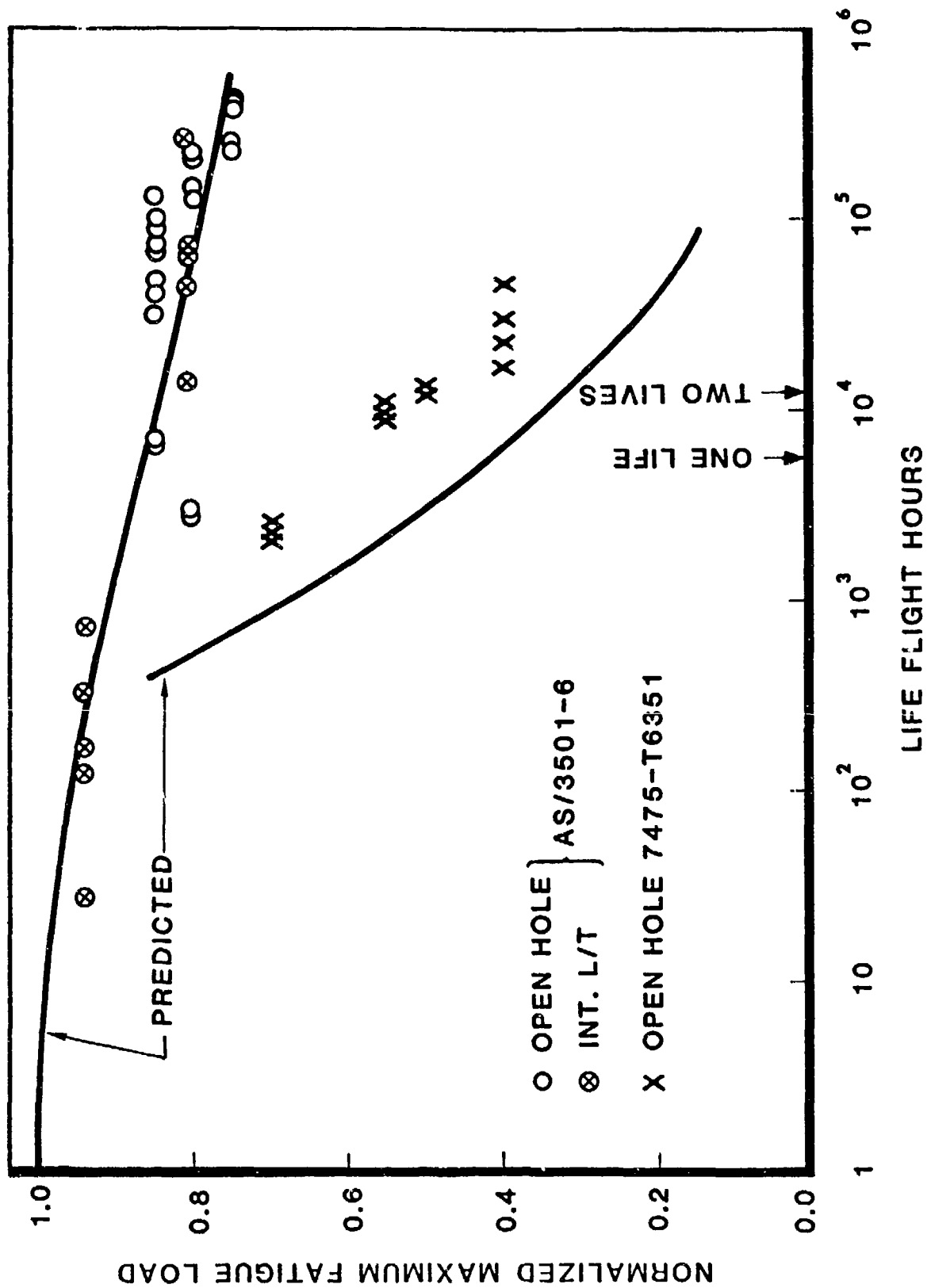
FIGURE 43. COMPARISON OF COMPOSITE AND METALLIC FATIGUE SPECTRUM FATIGUE BEHAVIOR FOR A WING SPECTRUM.
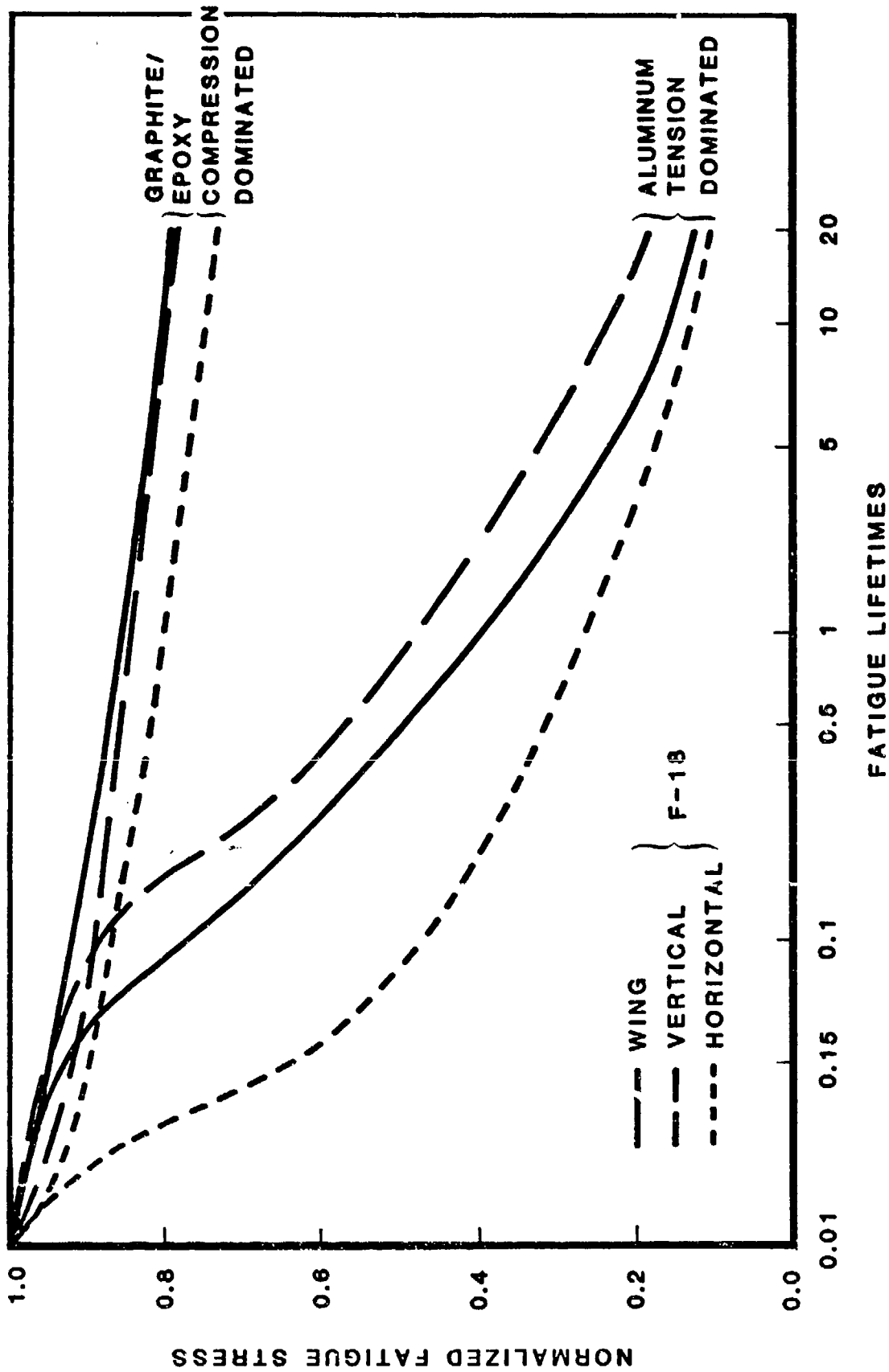
FIGURE 44. INFLUENCE OF SPECTRUM TYPE ON PREDICTED COMPOSITE
AND METAL FATIGUE LIFE.
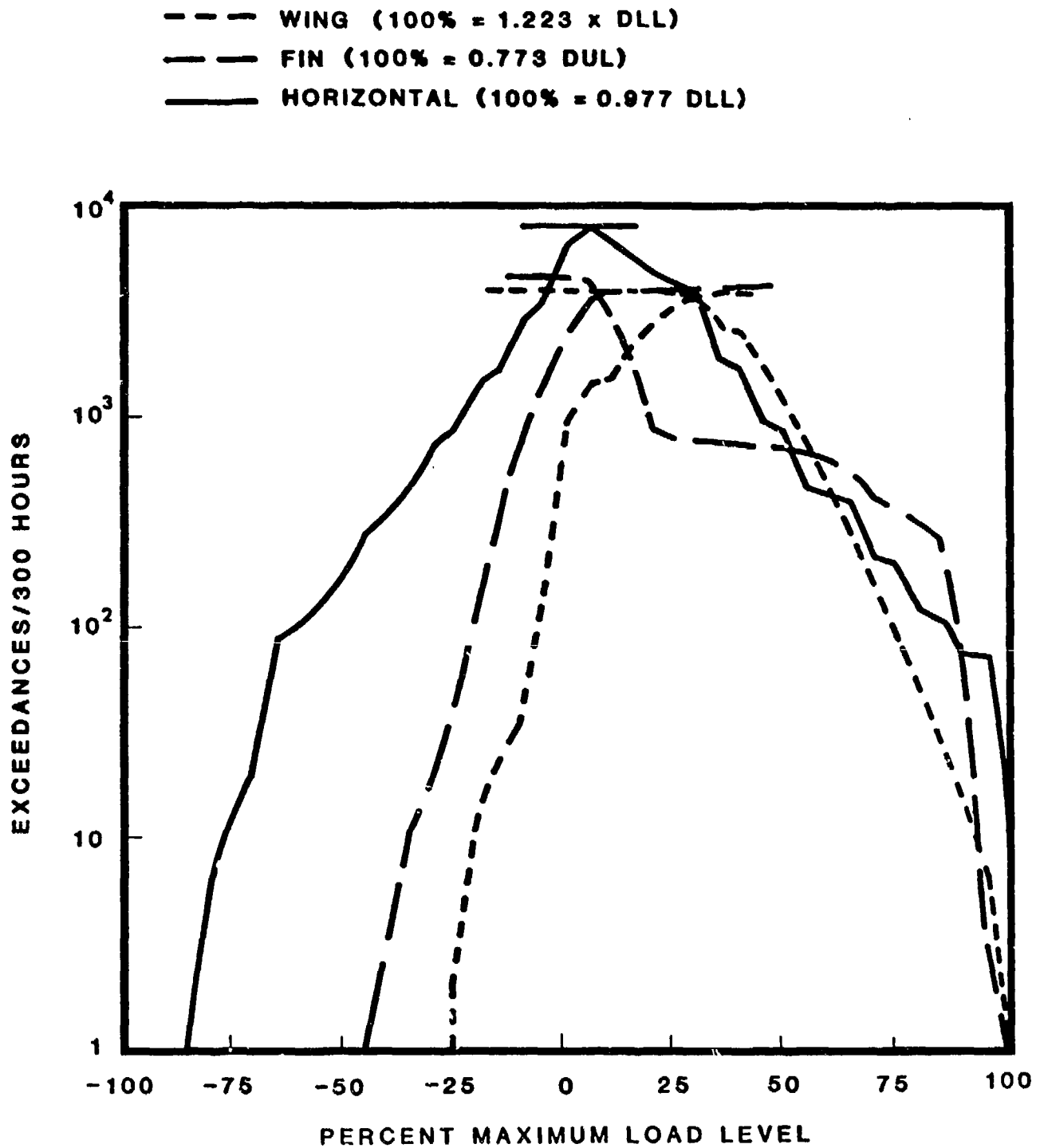
FIGURE 45. COMPARISON OF F-18 FATIGUE SPECTRA.
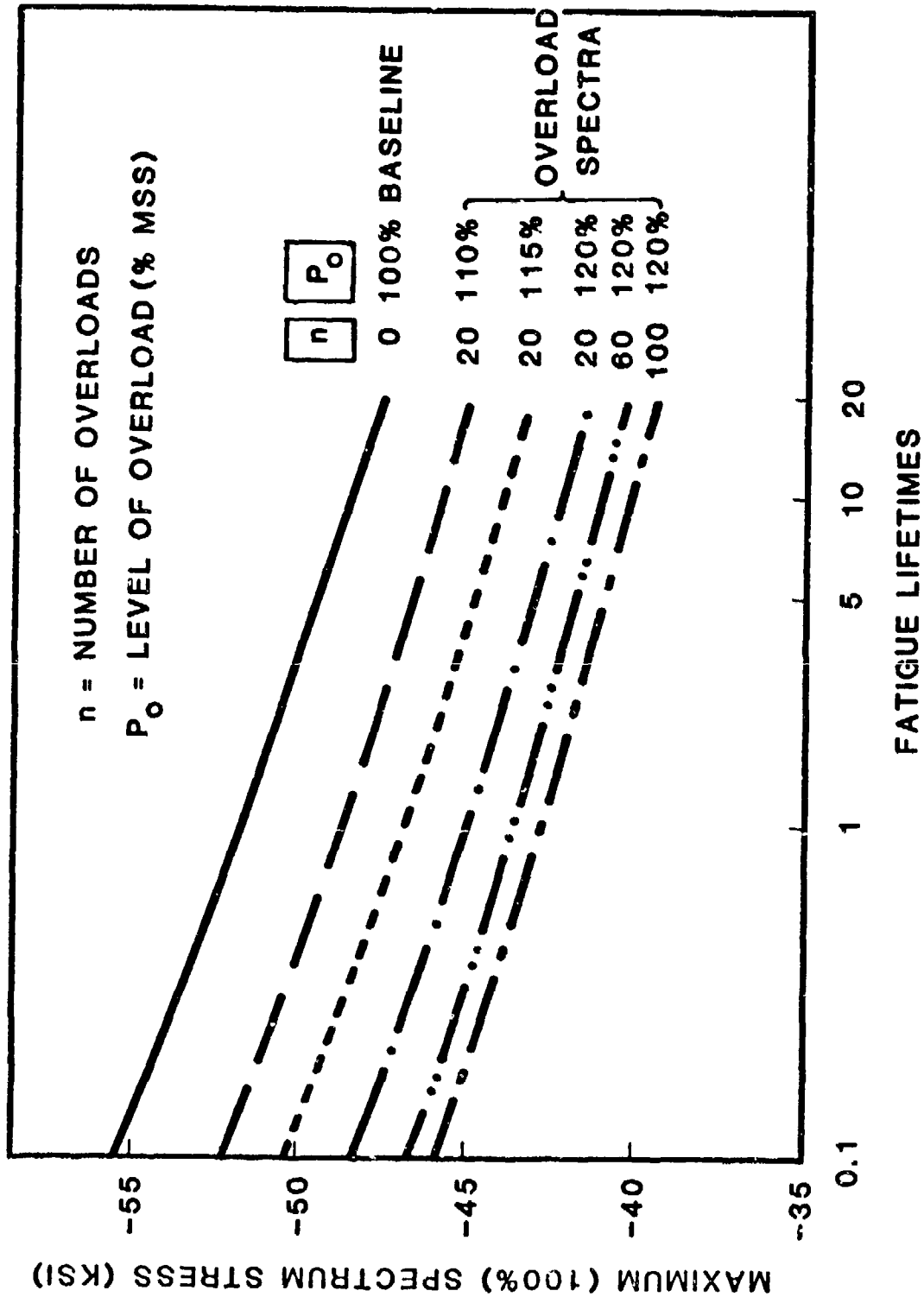
FIGURE 46. INFLUENCE OF OVERLOADS ON COMPOSITE FATIGUE LIFE FOR AN F-18 UPPER WING ROOT SPECTRUM.

aircraft lifetime. The data in Figure 46 show that overloads significantly reduce fatigue life for a constant maximum spectrum stress level (100 percent) for the allowable maximum spectrum stress level for a constant fatigue lifetime. This effect is observed for all three F-18 spectra type and is summarized in Figure 47. It can be seen that the influence of overloads on composite fatigue life is similar for all three spectrum types.

The overloads shown in Figure 47 are also checked for their influence on metallic fatigue life. For all over-load/occurrence combinations, metal fatigue life remains within approximately ±20 percent of the no overload baseline spectrum life. Thus, it is demonstrated analytically that overloads can be used to significantly reduce composite fatigue life, without significantly changing metal fatigue life. Figure 48 summarizes how overloads change the B-basis stress-life relationships for composites and metals. B-basis stress-life relationships are determined from mean values using the appropriate scatter factors for each material. These are determined from the Task I data analysis (see Volume I). For the metal data in Figure 48, the B-basis overload stress-life relationships is approximately equal to the B-basis baseline stress-life relationship. The B-basis overload stress-life relationship for composites includes the effects of both overload life sensitivity and increased fatigue life scatter (compared to aluminum). The data in Figure 48 show that the overloads significantly reduce the differences in life between metals and composites. However, the composite stress-life relationship remains very flat relative to the metal fatigue behavior. Thus, a common B-basis life for both materials is only established at one stress level (approximately 35 Ksi). At one fatigue lifetime, the allowable composite fatigue stress level is still significantly higher than that for aluminum.

It is, therefore, concluded that overloads can only cause an intersection (not overlap) of composite and metal B-basis stress-life relationships. This occurs because overloads
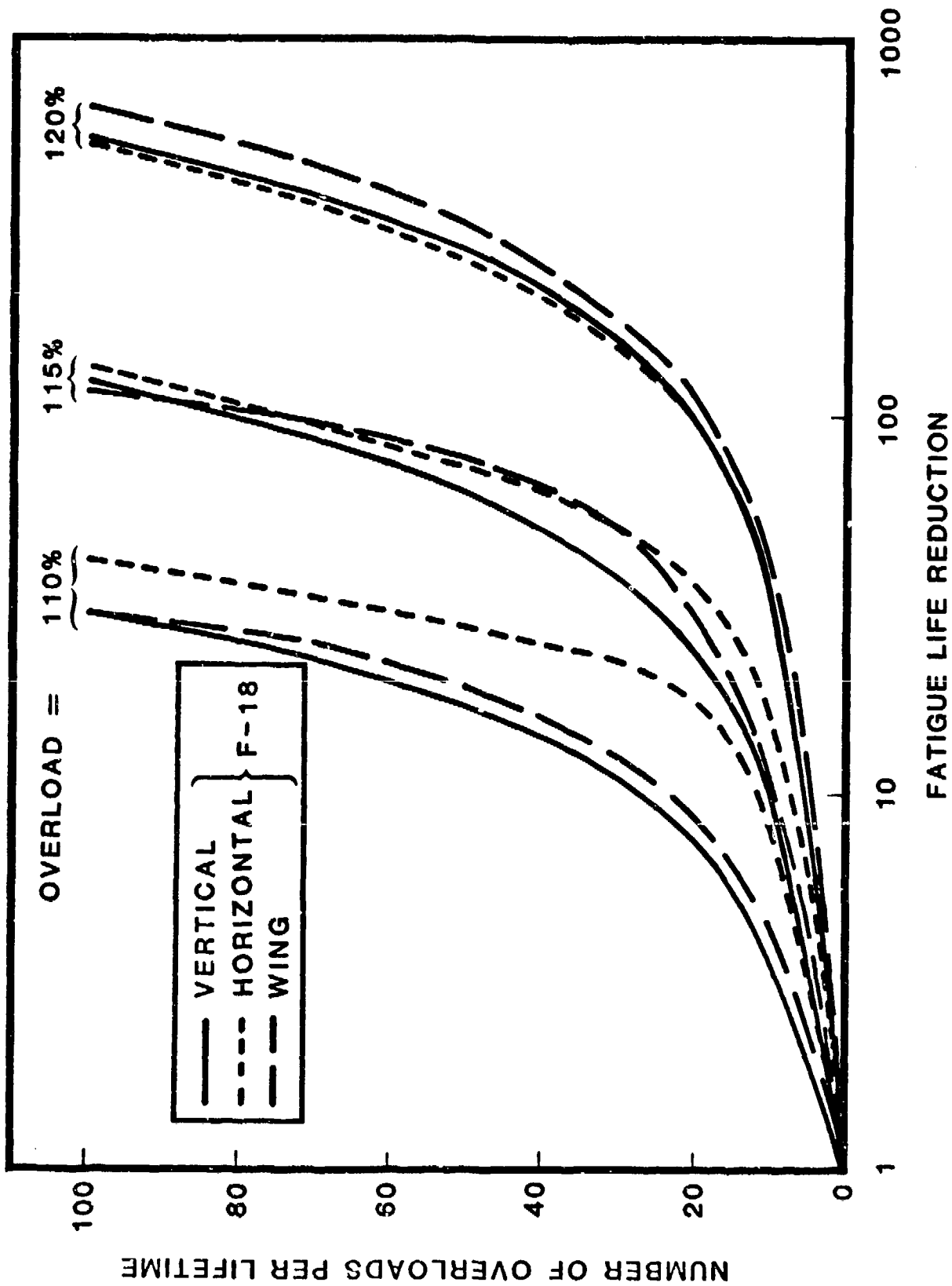
FIGURE 47. INFLUENCE OF SPECTRUM TYPE AND OVERLOADS ON COMPOSITE FATIGUE LIFE.
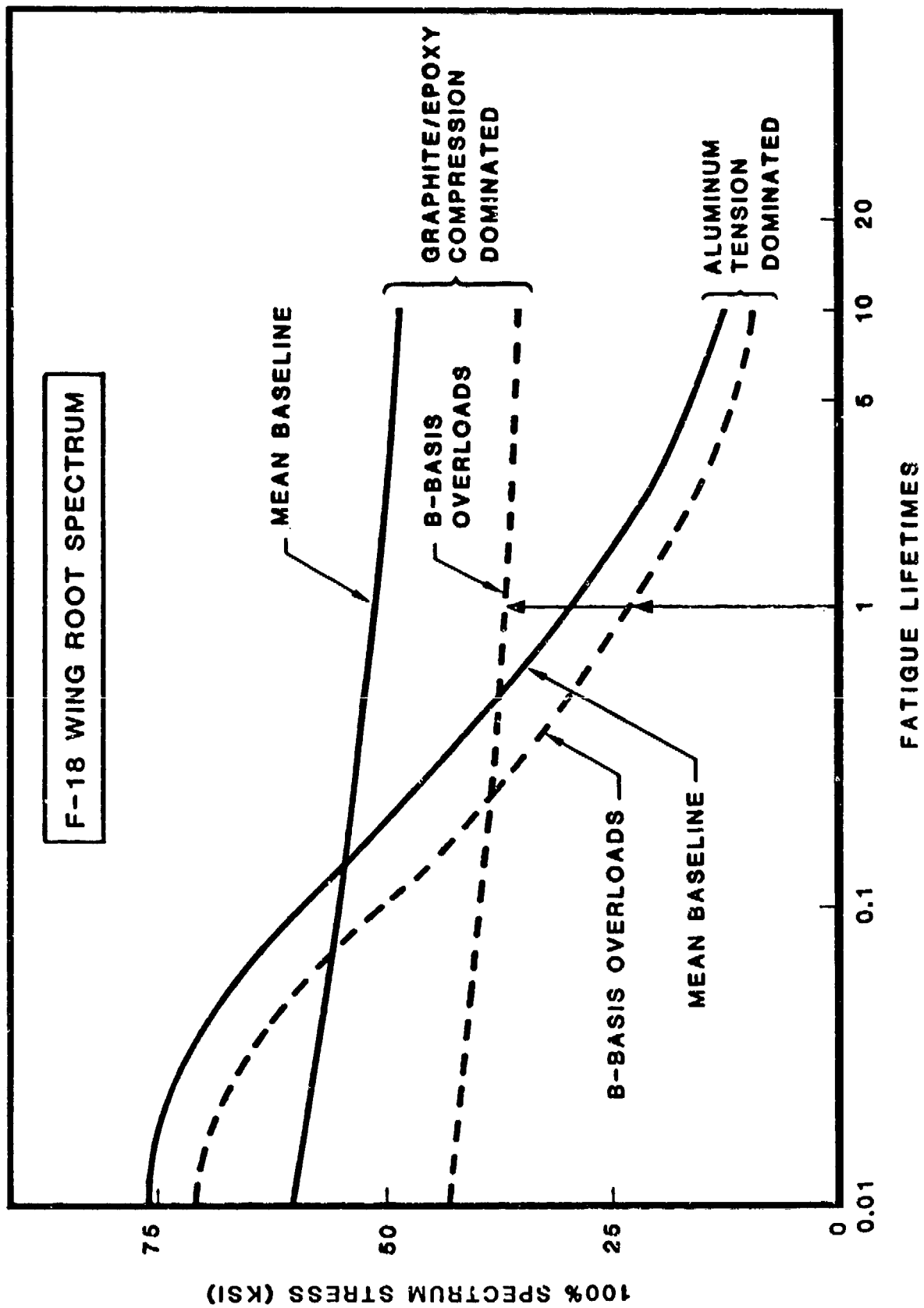
FIGURE 48. INFLUENCE OF OVERLOADS ON COMPOSITE AND METAL FATIGUE LIFE.

only reduce composite fatigue life. They do not alter the fatigue damage accumulation mechanisms, and thus cannot increase the slope of the S-N curve.

This observation does not necessarily limit the application of the overload approach to mixed structure fatigue testing. The influences of overloads on composite fatigue life (shown in Figure 46) can be used to calculate equivalent B-basis load enhancement factors. For example, Figure 25 shows that the B-basis LEF for a two lifetime fatigue test with n = 1 is 1.128. This value can be used to determine the number and magnitude of overloads which give a 1.128 times reduction in the allowable 100 percent stress level at two lifetimes. Thus, the overloads can be used to determine an equivalent test to the full load enhancement approach (all loads) for composites, without changing metal fatigue life. Thus, the disadvantage of the full LEF approach (Section 2.2) for mixed structure is overcome. Figure 49 summarizes the number and level of spectrum overloads required to demonstrate equivalent test severity to the full load enhancement approach. For the example discussed above, B-basis composite reliability for a two lifetime test can be achieved by using approximately ten 120 percent overloads/lifetime or approximately thirty-five 115 percent overloads/lifetime.

Thus, the change in spectrum approach can be used to provide demonstration of B-basis reliability for both composite and metals in a mixed structure without causing over severe metallic test. This significantly reduces the problems of mixed structure fatigue testing. It should be noted that the number and extent of overloads required will be spectrum type and stress-level dependent. In addition, transport or bomber type spectrum may be less amenable to this approach because of their relatively low number of high loads/lifetime. The promise of this approach should be explored in more detail and verified experimentally.

2.5     Summary of Certification Approach Evaluation

The evaluation of the Navy certification approach has
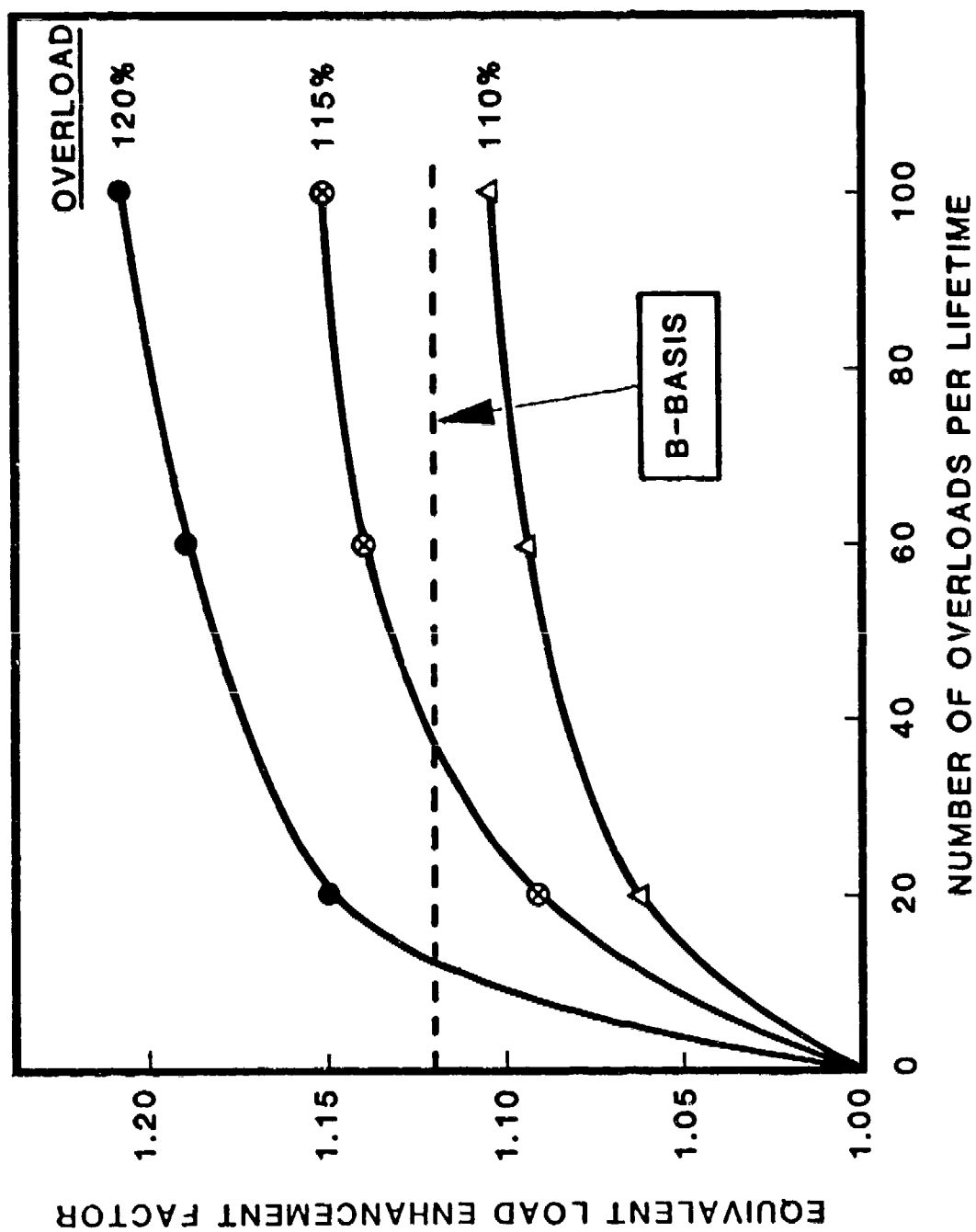
FIGURE 49. EQUIVALENT COMPOSITE LOAD ENHANCEMENT FACTORS FOR AN F-18 UPPER WING SPECTRUM.

shown that, in general, it is soundly based. In particular, it strikes a good balance between the technical requirements of the tests and cost effectiveness. For static testing, the main area of concern is the assumed ability to predict full-scale structure failure modes. Published data on full-scale composite static tests have shown that failure mode predictions are frequently incorrect. For fatigue testing, the two-lifetime fatigue test led to a concern about demonstrated fatigue reliability due to the high fatigue life scatter observed in composites. This is offset somewhat by the severe design spectra used. However, the degree of severity of the design spectra relative to the actual service spectra is uncertain at the time of certification.

Evaluation of the ultimate strength approach has established criteria for omission of a composite fatigue test. The overload requirement is shown to be a function of the fatigue spectrum type, specimen configuration, lay-up and test environment. In general, the ultimate strength approach has significant potential for reducing design development fatigue test requirements for composites.

The evaluation of the enhanced loads approach has shown that it has a sound theoretical basis and can be used with confidence for certification testing. However, some practical limits of this approach may exist. First, for asymmetric spectra, the degree of load enhancement may be limited because of a requirement not to exceed ultimate load. Second, for mixed structure, the enhanced load approach will provide an excessively severe fatigue test for the metal parts.

It is shown that the change in spectrum approach could not be used to overlap composite and metal B-basis stress-life relationships. This cannot be achieved because of the significant differences in the slopes of their stress-life relationships. However, it is shown that overloads can be introduced which permit demonstration of B-basis test reliability for both composite and metal parts in a mixed structure. This is achieved without making the test overly severe for metallic structures.

## SECTION 3

## METHODOLOGY DEVELOPMENT

The high cost of full-scale structural test prohibits generation of a sufficient number of data for statistical analysis. Thus, for a meaningful interpretation of full scale structural tests, a building-block approach is developed for certification testing of composite structures. This approach fully utilizes coupon, element, subcomponent and component level test data so that a limited number of full-scale structural test data can be interpreted statistically. The number of tests decreases from the coupon level to the component level. A relatively large number of tests is required at the coupon level to establish the data scatter and B-basis statistics for different loading modes, failure modes and environments for both static and fatigue tests. A smaller number of tests is required at the element and subcomponent level to determine the failure mode interaction and a sufficient number of component tests to demonstrate the variability in structure response. This information is then used for interpretation of the full-scale structural test data. The number and types of test specimens required for the building block approach are specified in Section 5. The variability in structural response is discussed in the following paragraphs.

3.1    Concept of Structural Response Variability (SRV)

In structural tests, aside from the scatter in the basic material properties, other factors will contribute to the scatter in structural response. The contributing factors (above the coupon level) are structural geometry, design tolerances, manufacturing and material nonuniformity and loading conditions. Because of this scatter in structural response, unexpected "hot-spot" failure can occur during static structural tests. Figure 50 shows schematically a potential static "hot-spot" failure in relation to the scatter in material strength and structural response. The shaded area where the two distributions intersect
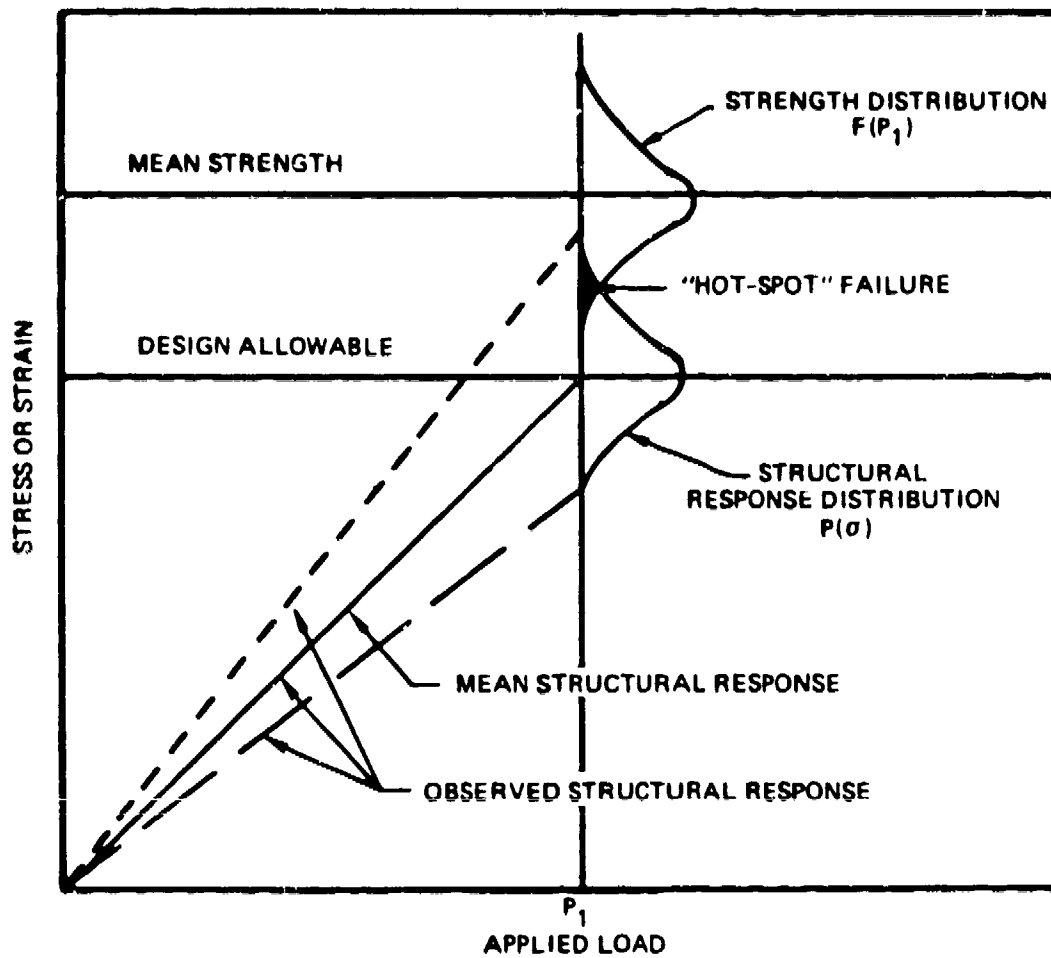
FIGURE 50. SCHEMATIC OF "HOT-SPOT" FAILURE IN RELATION TO
THE SCATTER IN STRENGTH AND IN STRUCTURAL
RESPONSE FOR A STATIC TEST.

represents the probability of failure at a certain location of the structure under the applied load $P_1$. In order to interpret the data statistically, both the material strength and structural response distributions must be obtained for different locations on the structure. This cannot be achieved from a limited number of full-scale structural tests.

The building block approach requires utilization of coupon level tension and compression static test data to establish the basic material strength distribution and element combination and subcomponent level test data to establish the structural response distribution. The structural reliability or the probability of structural failure (hot-spot failure) is then calculated from these two distributions as follows:

Using a joint probability function to combine the influence of material static strength scatter and the SRV, the probability of survival of a structure under a static load level of $P_1$ is given by

$$S(P_1) = \int_0^\infty F(P_1) \cdot p(\sigma) d\sigma \qquad (26)$$

where: $F(P_1)$ is the survivability of the structure at load level $P_1$ considering strength scatter only

$p(\sigma)$ is the probability of occurrence that the actual stress level in the structure is $\sigma$ due to SRV.

The survivability function $F(P_1)$ describes the static strength due to material scatter only. Therefore, the Weibull parameters obtained from the static strength data analysis in Task I can be used in Equation (26).

The scatter in structural response is a result of several factors, as discussed earlier. The proportion that each of these factors contribute to the total scatter in structural response is difficult to determine. The use of strain gages to measure structural response gives the combined effects of these

factors. To utilize the strain data in evaluating the scatter in structural response, a combined distribution, such as $p(\sigma)$ in Equation 26 is necessary.

The form of the distribution function $p(\sigma)$ can be assumed as a normal distribution or a two-parameter Weibull distribution. Within the range of observed SRV's for static testing, the two distribution yield approximately the same results when used in Equation (26). Numerical integration of Equation (26) has been carried out to evaluate the influence of SRV on the structural reliability. The results of these calculations are shown in Figure 51 through 54.

The influence of SRV on the design allowables depends on the static strength variability, which is characterized by the the static strength distribution Weibull shape parameter $(\alpha_S)$, and the sample size. It can be seen from Figures 51 and 52 that additional reduction factors (from the baseline values) on the design allowables are required to accommodate the structural response variability. However, within the range of the SRV values determined in Section 3.2, the additional reduction factor is fairly small. The numerical values of the reduction factors for mean and modal SRV with a single article test are given in Table 4. The B- and A-basis knockdown factors at mean and modal SRV values for various values of $\alpha_S$ are shown in Figures 53 and 54. In comparing these results with the baseline knockdown factors (no SRV), it is seen that the design allowables are dominated by the static strength variability, and that SRV is a secondary consideration.

The influence of structural response variability and the 95 percent reliability (R) for the baseline B- and A-basis design allowables is shown in Figures 55 and 56. The structural reliability decreases as the SRV increases (decreasing $\alpha_{SRV}$). The influence of SRV on reliability does not depend on the sample size but depends on the static strength variability $\alpha_S$. At high levels of static strength variability (low $\alpha_S$), the influence of SRV on reliability, R, is small. However, at low levels of
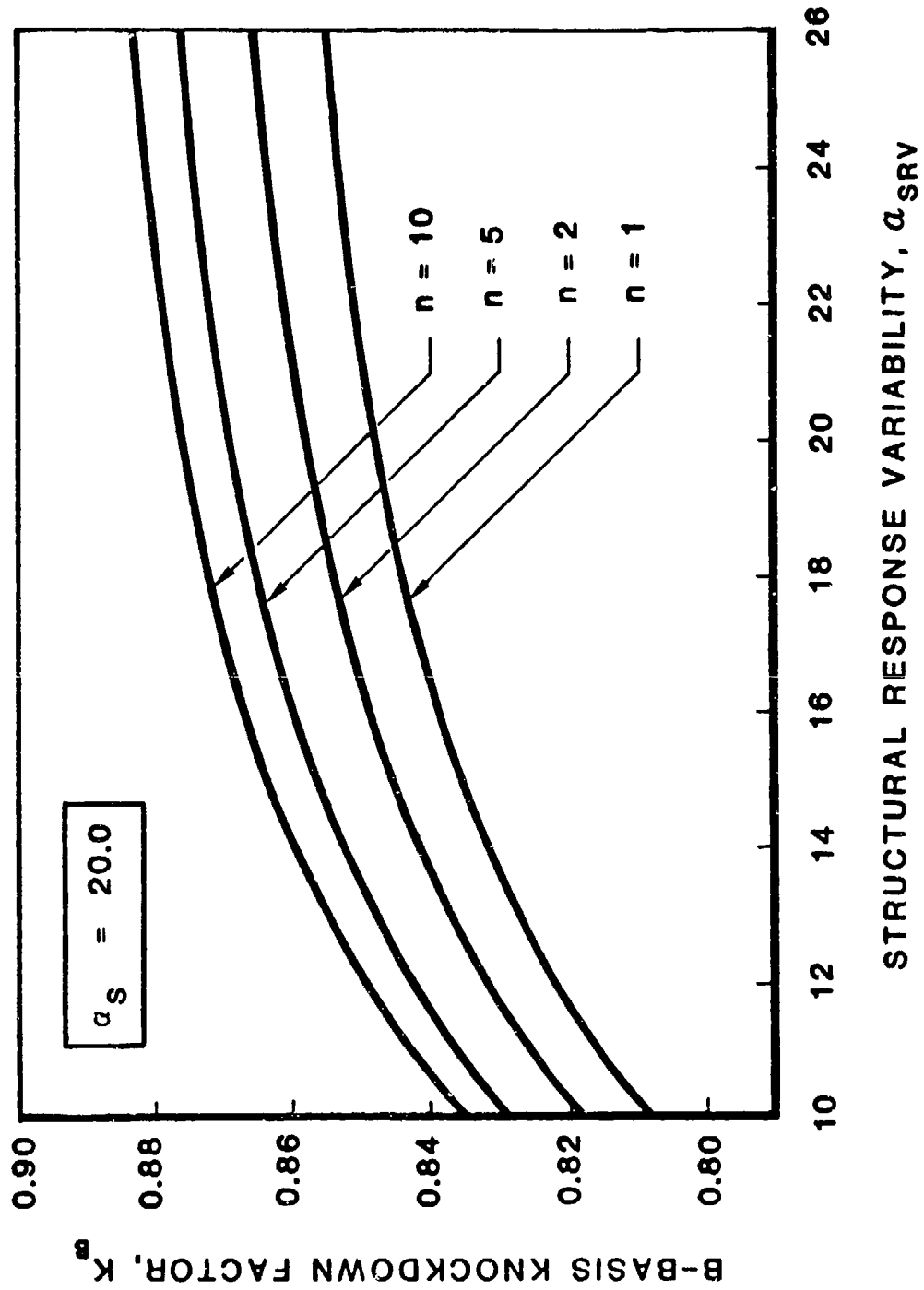
FIGURE 51. INFLUENCE OF SRV ON B-BASIS DESIGN ALLOWABLES, $\alpha_s = 20.0$.
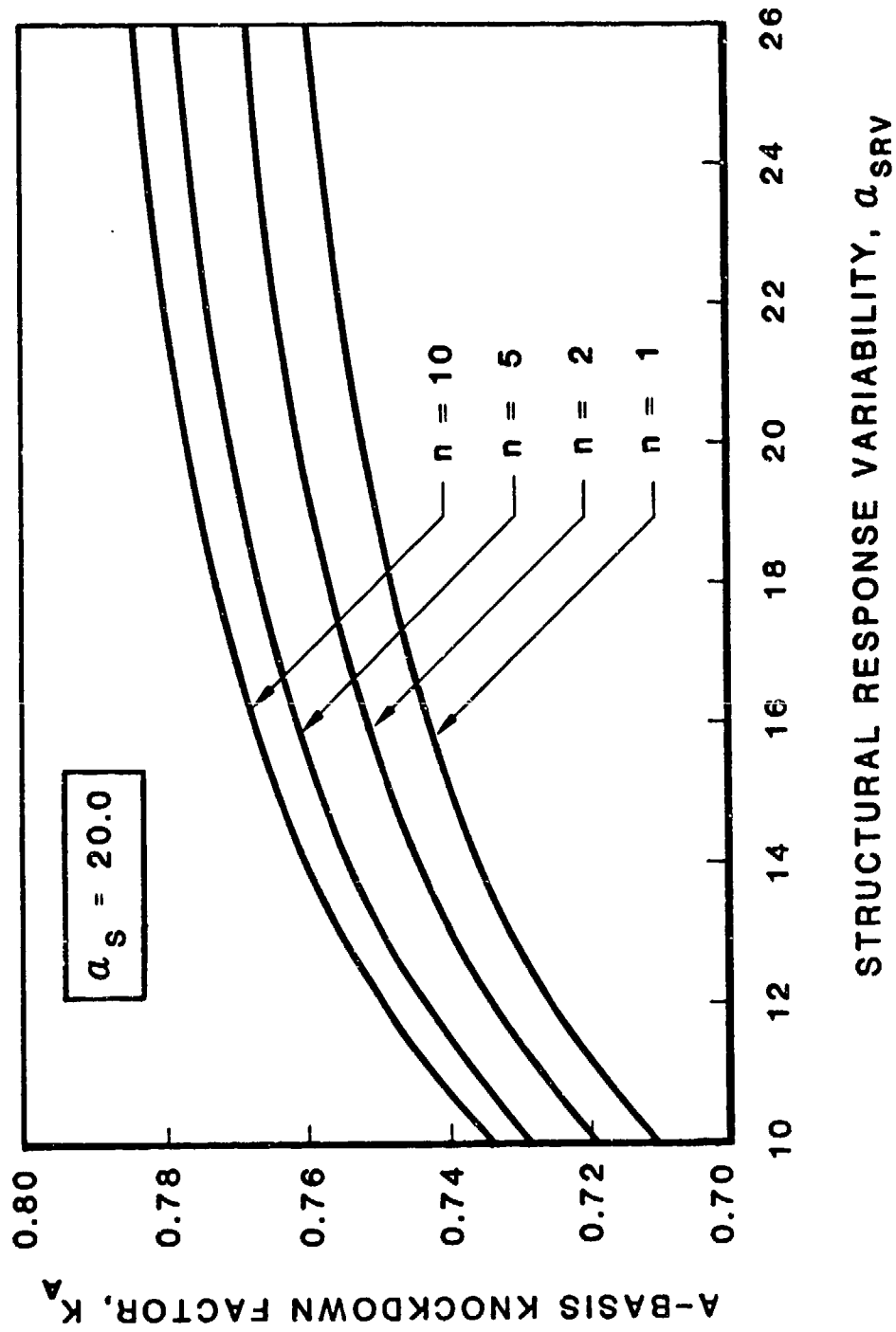
FIGURE 52. INFLUENCE OF SRV ON A-BASIS DESIGN ALLOWABLES, $\alpha_s = 20.0$.
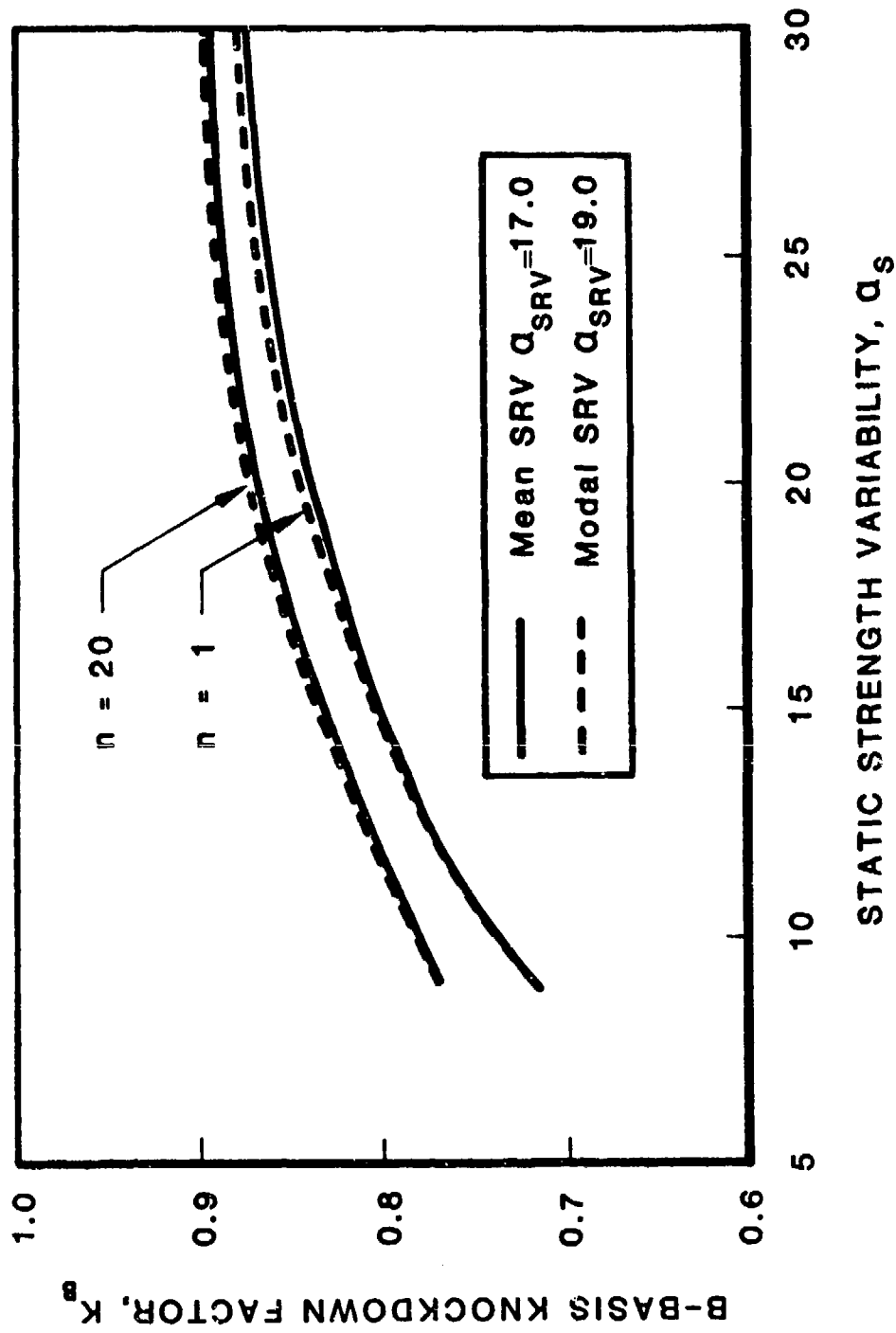
FIGURE 53. INFLUENCE OF STATIC STRENGTH VARIABILITY ON THE B-BASIS DESIGN ALLOWABLE AT MEAN AND MODAL SRV.
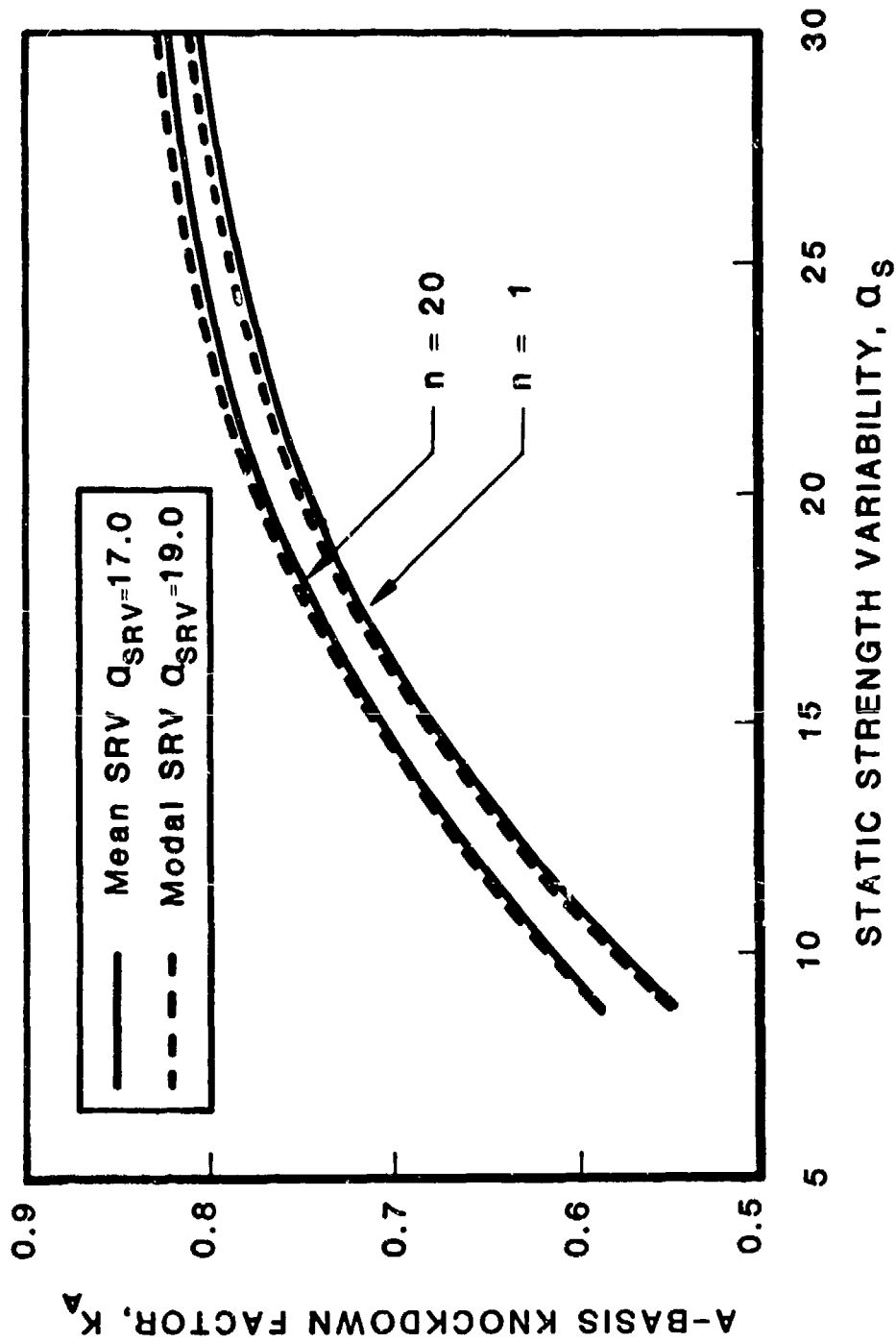
FIGURE 54. INFLUENCE OF STATIC STRENGTH VARIABILITY ON THE A-BASIS DESIGN ALLOWABLE AT MEAN AND MODAL SRV.

**TABLE 4.  INFLUENCE OF SRV ON KNOCKDOWN FACTORS, n = 1.**

| | NO SRV | MEAN SRV | MODAL SRV |
|---|---|---|---|
| $a_s$ | ∞ | 17.0 | 19.0 |
| CV | 0 | 0.071 | 0.065 |
| $a_m$ | B-BASIS KNOCKDOWN FACTOR | | |
| 8.8 (B-BASIS) | 0.723 | 0.711 | 0.713 |
| 20.0 (MODAL) | 0.869 | 0.842 | 0.846 |
| 23.2 (MEAN) | 0.886 | 0.855 | 0.860 |
| 30.0 (UPPER) | 0.911 | 0.874 | 0.879 |
| | A-BASIS KNOCKDOWN FACTOR | | |
| 8.8 (B-BASIS) | 0.553 | 0.544 | 0.545 |
| 20.0 (MODAL) | 0.773 | 0.746 | 0.750 |
| 23.2 (MEAN) | 0.801 | 0.775 | 0.770 |
| 30.0 (UPPER) | 0.842 | 0.803 | 0.810 |
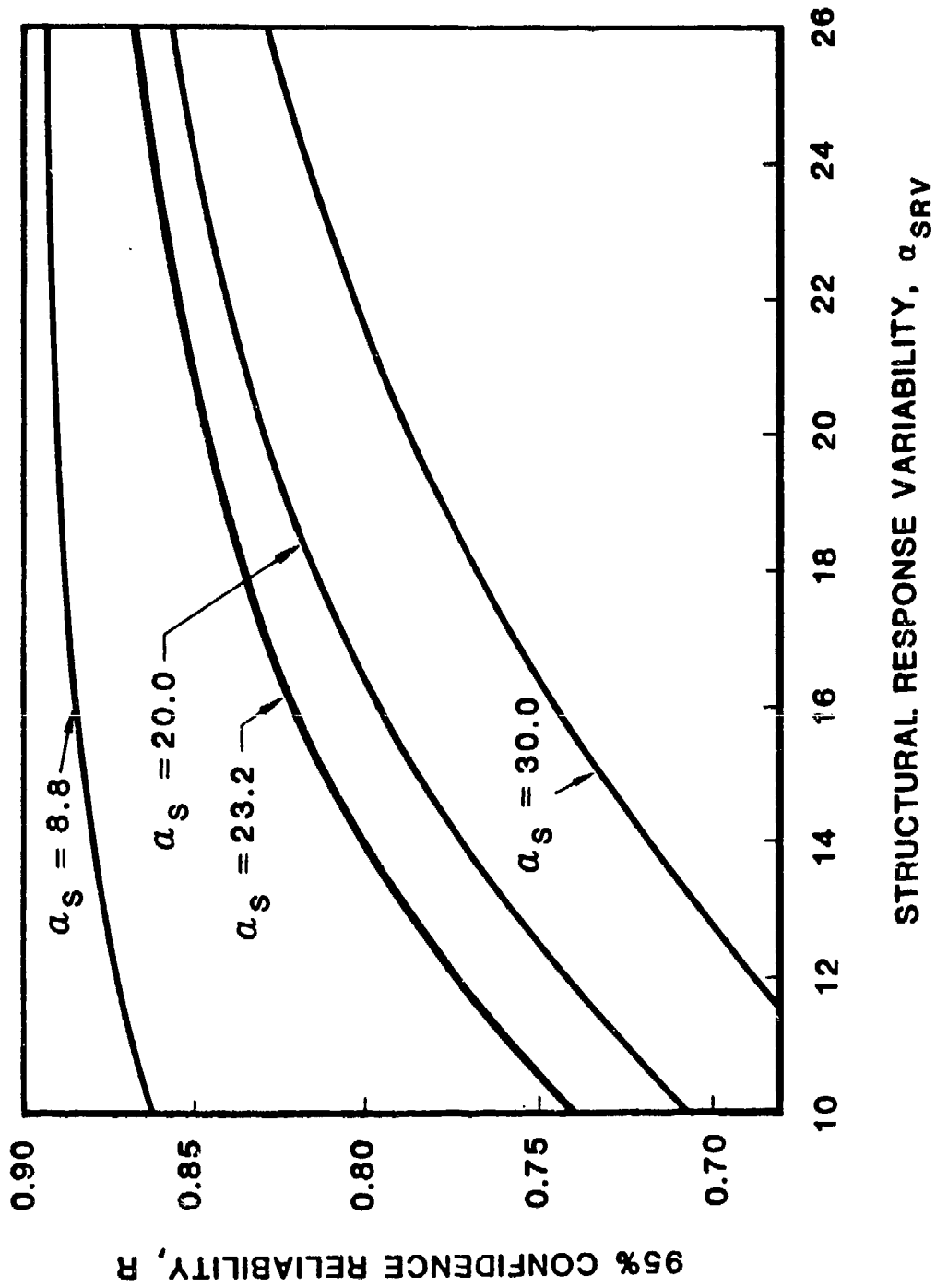
FIGURE 55. INFLUENCE OF STRUCTURAL RESPONSE VARIABILITY ON THE B–BASIS STRUCTURAL RELIABILITY.
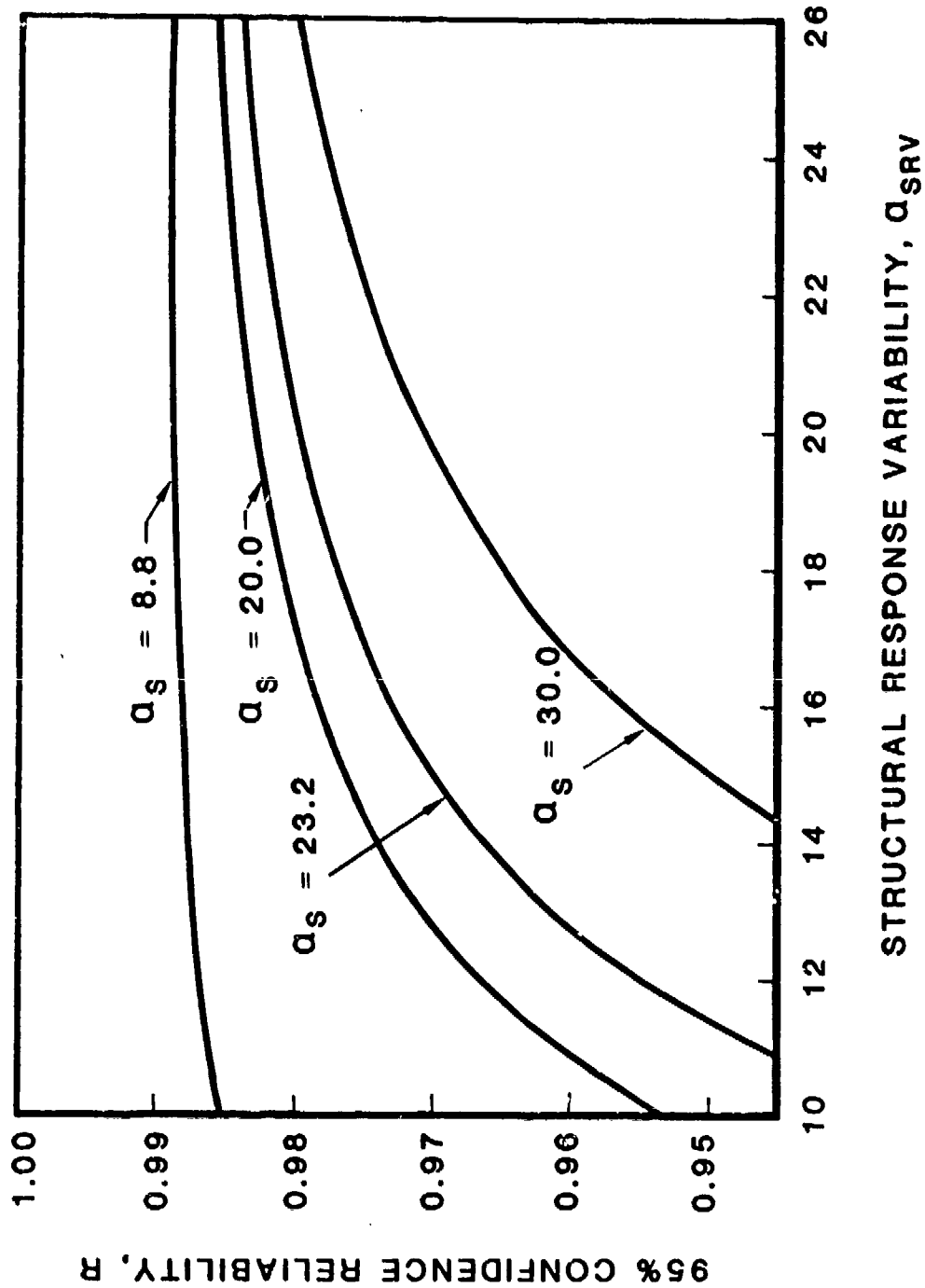
FIGURE 56. INFLUENCE OF STRUCTURAL RESPONSE VARIABILITY ON THE A-BASIS STRUCTURAL RELIABILITY.

static strength variability (high $\alpha_S$), the reliability decreases significantly as SRV increases. This can be seen in Figure 55. For B-basis reliability (95 percent confidence, 90 percent probability), at $\alpha_S$ = 8.8 the reliability reduces from 0.90 for $\alpha_{SRV}$ =∞ to 0.863 for $\alpha_{SRV}$ = 10.0. Whereas for $\alpha_S$ = 30.0, the reliability reduces to 0.655. Similar results can be observed for A-basis reliability shown in Figure 56.

3.2        Structural Response Variability Data

The extensive test data generated in Reference 5 have been used as a strain gage data source. The building block approach used in Reference 5 for the wing structure is shown in Figure 57. The design development test specimens are character-ized by four levels of structural complexity. The fifth level of complexity is assigned to the full-scale wing component. Further details of the wing specimens are discussed in Section 4. Table 5 summarizes the load-strain data available from Reference 5.

3.2.1        Determination of Structural Response Variability

To determine the structural response variability from the strain data, three or more nominally identical specimen tests are required. A typical data set from Reference 5 is shown in Figure 58. Strain data at the critical locations obtained from the structural test are normalized with respect to the mean strain at each load level. The distribution of the normalized strain is then fitted to a normal or Weibull distribution to evaluate the scatter in structural response. It should be noted that only the scatter of the structural response is of interest in this evaluation. The actual magnitude of the strain is not important. The structural response variability is characterized by the coefficient of variation in a normal distribution and by the shape parameter in the Weibull distribution. The combined structural response variability can then be obtained by inte-grating equation (26) with the given probability distributions.

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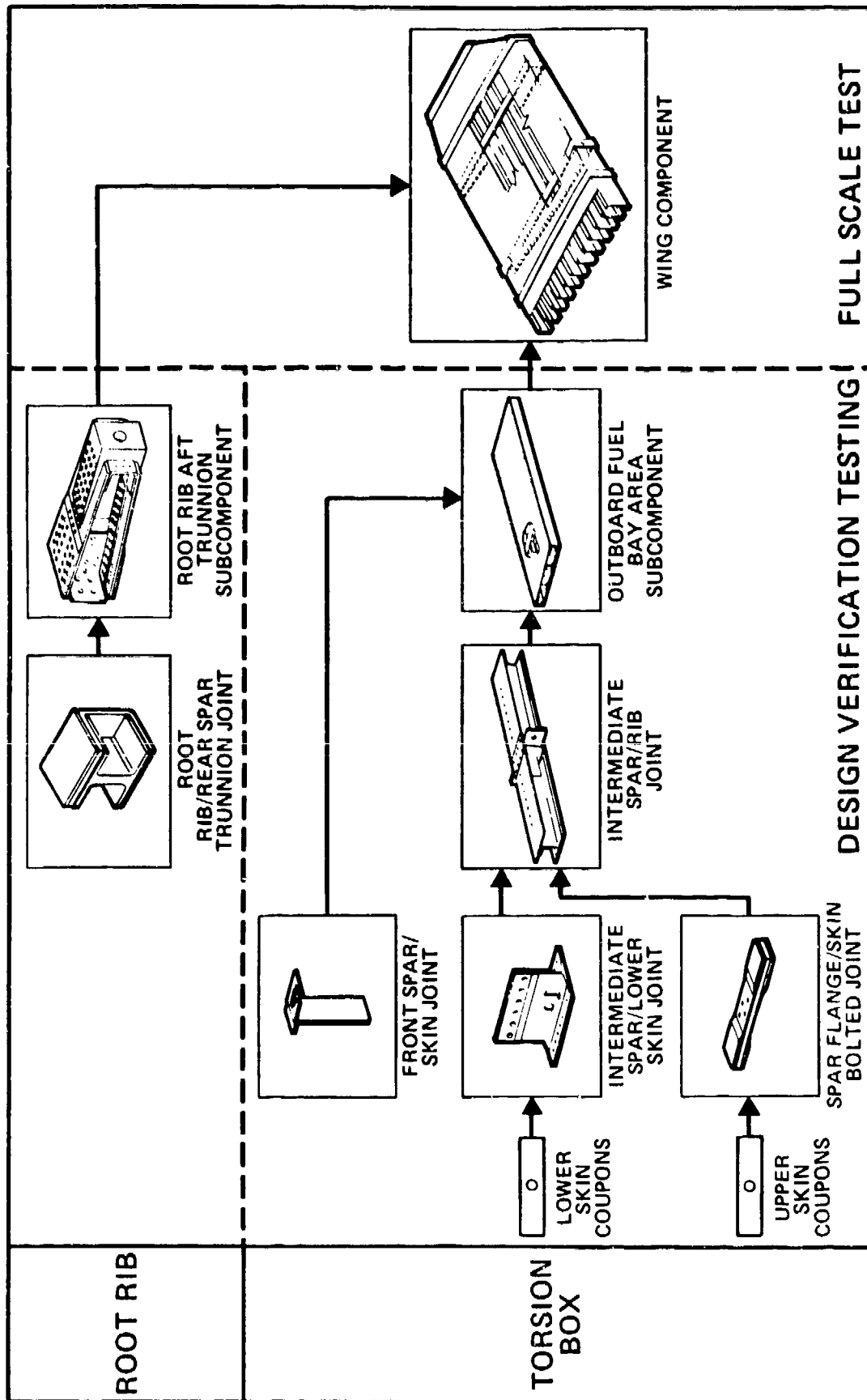
FIGURE 57. BUILDING BLOCK APPROACH FOR THE WING STRUCTURE IN THE COMPOSITE WING/FUSELAGE PROGRAM (REFERENCE 5).

**TABLE 5. SUMMARY OF STRAIN GAGE DATA AVAILABLE FROM THE COMPOSITE WING/FUSELAGE PROGRAM (REFERENCE 5).**

| COMPLEXITY LEVEL | SPECIMEN | NUMBER OF SPECIMENS | | | |
|---|---|---|---|---|---|
| | | STATIC | | RESIDUAL STATIC STRENGTH | |
| | | RT/AMBIENT | 250°F/WET | RT/AMBIENT | 250°F/WET |
| 1 | COMPRESSION COUPONS<br>TENSION COUPONS | 23<br>9 | 23<br>9 | 23<br>9 | 23<br>9 |
| 2 | WE-2<br>WEC-1 | 3<br>3 | 3<br>3 | 3<br>6 | 3<br>6 |
| 3 | WEC-3 | 3 | 3 | 6 | 6 |
| 4 | WS-1<br>WS-2 | 2<br>1 | 3<br>2 | ---<br>1 | 5<br>1 |
| 5 | WCC-1 | 1 | 1 | 1 | --- |
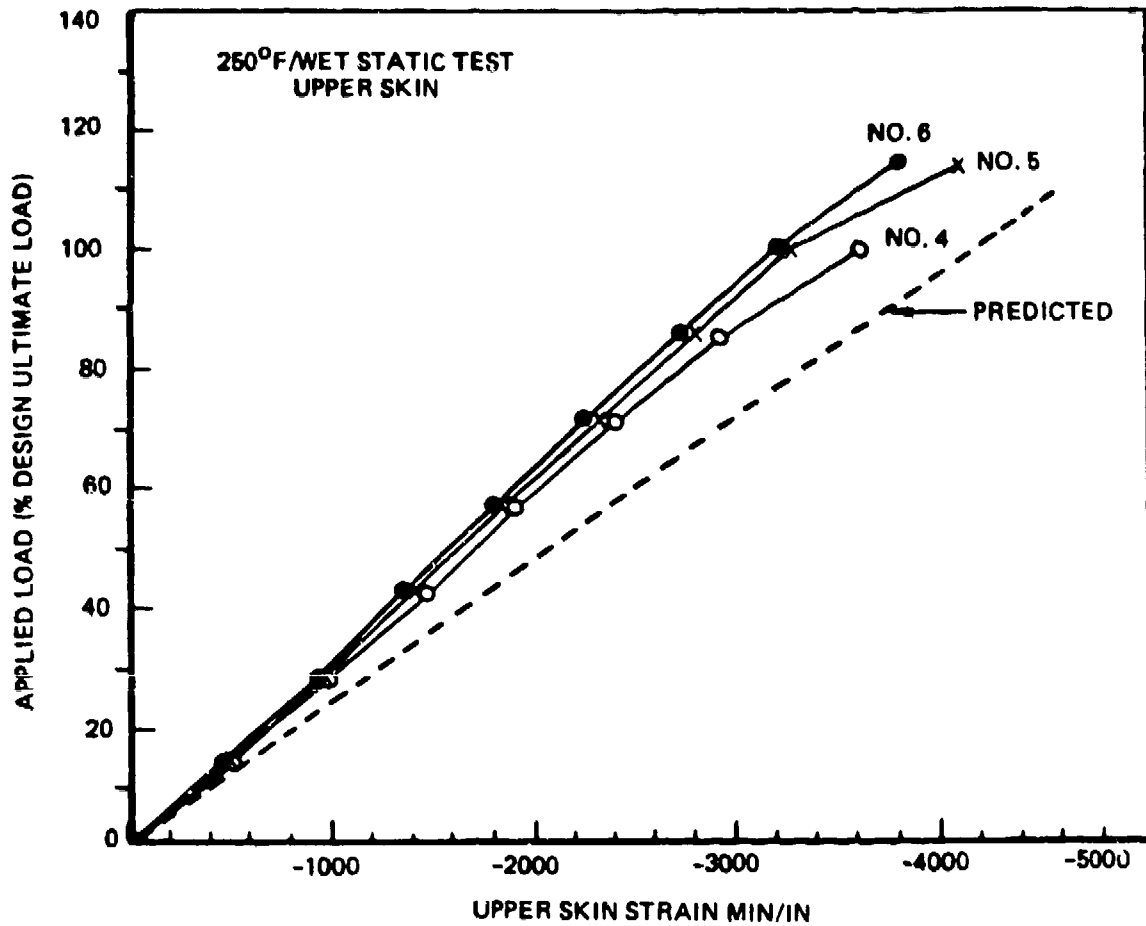
FIGURE 58.  TYPICAL APPLIED LOAD STRAIN RESPONSE FROM
A WS-1 SUBCOMPONENT TEST (REFERENCE 5).

## 3.2.2    Structural Response Variability Analysis Results

The results of the structural response variability analysis are presented in Table 6. The location of the strain gages on each specimen are given in the table.

For analysis purposes, the load-strain data are separated into three categories: RTA, ETW and the combined RTA and ETW data. In each category, static and residual static strength data are pooled. This significantly increased the amount of load-strain data available for analysis. Residual static strength load-strain data are included because these tests showed no fatigue degradation and exhibited no significant losses in static strength after two lifetimes of fatigue loading.

In general, the load-strain response of all specimens was essentially linear to failure. A typical example is presented in Figure 59, which shows the ETW lower skin load-strain response of the WS-1 box beam. Figures 60 through 64 show the upper skin load-strain response of the wing component close to the root rib. These data also show essentially linear load-strain response. The ETW upper skin load-strain response of the WS-1 box beam presented in Figure 60 shows nonlinear load-strain response and high structural response variability. The nonlinear load-strain response can be attributed to the severe test environment (250°F/1.3% moisture). This is the only load-strain data analyzed in Table 6, which exhibited significant nonlinear load-strain response. It should be noted that this type of load-strain response is not accounted for in the Navy scatter factor approach to certification.

The influence of specimen complexity on structural response variability is presented in Table 6 and summarized in Figures 65 through 67. The data show that structural response variability does not increase as specimen complexity increases. This observation holds true for the RTA, ETW and combined RTA/ETW load-strain data sets. In fact, the structural response variability of the wing component (complexity level 5) is less than the overall average for the RTA and combined RTA/ETW data sets.

# TABLE 6. STRUCTURAL RESPONSE VARIABILITY ANALYSIS RESULTS.

| SPECIMEN I.D. | SPECIMEN DESCRIPTION | FAILURE MODES | SPECIMEN COMPLEXITY LEVEL | GAGE NO. | GAGE LOCATION | RTA n | RTA C.V.% | ETW n | ETW C.V.% | RTA + ETW n | RTA + ETW C.V.% |
|---|---|---|---|---|---|---|---|---|---|---|---|
| WC-1 | Upper Wing Skin Coupon | NET SECTION AT FASTENER HOLE | 1 | — | GAGE SECTION | 20 | 3.4 | 19 | 4.5 | 39 | 5.1 |
| WC-2 | Lower Wing Skin Coupon | | 1 | — | | 6 | 5.2 | 5 | 10.4 | 11 | 7.9 |
| WC-3 | Upper Wing Skin Coupon | | 1 | — | | 20 | 3.4 | 15 | 3.4 | 35 | 3.5 |
| WC-4 | Lower Wing Skin Coupon | | 1 | — | | 6 | 2.7 | 5 | 3.8 | 10 | 5.4 |
| WC-5 | Lower Wing Skin Coupon | | 1 | — | | 4 | 3.1 | 5 | 7.1 | 9 | 5.5 |
| WC-6 | Upper Wing Skin Coupon | | 1 | — | | 6 | 1.4 | 6 | 2.6 | 12 | 5.3 |
| WE-2 | Upper Wing Skin/Spar Bolted Joint Element | FASTENER HOLE | 2 | — | UPPER SKIN | 6 | 1.9 | — | — | 6 | 1.9 |
| WEC-1 | Lower Wing Skin/Intermediate Spar Cocured Joint Element Combination | SPAR WEB FUEL DRAIN HOLE | 2 | 14 | SPAR WEB | 6 | 7.0 | 4 | 13.7 | 10 | 9.5 |
| WEC-3 | Wing Intermediate Spar/Rib Joint Element Combination | . COCURED JOINT / . UPPER SKIN | 3 | 7A | LOWER SKIN | 6 | 7.6 | 4 | 5.3 | 10 | 7.0 |
| | | | | 8A | UPPER SKIN | 6 | 9.7 | 4 | 4.9 | 10 | 8.6 |
| | | | | 15($\epsilon_\delta$) | SPAR WEB | 6 | 13.8 | 5 | 8.5 | 11 | 12.2 |
| WS-1 | Wing Outboard Fuel Bay Subcomponent | . UPPER SKIN / . LOWER SKIN / . COCURED JOINT | 4 | 19 | UPPER SKIN | 3 | 13.1 | 6 | 5.7 | 9 | 9.0 |
| | | | | 21R1 | | 2 | 10.8 | 6 | 17.5 | 8 | 15.9 |
| | | | | 21R2 | | 2 | 8.9 | 5 | 14.9 | 7 | 13.4 |
| | | | | 23 | | 2 | 4.9 | 6 | 7.7 | 8 | 7.1 |
| | | | | 20 | LOWER SKIN | 2 | 2.7 | 4 | 4.9 | 6 | 4.2 |
| | | | | 22R1 | | 2 | 11.6 | 7 | 7.0 | 9 | 6.3 |
| | | | | 22R2 | | 2 | 1.3 | 5 | 7.8 | 7 | 6.6 |
| | | | | 24 | | 3 | 9.6 | 4 | 2.4 | 7 | 6.6 |
| WS-2 | Wing Root Rib Aft Trunnion Subcomponent | . RIB WEB | 4 | 11 | RIB WEB | — | — | 2 | 6.3 | 3 | 8.1 |
| WCC-1 | Wing Component | . COCURED JOINT / . UPPER SKIN | 5 | 25 | UPPER SKIN | 2 | 3.6 | — | — | 3 | 9.7 |
| | | | | 30 | | 2 | 7.1 | — | — | 3 | 8.0 |
| | | | | 32 | | 2 | 5.9 | — | — | 3 | 5.6 |
| | | | | 40 | | 2 | 3.1 | — | — | 3 | 6.7 |
| | | | | 52 | | 2 | 6.0 | — | — | 3 | 7.0 |
| | | | | 65 | LOWER SKIN | 2 | 1.3 | — | — | 3 | 2.9 |
| | | | | 68 | | 2 | 1.2 | — | — | 3 | 6.5 |
| | | | | 86 | | 2 | 9.0 | — | — | 3 | 7.4 |
| | | | | 92 | | 2 | 7.0 | — | — | 3 | 4.7 |
| | | | | 98 | | 2 | 2.4 | — | — | 3 | 5.4 |

Notes: n = number of load-strain plots analyzed

$\epsilon_\alpha$ = spar web shear strain

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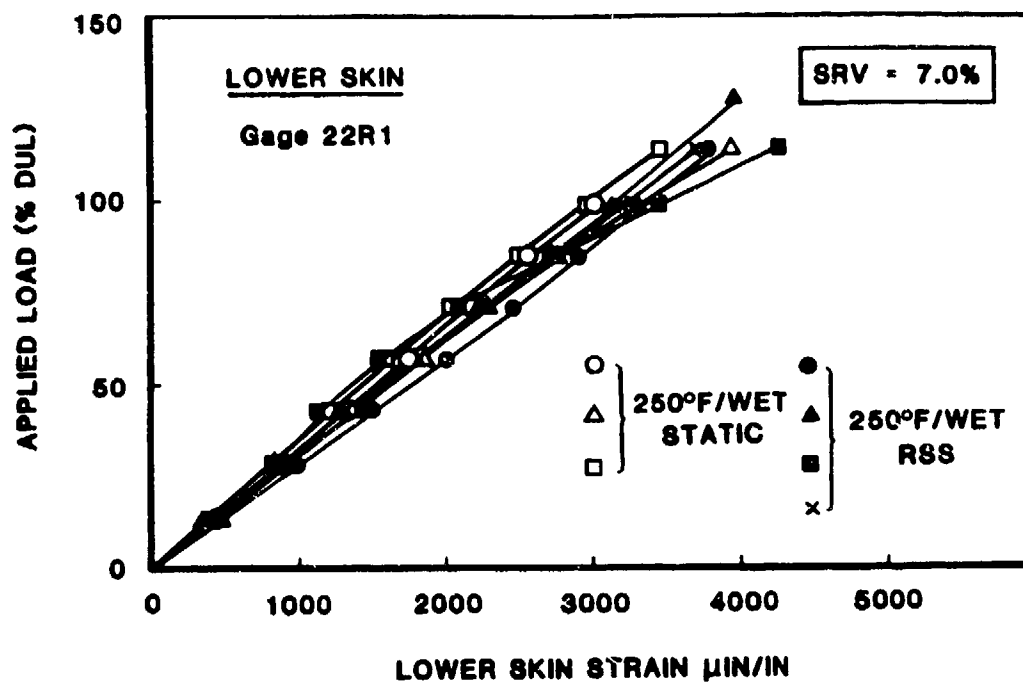
FIGURE 59. WS-1 LOWER SKIN ETW LOAD-STRAIN RESPONSE.
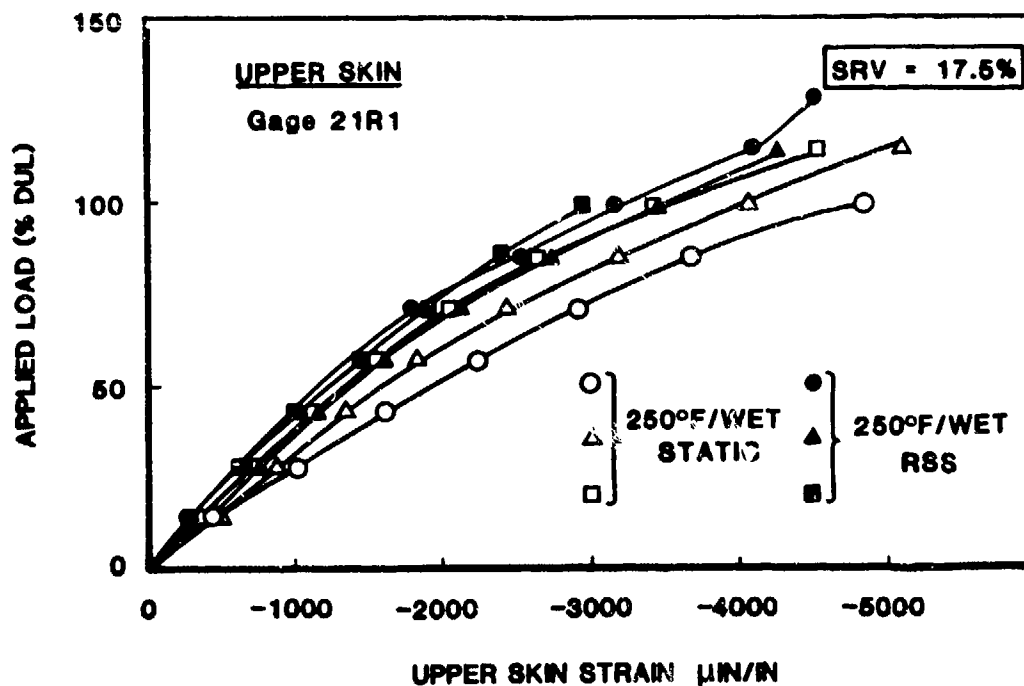


FIGURE 60. WS-1 UPPER SKIN ETW LOAD-STRAIN RESPONSE.
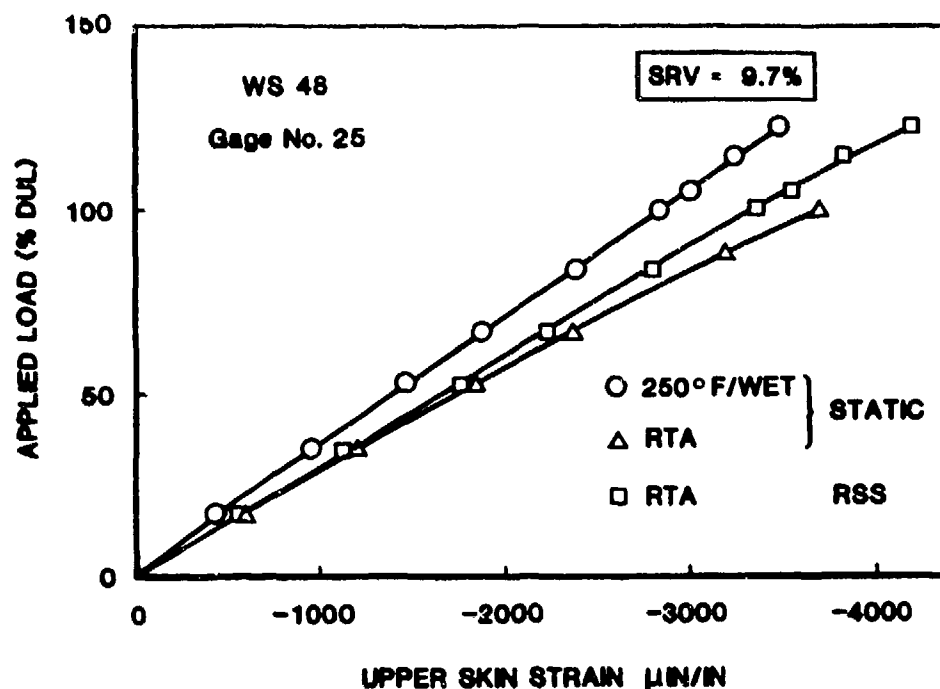
FIGURE 61. WCC-1 UPPER SKIN LOAD-STRAIN RESPONSE AT
WING STATION 48 (GAGE NUMBER 25) FOR LOAD
CASE 130.
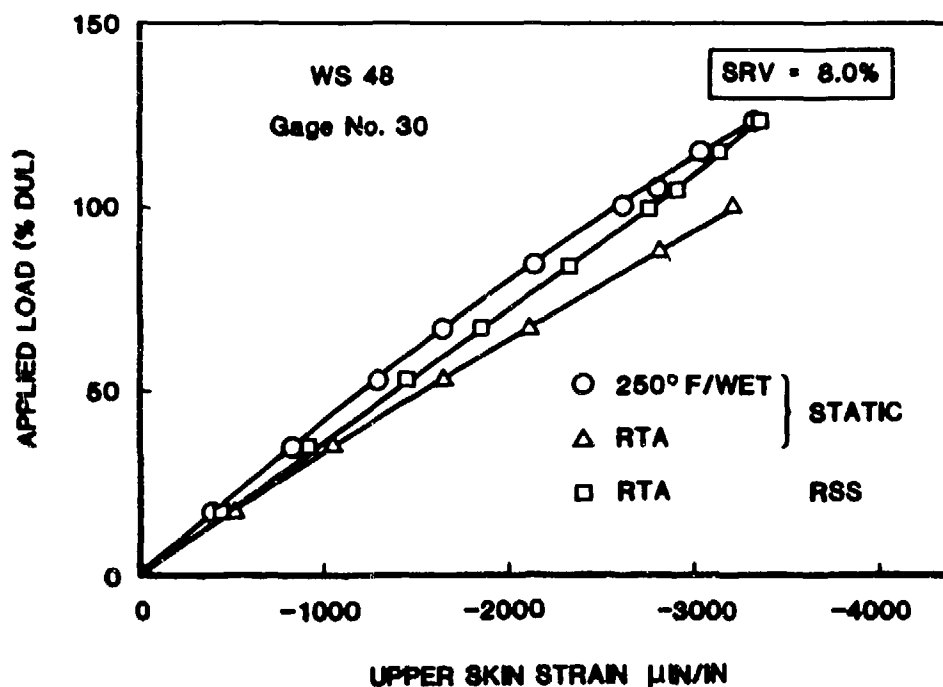


FIGURE 62. WCC-1 UPPER SKIN LOAD-STRAIN RESPONSE AT
WING STATION 48 (GAGE NUMBER 30) FOR LOAD
CASE 130.
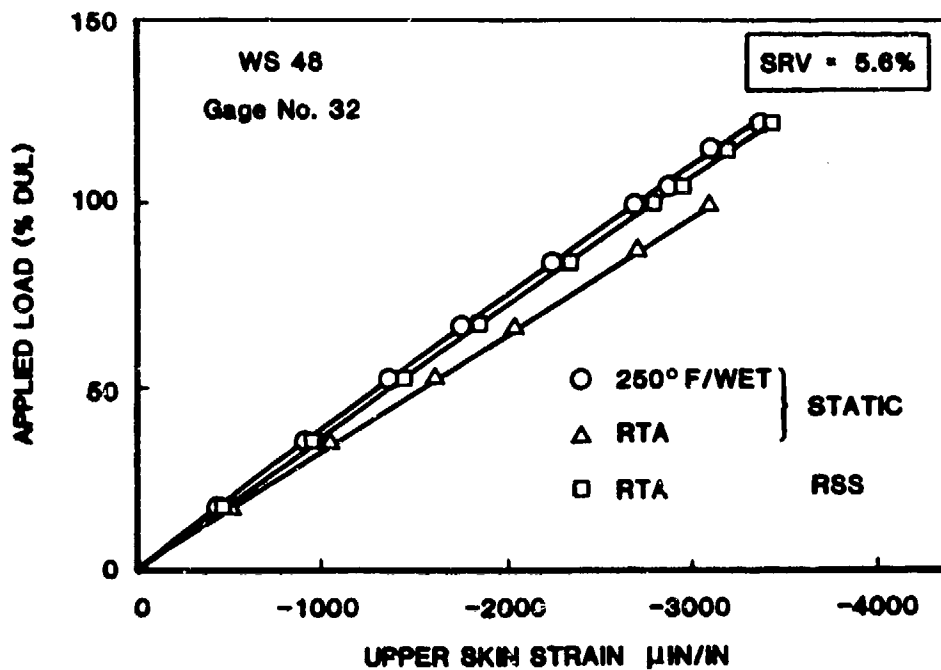
FIGURE 63. WCC-1 UPPER SKIN LOAD-STRAIN RESPONSE AT
WING STATION 48 (GAGE NUMBER 32) FOR LOAD
CASE 130.
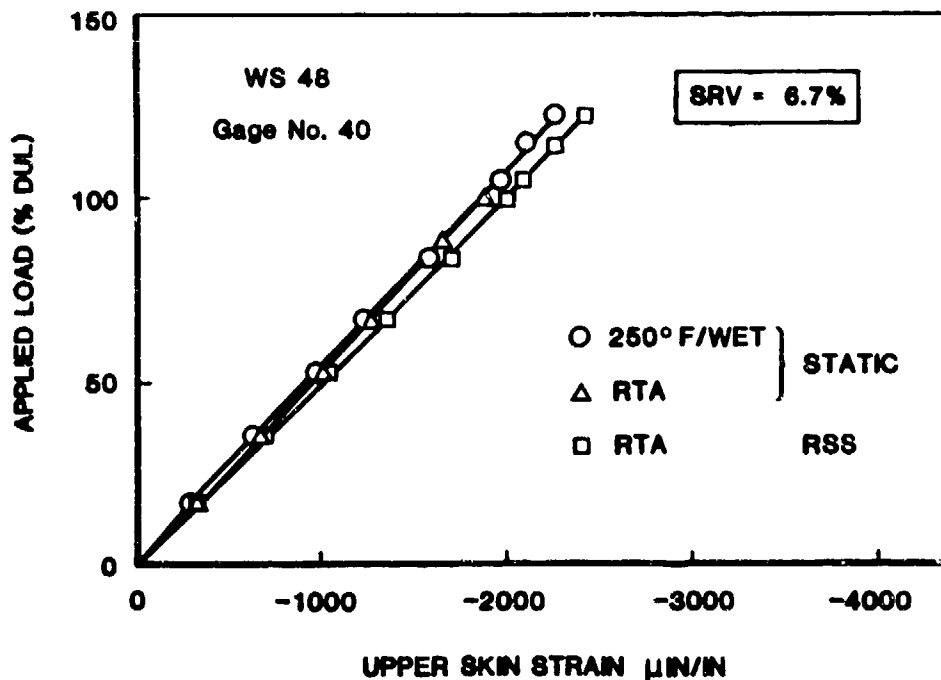


FIGURE 64. WCC-1 UPPER SKIN LOAD-STRAIN RESPONSE AT
WING STATION 48 (GAGE NUMBER 40) FOR LOAD
CASE 130.
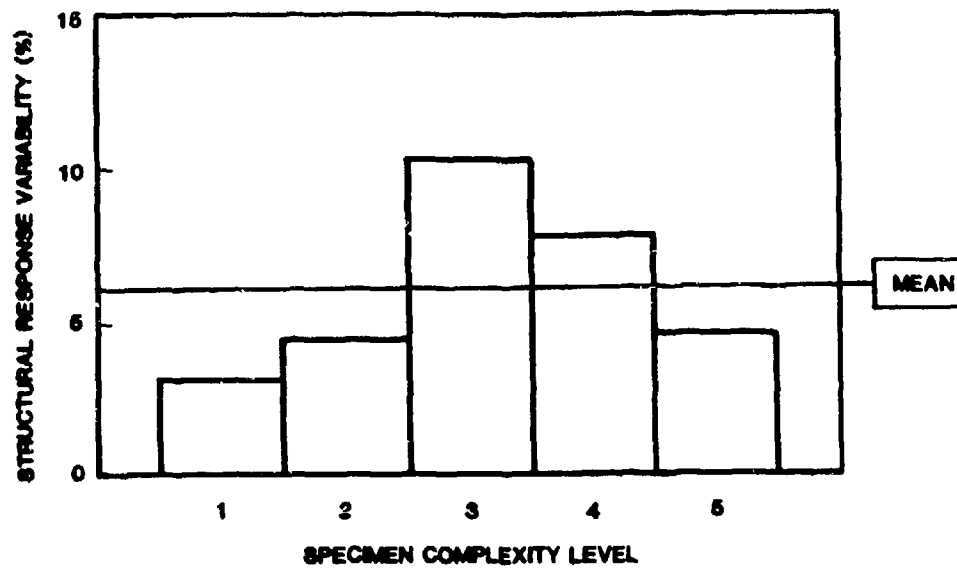
FIGURE 65. INFLUENCE OF SPECIMEN COMPLEXITY LEVEL CN
RTA STRUCTURAL RESPONSE VARIABILITY.
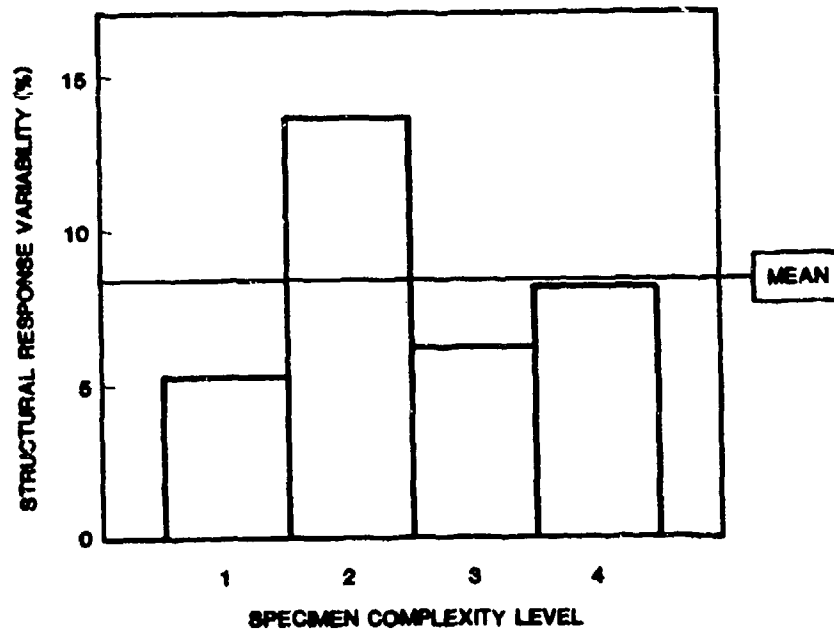


FIGURE 66. INFLUENCE OF SPECIMEN COMPLEXITY LEVEL ON
ETW STRUCTURAL RESPONSE VARIABILITY.

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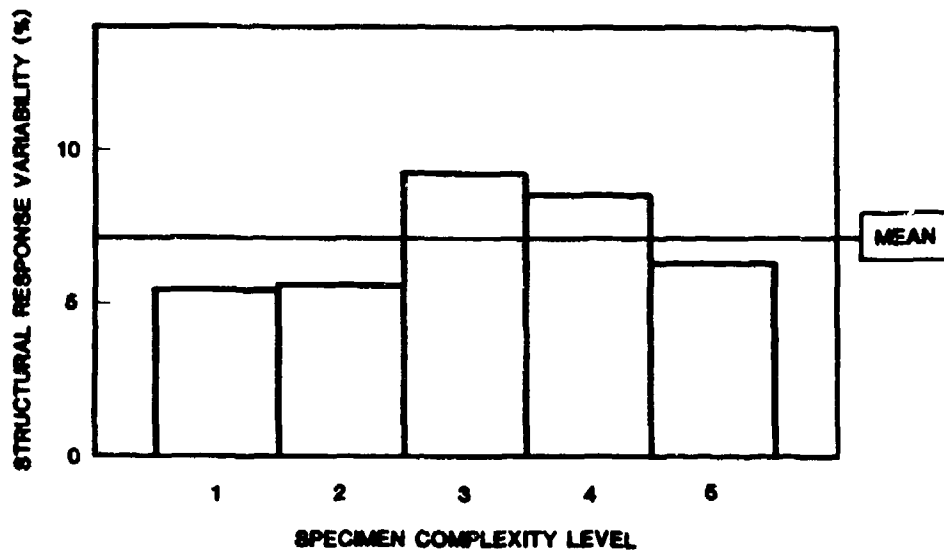
FIGURE 67. INFLUENCE OF SPECIMEN COMPLEXITY LEVEL ON
COMBINED RTA AND ETW STRUCTURAL RESPONSE
VARIABILITY.
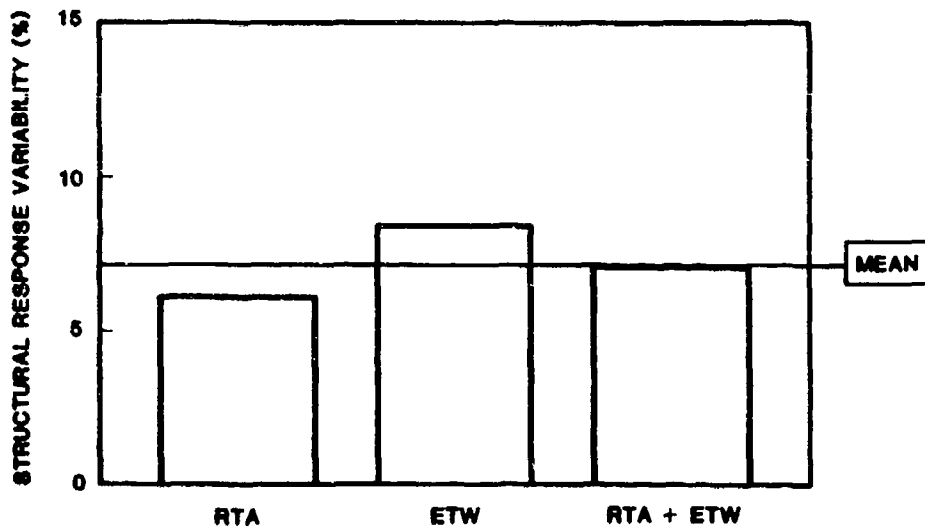


FIGURE 68. INFLUENCE OF ENVIRONMENT ON STRUCTURAL
RESPONSE VARIABILITY.

The observed differences between the individual structural response variability at each complexity level and the over all mean for all complexity levels are not statistically significant. Consequently, it can be concluded that structural response variability is independent of specimen complexity. This is an encouraging observation for the certification of full-scale structure. It is interesting to note that the scatter in static strength data in Reference 5 was also found to be independent of specimen complexity.

Figure 68 shows the influence of environment on structural response variability. It can be seen that structural response variability is also independent of test environment. The structural response variability of wing component sk n strain distributions is shown in Figures 69 through 71. A comparison of structural response variability of the load-strain distributions is shown in Table 7. The results show that the structural response variability in strain distribution is similar to that observed previously for load-strain response.

The WCC-1 structural response variability results presented in Table 6 were determined for static loading c 130. This is the most critical loading case for the wing. Pr to the five static and fatigue tests conducted on the wing component specimens in Reference 5, load-strain surveys were conducted to limit load for five critical loading cases. These are shown in Table 8. All the load-strain surveys were conducted u er RT/ambient conditions. Table 9 summarizes the wing component tests where these RT/Ambient load-strain surveys were conducted prior to test. The influence of loading case on structural response variability are shown in Figures 72 through 76 for a strain gage located on the upper skin close to the root rib.

Table 10 and Figure 77 present a summary of the influence of loading case on structural response variability. The results in Table 10 show that the influence of loading case on structural response variability is similar for all four strain gage locations. In addition, Figure 77 shows that the loading
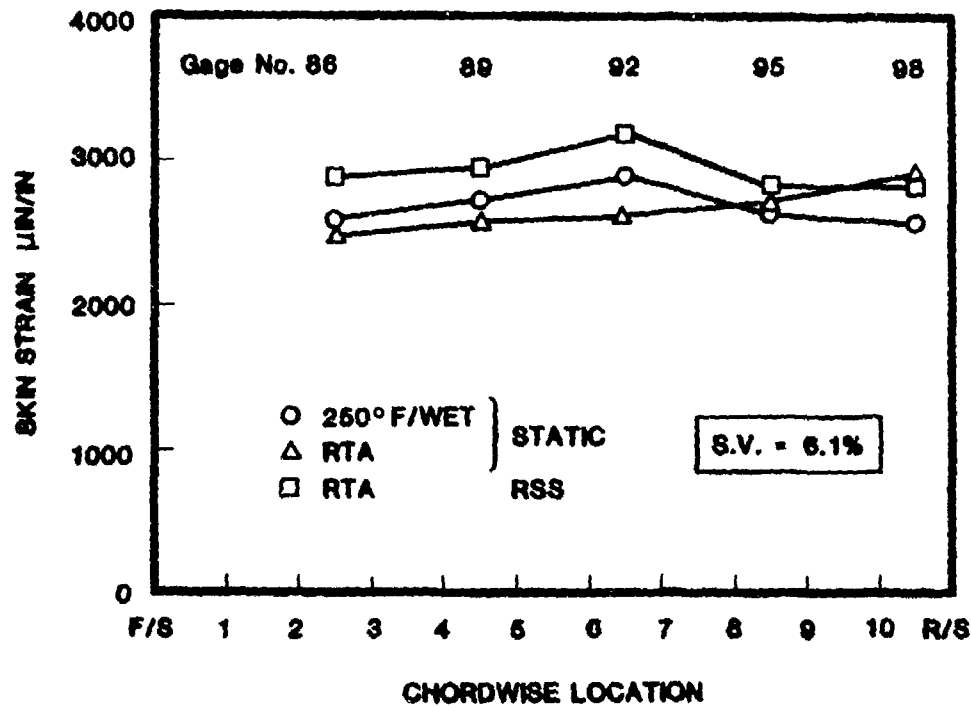
FIGURE 69. WCC-1 DESIGN ULTIMATE LOAD LOWER SKIN STRAIN
DISTRIBUTION AT WING STATION 43 FOR LOAD CASE 130.
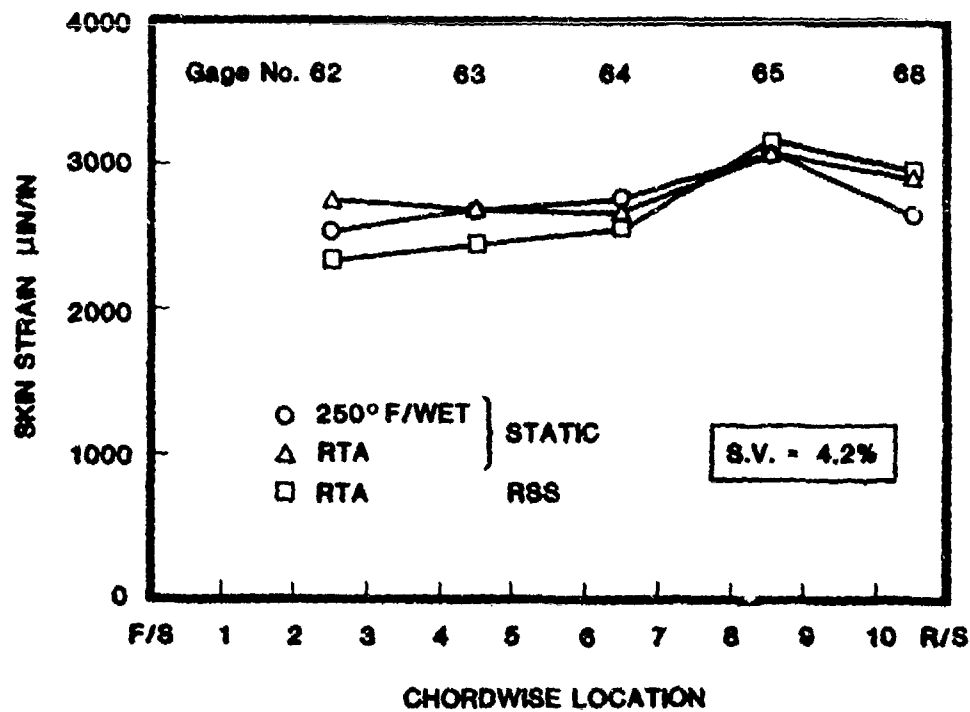


FIGURE 70. WCC-1 DESIGN ULTIMATE LOAD LOWER SKIN STRAIN
DISTRIBUTION AT WING STATION 85 FOR LOAD CASE 130.
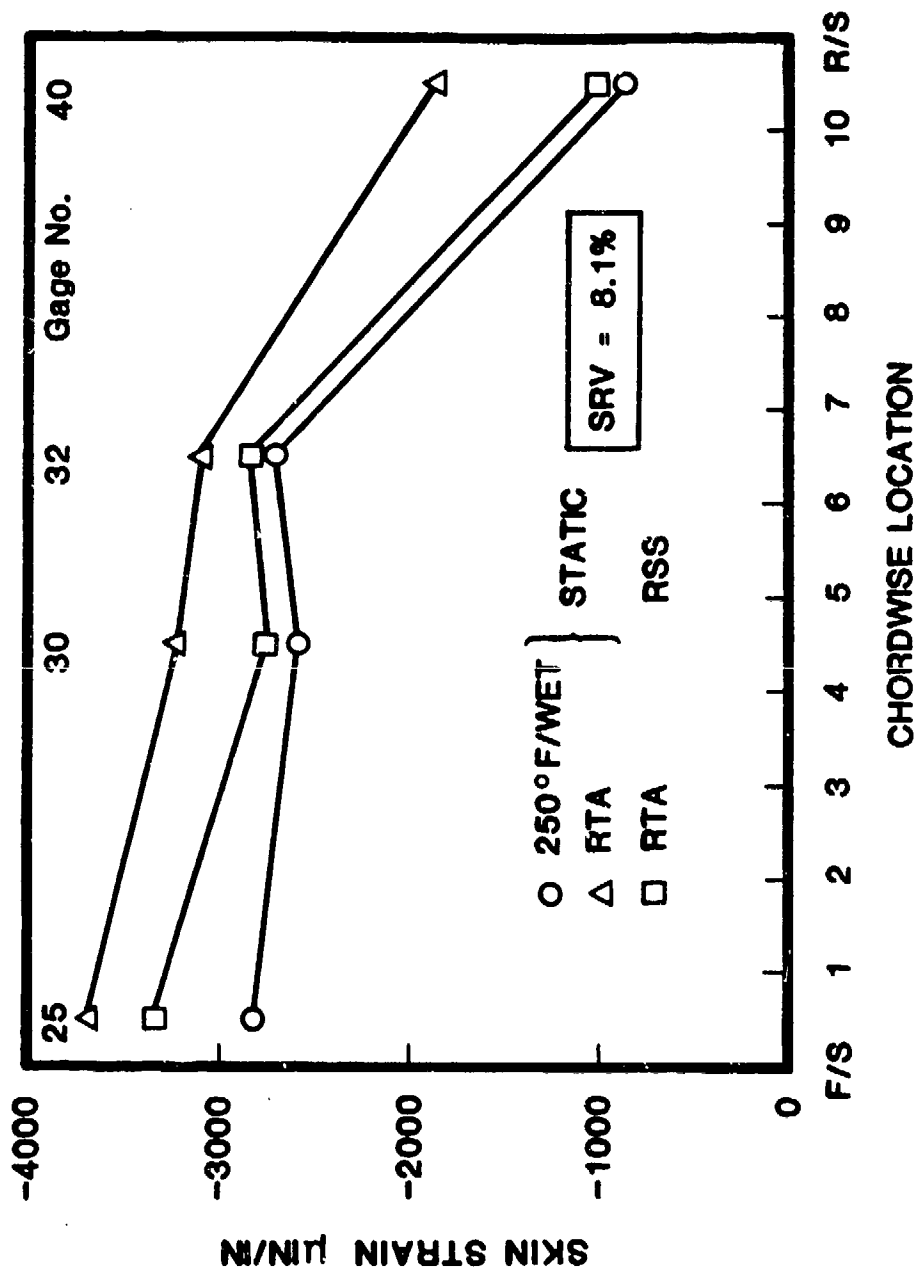
FIGURE 71. WCC-1 DESIGN ULTIMATE LOAD UPPER SKIN STRAIN DISTRIBUTION AT WING STATION 48 FOR LOAD CASE 130.

**TABLE 7. COMPARISON OF WCC-1 LOAD-STRAIN AND STRAIN DISTRIBUTION STRUCTURAL RESPONSE VARIABILITY.**

| LOCATION | STRUCTURAL RESPONSE VARIABILITY | |
|---|---|---|
| | LOAD-STRAIN AVERAGE | STRAIN DISTRIBUTION |
| LOWER SKIN WS 48 | 5.8 | 6.1 |
| LOWER SKIN WS 85 | 4.7 | 4.2 |
| UPPER SKIN WS 48 | 7.5 | 8.1 |

TABLE 8. SUMMARY OF WCC-1 ULTIMATE STATIC LOAD CASES.

| LOAD CASE | DESCRIPTION | TEMPERATURE °F | $S_z$ LB | $M_x \times 10^{-6}$ IN-LB | $M_y \times 10^{-6}$ IN-LB |
|---|---|---|---|---|---|
| 100 | SYMMETRICAL PULL-UP | 145 | 173,000 | 11.7 | -8.4 |
| 110 | SYMMETRICAL PULL-UP | 192 | 166,500 | 11.5 | -7.7 |
| 130 | SYMMETRICAL PULL-UP | 242 | 140,000 | 10.3 | -8.0 |
| 201 | ROLLING PULL-UP | 192 | 146,000 | 10.3 | -7.4 |
| 102 | SYMMETRICAL PUSHOVER | 145 | -57,800 | -4.3 | +1.8 |

NOTES:

(1) $S_z$ is the shear from the left wing tip to the root rib.

(2) $M_x$ is the left-hand wing moment at the root rib.

(3) $M_y$ is the wing torsion.

## TABLE 9. WING COMPONENT (WCC-1) TESTS IN REFERENCE 5.

| TEST SERIES | TEST TYPE |
|:---:|:---|
| 1 | 250°F/WET STATIC |
| 2 | RT/AMBIENT STATIC |
| 4 | RT/AMBIENT FATIGUE (2 LIFETIMES) |
| 6 | BASELINE ENVIRONMENTAL FATIGUE (2 LIFETIMES) |
| 9 | REAL TIME FATIGUE (1 LIFETIME) |

## TABLE 10. INFLUENCE OF LOADING CASE ON LOAD-STRAIN STRUCTURAL RESPONSE VARIABILITY.

| GAGE No. | LOCATION | STRUCTURAL VARIABILITY IN LOAD-STRAIN RESPONSE (C.V.%) | | | | | |
|:---:|:---:|:---:|:---:|:---:|:---:|:---:|:---:|
| | | CASE 100 | CASE 110 | CASE 130 | CASE 102 | CASE 201 | MEAN |
| 65 | LOWER SKIN | 6.1 | 5.4 | 4.3 | 14.5 | 5.4 | 7.1 |
| 86 | LOWER SKIN | 7.3 | 6.8 | 6.8 | 7.6 | 7.9 | 7.3 |
| 92 | LOWER SKIN | 6.5 | 6.8 | 7.2 | 7.0 | 6.9 | 6.9 |
| 25 | UPPER SKIN | 6.4 | 6.2 | 6.3 | 6.4 | 7.1 | 6.5 |
| MEAN | | 6.6 | 6.3 | 6.2 | 8.9 | 6.8 | 7.0 |
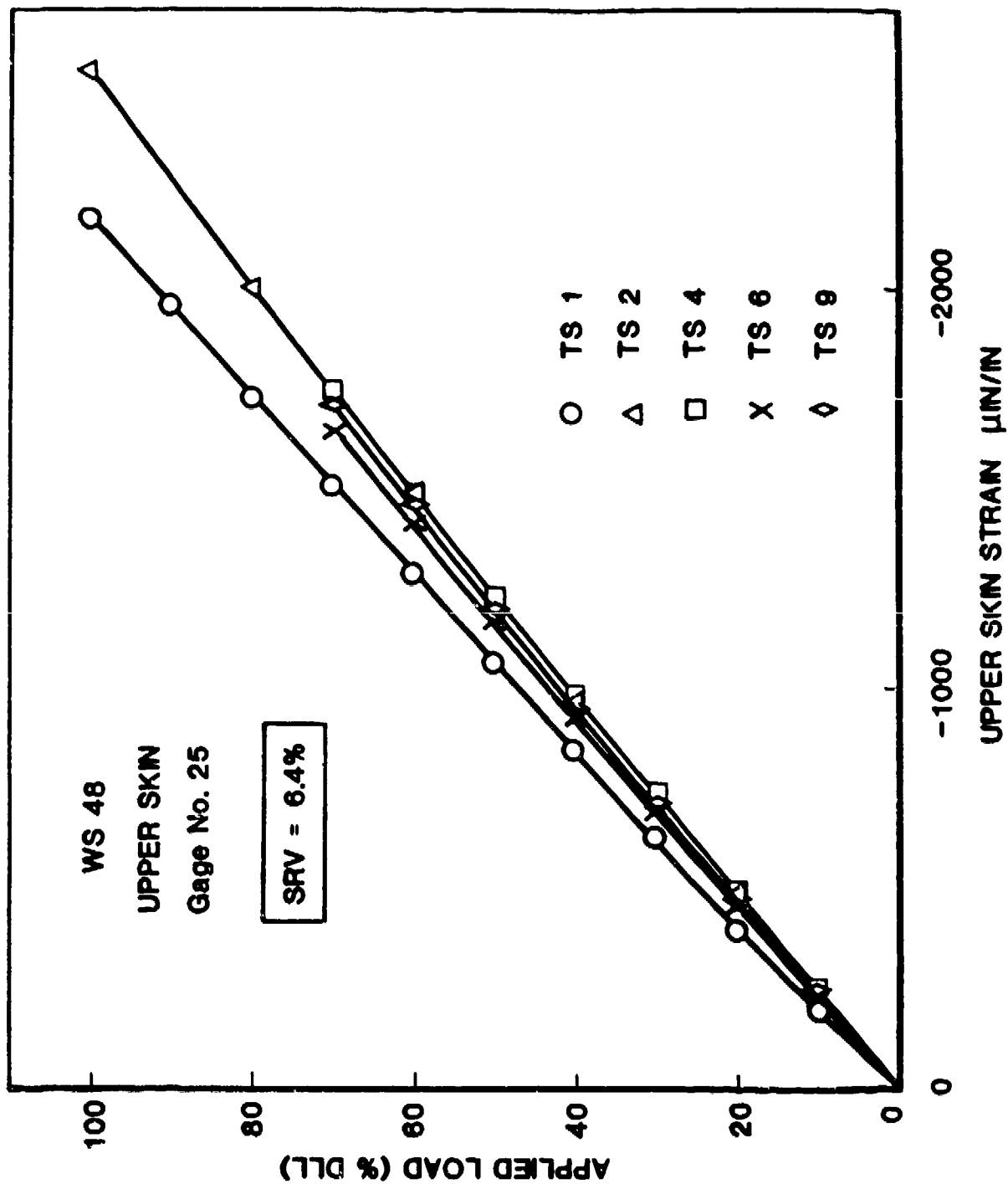
FIGURE 72. WCC-1 LOAD-STRAIN RESPONSE FOR LOAD CASE 100.
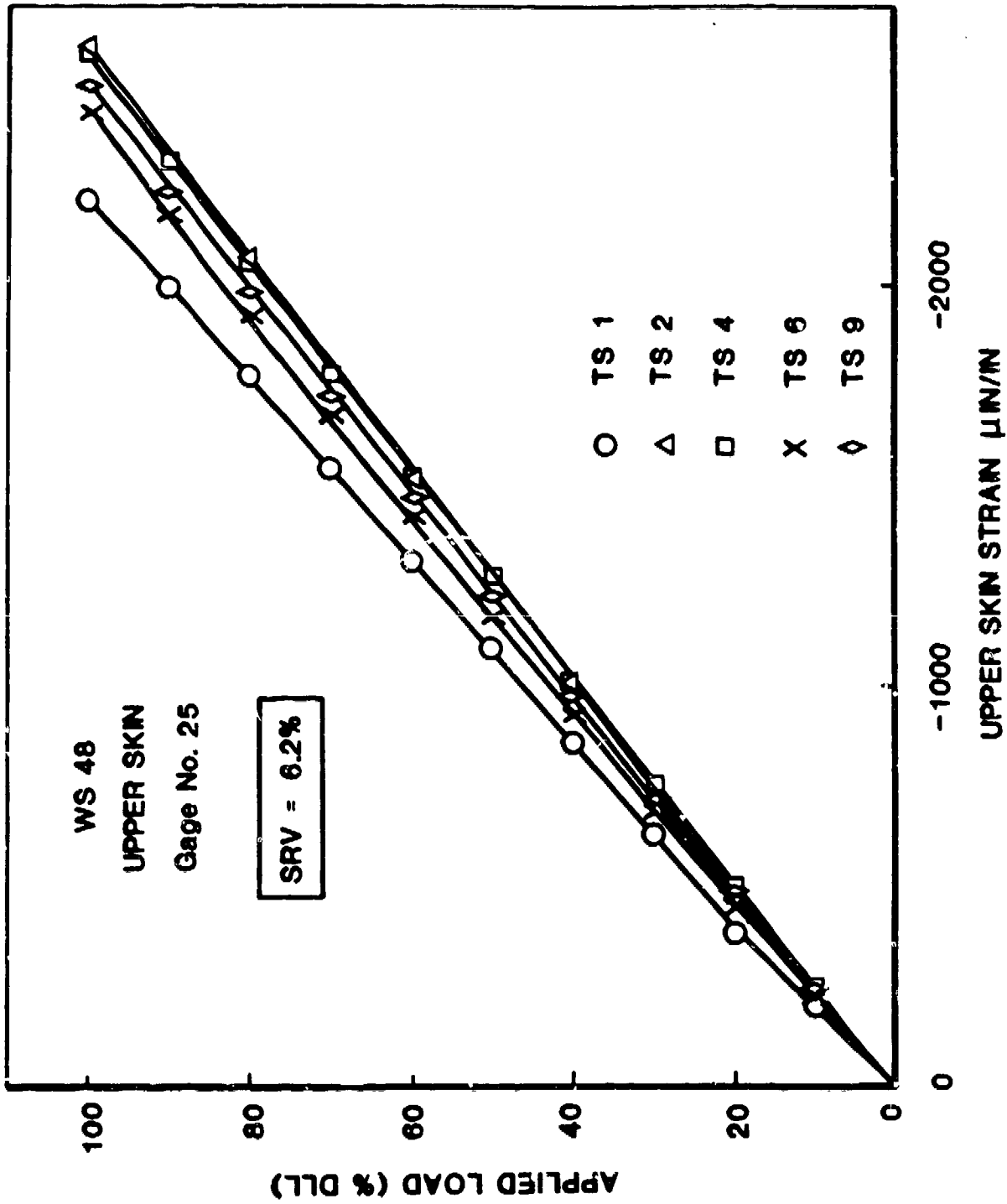
FIGURE 73. WCC-1 LOAD-STRAIN RESPONSE FOR LOAD CASE 110.
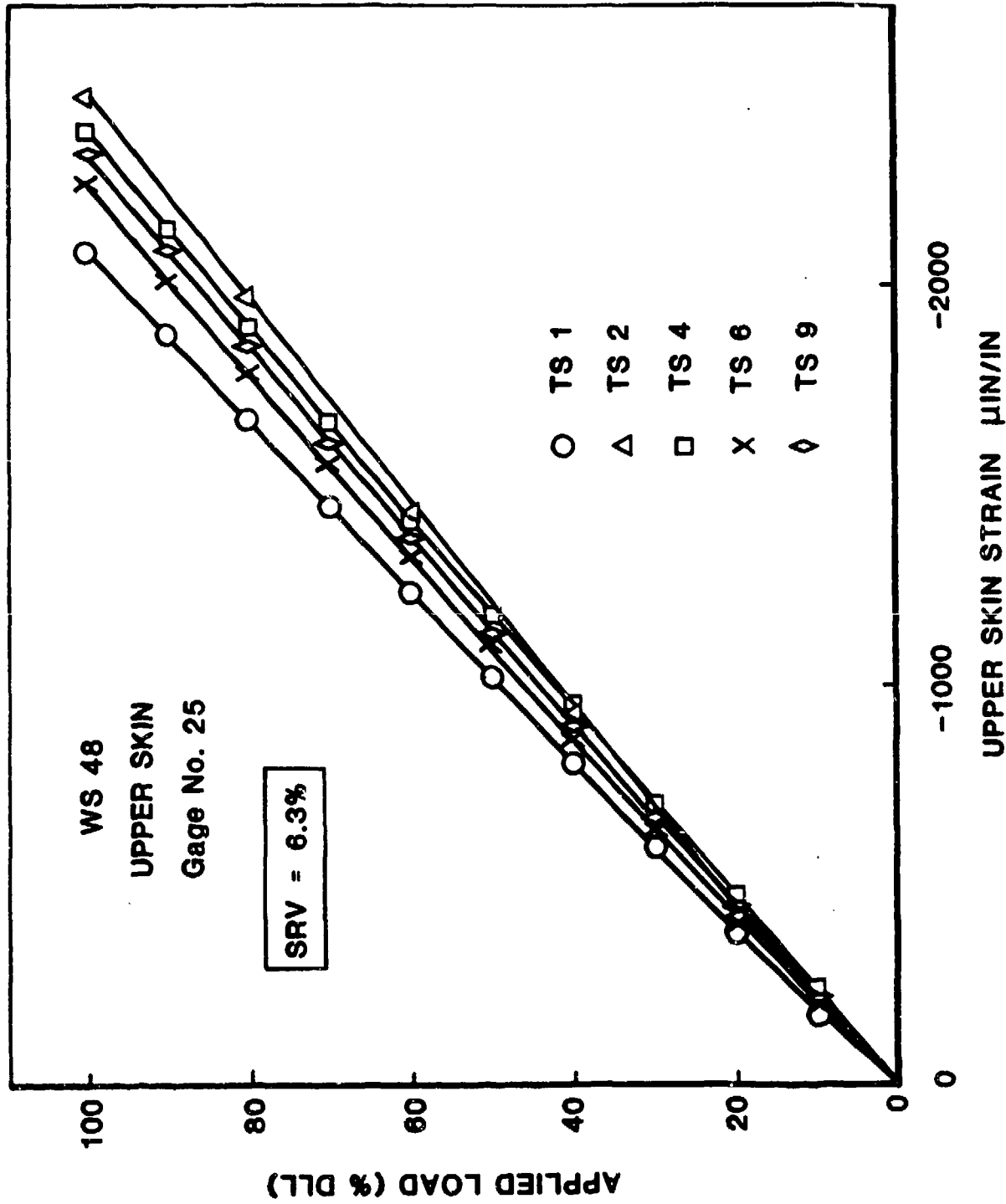
FIGURE 74. WCC-1 LOAD-STRAIN RESPONSE FOR LOAD CASE 130.
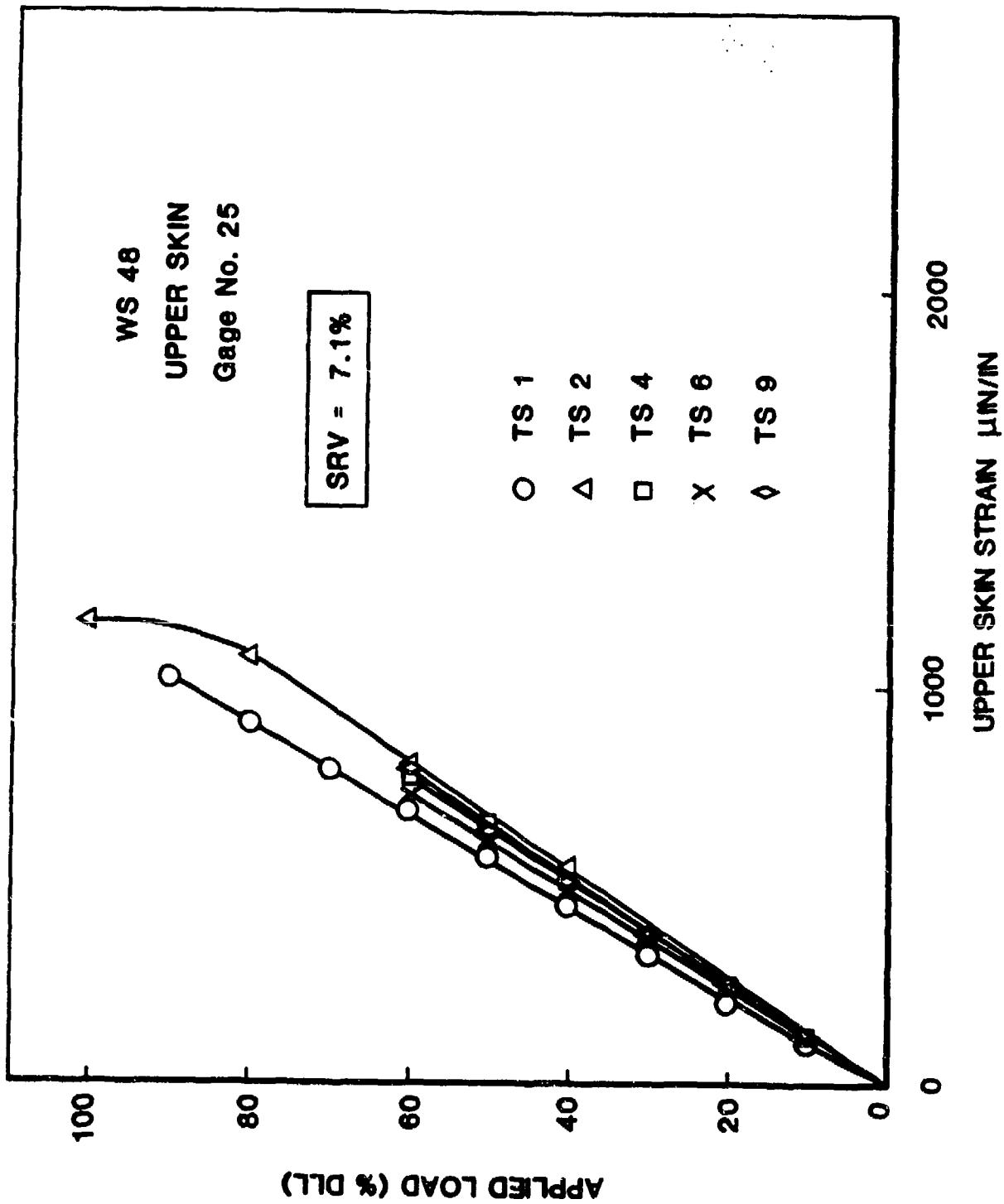
FIGURE 75. WCC–1 LOAD–STRAIN RESPONSE FOR LOAD CASE 102.
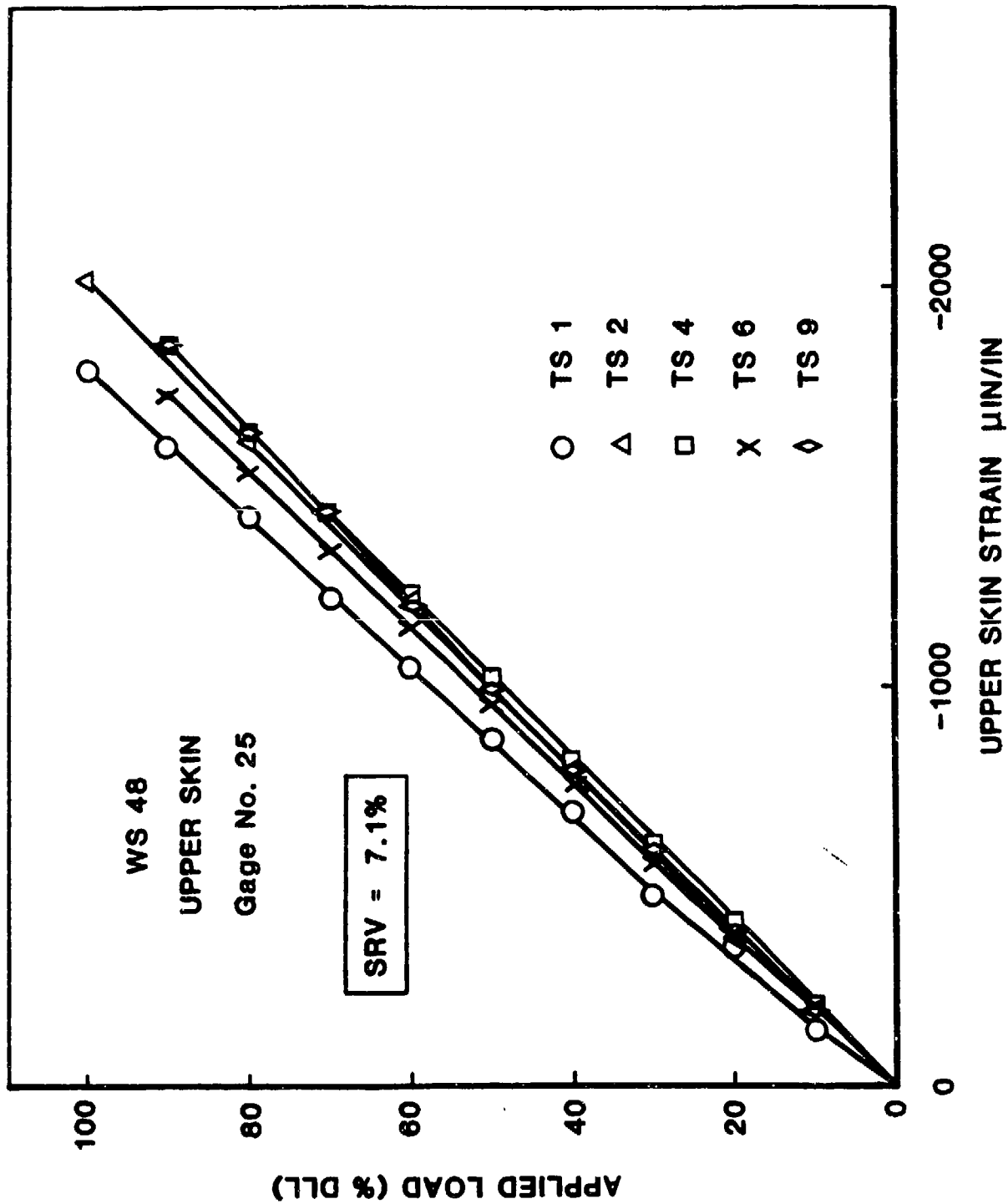
FIGURE 76. WCC-1 LOAD-STRAIN RESPONSE FOR LOAD CASE 201.
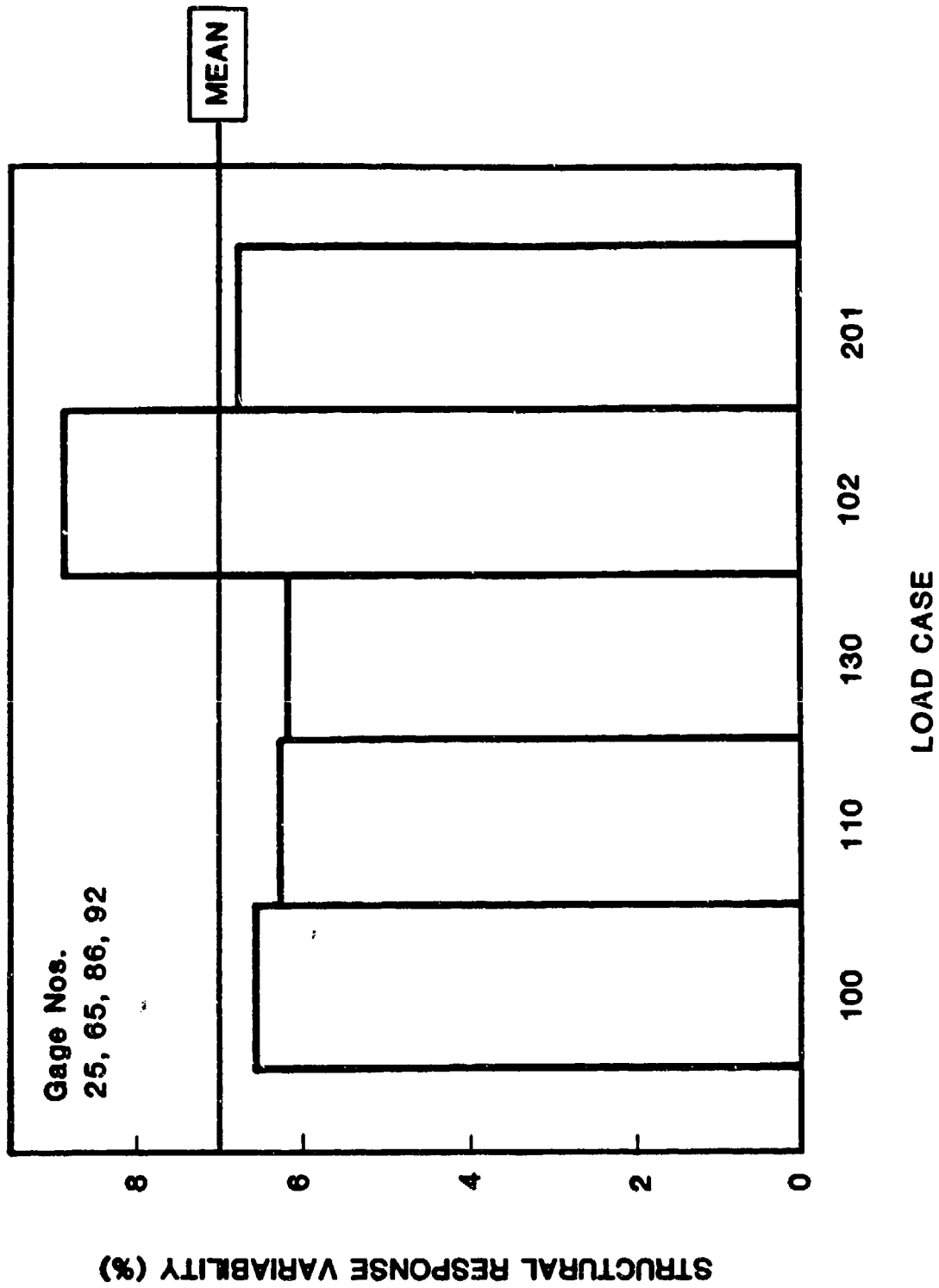
FIGURE 77. INFLUENCE OF LOADING CASE ON STRUCTURAL RESPONSE VARIABILITY.

case has little influence on structural response variability. Some increased variability is observed for the pushover (down-bending) load case 102. This is probably due to the low values of strain recorded at limit load. However, none of the five loading cases exhibited values of structural response variability that were significantly different from the overall mean value of 7.0 percent. This overall value is very similar to the overall mean value of /.1 percent determined for all specimens in Table 6.

An interesting observation can be made from the load-strain data in Figures 72 through 76. The scatter trend in the data is the same for all loading cases. That is, Test Series 1 gives the lowest strain at design limit load, while Test Series 2 and 4 give the highest strains at design limit load. This observation was true for all four strain gage locations that are analyzed.

Since the structural response variability was found to be independent of structural complexity, environment and loading case, all data in Table 6 are pooled. The pooled data are analyzed using a normal distribution and the resultant scatter distribution is shown in Figure 78. The following values of structural response variability are determined:

$$(SRV)_{MEAN} = 7.1\% \quad \text{or} \quad (\alpha_{SRV})_{MEAN} = 17.36$$

$$(SRV)_{MODAL} = 6.5\% \quad \text{or} \quad (\alpha_{SRV})_{MODAL} = 19.02$$

$$(SRV)_{B-BASIS} = 9.5\% \quad \text{or} \quad (\alpha_{SRV})_{B-BASIS} = 12.81$$

Following the philosophy adopted for static strength and fatigue life scatter values, the modal structural response variability of 6.5 percent will be used for incorporation in the certification testing methodology.

## Summary

A comprehensive analysis of a large load-strain data base has been conducted. A total of 360 separate load-strain plots containing in excess of 3000 individual data points are
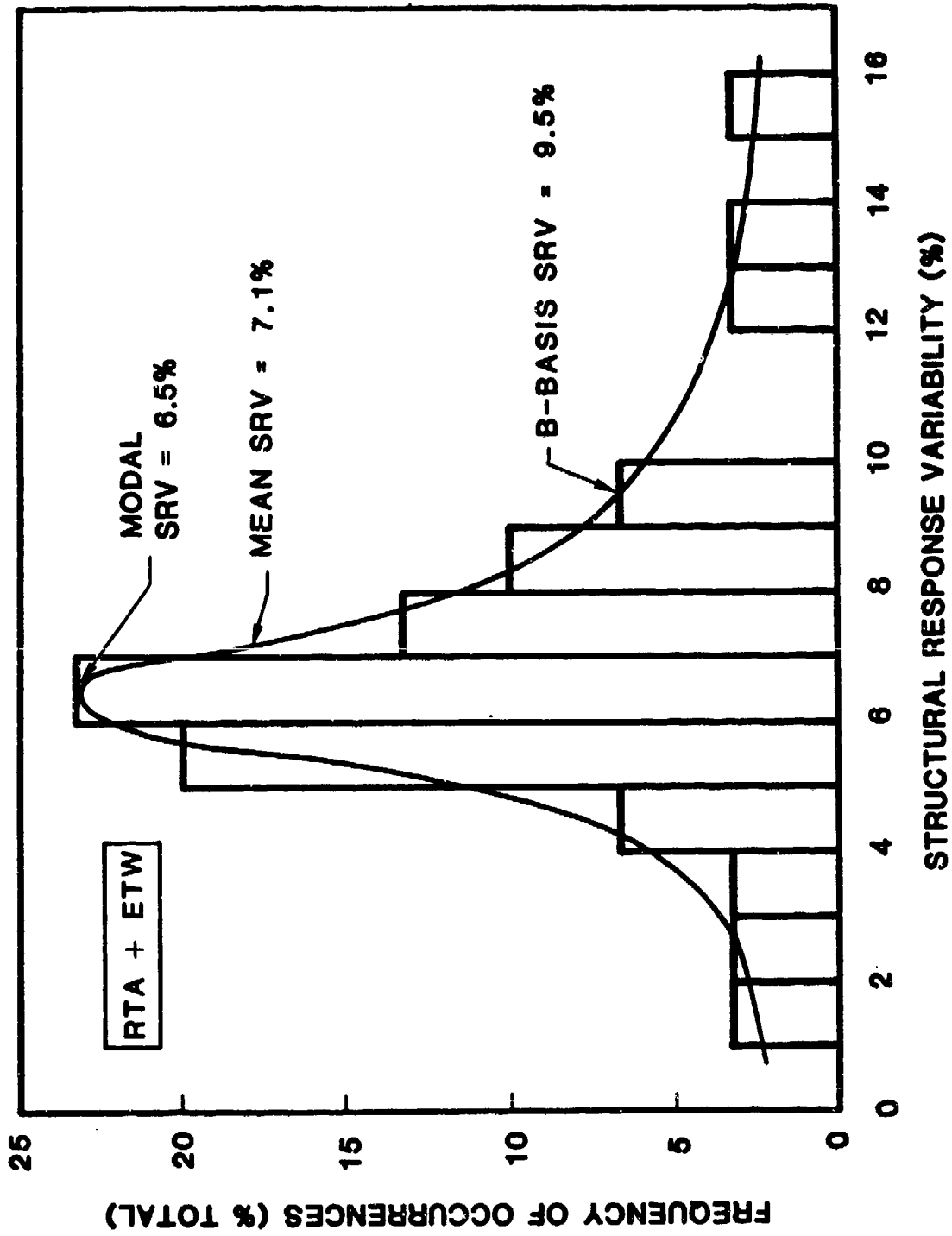
FIGURE 78. DISTRIBUTION OF STRUCTURAL RESPONSE VARIABILITY FOR COMBINED
RTD + ETW LOAD–STRAIN DATA.

analyzed. A novel analysis technique permitted the determination of structural response variability from the load-strain data for the first time. The following conclusions can be made from the study.

- A statistical analysis methodology is developed to determine structural response variability from load-strain data.

- Structural response variability is shown to be independent of specimen complexity, test environment and static loading cases.

- The structural response variability of load-strain and strain distributions is approximately equal.

- The mean and modal values of structural response variability are 7.1 and 6.5 percent, respectively.

3.3     Usage Variation

The differences between the fatigue response of metallic and composite materials were discussed in Section 2. It was concluded that current graphite-epoxy systems have a significantly superior fatigue response, as shown in Figure 43. This causes a different response to usage changes for the two materials. Consider, for example, a usage change, which is more severe than the design spectrum, and leads to an effective increase in operating stress level relative to the design stress level. Figure 79 summarizes the change in fatigue life as a function of this type of increased usage severity ($\sigma_u/\sigma_D$). It can be seen that increased usage severity causes a gradual reduction in metal design fatigue life. In contrast, increased usage severity initially has no effect on composite design fatigue life. Further increase in usage severity eventually causes very rapid reductions in composite design fatigue life. This behavior is caused by the flat S-N behavior of composites. The design stress level is significantly below the composite fatigue limit (defined as a very long life) such that increased usage severity does not initially change the design life. When severe usage changes

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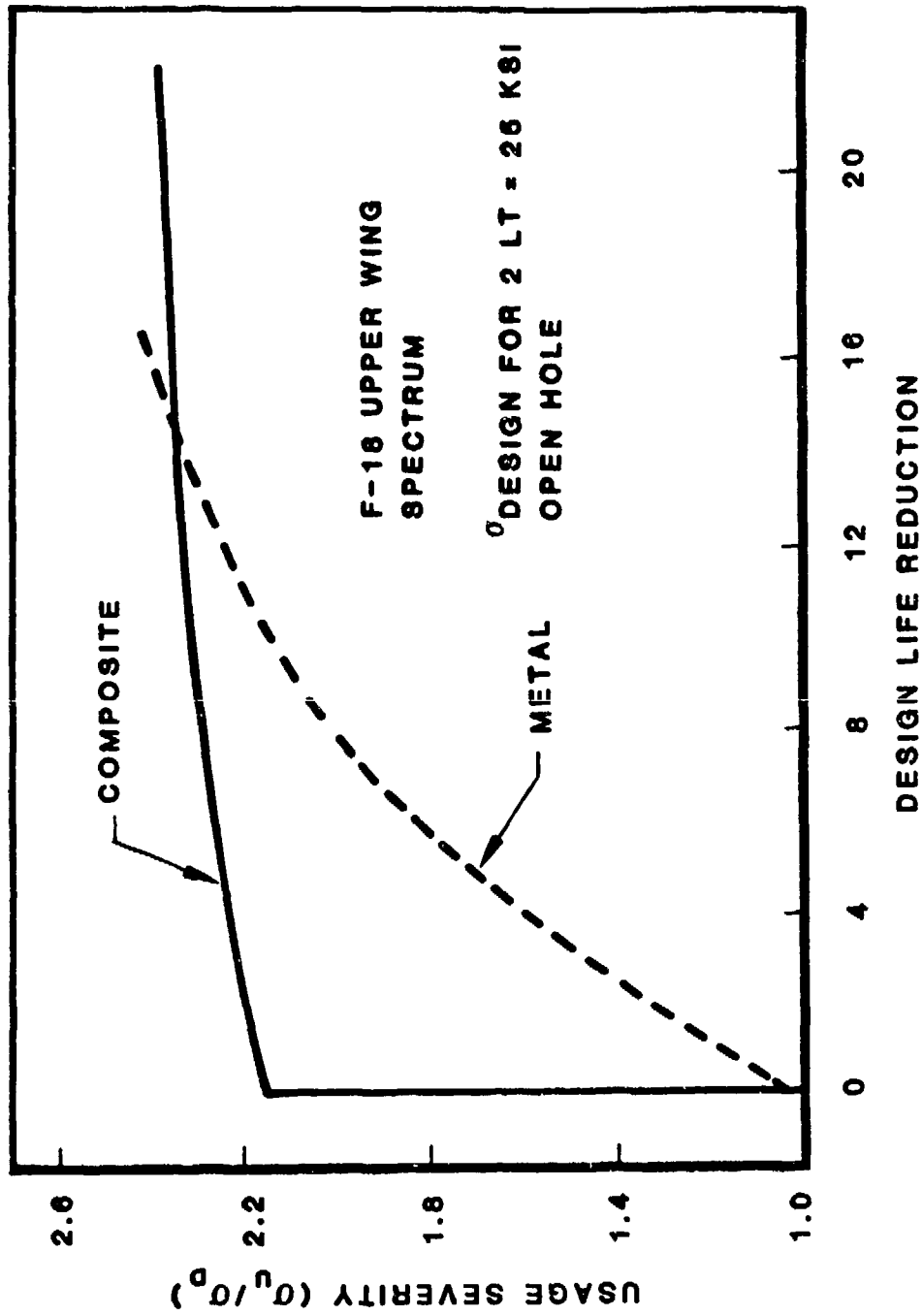
FIGURE 79. COMPARISON OF USAGE SEVERITY CHANGE FOR COMPOSITES AND METALS.

cause the operating stress level to intersect the composite S-N curve rapid life changes occur.

The maximum spectrum load in the actual service environment often deviates from the design load. The maximum load in a fatigue spectrum is a random variable in nature, and its probalistic distribution can be described by a distribution function. This distribution function in onjunction with the fatigue life distribution and the fatigue wear-out law can be used to assess the influence of usage change on structural reliability. One approach is to assume the forms of the distribution function for the service load variation and the fatigue life distribution, together with an assumed wear-out equation to form a joint probability function. This is illustrated in Figure 80. These functions are selected and given below:

(1) The composite fatigue life scatter is described by a Weibull distribution with probability of survival given by

$$P_L = \exp\left[-\left(\frac{x}{\beta_L}\right)^{\alpha_L}\right] \tag{27}$$

(2) The stress-life relationship can be described by the wear-out equation from Sendeckyj analysis

$$\left(\frac{\sigma_u}{\sigma_a}\right)^{1/S} + C - 1 = CN \tag{28}$$

where   $\sigma_u$   is the static strength

   $\sigma_a$   is the maximum applied stress

   C, S   are wear-out parameters

   N   is the fatigue life at stress level $\sigma_a$

(3) Two types of service load variation distribution, $f(\sigma)$, are assumed

   a. Normal distribution with mean $\bar{x} = \sigma_a$ and standard deviations

   b. Weibull distribution with shape and scale parameters $\alpha_u$ and $\beta_u$ so that $\bar{x} = \sigma_a$
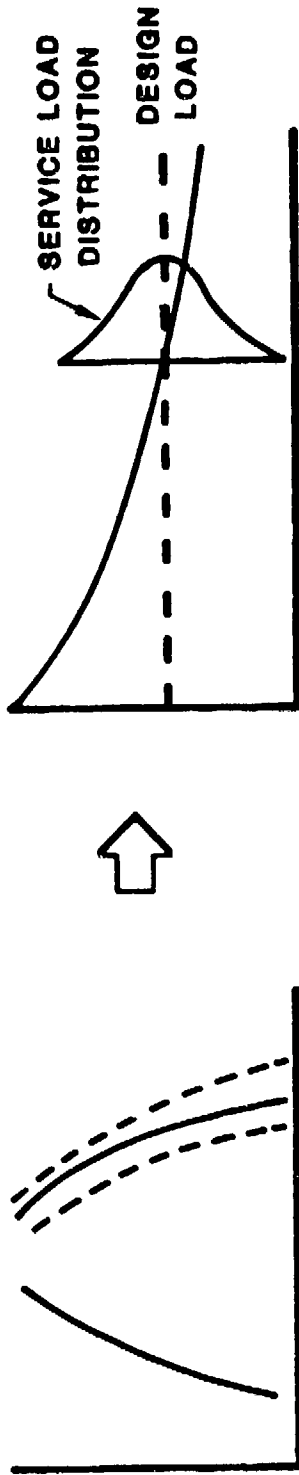
FIGURE 80.  METHODOLOGY FOR INCORPORATING USAGE VARIATION INTO FATIGUE LIFE RELIABILITY CALCULATIONS.

or

$$\text{or } \beta_u = \frac{\sigma_a}{\Gamma\left(\dfrac{\alpha_u + 1}{\alpha_u}\right)} \tag{29}$$

The joint probability of survival for the structure is then given by

$$P_s = \int_0^\infty P_0(\sigma) f(\sigma) d\sigma \tag{30}$$

where $f(\sigma)$ is the probability density function of the service load distribution. The function $p_L(\sigma)$ is the probability of survival at stress level $\sigma$. Numerical integrations are conducted to evaluate the influence of usage change on fatigue reliability.

The parameters used in the numerical evaluation are given below:

1.  Fatigue life variability, $\alpha_L = 1.25$

2.  Design applied stress level, $\sigma_a = 0.8 \ \sigma_u$

3.  Wear-out equation parameters, $S = 0.0625$ and $C = 0.03544$

4.  Usage change variability, $\alpha_u = 10$ to $26$

The influence of the usage variation on the structural life reliability is shown in Figure 81. The 95 percent confidence reliability with no usage variation is assumed to be 0.90 (B-basis). This figure shows that the fatigue reliability is significantly reduced due to usage variation. For a sample size of $n = 1$, the reliability reduces to 0.68 for $\alpha_u$ of 16. The results using a normal distribution to describe the usage variation are also shown in the figure for $n = 20$. As can be seen from the figure, the selection of distribution function has a very small effect on the resulting reliability. The 95 percent confidence reliability plotted against the coefficient of variation of usage change is shown in Figure 82.

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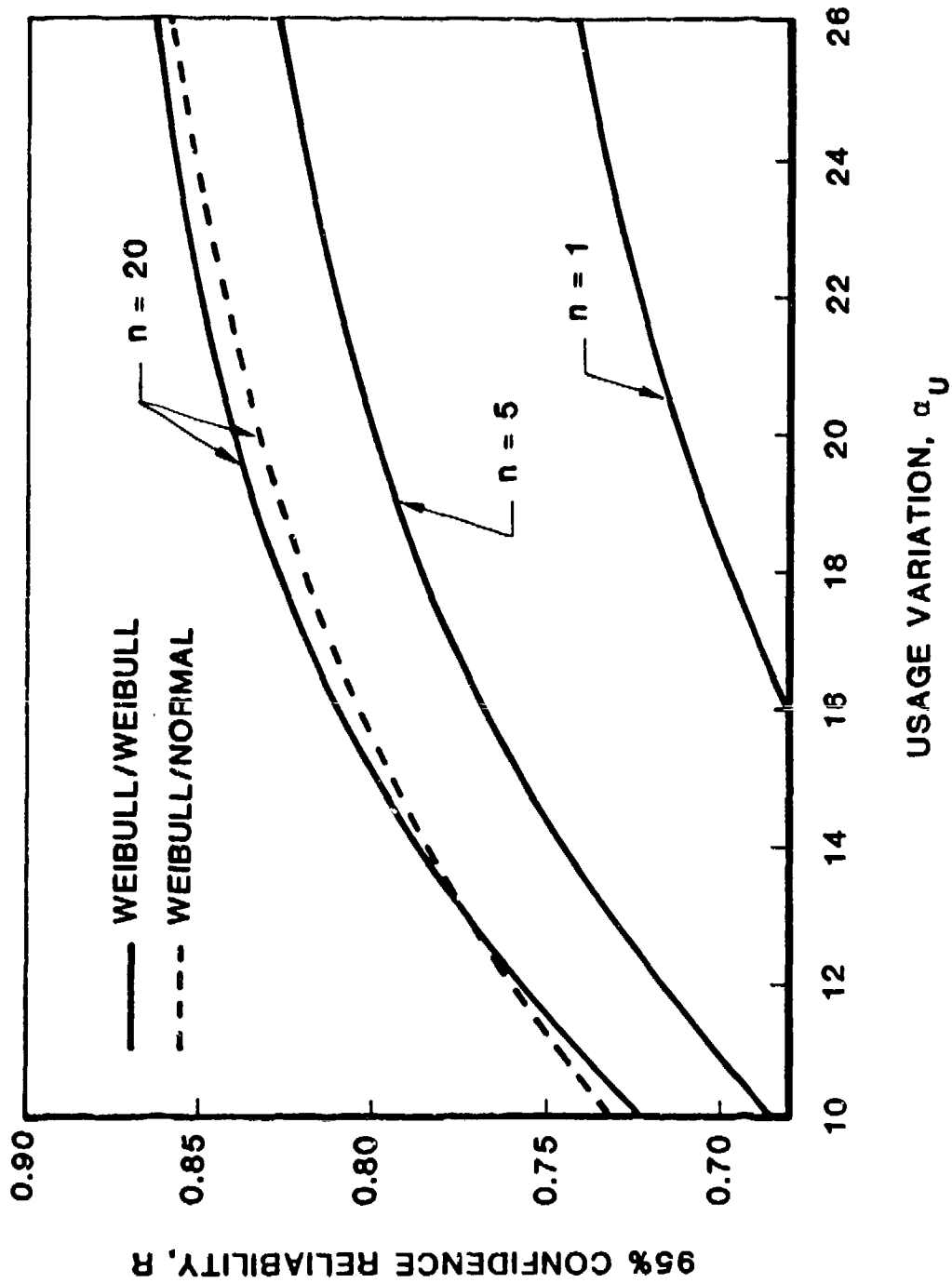
FIGURE 81. INFLUENCE OF USAGE VARIATION ON FATIGUE RELIABILITY ($\sigma_a / \sigma_u = 0.8$).
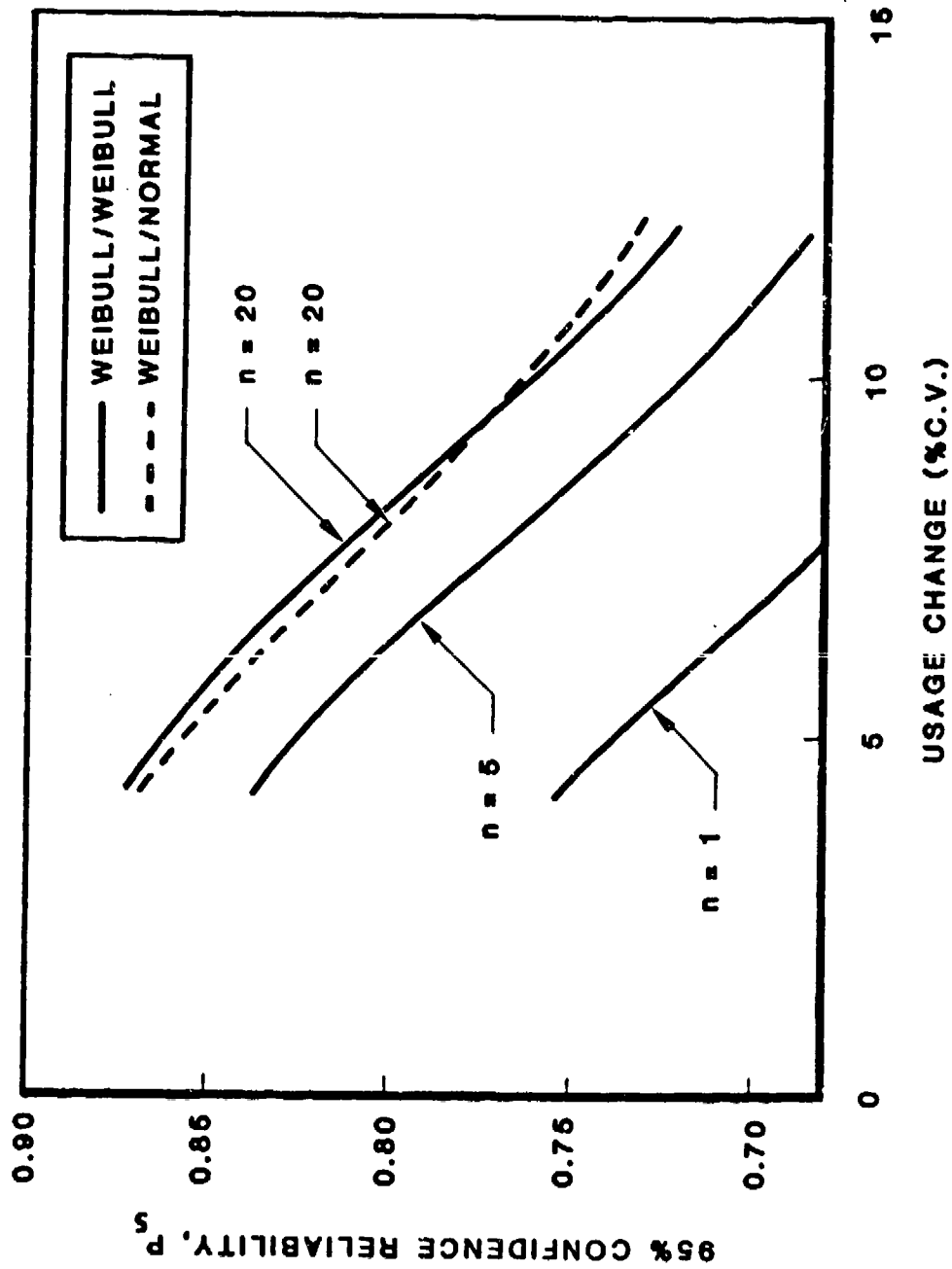
FIGURE 82. CHANGE OF RELIABILITY WITH USAGE VARIATION.

The influence of usage variation on the B-basis fatigue life is shown in Figure 83. The figure shows that the B-basis fatigue life is significantly reduced by usage variation. For a sample size of $n = 1$, the B-basis life is computed to be 133 hours when usage variation is not considered. The B-basis life becomes 54 hours for $\alpha_u = 26$ and is reduced to 17 hours when $\alpha_u = 10$. The mean life to B-basis life ratio (life factor) is shown in Figure 84. This figure shows that the life factor is increased significantly as the usage variation increases (decreasing $\alpha u$). The life factor for $n = 1$ with no usage variation is 13.55. It becomes 18 when $\alpha_u = 26$ and increased to 38.3 as $\alpha_u = 12$.

It should be noted that the methodology developed for usage variation effects can also be used to determine the effects of structural response variability on the fatigue life reliability. In this case, the probability density function $f(\sigma)$ in Equation 27 is replaced by the density function for SRV and the analysis procedure is identical. This analysis technique is used for methodology demonstration discussed in Section 4.
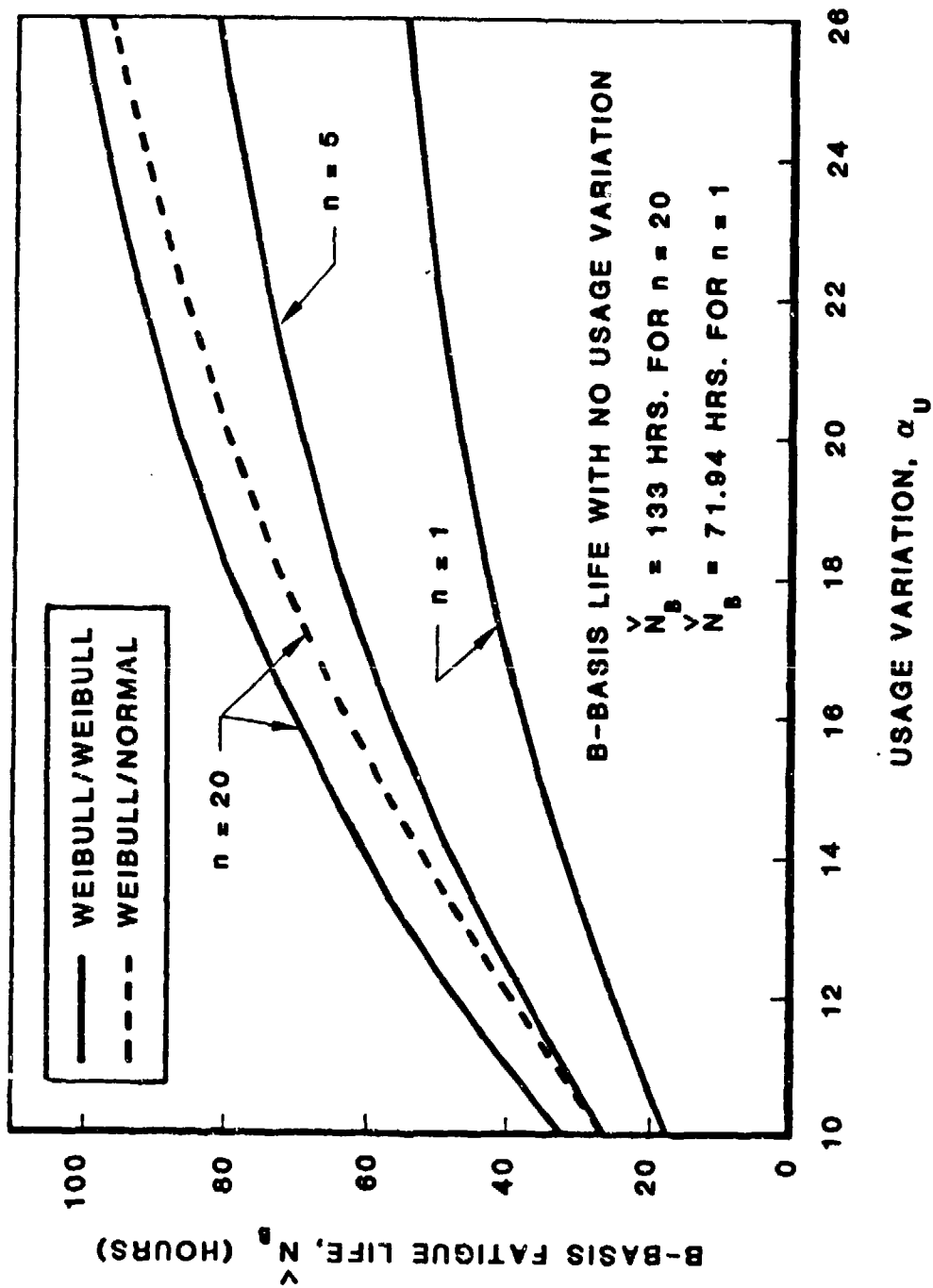
FIGURE 83. EFFECTS OF USAGE VARIATION ON THE B-BASIS LIFE.
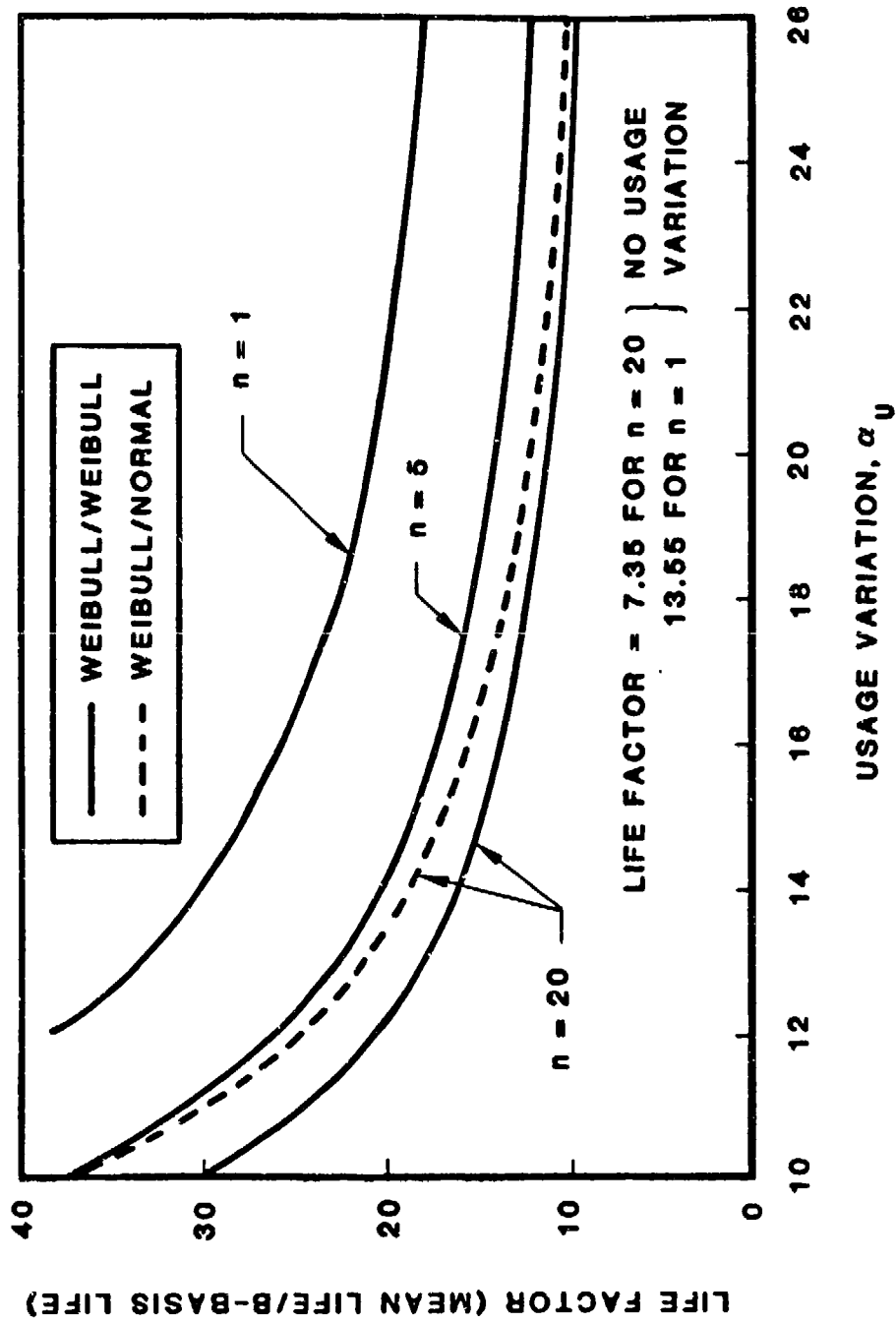
FIGURE 84. EFFECTS OF USAGE VARIATION ON B-BASIS FATIGUE LIFE FACTOR.

# SECTION 4

## METHODOLOGY DEMONSTRATION

The full scale wing and fuselage components from the Composite Wing/Fuselage Program (Reference 5) are selected as the primary demonstration articles. An extensive design development and full scale test data base was generated in this program. These data are reevaluated using the methodology developed in Sections 2 and 3. Static and fatigue margins of safety and reliabilities are established from this evaluation. Results of this evaluation are summarized in this section. Evaluations of the wing component (WCC-1) and fuselage component (FCC-1) are discussed in full detail.

4.1     Composite Wing/Fuselage Program Data Base

Figure 57 summarizes the wing test specimens used in Reference 5. This figure also shows the building block approach used for design development testing. The design development testing is characterized by four levels of complexity. The fifth level of complexity is assigned to the full-scale component.

The wing skin coupon specimens represent the first complexity level in the building block approach and simulate single tension and compression failure modes. The second complexity level in the building block approach contains specimens WE-2 and WEC-1. Each of these specimens has two potential failure modes. WE-2 is an upper skin/rear spar mechanical joint designed to check the influence of load transfer on compression strength. Potential failure modes for this specimen are laminate failure or bearing failure at a fastener hole. The WEC-1 specimen is a lower skin/intermediate spar cocured joint designed to check spar web strength in the presence of a fuel drain hole and the cocured bonded joint under combined shear, fuel pressure and chordwise loading. Potential failure modes are web failure at the fuel drain hole and bondline failure in the cocured joint. The third complexity level in the building block approach is

represented by WEC-3, which is an intermediate spar/pylon rib transfer joint, and is designed to check load transfer from the discontinuous spar into the rib and back to the spar. This specimen combines the potential failure modes of the wing coupons and WEC-1, i.e., upper and lower skin failure at a rib attachment fastener hold, spar web failure and intermediate spar/lower skin failure in the cocured joint. The fourth and final level of complexity in the torsion box design development testing is represented by the wing subcomponent WS-1 which is a three bay box beam and WS-2 which represents the highly loaded root rib/aft trunnion area. All of the failure modes of the wing coupons, WE-2, WE-1 and WEC-3 are represented in the WS-1 specimen. In addition, an upper skin access hole provides a further potential failure mode. The fifth level of complexity is the wing component WCC-1 which is fully representative of the actual wing structure.

Table 11 summarizes the test matrix for the wing specimens. Two static tests were conducted under RT/ambient and 250°F/wet conditions, where wet was defined as end-of-lifetime moisture level. A similar building block approach was used for the fuselage structure. Figure 85 summarizes the fuselage test specimens. The test matrix for the fuselage specimens is shown in Table 12.

Five fatigue test schemes were also used; these are summarized in Figure 86. All fatigue tests were conducted to two lifetimes followed by a residual static strength test. Test series 4 was a conventional RT/ambient accelerated fatigue test. Test series 6, 10 and 11 were accelerated environmental fatigue tests of varying complexity. All specimens in these test series were moisture conditioned prior to fatigue testing. Test series 12 was the least complex of the three test schemes with a constant temperature of 145°F imposed along with the accelerated flight loads. Test series 11 was the next most complex scheme with thermal spikes to 250°F and 218°F added to the 145°F baseline temperature profile. Test series 6 was the most complex

TABLE 11. SUMMARY OF WING TEST MATRIX FROM REFERENCE 5.

| SPECIMEN TYPE | STATIC TESTS | | FATIGUE TESTS (TWO LIFETIMES) | | | | |
| --- | --- | --- | --- | --- | --- | --- | --- |
| | RTA (TS2) | 250°F/WET (TS1) | RTA (TS4) | BASELINE (TS6) | ENVIRONMENTAL | | |
| | | | | | ALTERNATE No. 1 (TS11) | ALTERNATE No. 2 (TS12) | REAL FLIGHT TIME (TS10) |
| WC-1 | 10 | 10 | 10 | 10 | 3 | 3 | - |
| WC-2 | 3 | 3 | 3 | 3 | - | - | - |
| WC-3 | 10 | 10 | 10 | 10 | 3 | 3 | - |
| WC-4 | 3 | 3 | 3 | 3 | - | - | - |
| WC-5 | 3 | 3 | 3 | 3 | - | - | - |
| WC-6 | 3 | 3 | 3 | 3 | 3 | 3 | - |
| WE-1 | 3 | 3 | 3 | 3 | - | - | - |
| WE-2 | 3 | 3 | 3 | 3 | - | - | - |
| WEC-1 | 3 | 3 | 3 | 3 | 3 | 1 | 2 |
| WEC-2 | 3 | 3 | 3 | 3 | - | - | - |
| WEC-3 | 3 | 3 | 3 | 3 | 2 | 1 | 2 |
| WS-1 | 3 | 3 | 3 | 3 | 2 | 1 | 2 |
| WS-2 | 1 | 2 | - | 1 | - | - | - |
| WCC-1 | 1 | 1 | 1 | 1 | - | - | 1 |

FUSELAGE COMPONENT FCC-1

FUSELAGE HONEYCOMB PANEL TRANSITION AREA ELEMENT FE-1

FUSELAGE TRUNNION BULKHEAD CRITICAL ELEMENT COMBINATION FEC-2

FUSELAGE INLET DUCT/ BULKHEAD JOINT ELEMENT COMBINATION FEC-3

FUSELAGE HAT STIFFENED PANEL ELEMENT COMBIN- ATION FEC-4
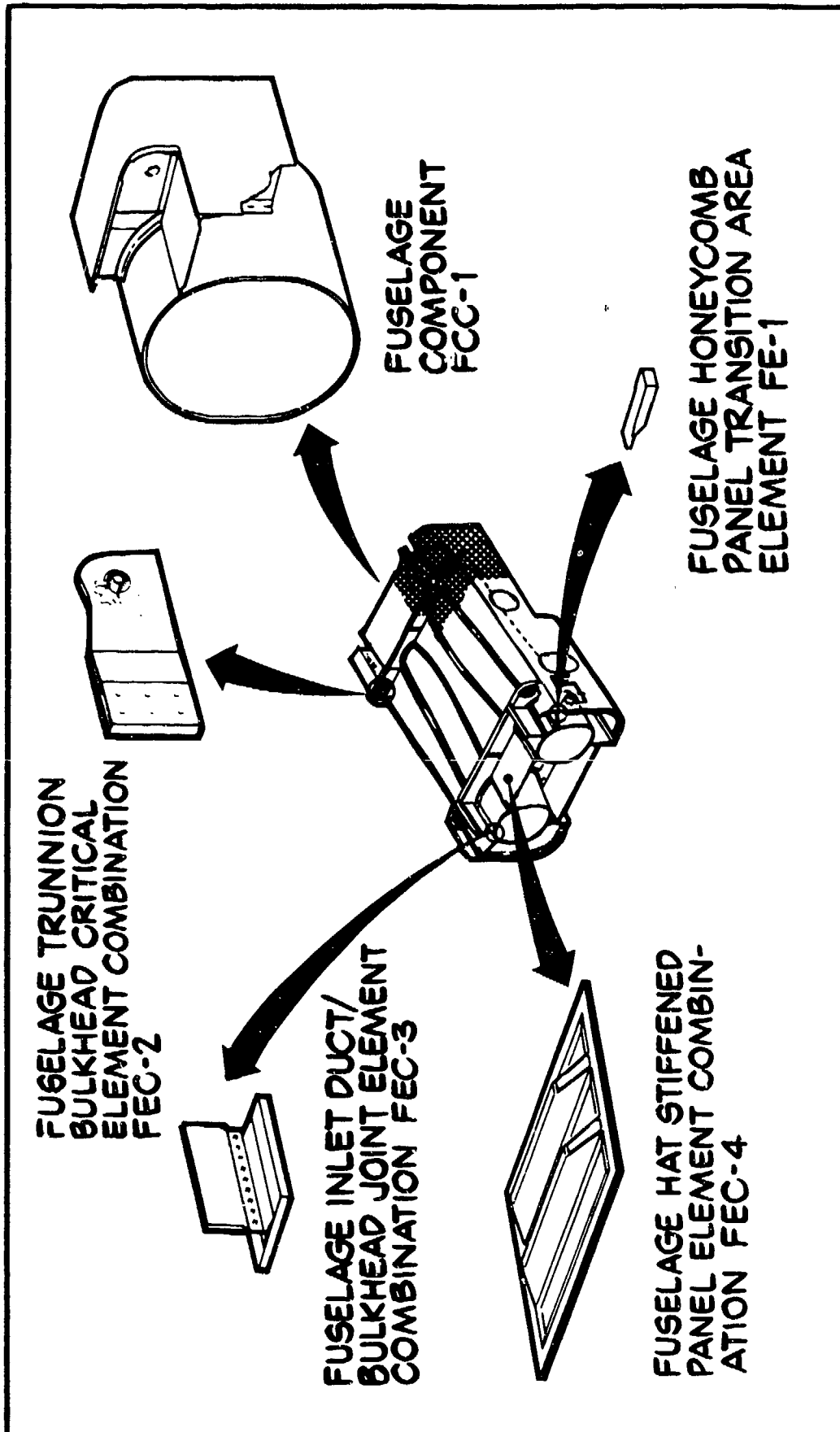
FIGURE 85. BUILDING BLOCK APPROACH FOR THE FUSELAGE STRUCTURE IN REFERENCE 5.

TABLE 12.  SUMMARY OF FUSELAGE TEST MATRIX FROM REFERENCE 5.

| SPECIMEN TYPE | STATIC TESTS | | FATIGUE TESTS (TWO LIFETIMES) | | | |
|---|---|---|---|---|---|---|
| | | | RTA | ENVIRONMENTAL | | |
| | RTA (TS 2) | 250 F/WET (TS 1) | (TS 4) | BASELINE (TS 6) | ALTERNATE No. 1 (TS 11) | ALTERNATE No. 2 (TS 12) |
| FE-1 | 3 | 3 | 3 | 3 | 3 | 3 |
| FEC-2 | 3 | 3 | 3 | 3 | – | – |
| FEC-3 | 3 | 2 | 3 | 2 | – | – |
| FEC-4 | 2 | 3 | 3 | 3 | – | – |
| FCC-1 | 1 | 1 | 1 | – | – | – |
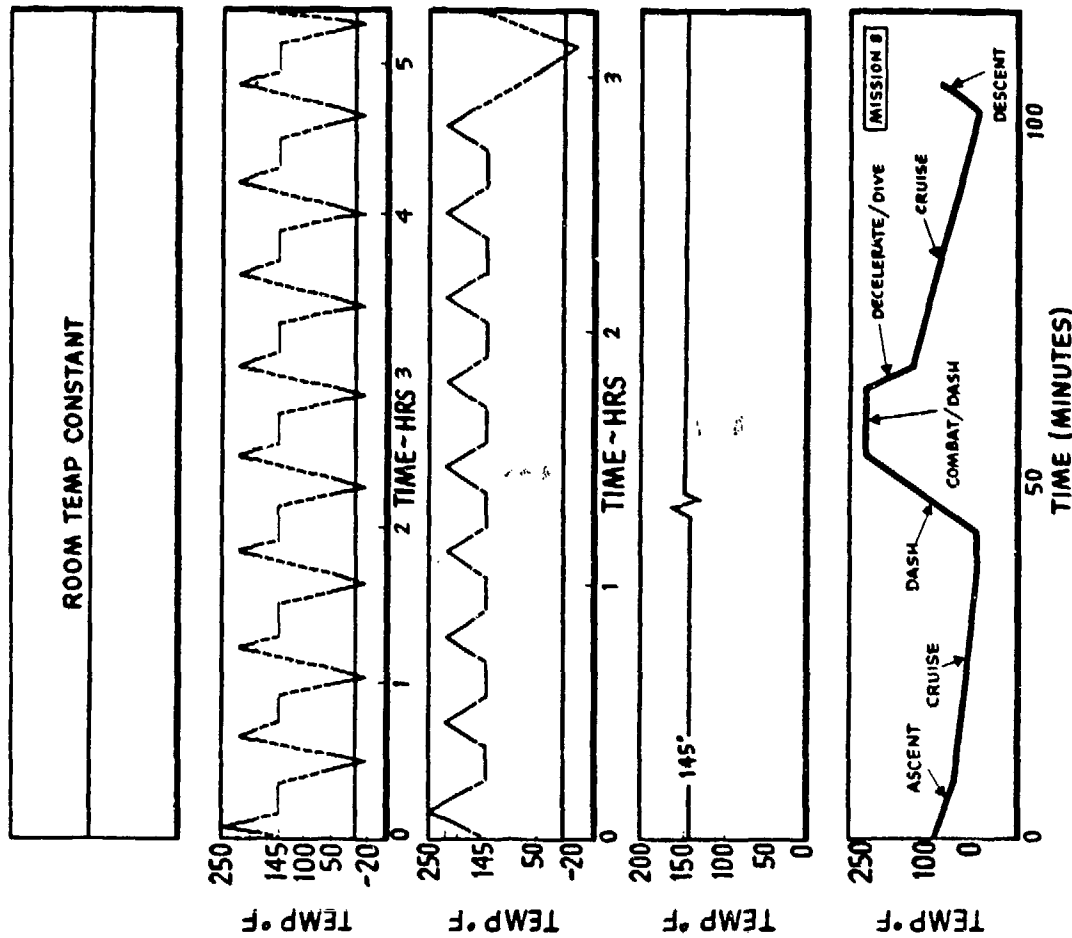
FIGURE 86. FATIGUE TEST SCHEMES USED IN REFERENCE 5.

accelerated environmental test scheme with thermal cycles ranging from -20 to 250°F. In addition, the complexity of test series 6 was increased by reconditioning test specimens during testing to replace moisture lost during thermal cycling. Test series 10 was a real flight time fatigue test, where flight loads and the associated temperature were applied to the actual rates seen by an aircraft. In addition, the moisture history of the test specimens was carefully simulated to match that seen in a 20 year service life.

Residual static strength tests after environmental fatigue loading (Test series 6, 10, 11 and 12) were conducted under the same 250°F/wet conditions as test series 1.

All testing shown in Table 11 is complete except for the real flight test on the Wing Component (WCC-1, Test Series 10). This test is in progress and has reached approximately one lifetime.

## 4.2 Wing Component (WCC-1) Data Evaluation

A detailed discussion of the wing component test results is given in Reference 20. The wing component with load introduction structure is shown in Figure 87 and the WCC-1 test set-up is shown in Figure 88. The most critical static design ultimate loads are shown in Table 13. A total of two static and two fatigue wing components were tested. The results are presented in Table 14. The static strength and fatigue life reliabilities of the component are determined based on the actual test data shown in Table 14. The results are discussed in the following paragraphs.

## 4.2.1 WCC-1 Static Test Data Evaluation

The static strength reliability is evaluated using the two-parameter Weibull distribution. The 95% confidence reliability of the component at DUL and DLL and the A- and B-basis static load are determined using the modal Weibull shape parameter $\alpha_S = 20.0$, for composite static data scatter (see Volume I). The effects of the structural response variation (SRV) are also
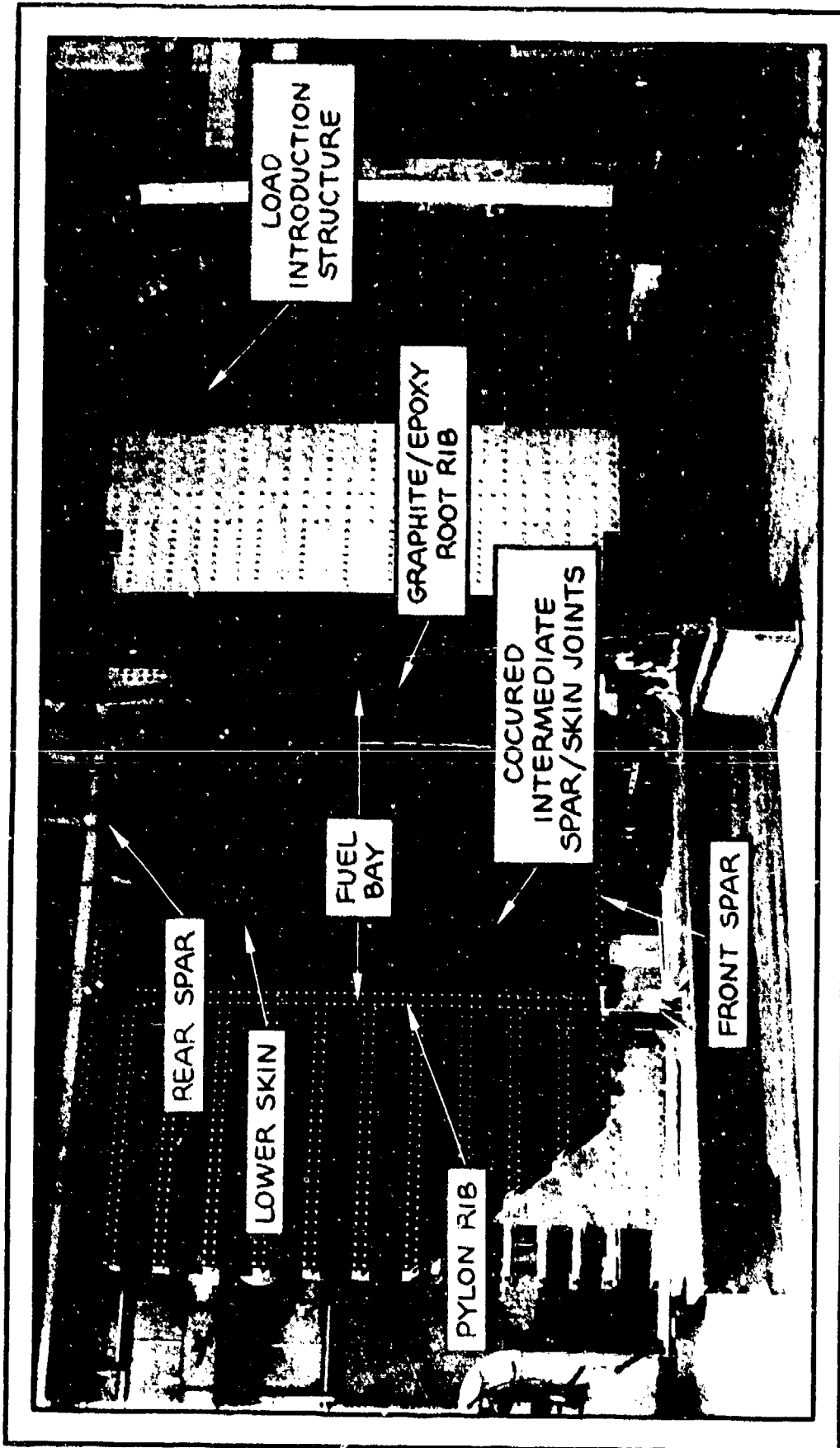
FIGURE 87. WING COMPONENT WITH LOAD INTRODUCTION STRUCTURE (REFERENCE 5)
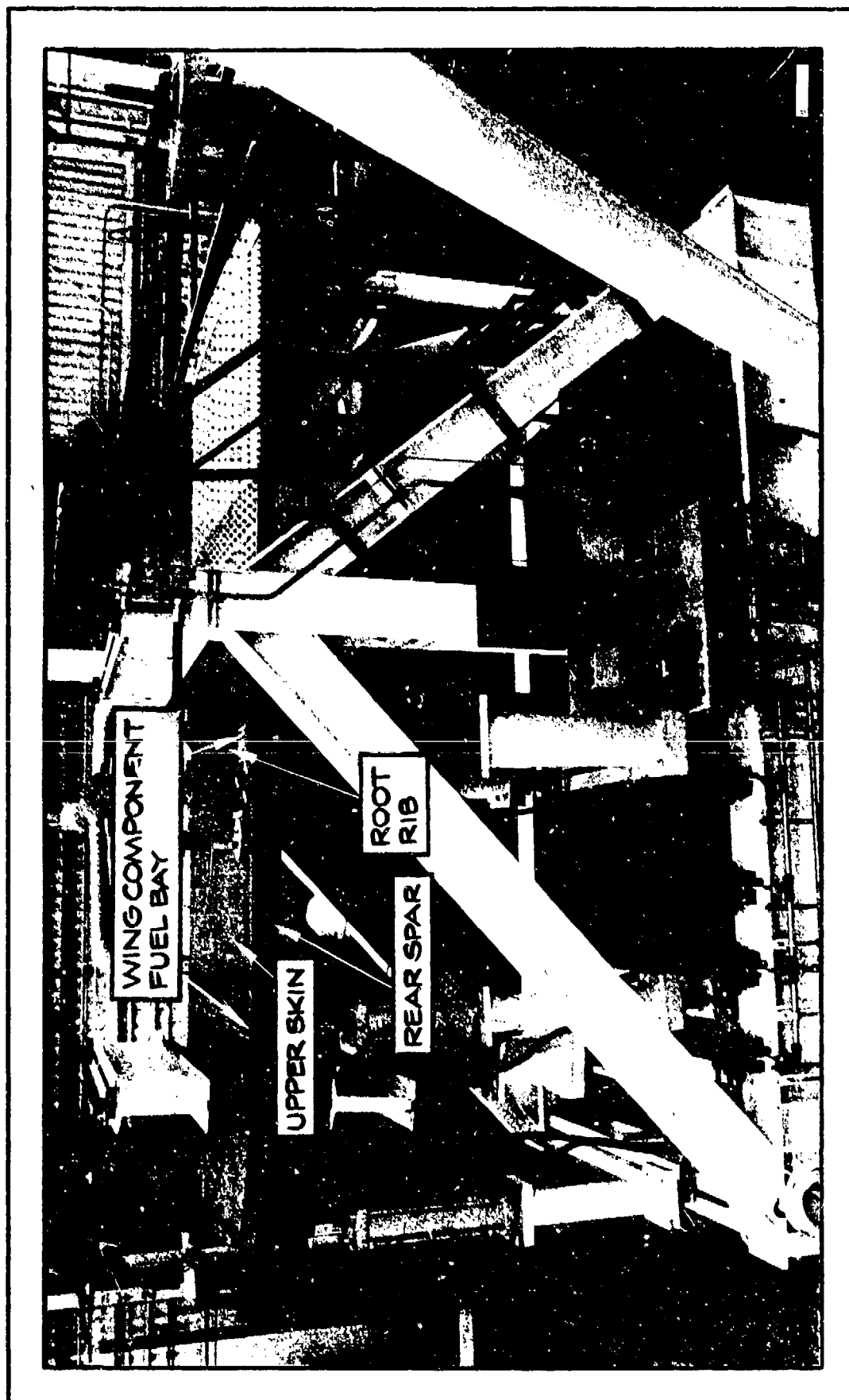
FIGURE 88. WCC-1 TEST SET-UP (REFERENCE 5).

TABLE 13. SUMMARY OF WING COMPONENT (WCC-1) STATIC DESIGN ULTIMATE LOADS.

| CONDITION NUMBER | LOAD FACTOR $N_z$ | TEMPERATURE °F | BENDING MOMENT IN-LB | TORQUE IN-LB | VERTICAL SHEAR LBS |
|---|---|---|---|---|---|
| 130 * | 9.75 Supersonic Pull-up | 242 | $10.37 \times 10^6$ | $6.01 \times 10^6$ | 140,400 |
| 100 | 11 Subsonic Pull-up | 145 | $11.70 \times 10^6$ | $6.25 \times 10^6$ | 172,500 |

* CASE 130 IS CRITICAL STATIC DESIGN CONDITION

TABLE 14. SUMMARY OF WING COMPONENT (WCC-1) FAILURE LOADS.

| TEST SERIES | TEST CONDITION | FAILURE LOAD % DUL | FAILURE MODE |
|---|---|---|---|
| 2 | RTA STATIC | 122 | SUBSTRUCTURE |
| 1 | 250°F/WET STATIC | 126 | UPPER SKIN |
| 4 | TWO LIFETIME RTA FATIGUE + RTA RSS | 118 | SUBSTRUCTURE |
| 6 | TWO LIFETIME ACCELERATED ENVIRONMENTAL FATIGUE + 250°F/WET RSS | 124 | UPPER SKIN |

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accounted for in the reliability evaluation. The overall mean SRV ($\alpha_{SRV}$ = 17.0) is used in determining the reliability. The numerical procedure is illustrated below for the RTA data.

Mean Failure Load $\quad\quad\quad\quad \bar{x}$ = 1.22 DUL

Sample Size $\quad\quad\quad\quad\quad\quad\quad$ n = 1

Strength Shape Parameter $\quad \alpha_S$ = 20.0

SRV Shape Parameter $\quad\quad\quad \alpha_{SRV}$ = 17.0

Scale Parameter

$$\hat{\beta} = \bar{x}/\Gamma(1+\frac{1}{\alpha_S}) = 1.22/0.9735 = 1.253$$

95% Confidence Scale Parameter $\check{\beta} = \hat{\beta}/[\chi^2_{0.95}\ (2n)/2n]^{1/\alpha_S}$

$$= 1.253/(5.991/2)^{0.05}$$

$$= 1.186$$

The reliability at any load level is then evaluated by numerical integration of Equation (26). The probability function p($\sigma$) is assumed to be a two-parameter Weibull distribution with a shape parameter of 17.0.

The reliability at the 250°F/wet environment is evaluated by two different methods. The first method is a direct assessment from the ETW test data and the second method uses the RTA test data and an environmental knockdown factor (k) determined from design allowable data. For the direct method, the average failure of 1.26DUL is used in the evaluation. An environmental knockdown factor of k = 1.30 is used in the indirect method.

The 95% static strength distributions are presented in Figure 89. Static strength reliabilities and allowable operating loads are given in Table 15. The RTA static strength reliability of 0.94 is better than B-basis at DUL and higher than A-Basis at DLL. ETW static strength reliability is similar to RTA reliability at DUL and DLL when calculated directly from the ETW test data. Table 15 shows that ETW static strength reliability calculated indirectly from the RTA test data are lower than those
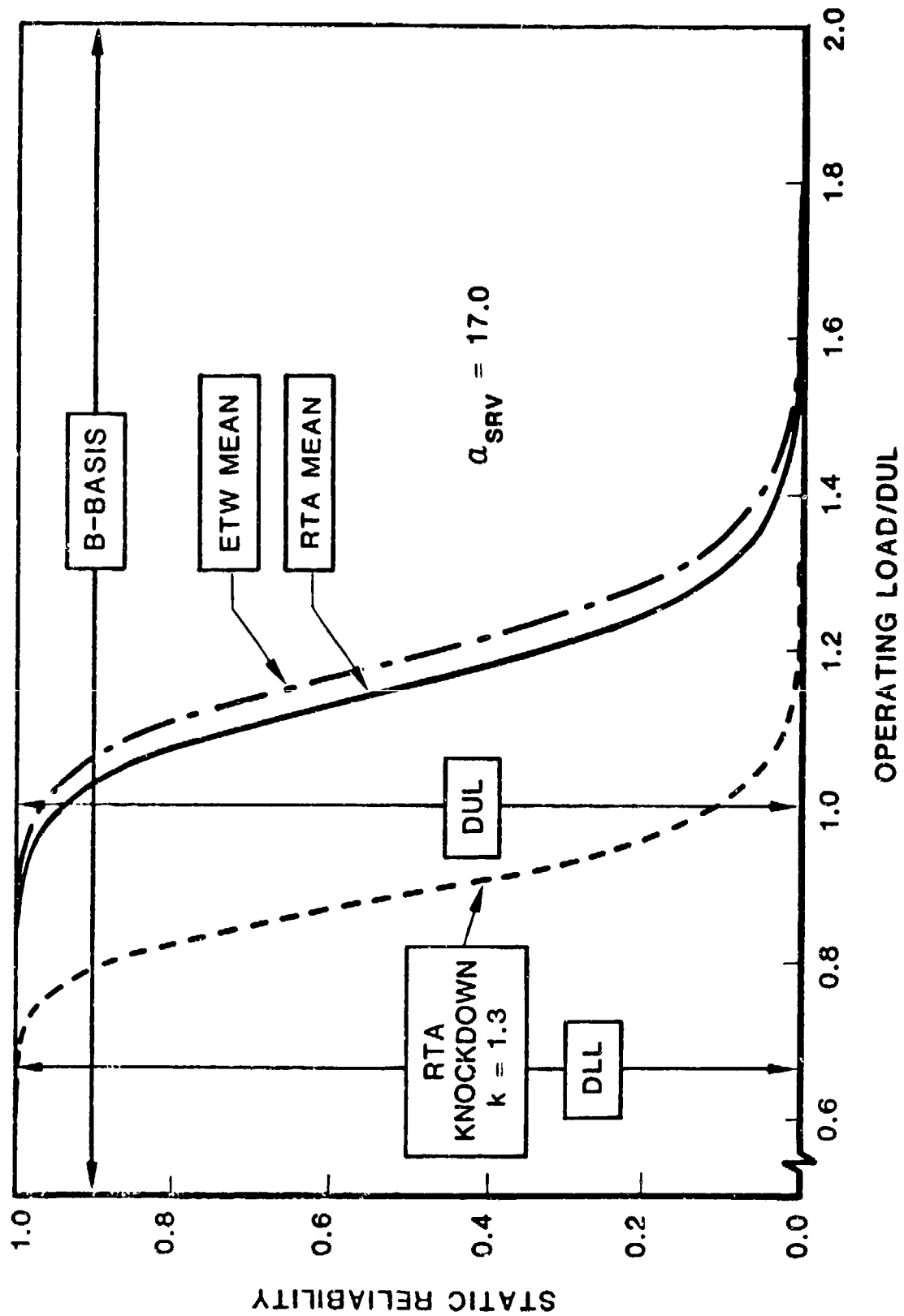
FIGURE 89. WCC-1 STATIC STRENGTH RELIABILITY DISTRIBUTIONS.

TABLE 15. SUMMARY OF WCC–1 STATIC STRENGTH RELIABILITY AND ALLOWABLE OPERATING LOADS.

| CONDITION | RELIABILITY | | ALLOWABLE (% DUL) | |
| --- | --- | --- | --- | --- |
| | DLL | DUL | B–BASIS | A–BASIS |
| RTA | 1.000 | 0.938 | 103 | 91 |
| ETW (MEAN DATA) | 1.000 | 0.966 | 106 | 94 |
| ETW (RTA KNOCKDOWN) (k = 1.30) | 0.996 | 0.107 | 79 | 70 |

calculated directly from ETW data. This indicates that the environmental knockdown factor approach produces a conservative estimate of the wing component ETW static strength reliability. This occurs because RTA and ETW static strengths of the wing component are similar. This was caused by a change in failure mode at ETW test conditions, which was attributed to a difference in the most critical static design case for the two test environments.

Table 15 shows that the A-basis RTA and ETW allowable operating loads are 91% DUL and 94% DUL, respectively. This compares with a maximum service load of 87% DUL. It can be concluded, therefore, that the static component tests verified the component design to a high level of confidence for the required in-service loading.

## 4.2.2    WCC-1 Fatigue Test Data Evaluation

Table 16 summarizes the WCC-1 fatigue test loads. Test enhancement load factor for both RTA and ETW tests was 1.025.

Three methods are used for the fatigue data evaluation. These methods are the load enhancement factor approach, the ultimate strength approach and the residual strength approach. The theoretical background of these approaches was discussed in Section 2. The numerical procedure for the three methods are given below.

## 4.2.2.1    Load Enhancement Factor Approach

The load enhancement factor approach was discussed in Section 2.2. The one lifetime fatigue reliability, without SRV, is obtained from Equations (16), (17) and (18), and is given by:

$$\ell = EXP\left\{ - \frac{\chi^2(2n)/2n}{P_L^{\alpha}R} \left[ \frac{\Gamma(1 + \frac{1}{\alpha_L})}{N} \right]^{\alpha_L} \right\} \tag{31}$$

With SRV, Equation (31) together with the SRV distribution function are used in Equation (26). Results are obtained by carrying out the numerical integration of Equation (26).

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**TABLE 16. SUMMARY OF WCC-1 FATIGUE TEST LOADS.**

| TEST | MAXIMUM SPECTRUM LOAD | | |
|---|---|---|---|
| | $P_{MSL}^{TEST}$ LBS | $P_{MSL}^{DES}$ LBS | $P_{MSL}^{TEST}/P_{MSL}^{DES}$ |
| RTA (TS4) | 115,100 | 118,000 | 1.025 |
| ETW (TS6) | 115,100 | 118,000 | 1.025 |

**MAXIMUM SPECTRUM LOAD = 82% DUL**

The values of various parameters use in the WCC-1 fatigue evaluation are given below:

$$\alpha_L = 1.25 \qquad \Gamma(1 + 1/\alpha_L) = 0.93138$$

$$\alpha_R = 20.0 \qquad \Gamma(1 + 1/\alpha_R) = 0.97350$$

$$\alpha_{SRV} = 17.0 \qquad \Gamma(1 + 1/\alpha_{SRV}) = 0.96930$$

$$N = 2.0 \text{ LIFETIME}$$

$$P_L = 1.025$$

$$n = 1 \qquad \chi^2_{0.95}(2n)/2n = 2.9955$$

The one lifetime fatigue reliability is plotted against the load enhancement factor ($P_L$) and is shown in Figure 90. At $P_L = 1.025$ the 95% confidence reliability at one lifetime is 0.493. The ETW fatigue reliability, obtained from indirect method using k = 1.30, is 0.091. The relatively low reliability is because of the low load factor used in the test. As shown in Figure 90, a minimum load factor of 1.18 is required in order to achieve the B-basis reliability. The required load factor for the A-basis reliability is 1.39. The reliability provided by the test data and the required A- and B-basis factors are summarized in Tables 17 and 18. Table 17 summarizes the RTA fatigue reliability and Table 18 shows the ETW reliability. The ETW fatigue reliabilities are calculated from both the direct method and the indirect method using environmental knockdown factor.

## 4.2.2.2 Ultimate Strength Approach

The fatigue reliability determined by the ultimate strength approach is given by

$$R = EXP\left\{-\left[\frac{x}{F}\Gamma\left(1 + \frac{1}{\alpha}\right)\right]^\alpha \frac{\chi^2(2n)}{2n}\right\} \qquad (32)$$

where F is the static failure load

$\alpha$ is the static strength scatter parameter

x is the constant determined from the threshold stress
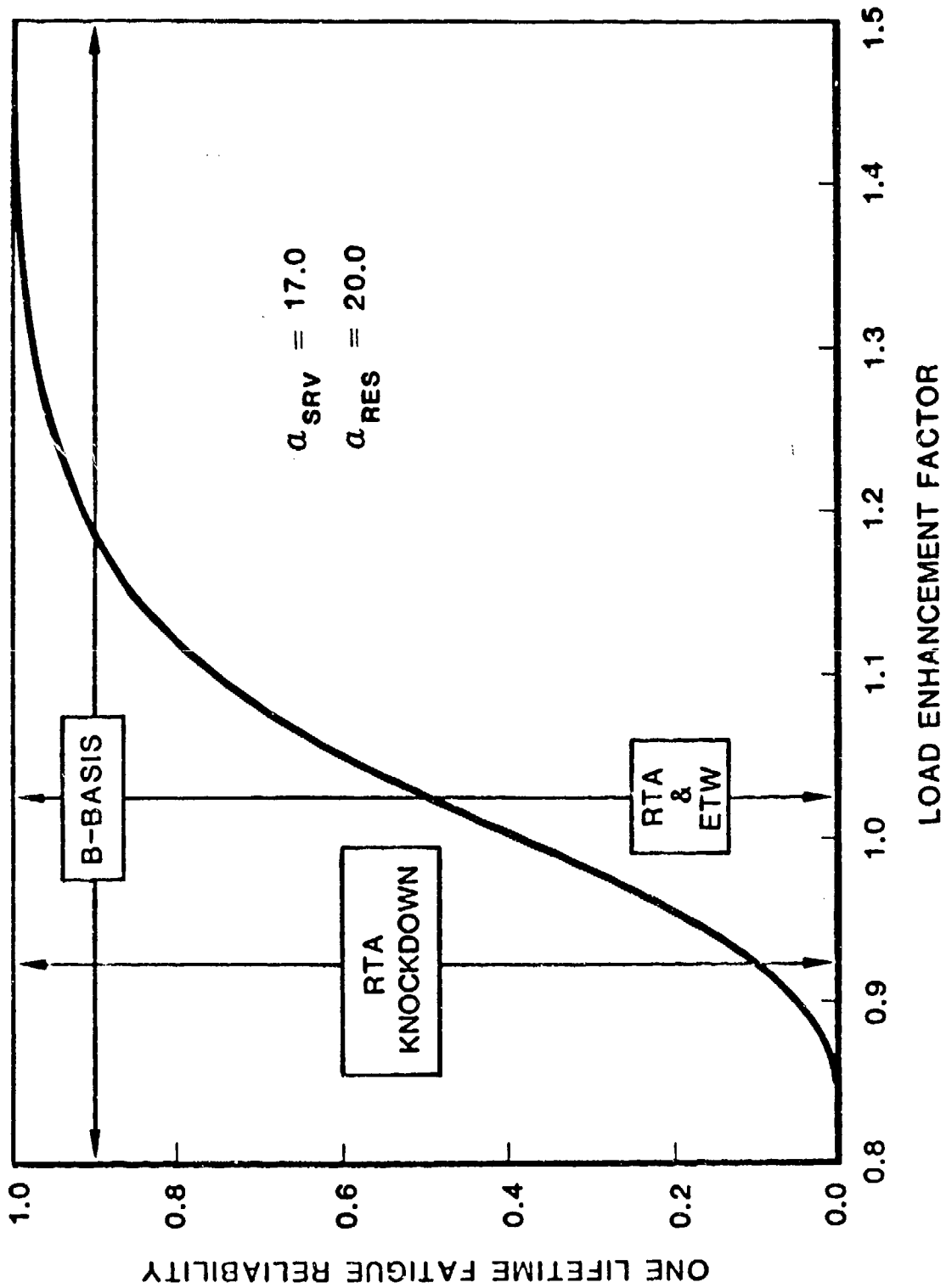
FIGURE 90.  WCC-1 FATIGUE RELIABILITY DISTRIBUTION DETERMINED BY
THE LOAD ENHANCEMENT FACTOR APPROACH.

TABLE 17. SUMMARY OF WCC–1 RTA ONE LIFETIME FATIGUE RELIABILITY.

| APPROACH | ONE LIFETIME FATIGUE RELIABILITY | REQUIRED FACTOR | |
|---|---|---|---|
| | | B-BASIS | A-BASIS |
| LOAD ENHANCEMENT FACTOR | 0.493 | $1.18\ P_{MSL}$ | $1.39\ P_{MSL}$ |
| ULTIMATE STRENGTH | 0.380 | 1.45 DUL | 1.70 DUL |
| RESIDUAL STRENGTH | 0.693 | 1.21 DUL | 1.22 DUL |

TABLE 18. SUMMARY OF WCC-1 ENVIRONMENTAL ONE LIFETIME FATIGUE RELIABILITY.

| APPROACH | DATA BASE | ONE LIFETIME FATIGUE RELIABILITY | REQUIRED FACTOR | |
|---|---|---|---|---|
| | | | B–BASIS | A–BASIS |
| LOAD ENHANCEMENT FACTOR | ETW RTA | 0.493 0.091 | 1.18 $P_{MSL}$ 1.31 $P_{MSL}$ | 1.39 $P_{MSL}$ 1.54 $P_{MSL}$ |
| ULTIMATE STRENGTH | ETW RTA | 0.521 0.041 | 1.45 DUL 1.61 DUL | 1.70 DUL 1.88 DUL |
| RESIDUAL STRENGTH | ETW RTA | 0.823 0.482 | 1.25 DUL 1.34 DUL | 1.26 DUL 1.35 DUL |

The value of x is determined by solving equation (32) for $F = F_1$ and $R = 0.9$, where

$$F_1 = \frac{P_{MSL}}{P_{DUL}} \cdot \frac{\sigma_S^M}{\sigma_{TH}^M} \cdot \frac{\sigma_{TH}^M}{\sigma_{TH}^B} \tag{33}$$

The definitions and values of $\sigma_S^M/\sigma_{TH}^M$ and $\sigma_{TH}^M/\sigma_{TH}^B$ were given in Section 2. In the present data evaluation, the modal values, $\sigma_{TH}^M/\sigma_S^M = 0.71$ and $\sigma_{TH}^B/\sigma_{TH}^M = 0.895$ are used.

The results of WCC-1 component fatigue reliability evaluation are summarized in Tables 17 and 18. The one lifetime fatigue reliability is plotted against static strength and shown in Figure 91. Based on the RTD static test failure load of .22 DUL, the one lifetime fatigue reliability is 0.380. The required static failure load is 1.45 DUL in order to achieve the B-basis fatigue reliability at one lifetime.

## 4.2.2.3   Residual Strength Approach

The residual strength approach is an application of the Sendeckyj's fitting model (Reference 14). The original model proposed by Sendeckyj required static strength, fatigue life and residual strength data. However, the available data only includes static and residual strength data, the fatigue reliability can only be approximately evaluated. In the Sendeckyj mode, the wear-out equation is given by

$$\sigma_e = \sigma_a \left[ (\sigma_r / \sigma_a)^{1/S} + NC \right]^S \tag{34}$$

where, $\sigma_e$ is the equivalent static strength

$\sigma_a$ is the maximum applied stress

$\sigma_r$ is the residual strength

N  is the test duration

S and C are fitting constants.
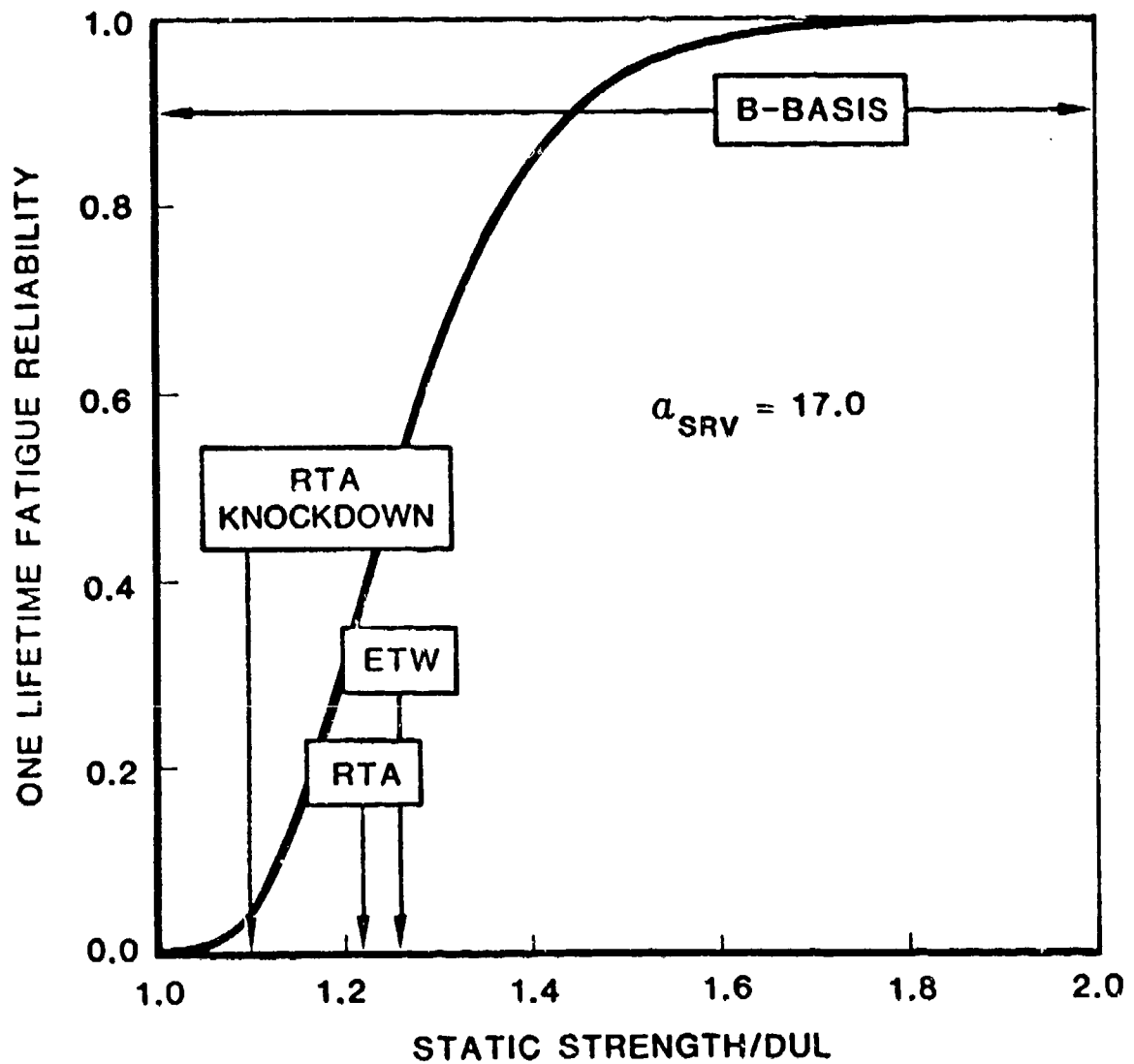
FIGURE 91. WCC-1 FATIGUE RELIABILITY DISTRIBUTION DETERMINED
BY THE ULTIMATE STRENGTH APPROACH.

As was discussed in Volume I, it can be shown that the fatigue life can be approximately described by a two-parameter Weibull distribution, with shape parameter $\alpha_L$. Let $\alpha_e$ be the shape parameter of the equivalent strength distribution. Theoretically, $\alpha_e$ should be equal to the static strength scatter parameter $\alpha_S$. The relation of $\alpha_L$ and $\alpha_e$ can be approximately given by

$$\alpha_L = S\alpha_e = S\alpha_S$$

(35)

Using the modal values of $\alpha_L = 1.25$ and $\alpha_S = 20.0$, the value of S is then 0.0625. From the static and residual strength data, the value of the other fitting constant C can be obtained by solving equation (34). The fatigue life reliability is then computed by first estimating the fatigue life and then substituting it into the Weibull distribution. The following example outlines the numerical procedure

| | |
|---|---|
| Static strength | $\sigma_n = 1.31$ DUL |
| Residual strength | $\sigma_r = 1.25$ DUL |
| Test duration | $N = 2.0$ |
| Applied stress | $\sigma_a = 1.075 \times 0.87 = 0.934$ DUL |

From equation (34)

$$1.31 = 0.934 \left[ \left(\frac{1.25}{0.934}\right)^{1/0.0625} + 2C \right]^{0.0625}$$

Estimate mean fatigue life at $\sigma_a = 0.87$ DUL

$$\left(\frac{\sigma_u}{\sigma_a}\right)^{1/S} = \left(\frac{\sigma_r}{\sigma_a}\right)^{1/S} + NC$$

At fatigue failure, $N = \bar{N}_F$, $\sigma_r = \sigma_a$, then

$$\bar{N}_F = \frac{\left(\frac{1.31}{0.87}\right)^{1/0.0625} - 1}{59.176} = 11.78 \text{ LT}$$

$$\hat{\beta} = N_F / \Gamma (1 + \frac{1}{\alpha_L}) = 11.78/0.93138 = 12.65 \ LT$$

$$\check{\beta} = \hat{\beta} / [\chi(2n)/2n]^{1/\alpha_L} = 12.65/(2.995)^{1/1.25} = 5.26 \ LT$$

The 95% confidence reliability at one lifetime under the maximum spectrum load is

$$R = EXP \ [-(1.0/5.26)^{1.25}] = 0.882$$

The reliability including SRV effects is computed by numerical integration of an expression similar to equation (26).

The results of fatigue reliability evaluation are summarized in Tables 17 and 18. The one lifetime fatigue reliability is shown as a function of the residual strength in Figure 92. The RTA one lifetime fatigue reliability obtained from the test data is 0.693 and the required residual strength for the B-basis fatigue reliability is 1.21 DUL.

4.3      Wing Specimens Test Data Evaluation Summary

All the wing specimens test data are evaluated for static and fatigue reliability in the same manner as for the WCC-1 component. This section summarizes the results of these evaluations.

Table 19 summarizes the static strength reliabilities and maximum allowable operating loads for all the wing test specimens.

Table 19 shows that under RTA conditions, average test failure loads ranged from 122% DUL to 244% DUL. Static strength reliabilities calculated at RTA conditions exceed A-basis at DLL for all the test specimens. In addition, B-basis allowable operating loads exceed DUL and A-basis allowable operating loads exceed 90% DUL for all test specimens. Thus, it can be concluded that very high RTA static reliability has been demonstrated for the composite wing structure.
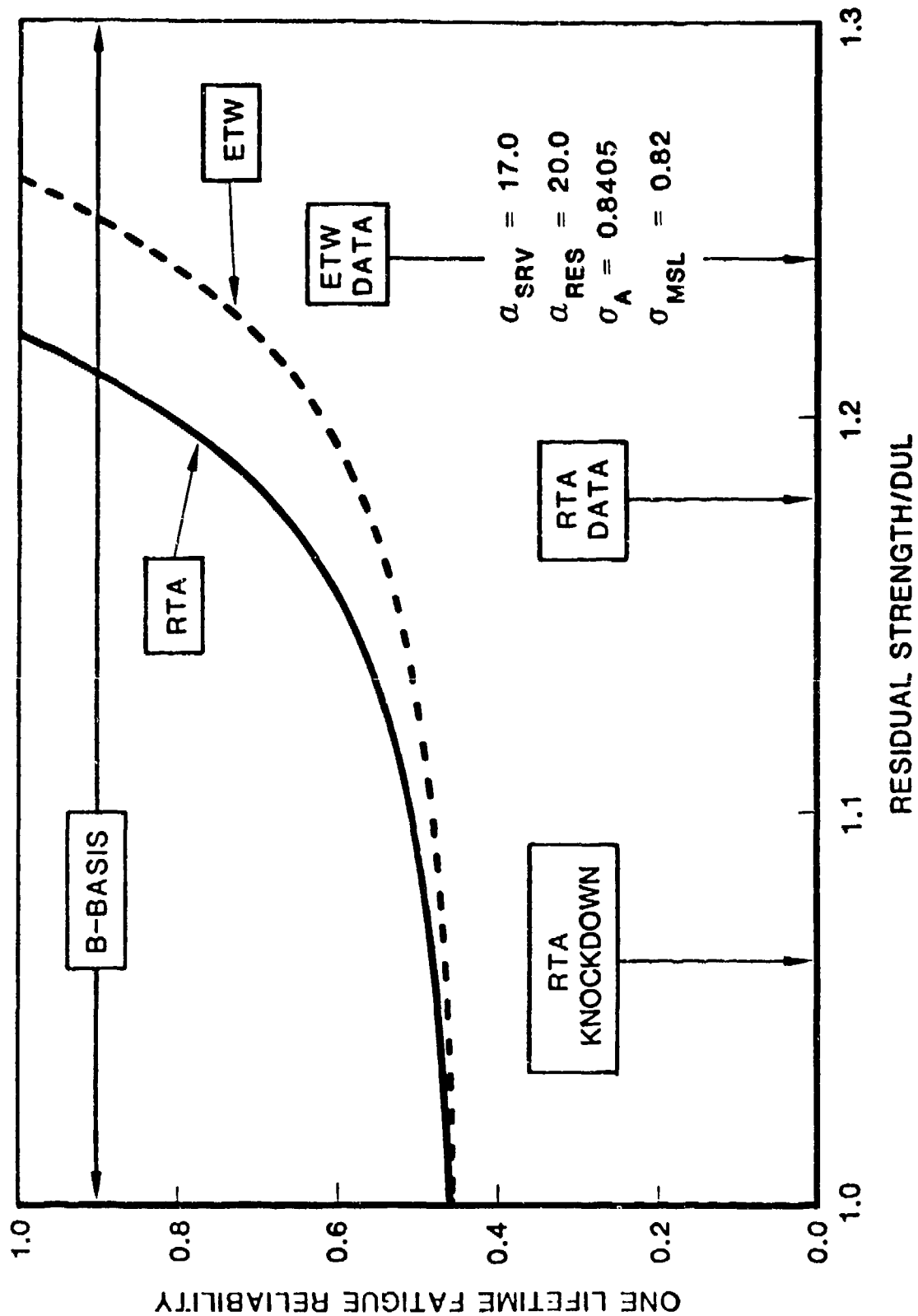
FIGURE 92. WCC-1 FATIGUE RELIABILITY DISTRIBUTION DETERMINED BY THE RESIDUAL STRENGTH APPROACH.

TABLE 19. SUMMARY OF STATIC STRENGTH RELIABILITIES AND ALLOWABLE OPERATING LOADS FOR WING TEST SPECIMENS.

| SPECIMEN I.D. | FAILURE LOAD %DUL | | STATIC RELIABILITY | | | | ALLOWABLE OPERATING LOAD (%DUL) | | | |
|---|---|---|---|---|---|---|---|---|---|---|
| | | | RTA | | ETW* | | RTA | | ETW* | |
| | RTA | ETW | DLL | DUL | DLL | DUL | B-BASIS | A-BASIS | B-BASIS | A-BASIS |
| WE-1 | 244 | 76 | 1.000 | 1.000 | 0.863 | 0.005 | 209 | 185 | 66 | 58 |
| WE-2 | 173 | 123 | 1.000 | 1.000 | 1.000 | 0.961 | 149 | 132 | 105 | 93 |
| WEC-1 | 186 | 160 | 1.000 | 1.000 | 1.000 | 1.000 | 159 | 141 | 137 | 122 |
| WEC-2 | 197 | 130 | 1.000 | 1.000 | 1.000 | 0.981 | 166 | 147 | 108 | 97 |
| WEC-3 | 128 | 128 | 1.000 | 0.975 | 1.000 | 0.978 | 108 | 95 | 108 | 98 |
| WS-1 | 131 | 102 | 1.000 | 0.985 | 0.999 | 0.362 | 111 | 98 | 85 | 75 |
| WS-2 | 197 | 180 | 1.000 | 1.000 | 1.000 | 1.000 | 164 | 145 | 150 | 133 |
| WCC-1 | 122 | 126 | 1.000 | 0.838 | 1.000 | 0.976 | 103 | 91 | 106 | 94 |

* 250°F/WET CONDITIONS

Table 19 shows that under 250°F/wet test conditions all test specimens except WE-1, had average failure loads that exceed DUL. Failure loads were in the range of 102% DUL to 180% DUL. Average failure load for specimen WE-1 was 76% DUL. 250°F/wet static strength reliabilities for WE-1 are 0.863 at DLL and 0.005 at DUL. These low reliabilities are simply a reflection of the low average failure load (76% DUL) of specimen WE-1. However, it should be noted that the temperature associated with the WE-1 critical static design case is 192°. Therefore, the reliabilities calculated from the 250°F/wet test data are conservative. The static strength reliability at 192°F was recalculated from the 250°F/wet test data using the environmental knockdown factor (k). The resultant reliabilities at DLL and DUL increased to 1.000 and 0.498, respectively. For the remaining wing test specimens, the 250°F/wet static strength reliability exceeded A-basis at DLL. The static strength reliabilities at DUL exceeded B-basis for all remaining specimens except WS-1 (0.362). The low 250°F/wet static strength reliability calculated for WS-1 can be attributed to the mixed failure modes observed in the 250°F/wet static tests. Average failure load for all tests was 116% DUL, however, two failure modes were observed, these were, upper skin failure and intermediate spar/lower skin cocured joint failure. The average failure load for upper skin failure was 123% DUL, while the average failure load for intermediate spar/lower skin failure was 102% DUL. Since mixed failure modes were observed, the static strength reliability was calculated using the failure load which gave the lowest static strength. Thus, an average failure load of 102% DUL was used for the reliability calculations. This resulted in a low static strength reliability at DUL. However, it should be noted that Table 19 does show that the A-basis allowable operating load exceeds DLL.

Table 20 summarizes the one lifetime fatigue reliabilities for all wing test specimens. The fatigue reliabilities calculated using the load enhancement approach assume that fatigue failures occurred after two lifetimes. However, the major-

TABLE 20. SUMMARY OF ONE LIFETIME FATIGUE RELIABILITIES FOR WING TEST SPECIMENS.

| SPECIMEN I.D. | FATIGUE TEST LOAD ENHANCEMENT FACTOR | | FATIGUE RELIABILITY AT 1 LT | | | | | |
| --- | --- | --- | --- | --- | --- | --- | --- | --- |
| | | | RTA | | | ETW | | |
| | RTA | ETW | LEF | US | RSS | LEF | US | RSS |
| WE-1 | 1.87 | 0.58 | 1.000 | 1.000 | 1.000 | 0.000 | 0.000 | 0.000 |
| WE-2 | 1.55 | 1.10 | 0.999 | 0.898 | 1.000 | 0.797 | 0.001 | 0.976 |
| WEC-1 | 1.53 | 1.32 | 0.998 | 0.996 | 1.000 | 0.971 | 0.957 | 0.993 |
| WEC-2 | 1.66 | 1.10 | 0.998 | 0.995 | 1.000 | 0.764 | 0.532 | 0.814 |
| WEC-3 | 0.94 | 0.94 | 0.245 | 0.060 | 1.000 | 0.282 | 0.069 | 0.484 |
| WS-1 | 1.08 | 0.96 | 0.827 | 0.775 | 0.888 | 0.132 | 0.055 | 0.775 |
| WS-2 | -- | 1.00 | -- | 0.462 | -- | 0.400 | 0.969 | 0.462 |
| WCC-1 | 1.03 | 1.03 | 0.493 | 0.380 | 0.693 | 0.493 | 0.521 | 0.832 |

LEF = LOAD ENHANCEMENT FACTOR APPROACH
US = ULTIMATE STRENGTH APPROACH
RSS = RESIDUAL STRENGTH APPROACH

ity of the wing specimen fatigue tests were run-outs at two fatigue lifetimes, which were subsequently residual static strength tested. Thus, the fatigue reliabilities calculated by the load enhancement approach are conservative. Fatigue reliabilities calculated using the ultimate strength approach are estimates based on static failure loads. The fatigue reliabilities calculated using the residual strength approach provide the best estimates of one lifetime fatigue reliabilities because it accounts for the specimen residual strength after two lifetimes of fatigue loading.

The fatigue reliabilities calculated using the residual strength approach, which are presented in Table 20, show significant scatter. The reason for this scatter in calculated fatigue reliability is the variation in test load enhancement factors used in fatigue tests. For RTA tests, test load enhancement factors ranged from 0.94 to 1.87 and for environmental tests, test load enhancement factors ranged from 0.58 to 1.32. These variations in test load enhancement factors were caused by variations in static test failure load as a function of DUL for each of the test specimens. Because of the higher scatter in composite fatigue test data, it is necessary to increase test load severity in order to demonstrate adequate fatigue reliability in a two lifetime fatigue test.

The following conclusions can be drawn from the results of the wing test data evaluation.

(1)  A high degree of static strength reliability has been demonstrated by the wing test data.

(2)  One lifetime fatigue reliabilities demonstrated by the wing test data were inadequate because insufficient test load severity was used for the two lifetime fatigue tests.

## 4.4    Fuselage Component FCC-1 Data Evaluation

The FCC-1 fuselage component is shown in Figure 93 and the loading of the component is shown in Figure 94. The criti-

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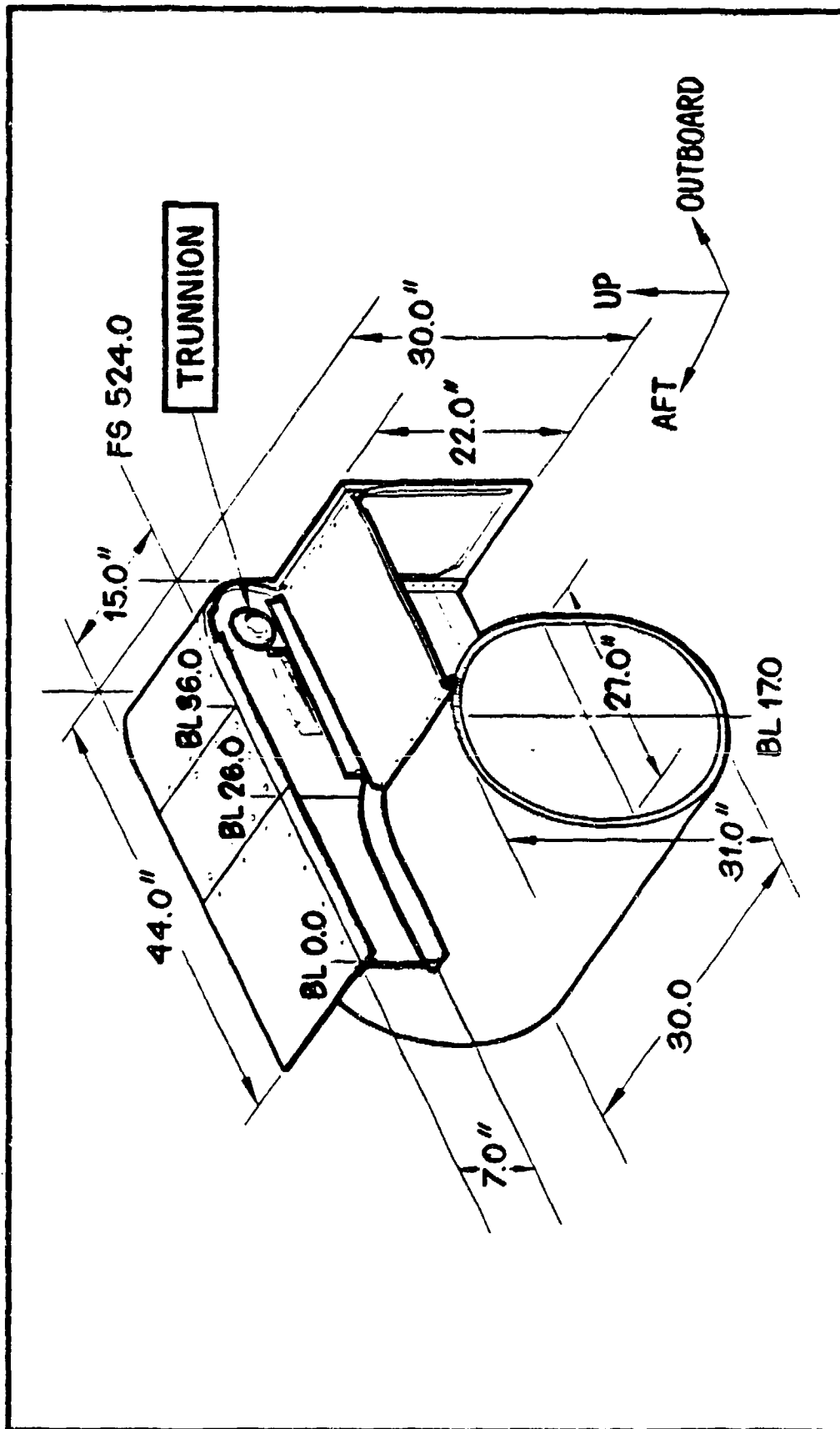
FIGURE 93. FUSELAGE COMPONENT (FCC-1).
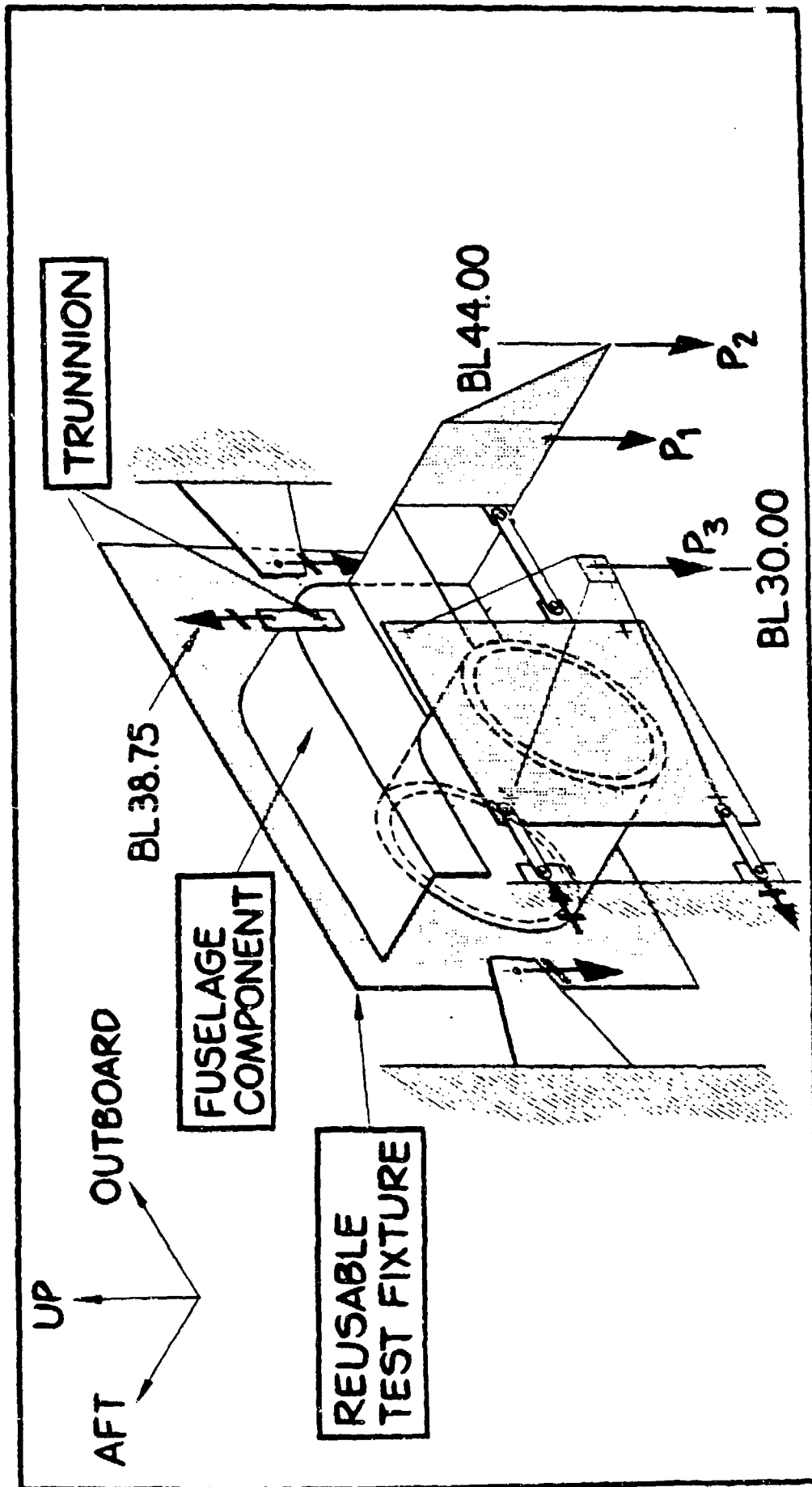
FIGURE 94. FCC-1 LOADING.

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cal static design conditions is given in Table 21. A total of two static and one fatigue components were tested. The results are presented in Table 22.

### 4.4.1    FCC-1 Static Test Data Evaluation

The 95% confidence static strength reliability distributions are presented in Figure 95. Static strength reliabilities and allowable operating loads are summarized in Table 23. As can be expected, the RTA static strength reliability at DUL is very low (0.340), because of the low RTA static failure load (102% DUL). As a result, the static strength reliability at the design environment (242°F/wet) determined from knockdown of RTA data is also low. This reliability is 0.882 at DLL and 0.006 at DUL (Table 23). The reliability determined from the RTA test result is over-conservative. This is because the failure load of 102% DUL does not represent the total structure strength. The test failure was fixture failure as indicated in Table 22. The static strength determined from the 250°F/wet test data exceeded A-basis as shown in Table 23.

### 4.4.2    FCC-1 Fatigue Test Data Evaluation

FCC-1 fatigue test loads is summarized in Table 24. The test load enhancement factor was 1.34. Only a RTA fatigue test was conducted.

Table 25 shows the one lifetime RTA fatigue reliability determined from the three analytical approaches. The low fatigue reliability (0.001) obtained from the ultimate strength approach, reflects the low RTA static failure load. This value is not accurate because of the test fixture failure during test. The load enhancement factor approach gives a reliability of 0.983 and the residual strength approach gives a fatigue reliability of 1.000. The reliability distributions are shown in Figures 96 and 97.

The environmental fatigue reliability is shown in Table 26. Because only RTA fatigue test was conducted, the fatigue reliability is evaluated based only on RTA test data. At

TABLE 21. SUMMARY OF FUSELAGE COMPONENT (FEC-1) STATIC ULTIMATE DESIGN LOADS.

| CONDITION NUMBER | LOAD FACTOR $N_z$ | TEMPERATURE °F | $P_1$ LBS | $P_2$ LBS | $P_3$ LBS | TRUNNION LOAD LBS |
|---|---|---|---|---|---|---|
| 120 * | 9.75 | 242 | 3410 | -16370 | -6480 | 76800 |

* CASE 120 IS CRITICAL STATIC DESIGN CONDITION

TABLE 22. SUMMARY OF FCC–1 FAILURE LOADS.

| TEST SERIES | TEST CONDITION | FAILURE LOAD % DUL | FAILURE MODE |
|---|---|---|---|
| 2 | RTA STATIC | 102 | TEST FIXTURE FAILURE |
| 1 | 250°F/WET STATIC | 145 | FUSELAGE SIDE PANEL |
| 4 | TWO LIFETIME RTA FATIGUE + RTA RSS | 145 | * |

FIGURE 95. FCC-1 STATIC STRENGTH RELIABILITY DISTRIBUTIONS.
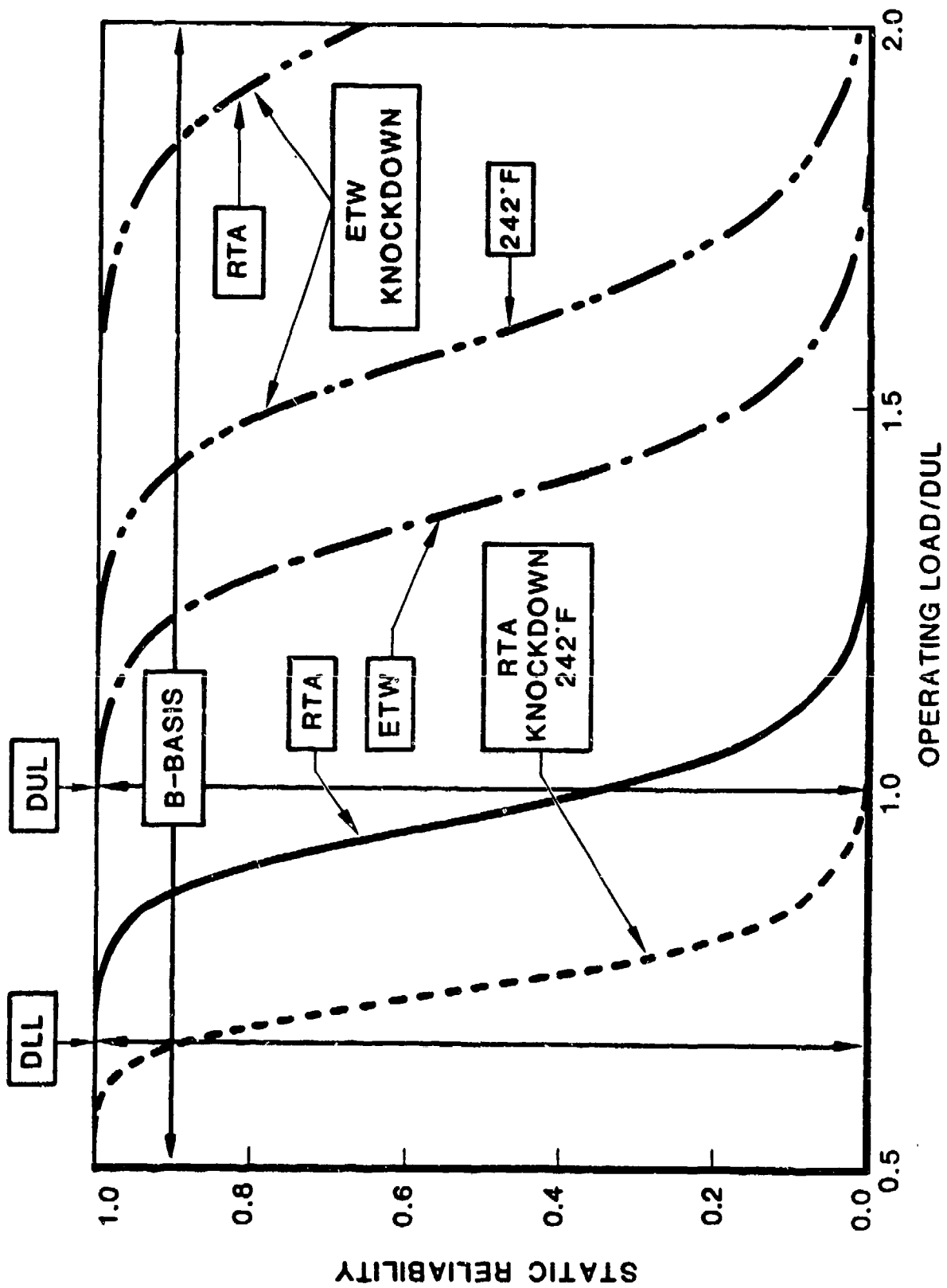
**TABLE 23. SUMMARY OF FCC-1 STATIC RELIABILITY.**

| CONDITION | RELIABILITY | | ALLOWABLE % DUL | |
|---|---|---|---|---|
| | DLL | DUL | B-BASIS | A-BASIS |
| RTA | 0.999 | 0.340 | 86 | 76 |
| ETW<br>250°F – DATA<br>242°F – KNOCKDOWN<br>RT – KNOCKDOWN | 1.000<br>1.000<br>1.000 | 0.998<br>1.000<br>1.000 | 122<br>142<br>184 | 108<br>126<br>163 |
| ETW<br>RTA KNOCKDOWN<br>242°F | 0.882 | 0.006 | 66 | 59 |

**TABLE 24. SUMMARY OF FCC-1 FATIGUE LOADS.**

| TEST | MAXIMUM SPECTRUM LOAD | | |
|------|------|------|------|
| | $P_{MSL}^{TEST}$ LBS | $P_{MSL}^{DES}$ LBS | $P_{MSL}^{TEST}/P_{MSL}^{DES}$ |
| RTA (TS4) | 84400 | 63000 | 1.34 |

**MAXIMUM SPECTRUM LOAD = 82% DUL**

TABLE 25. SUMMARY OF FCC-1 RTA FATIGUE RELIABILITY AT ONE LIFETIME.

| APPROACH | ONE LIFETIME FATIGUE RELIABILITY | REQUIRED FACTOR | |
|---|---|---|---|
| | | B-BASIS | A-BASIS |
| LOAD ENHANCEMENT FACTOR | 0.983 | 1.18 $P_{MSL}$ | 1.39 $P_{MSL}$ |
| ULTIMATE STRENGTH | 0.001 | 1.47 DUL | 1.70 DUL |
| RESIDUAL STRENGTH | 1.000 | 0.82 DUL | 0.82 DUL |
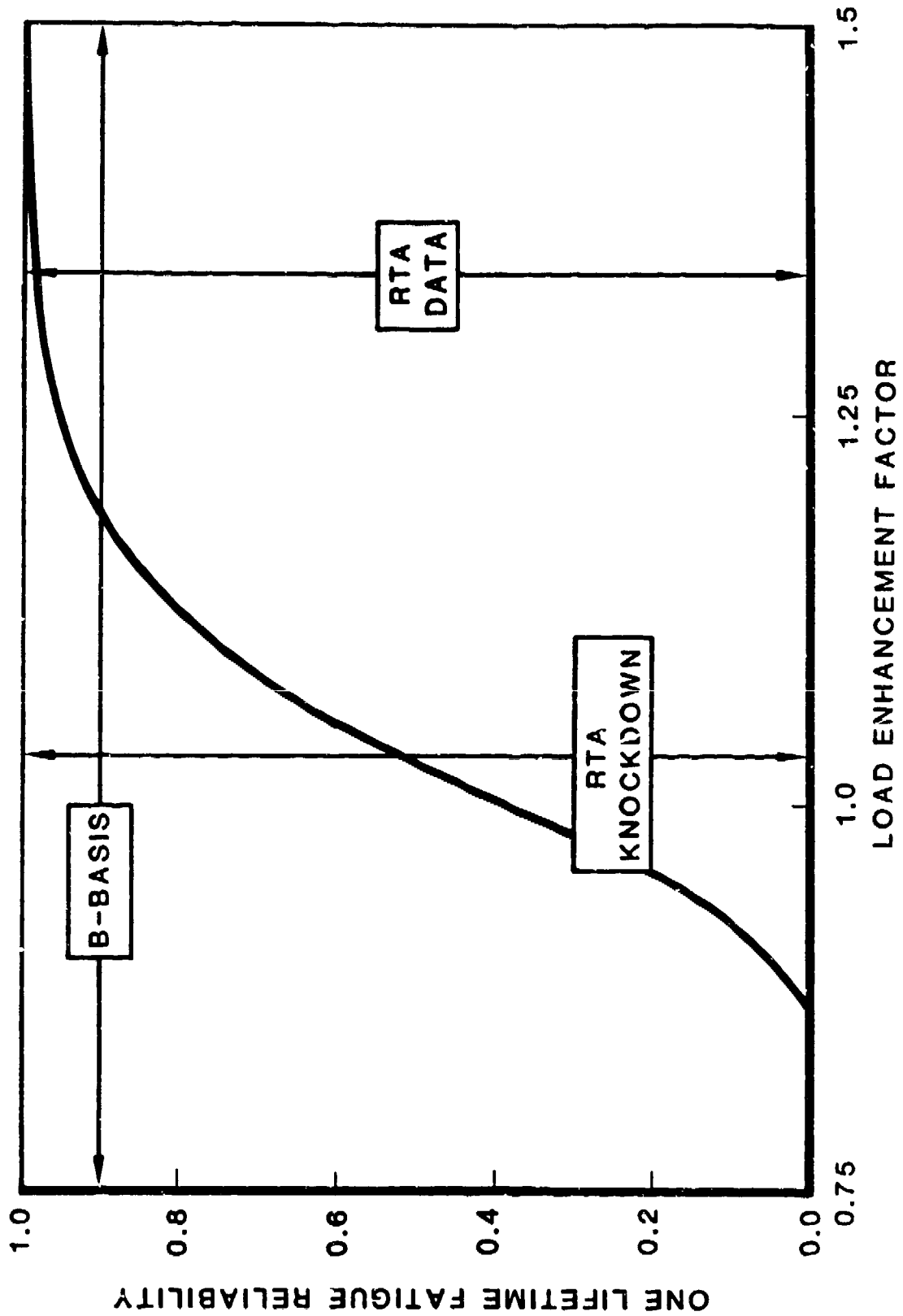
FIGURE 96. FCC-1 FATIGUE RELIABILITY - LOAD ENHANCEMENT FACTOR APPROACH.
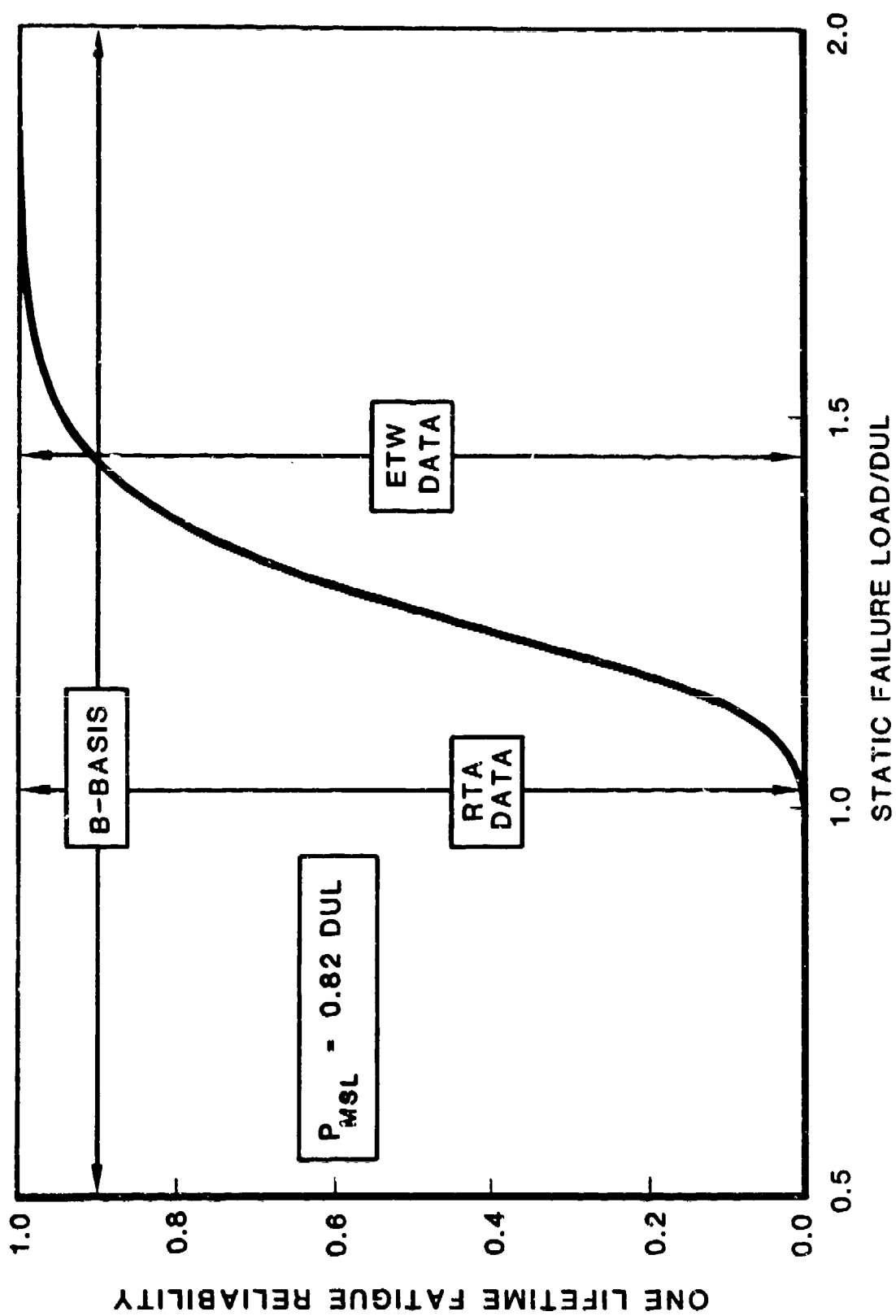
FIGURE 97. FCC-1 FATIGUE RELIABILITY – ULTIMATE STRENGTH APPROACH.

TABLE 26. SUMMARY OF FCC-1 ENVIRONMENTAL FATIGUE RELIABILITY AT ONE LIFETIME.

| APPROACH | DATA BASE | ONE LIFETIME FATIGUE RELIABILITY | REQUIRED FACTOR | |
|---|---|---|---|---|
| | | | B-BASIS | A-BASIS |
| LOAD ENHANCEMENT FACTOR | RTA | 0.516 | 1.31 $P_{MSL}$ | 1.54 $P_{MSL}$ |
| ULTIMATE STRENGTH | ETW | 0.903 AT 250°F | 1.45 DUL | 1.70 DUL |
| | ETW | 0.989 AT 242°F | 1.25 DUL | 1.46 DUL |
| | ETW | 1.000 AT RT | 0.96 DUL | 1.12 DUL |
| | RTA | 0.000 AT 242°F | 1.88 DUL | 2.21 DUL |
| RESIDUAL STRENGTH | RTA | 1.000 AT 242°F | 1.07 DUL | 1.07 DUL |

242°F/wet condition, the one lifetime fatigue reliability is 0.516, 0.000 and 1.000 from the load factor, ultimate strength and residual strength, respectively, based on RTA data. Based on the ETW static test data, the fatigue reliability at 242° is 0.989 as determined from the ultimate strength approach.

4.5        Fuselage Specimen Test Data Evaluation Summary

Table 27 summarizes the static strength reliabilities and allowable operating loads for all the fuselage test specimens.

Table 27 shows that the average RTA static failure loads ranged from 102% DUL to 286% DUL. Static strength reliabilities calculated at RTA conditions exceed A-Basis at DLL for all the test specimens, except FEC-4. The low strength reliability of FEC-4 specimen is a reflection of the high data scatter associated with the particular failure mode (stiffener disbond) of this specimen type. The static strength reliability at DUL exceeds B-Basis for specimens FEC-1, FEC-2 and FEC-3. The reliability at DUL for FEC-4 specimens is 0.402. This is also because of the high strength data scatter of the stiffener disbond failure mode observed in these specimens. The reliability of the FCC-1 specimen at DUL is 0.340; this is because of the low failure load (102% DUL) of the specimen. However, as was pointed out earlier, the failure was test fixture failure and thus did not represent the actual strength of the specimen. The RTA static strength reliability at DUL calculated from the 250°F/wet data was 1.000. The B-Basis RTA allowable operating loads for all fuselage specimens exceed DLL. The A-Basis RTA allowable operating loads for all specimens, except FEC-4, also exceed DLL.

Table 27 shows that under 250°F/wet test conditions, all test specimens except FEC-4 had average failure loads exceed DUL. The average failure load ranged from 125% DUL to 359% DUL. The average failure load for specimen FEC-4 was 86% DUL. Except for specimen FEC-4, the 250°F/wet static strength reliabilities at DLL all exceed A-Basis and they all exceed B-Basis at DUL. The 250°F/wet reliability for specimen FEC-4 is 0.703 at DLL and

TABLE 27. SUMMARY OF STATIC STRENGTH RELIABILITIES AND ALLOWABLE OPERATING LOADS FOR FUSELAGE TEST SPECIMENS.

| SPECIMEN I.D. | FAILURE LOAD %DUL | | STATIC RELIABILITY | | | | ALLOWABLE OPERATING LOAD (%DUL) | | | |
|---|---|---|---|---|---|---|---|---|---|---|
| | | | RTA | | ETW* | | RTA | | ETW* | |
| | RTA | ETW | DLL | DUL | DLL | DUL | B-BASIS | A-BASIS | B-BASIS | A-BASIS |
| FE-1 | 286 | 216 | 1.000 | 1.000 | 1.000 | 1.000 | 245 | 217 | 185 | 164 |
| FEC-2 | 125 | 125 | 1.000 | 0.972 | 1.000 | 0.974 | 107 | 95 | 108 | 95 |
| FEC-3 | 281 | 359 | 0.995 | 0.977 | 0.998 | 0.987 | 146 | 81 | 170 | 95 |
| FEC-4 | 110 | 86 | 0.802 | 0.402 | 0.703 | 0.067 | 67 | 44 | 53 | 35 |
| FCC-1 | 102 | 145 | 0.999 | 0.340 | 1.000 | 0.998 | 86 | 76 | 122 | 108 |

* 250°F/WET CONDITIONS

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0.067 at DUL. The ETW B-basis allowable exceed DLL for all specimens with the exception of FEC-4. The ETW B-basis allowable for FEC-4 specimens is 53% DUL and A-basis is 35% DUL.

From the results of the static data evaluation, it can be concluded that the composite fuselage structures in Reference 5 demonstrated very high RTA and ETW static reliability for the typical in-plane composite failures. For specimens that exhibited out-of-plane structural failure modes, such as specimens FEC-3 and FEC-4, the static strength scatter is higher and the test data of Reference 5 demonstrated B-basis reliability at DLL under RTA conditions.

Table 28 summarizes the one lifetime fatigue reliabilities for all fuselage test specimens. The fatigue reliabilities calculated using the load enhancement approach assume that fatigue failures occurred after two lifetimes. However, all of the fuselage specimen fatigue tests were run-outs at the two fatigue lifetimes, which were subsequently residual static strength tested. Thus, the fatigue reliabilities calculated by the load enhancement approach are conservative. Fatigue reliabilities calculated using the ultimate strength approach are estimates based on static failure loads. The fatigue reliabilities calculated using the residual strength approach provide the best estimates of one lifetime fatigue reliabilities because it accounts for the specimen residual strength after two lifetimes of fatigue loading.

The fatigue reliabilities calculated using the residual strength approach exceed B-basis for all specimens. As shown in Table 28, the one lifetime fatigue reliability for all specimens, except FEC-4, exceed B-basis using all three approaches. The low fatigue reliability of FEC-4 calculated using load enhancement factor (LEF) approach is a reflection of the low LEF (1.08 at RTA and 0.94 at ETW) used in the test. The low fatigue reliability for FEC-4 calculated using the ultimate strength approach is because of the low average strength and high scatter. The low fatigue reliability for FCC-1 calculated using the ultimate

TABLE 28. SUMMARY OF ONE LIFETIME FATIGUE RELIABILITIES FOR FUSELAGE TEST SPECIMENS.

| SPECIMEN I.D. | FATIGUE TEST LOAD ENHANCEMENT FACTOR | | FATIGUE RELIABILITY AT 1 LT | | | | | |
|---|---|---|---|---|---|---|---|---|
| | | | RTA | | | ETW | | |
| | RTA | ETW | LEF | US | RSS | LEF | US | RSS |
| FE-1 | 2.81 | 2.13 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 | 1.000 |
| FEC-2 | 1.50 | 1.50 | 0.998 | 0.994 | 1.000 | 0.998 | 0.994 | 1.000 |
| FEC-3 | 2.38 | 2.38 | 0.974 | 0.996 | 0.993 | 0.970 | 0.999 | 0.979 |
| FEC-4 | 1.08 | 0.94 | 0.578 | 0.826 | 0.941 | 0.319 | 0.508 | 1.000 |
| FCC-1 | 1.34 | -- | 0.983 | 0.001 | 1.000 | -- | 0.903 | -- |

LEF = LOAD ENHANCEMENT FACTOR APPROACH
US = ULTIMATE STRENGTH APPROACH
RSS = RESIDUAL STRENGTH APPROACH

strength approach is simply a reflection of the fixture failure of the test.

The following conclusions can be drawn from the fuselage data evaluation:

1.  A high degree of static strength reliability has been demonstrated by the fuselage test data with typical composite in-plane failures.

2.  Static strength reliability for structures with out-of-plane failures may be inadequate because of the high strength scatter.

3.  One lifetime fatigue reliabilities demonstrated by the fuselage test data were adequate.

# SECTION 5

## CERTIFICATION TESTING RECOMMENDATIONS

### 5.1  Testing Requirements

Specific certification testing requirements are detailed in the following paragraphs. The design allowable tests are coupon level tests. The results of these tests are used to establish allowable strengths. Test variables at this level should include loading mode (tension, compression, shear) and environment (temperature, moisture). The design development testing should be planned based on the building block approach. Structural element, element combinations and subcomponents are tested to verify the design concepts. Sufficient number of tests should be conducted to identify different failure modes. The worst case environment should be included in the test plan. Full-scale tests are used to verify the overall reliability of the structure and to idenfiy any unanticipated hot spots. Separate tests should be conducted for static strength and fatigue life.

### 5.1.1  Design Allowables

The purpose of design allowable tests is to evaluate the material scatter and to establish strength and life parameters for structure design. Because composites are environmental sensitive, design allowables should be obtained for the entire range of the environmental service envelope of an aircraft. Statistical analysis methods must be used to compute the design allowables. Sufficient number of tests are required to sustain the specific level of confidence of the allowables. Key elements of test planning and data analysis are discussed below.

### 5.1.1.1  Static Design Allowables

Static design allowables include static tension, compression and shear strengths of composite materials. The Weibull distribution is recommended to describe the test data variation. The maximum likelihood estimate (MLE) method discussed in Sec-

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tion 2 of Volume I is recommended for computation of the Weibull parameters. The shape and scale parameters ($\alpha$ and $\beta$) are estimated by solving equations (5) and (6) and the A- or B-Basis allowables are computed from equation (9) in Volume I. The computer program WEIBULL in the Appendix is written for the computation of design allowables from test data.

From the extensive data analysis, documented in Volume I, the modal value of $\alpha=20.0$ is recommended for determination of the static strength allowables for typical graphite/epoxy composites. Based on this value of $\alpha$, with 15 data points, the B-basis to mean strength ratio is 0.901 and the A-basis to mean strength ratio is 0.801. For a sample size of 30, the B-basis and A-basis knockdown factors are 0.905 and 0.805, respectively. The influence of sample size on the B-basis static strength was shown in Figure 7. It can be seen from the figure that for typical graphite/epoxy composites, $\alpha=20.0$, the B-basis knockdown factor exceeds 0.85 for any sample size. Therefore, for typical graphite/epoxy composites, a knockdown factor of 0.8 times mean strength (as sometimes used in the industry) is a conservative approach to determination of B-basis allowable. Figure 7 also indicates that for the Weibull shape parameter of $\alpha<10.0$, the knockdown factor of 0.8 becomes an unconservative estimate of the B-basis allowable.

In planning a design allowable testing, it is important that sufficient number of tests be conducted to generate meaningful statistical parameters. In general, the number of specimens required depends on the scatter of the data. The higher the data scatter, the larger number of specimens are required. Based on the scatter analysis performed in Task I, the number of specimens recommended for B-basis allowables is 15 and for A-basis is 30. This is because within the range of $\alpha$ for typical composites ($\alpha=20 - 30$) the B-basis knockdown factor remains approximately constant for sample size greater than 15. The A-Basis knockdown factor stabilizes for sample sizes larger than 30. The B- and A-basis knockdown factors at these recommended samples sizes for typical $\alpha$ values are provided in Table 29.

**TABLE 29.  DESIGN ALLOWABLE TO MEAN STRENGTH RATIO FOR THE RECOMMENDED SAMPLE SIZE.**

| $a$ | B–BASIS n=15 | A–BASIS n=30 | |
|---|---|---|---|
| 10.0 | 0.808 | 0.645 | |
| 12.0 | 0.838 | 0.695 | |
| 14.0 | 0.860 | 0.733 | |
| 16.0 | 0.877 | 0.762 | |
| 18.0 | 0.890 | 0.786 | |
| 20.0 | 0.901 | 0.805 | MODAL $a$ |
| 22.0 | 0.909 | 0.821 | |
| 24.0 | 0.917 | 0.835 | |
| 26.0 | 0.923 | 0.847 | |
| 28.0 | 0.928 | 0.857 | |
| 30.0 | 0.933 | 0.866 | |

The design allowable tests should be planned to develop the strength to temperature envelope relationship for the full range of the service temperature of aircraft. The moisture level for the test specimens should be either end-of-lifetime level or the maximum level in the design lifetime, which ever is higher. The tests should also provide data for each failure mode. Tension, compression and shear strengths test should be conducted at each environment. The design allowable tests should be conducted at both lamina and laminate levels. The purpose of the lamina test is to establish the mechanical properties such as Young's moduli, shear modulus and Poisson ratio. These tests should include longitudinal and transverse tension and compression and shear tests. At the laminate level, two laminates representing the practical fiber dominated and matrix dominated lay-ups should be selected. The test specimens should include unnotched specimens to determine the laminate design strain. Other tests such as open and filled holes, bearing and bolt bearing by-pass should also be included.

Typical design allowable test matrices are shown in Tables 30 and 31. These tables show that design allowables will be obtained directly from the test data for the key test environments. For the secondary test conditions, the allowables can be computed using the statistical parameters established from the key tests.

The results of the design allowable tests should be analyzed individually for each key test condition. The design allowables are recommended to be generated using the worst scatter parameter (lowest $\alpha$) among all the tests. This procedure is shown schematically in Figure 98.

An alternate approach for determining design allowables is to use the pooling technique discussed in Section 2 of Volume I. The joint Weibull analysis can be used to pool the test data. The joint Weibull shape parameter is then used to compute the allowables. The computation procedure is automated in the computer program WEIBJNT given in the Appendix. The results obtained from the joint Weibull analysis are less conservative

### TABLE 30. LAMINA STATIC STRENGTH ALLOWABLE TEST MATRIX.

| TEST CONDITION | TENSION | | COMPRESSION | | SHEAR |
|---|---|---|---|---|---|
| | 0° | 90° | 0° | 90° | ±45° |
| LTD | 15(30) | 15(30) | 6 | 6 | 6 |
| RTW | 15(30) | 15(30) | 15(30) | 15(30) | 15(30) |
| ETW-1 | 6 | 6 | 6 | 6 | 6 |
| ETW-2 | 6 | 6 | 6 | 6 | 6 |
| ETW-3 | 6 | 6 | 6 | 15(30) | 15(30) |

NUMBERS IN ( ) ARE FOR A-BASIS ALLOWABLES.
ETW-3 IS THE WORST CASE ENVIRONMENT.

## TABLE 31. LAMINATE STATIC STRENGTH ALLOWABLE TEST MATRIX.

| TEST TYPE | LOADING MODE | | LAMINATE | TEST ENVIRONMENT | | | | |
|---|---|---|---|---|---|---|---|---|
| | | | | LTW | RTW | ETW-1 | ETW-2 | ETW-3 |
| UNNOTCHED | TENSION | | FD | 6 | 6 | – | – | 6 |
| | | | MD | 6 | 6 | – | – | 6 |
| | COMPRESSION | | FD | – | 6 | – | 6 | 6 |
| | | | MD | – | 6 | – | 6 | 6 |
| OPEN HOLE | TENSION | | FD | 15(30) | 15(30) | 6 | 6 | 6 |
| | | | MD | 6 | 15(30) | 6 | 6 | 6 |
| | COMPRESSION | | FD | – | 6 | – | 6 | – |
| FILLED HOLE | COMPRESSION | | FD | 6 | 15(30) | 6 | 6 | 15(30) |
| | | | MD | 6 | 6 | 6 | 6 | 6 |
| BEARING | TENSION | | FD | 6 | 6 | 6 | 6 | 15(30) |
| | | | MD | 6 | 15(30) | 6 | 6 | 15(30) |
| BOLT BEARING BY-PASS | TENSION | 20% LT | FD | 15(30) | 15(30) | – | – | – |
| | | | MD | 15(30) | 15(30) | – | – | – |
| | | 30% LT | FD | 15(30) | 15(30) | – | – | – |
| | | | MD | 15(30) | 15(30) | – | – | – |
| | | 50% LT | FD | 15(30) | 15(30) | – | – | – |
| | | | MD | 15(30) | 15(30) | – | – | – |
| | COMPRESSION | 20% LT | FD | – | 15(30) | – | – | 15(30) |
| | | | MD | – | 15(30) | – | – | 15(30) |
| | | 30% LT | FD | – | 15(30) | – | – | 15(30) |
| | | | MD | – | 15(30) | – | – | 15(30) |
| | | 50% LT | FD | – | 15(30) | – | – | 15(30) |
| | | | MD | – | 15(30) | – | – | 15(30) |

NOTE: FD = Fiber Dominate Laminate
MD = Matrix Dominate Laminate
ETW-3 is the Worst Case Environment
Numbers in ( ) are for A-Basis Allowables

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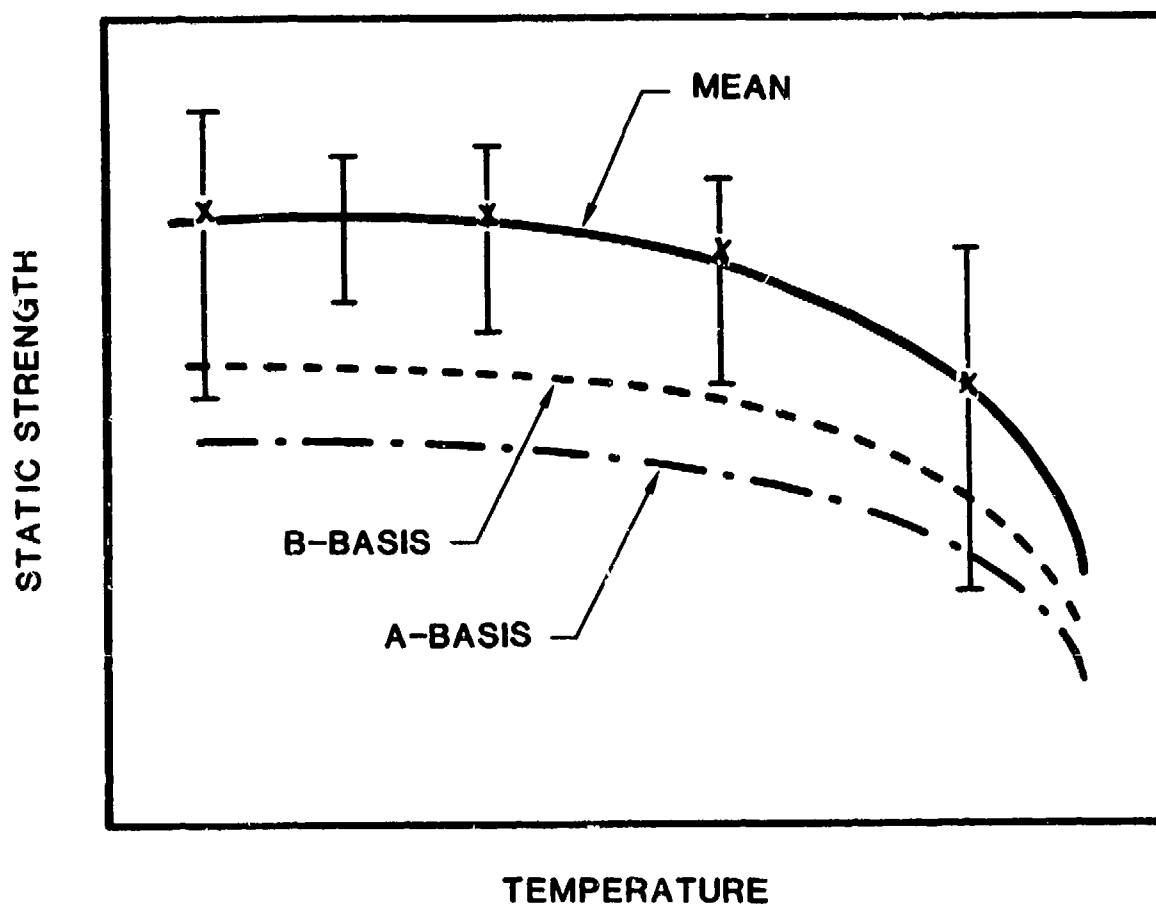
FIGURE 98. DETERMINATION OF DESIGN ALLOWABLES
OVER A RANGE OF TEMPERATURE.

than those obtained from the worst case scatter analysis method. However, because the inherent assumption in the analysis that the scatter parameter in each data group is approximately equal, this method must be applied with care. A significance test of the equality of $\alpha$'s is recommended prior to the application of the joint Weibull analysis. The method of statistical significant test was discussed in Section 2 of Volume I.

The modal value of Weibull shape parameter was determined based on a large amount of static test data over various test parameters. The value of $\alpha(20.0)$ is recommended to be used in computing the design allowables whenever the $\alpha$ value obtained from the allowable tests is higher than 20. When the value of $\alpha$ for the test data is below 20, worst case $\alpha$ should be used.

It is also recommended that the structural response variation (SRV), discussed in Section 3, is incorporated in the computation of design allowables. This can be accomplished by including equation (26) in the design allowable computation. The computer program BSRV or CSRV given in the Appendix can be used for such computations. The incorporation of SRV in the allowable computation will pose an additional penalty to the strength. However, within the range of typical SRV observed in Reference 5, the A- and B-basis knockdown factors are reduced only by a small amount. For a sample size of 15 with the average SRV ($\alpha_{SRV} = 17.0$), the B-basis knockdown factor is reduced from 0.901 to 0.872 for static strength $\alpha$ of 20.0. For the B-basis value of SRV ($\alpha_{SRV} = 13.0$) the knockdown factor reduced from 0.901 to 0.858. With a sample size of 30, the A-basis knockdown factor is reduced from 0.805 to 0.777 when $\alpha_{SRV} = 17.0$ and to 0.762 when $\alpha_{SRV} = 13.0$. The A- and B-basis knockdown factors for $n = 1$ at various values of $\alpha_{SRV}$ are shown in Figure 99.

## 5.1.1.2 Fatigue Design Allowables

The fatigue design allowables may be determined by the load factor approach, life factor approch or the ultimate strength approach. The individual or joint Weibull analyses are recommended for computation of design allowables. These approaches for fatigue allowable determination are schematically
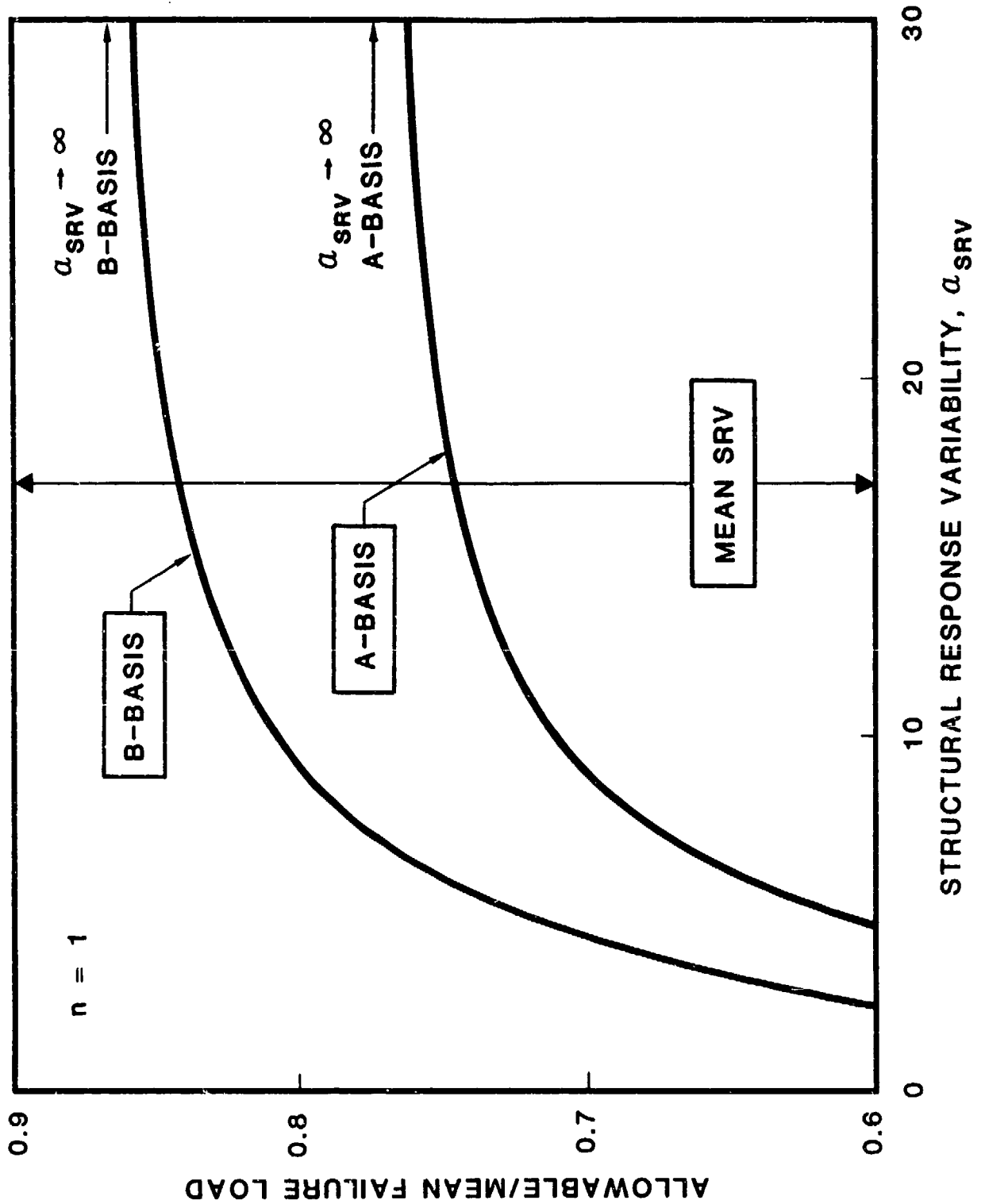
FIGURE 99. INFLUENCE OF SRV ON THE A- AND B-BASIS KNOCKDOWN FACTOR.

shown in Figure 100. Both the load factor approach and the life factor approach require the computation of design fatigue stress level for B- or A-basis fatigue life at one lifetime. For these approaches the joint Weibull analysis is used to define the B- or A-basis stress-life curve. The fatigue allowable strength is then defined as the stress level on the B- or A-basis curve at one fatigue lifetime. The ultimate strength approach is a more conservative approach. In this approach, the fatigue allowable strength is defined as the B- or A-basis stress level below which no fatigue failure will occur. The value $\sigma_{TH}^{B}$ shown in Figure 100 is the B-Basis fatigue allowable strength. Although the ultimate strength approach is more conservative, it is recommended that this approach be used to define the fatigue allowable. This is because the flatness of the S-N curve for typical composites, and the high scatter observed in fatigue test data. The value of $\sigma_{TH}^{B}$ or $\sigma_{TH}^{A}$ is more clearly defined as compared to $\sigma_{F}^{B}$ or $\sigma_{F}^{A}$. Furthermore, for typical fighter aircraft fatigue load spectra the penalty imposed by this approach is negligible.

The data analysis in Task I have shown that the fatigue life scatter has a modal $\alpha$ factor of 1.25 for typical graphite/epoxy composites. This indicates that fatigue life scatter for composites is considerably higher than that of aluminum. The modal value of $\alpha$ for aluminum under spectrum loading is 7.5. The B-basis to mean life ratio for composites with a sample size of 15 is 0.131 as compared to 0.750 for aluminum.

In planning the fatigue allowable tests, the main consideration is the test environment. The test environment depends on the relationship between the load/temperature spectrum and the MOL. The recommended approach is to use simple conservative constant temperature tests with a constant moisture level. The stress levels used in the fatigue tests should be selected so that the $\sigma_{TH}$ can be established. For typical graphite/epoxy composites under typical fighter aircraft spectra, the threshold stress level would be approximately 60% of the mean static strength. This would require a minimum of four stress levels for each test condition. From these considerations, using the same
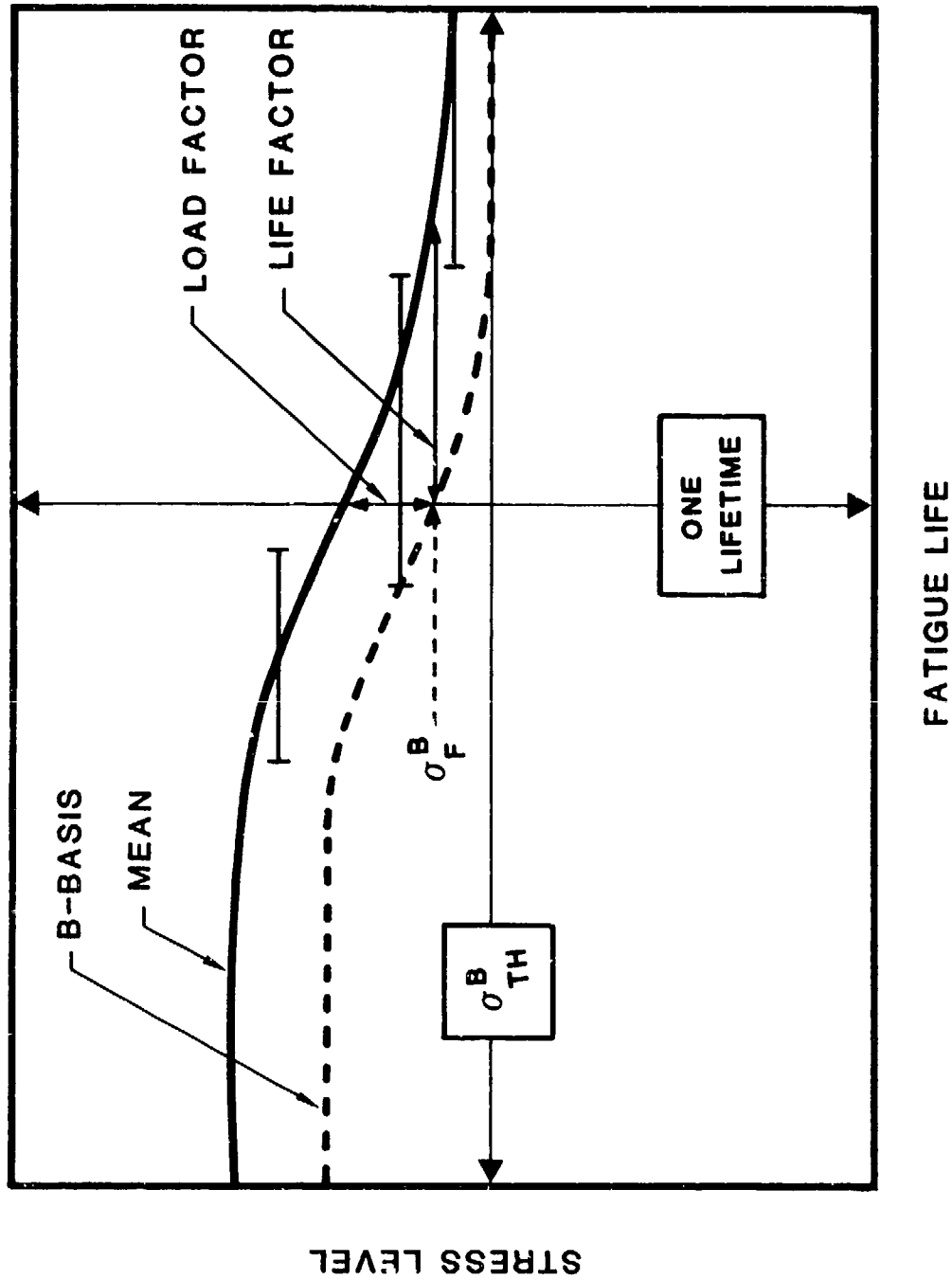
FIGURE 100. FATIGUE ALLOWABLE APPROACHES.

number of specimens required for the static allowable tests (15 for B-basis and 30 for A-basis) a large test matrix would result. However, as discussed in Volume I, the fatigue life scatter does not depend on the stress level for a given test condition. Therefore, the pooling techniques for statistical data analysis are justified. The number of tests at each test condition can therefore be reduced. The recommended number of tests for each test condition is 6 for B-basis and 10 for A-basis. A typical test matrix is shown in Table 32.

It is recommended that all fatigue tests are tested until fatigue failure occurs, except at the lowest stress level. At this stress level, because the fatigue threshold is approached, long life is expected. To reduce the test time, fatigue tests may be censored at a specified lifetime. Based on a life scatter of $\alpha = 1.25$, the test should last for a minimum of 8 lifetimes for B-basis tests and 50 lifetimes for A-basis.

As in the calculation of static strength allowables, it is recommended that SRV be incorporated in the fatigue allowable computations. Because of the large scatter observed in the fatigue life data, the life factor would be impractical for this purpose. The penalty on the load factor imposed by incorporation of SRV is small. The load factor approach is therefore recommended. Typical load enhancement factors required for graphite/epoxy composites ($\alpha_S=20.0$ and $\alpha_L=1.25$) are shown in Table 33. It can be seen from the table that for the average value of SRV ($\alpha_{SRV}=17.0$) the required load factor is increased by approximately 5% for a B-basis allowables. The increase is approximately 10% for A-basis allowables. It may be noted that because of high scatter in fatigue life for composites, B-basis allowable is the most appropriate statistic for fatigue design. This is also consistent with the approach adopted for metal structures, where the average fatigue property with a safety factor of two or four is used for fatigue design.

## 5.1.2    Design Development Testing

A building block approach to design development testing is essential for the certifciation of composite structures. This

## TABLE 32. RECOMMENDED FATIGUE ALLOWABLE TEST MATRIX.

| STRESS LEVEL | TENSION | | | COMPRESSION | | |
|---|---|---|---|---|---|---|
| | RTA | ETW | | RTA | ETW | |
| | NO HOLE | NO HOLE | HOLE | NO HOLE | NO HOLE | HOLE |
| $\sigma_1$ | 6(10) | 6(10) | 5 | 6(10) | 6(10) | 5 |
| $\sigma_2$ | 6(10) | 6(10) | 5 | 6(10) | 6(10) | 5 |
| $\sigma_3$ | 6(10) | 6(10) | 5 | 6(10) | 6(10) | 5 |
| $\sigma_4$ | 6(10) | 6(10) | 5 | 6(10) | 6(10) | 5 |

NOTE: Stress Levels are the Maximum Spectrum Stress.

Numbers in ( ) are the Required Number of Specimens for A-Basis Allowable, Other Numbers are for B-Basis Allowable.

## TABLE 33. TYPICAL LOAD FACTORS.

| FATIGUE LIFE (Spectrum Lifetimes) | REQUIRED LOAD FACTOR | | | | | |
| | NO SRV | | $a_{SRV} = 17.0$ (Mean SRV) | | $a_{SRV} = 13.0$ (B-Basis SRV) | |
| | A | B | A | B | A | B |
| 0.60 | 1.338 | 1.197 | 1.466 | 1.257 | 1.556 | 1.291 |
| 0.75 | 1.305 | 1.167 | 1.430 | 1.226 | 1.517 | 1.258 |
| 1.00 | 1.282 | 1.146 | 1.404 | 1.204 | 1.490 | 1.236 |
| 1.50 | 1.250 | 1.117 | 1.369 | 1.174 | 1.453 | 1.205 |
| 2.00 | 1.227 | 1.097 | 1.344 | 1.153 | 1.427 | 1.183 |
| 2.50 | 1.210 | 1.082 | 1.326 | 1.137 | 1.407 | 1.167 |
| 3.00 | 1.197 | 1.070 | 1.311 | 1.124 | 1.391 | 1.154 |
| 3.50 | 1.185 | 1.060 | 1.298 | 1.113 | 1.378 | 1.143 |
| 4.00 | 1.175 | 1.051 | 1.287 | 1.104 | 1.367 | 1.133 |
| 4.50 | 1.167 | 1.043 | 1.278 | 1.096 | 1.357 | 1.125 |
| 5.00 | 1.159 | 1.036 | 1.269 | 1.089 | 1.348 | 1.118 |

NOTE:  A for A-Basis with n = 10

B for B-Basis with n = 6

is because of the inherent sensitivity of the composite structure to out-of-plane loads and their multiplicity of potential failure modes. Details of the building block approach are discussed in Reference 13 and 21. The essence of the building block approach for composites is as follows. First, use the design/analysis of the aircraft structure to select critical areas for test verification. Second, determine the most strength-critical failure mode for each design feature. Third, select the test environment which will produce the strength critical failure mode. Special attention should be given to matrix sensitive failure modes (such as compression and bondline) and potential stress "hot spots" caused by out-of-plane loads. Following selection of the critical failure modes, a series of specimens is designed, each one to simulate a single failure mode. The specimens will generally be low complexity specimens. However, the crux of the building approach is to also design test specimens which simulate progressive design complexity. In this way, multiple potential failure modes are simulated.

This building block method to design development testing provides a step-by-step approach to composite design development testing which has several advantages.

- The influence of the environment on individual failure mode is determined.

- The interaction of failure modes is established from the known behavior of individual failure modes.

- Scale-up effect is determined from data on smaller scale specimens.

- "Hot spots" induced in complex structure can be analyzed relative to the known behavior of smaller specimens.

Several factors determine the test complexity of composite design development tests. These are: structural geometry complexity, hygrothermal environment simulation, fatigue load spectrum simulation and mixed composite/metal structure.

The levels of complexity in the design development testing should be functions of the design feature being validated and the predicted failure mode. Special attention should be given to correct failure mode simulation since failure modes are frequently dependent on the test environment. In particular, the influence of complex loading on the local stress at a given design feature must be evaluated. In composites, out-of-plane stresses can be detrimental to structural integrity and, therefore, require careful evaluation.

The sensitivity of composite matrix dominated failure modes to the aircraft hygrothermal environment makes environmental test simulation a key issue. Environmental test simulation should be considered separately for static and fatigue testing. However, the static test philosophy will form an integral part of the overall certification philosophy. The static and fatigue testings are discussed in the following paragraphs.

5.1.2.1   Static Tests

The philosophy for design development testing should be that the test environment used is the one that produces the failure mode which gives the lowest static strength. That is the worst case environment, or the temperature associated with most critical load should be used.

The extent of the static test effort will be different from aircraft to aircraft and also from component to component. The number of replicates for each test should be sufficient to identify the critical failure mode and provide a reasonable estimate of the mean strength of the element. The test effort should be concentrated on the most critical design feature of the structure. The number of replicates should be increased for the critical design features. A cost trade-off is usually involved in deciding the levels of complexity and the number of replicates.

If mixed failure modes are observed in a certain specimen type, more tests are required to establish the worst failure mode and the associated mean strength. An example of

mixed failure mode in the design development testing of Reference 5 was discussed in Section 4.3. Two failure modes were observed in the Wing Outboard Fuel Bay Subcomponent (WS-1) tests under 250°F/wet condition. The observed failure modes were upper skin failure and intermediate spar/lower skin cocured joint failure. For WS-1 the intermediate spar/lower skin cocured joint failure mode gave the lower failure load. Thus, the mean strength associated with this failure mode would need to be established.

An example of the building block approach for specimen complexity was given in Figure 57, which shows the approach used for the wing structure in Reference 5. There, the wing structure was broken down into critical areas. Each critical area was simulated in a test specimen whose complexity is governed by the necessity to simulate the predicted failure mode(s). Particular attention was given to matrix critical failure modes. The following recommendations are made for specimen complexity simulation in design development testing:

1.  Use the design/analysis of the aircraft structure to select critical areas for test verification.

2.  Specimen complexity should be controlled by the requirement to simulate the correct (full-scale structure) failure mode(s) in the specimen.

3.  Special attention should be given to matrix sensitive failure modes, such as compression, bondline and hole wear.

4.  Potential "hot spots" caused by out-of-plane loads should be carefully evaluated.

5.1.2.2  Fatigue Tests

The environmental complexity necessary for fatigue design development testing will depend on the aircraft hygrothermal history. Three factors must be considered. These are: structural temperature for each mission profile, the load/temperature relationships for the aircraft, and the moisture content as a function of the aircraft usage and structure thickness. In order to obtain these data, it is necessary to derive

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the real time load-temperature profiles for each mission in the aircraft's history. These relationships will have a significant influence on the environmental fatigue test requirements.

An example of this approach was given in Reference 5. The aircraft temperature spectrum and load factor/temperature relationships derived from the mission profiles were shown in Figures 20 and 21. These relationships strongly depend on the aircraft type, configuration and mission requirements and must be carefully developed on a case by case basis. The structural material should be selected to meet these mission requirements without exceeding the MOL. If this is accomplished, hot/wet testing would not be required. Material sele tions which lead to significant environmental fatigue test requirements should be a last resort.

In composite materials, no significant load sequence effect on fatigue life has been observed. However, studies on load spectrum variations have shown that composites are extremely sensitive to variation in the number of high loads in the fatigue spectrum. In contrast, truncation of low loads does not significantly affect fatigue life. Therefore, the following recommendations are made for load spectrum simulation in composite fatigue testing:

1.  High loads in the fatigue spectrum must be carefully simulated.

2.  Low loads (<30 percent limit load stress) may be truncated to save test time without significantly affecting fatigue life.

3.  Fatigue testing of mixed metal/composite structure may introduce conflicting requirements and should be evaluated on an individual basis.

As was discussed in Section 4, the use of fatigue test data to verify fatigue life on subcomponents require long test duration because of the high fatigue life scatter observed in composite structures. The load enhancement factor approach or the ultimate strength approach is recommended in planning the

fatigue design development testing. In the load enhancement factor (LEF) approach, the B- or A-basis LEF is applied to the design load spectrum. A two lifetime fatigue test plus static residual strength test is recommended. Typical LEF for two lifetime tests is approximately 1.15 for a B-Basis reliability with average SRV considered (see Table 33).

In applying the ultimate strength approach, the maximum spectrum load should be kept below the B- or A-basis fatigue threshold. The fatigue life requirements are verified from the static test results. No fatigue tests are required for this approach. This approach provides a conservative estimate of the fatigue reliability.

An alternate for environmental fatigue tests is to use increased loads at RTA condition to account for the environmental effects. An environmental knockdown factor, as discussed in Section 2.1 together with LEF is applied to the fatigue spectrum for RTA fatigue tests. The environmental knockdown factor can be calculated from the results of the design allowable tests. This approach eliminates the environmental fatigue tests. However, in some cases a high load factor may be required and result in a quasi-static failure at the first appearance of the maximum spectrum load. For example, the environmental knockdown factor obtained in Reference 5 for the 242°F/wet condition was 1.31. The B-Basis two lifetime test LEF at the mean SRV is 1.153. From these values, the resulting load factor is 1.51. Therefore, this approach is feasible only if the maximum design spectrum load is below 0.66 of the mean static strength, otherwise quasi-static failure will occur at the peak fatigue load during the RTA fatigue tests. The fatigue reliability will not be verified when this happens.

The number of replicates to be used in the fatigue design development testing should be determined using the same philosophy as in the static tests. A sufficient number must be used to verify the critical failure modes and to reasonably estimate the required fatigue reliability.

5.1.2.3    Mixed Composite/Metal Structures

An analysis of mixed composite/metallic full-scale testing in Reference 8 has shown that:

- The increased environmental sensitivity and material property variability in composites (compared to metals) can lead to inadequate assurance of the static strength integrity of composite structure.

- Because of the superior fatigue performance of composites, a mixed composite/metal structure fatigue will essentially interrogate only the metal structure. Thus, any potential "hot spots" in the composite structure may not be found.

Because of the potential inadequacy of full-scale tests on mixed composite/metal structure and also the natural reluctance to overdesign metal parts in a full-scale test structure, it will be necessary to validate the composite structure during the design development testing phase. However, the specimen complexity should be adequate to enable the performance of the full-scale structure to be correctly simulated. Validation of the composite structure using subcomponent tests can offer the following advantages:

- The components may be chosen for test purposes to interrogate the composite structure only.

- If environmental test conditions are required it will be easier and cheaper to achieve in a component.

- It may be possible to test more than one replicate and thus increase confidence in the data base.

- The results can be utilized in qualification of the full-scale structure.

In order that component tests achieve their objective, great care must be exercised in getting the boundary conditions correct. In addition, eliminating metal failure modes

by overdesign or replacement must be carefully evaluated so that relative effects such as differential thermal expansion are not masked.

It is concluded, therefore, that certification of combined composite/metal structures offer special problems which must be addressed carefully in the design analysis and test phase.

An alternative approach for the certification of mixed structures is the change in spectrum approach, discussed in Section 2.4. The application of this approach involves insertion of overload cycles in the fatigue spectrum in order to reduce the differences in life between metals and composites. The magnitude and frequency of the overload cycles can be determined from life analysis of the metal and composite parts. The introduction of the overload cycles permit demonstration of the B- or A-basis test reliability for both composite and metal parts in a mixed structure. This is achieved without making the test overly severe for metallic structures.

## 5.1.3 Full-Scale Testing

Following successful conclusion of the design development testing program, qualification culminates in full-scale static and durability tests.

## 5.1.3.1 Full-Scale Static Test

The full scale static test is the most crucial qualification test for composite structures for the following reasons. Secondary loads are virtually impossible to eliminate from complex built-up structure. Such loads can be produced by eccentricities, stiffness changes, discontinuities, fuel pressure loading and loading in the post-buckled range. Some of these sources of secondary loads are represented for the first time in the full-scale structural test article. These loads are not a significant design driver in metallic structures. However, the poor interlaminar strength of composites makes them extremely

susceptible to out-of-plane secondary loads. It is very important, therefore, to carefully account for these loads in the design of composite structures. Unfortunately, there is a general state of uncertainty as to the source, magnitude and effects of secondary loads in complex built-up full-scale composite structures. This has been confirmed by several documented examples of unanticipated secondary loads leading to unexpected failure modes in full scale composite structural static tests.

Work in Reference 12 has shown that the RTA static test plays the most significant role in revealing unexpected hot spot failures from secondary out-of-plane loads. A room temperature enviroment is therefore recommended for full scale static test, which should be conducted to failure.

In addition, a detailed correlation in terms of measured load and strain distributions, structural analysis data and environmental effects between the design development and full scale test data will be necessary to provide assurance of composite static strength. Static test environmental degradation must be accounted for separately either by adverse condition testing, by additional test design factors or by correlation with environmental design development test data.

5.1.3.2    Full-Scale Durability Test

Current practice for metallic structures is a two lifetime test using the design load spectrum under RT/ambient test conditions.

The work in Reference 12 and other USAF sponsored programs have shown that composites possess excellent durability. In particular, the extensive data base developed in Reference 12 showed that composite structures, which demonstrated adequate static strength, were fatigue insensitive.

Therefore, it is recommended that no durability full scale test is required for all composite structures or mixed

composite/metal structures with non-fatigue critical metal parts, provided the design development testing and full scale static test are successful. For mixed structure, with fatigue critical metal parts, a two lifetime ambient test will be required to demonstrate durability validation of the metal parts.

5.2     Test Result Interpretation

The certification testing methodology developed in Section 3 is recommended to be used in test result interpretation. The static strength and fatigue life reliabilities should be assessed based on the test data. The procedure for static and fatigue data evaluation was demonstrated in Section 4. This procedure is summarized below.

1.    Establish baseline reliability parameters: These parameters are determined from the results of the allowable tests or obtained from the existing data base. The parameters include the static and fatigue life Weibull scatter parameters ($\alpha_S$ and $\alpha_L$), the structural response variability ($\alpha_{SRV}$) and environmental knockdown factor ($k_{ENV}$).

2.    Determine the static strength reliability: Both the RTA and worst case environment strength reliabilities should be computed at both DLL and DUL. The reliability is computed based on the test failure load. The reliability calculation methods described in Sections 3 and 4 or the computer programs BSRV or CSRV can be used for any composite structures. However, case should be exercised in computing reliabilities for structures with multiple failure modes because work in Reference 5 has shown that these failure modes can exhibit high scatter in static strength. Reliability analysis of these failures should be conducted on a case-by-case basis using experimentally determined scatter factor ($\alpha_S$) for the particular failure mode.

3. Determine the required test failure load: The required (minimum) test failure load for each test environment for a specified static strength reliability can be determined using the method developed in Section 3 and 4.

4. Determine the one lifetime fatigue reliability from static strength data: The one lifetime fatigue reliability can be computed using the ultimate strength approach discussed in Section 4 (or computer program BSRV or CSRV). The reliability calculated from this approach is a conservative estimate of the fatigue life reliability. A low reliability based on this approach does not imply that the design feature has poor fatigue strength. It just means that no significant reliability can be inferred from the static strength data. The reliability depends on the maximum spectrum load, the spectrum characteristics, SRV and the static failure load.

5. Determine the fatigue life reliability, using the load enhancement factor approach: The load enhancement factor approach was described in Section 2.2 and the application of this approach, accounting for SRV, was demonstrated in Section 4. Where fatigue failure occurred during test, the actual fatigue life is used to directly assess the fatigue reliability. When fatigue test is terminated without fatigue failure, the test duration (e.g. two lifetimes) is used to conservatively estimate the fatigue reliability.

6. Determine the fatigue life reliability based on the residual strength approach: In the case that the test LEF is not sufficient and the fatigue test is terminated after a short duration, the fatigue reliability demonstrated by the LEF

approach would usually be low. This is typical for composite structures, because of the high fatigue life scatter. For situations like this, residual static strength test to failure is recommended to provide further information on fatigue reliability. The residual static strength data is then used to determine the fatigue life reliability. This approach was demonstrated in Section 4.

## SECTION 6

## SUMMARY AND CONCLUSIONS

6.1    Summary

The results of this research program are summarized below:

1.  An extensive data analysis has been conducted to establish the static strength and fatigue life data scatter for typical gi .phite/epoxy composites.

2.  Various approaches to composite structure certification were analytically evaluated. The capability, advantages and disadvantages of each approach were fully discussed.

3.  A methodology for certification testing of composite structures has been developed.

4.  The methodology was demonstrated on an existing composite structure. The static and fatigue reliabilities of the structure were assessed based on test data.

5.  A detailed certification testing approach was presented. Guidelines for test planning and test data interpretation were recommended.

6.2    Conclusions

The following conclusions may be drawn from the investigation undertaken in this program:

1.  The Navy approach to static strength certification is, in general, soundly based. In particular, it strikes a good balance between the technical requirements of the test and cost effectiveness. The main area of concern in this approach is the assumed ability to predict full-scale structure failure mode. Because of the inherent sensitivity of composite structures to out-of-plane

secondary loads, a certification process which assumes correct prediction of a full-scale structural failure mode must carry some degree of risk.

2. The high scatter of composite fatigue life data makes life factor approach for fatigue certification very difficult. A practical approach is to use the load enhancement factor approach. In principle, this is equivalent to the Navy Approach of applying severe spectrum in fatigue tests.

3. The incorporation of the structural response variability in the reliability calculations is important in reducing the possibility of "hot spots" failure.

4. Environmental sensitivity is a key issue in design of composite structures. This problem can be avoided by careful selection of materials so that the structure only operates within the material operating limit.

5. Success of certification testing depends on test planning and soundly based data interpretation method.

## REFERENCES

1.    Weinberger, R.A.; Somoroff, A.R.; and Riley, B.L.; "U.S. Navy Certification of Composite Wings for the F-18 and Advanced Harrier Aircraft," Paper No. 77-466 in AIAA Aircraft Composites: The Emerging Methodology for Structural Assurance, San Diego, California (March 1977).

2.    Murray, J.E., "F-18 Composites Development Tests," N00019-79-C-0044 (January 1981).

3.    Stenberg, K.V., et al., "YAV-8B Composite Wing Development," Volumes I-III N00019-76-C-0242, 0666 (October 1980).

4.    Riley, B.L.; Rintoul, R.D. and Roberts, L.A.; "Certification of Composite Structures," presented at the 5th DOD/NASA Conference on Fibrous Composites in Structural Design, New Orleans, Louisiana (January 1981).

5.    Whitehead, R.S. et al "Composite Wing Fuselage Program," Contract F33615-79-C-3203, Interim Reports 1-11, October 1979 - October 1984.

6.    Cardrick, A.W., "Proposals for a Common International Approach to Structural Airworthiness Certification," RAE (August 1983).

7.    Bohon, H.L., et al., "Ground Test Experience With Large Composite Structures For Commercial Transports," NASA TM 84627 (March 1983).

8.    Guyett, P.R. and Cardrick, A.W., "The Certification of Composite Airframe Structures," Paper No. 830/3 Aeronautical Journal (July 1980), pp. 188-203.

9.    Whitehead, R.S. and Deo, R.B., "Wing/Fuselage Program Durability Methodology," presented at 6th Mechanics of Composites Review, Dayton, Ohio, November 1980.

10.   Whitehead, R.S., "A Review of the Rationale for Durability Validation of Composite Structures," proceedings of the 5th DOD/NASA Conference on Fibrous Composites in Structural Design, New Orleans, Louisiana (January 1981).

11.   Whitehead, R.S. and Eves, J.J., "Durability Certification Data for Composite Structures," presented at the 6th DOD/NASA Conference on Fibrous Composites in Structural Design, New Orleans, Louisiana (January 1983).

12.   Whitehead, R.S.; Ritchie, G.L. and Mullineaux, J.L.; "Durability of Composites" presented at 9th Mechanics of Composites Review, Dayton, Ohio (October 1983).

13. Whitehead, R.S.; Ritchie, G.L. and Mullineaux, J.L.; "Quali-
    fication of Primary Composite Aircraft Structures" to be
    presented at USAF "Structural Integrity Conference" Macon,
    Georgia, 27-29 November 1984.

14. Sendeckyj, G.P., "Fitting Models to Composite Materials
    Fatigue Data," ASTM STP 734, 1981, p. 245-260.

15. Whitehead, R.S. and Schwarz, M.G., "The Role of Fatigue
    Scatter in the Certification of Composite Structures,"
    presented at ASTM Symposium on the Long Term Behavior of
    Composites, Williamsburg, Virginia (March 1982).

16. Badaliance, R and Dill, H.D., "Compression Fatigue Life
    Prediction Methodology for Composite Structures," NADC-
    83060-60 Volumes 1 and 2, September 1982.

17. Ratwani, M.M. and Kan, H.P., "Compression Fatigue Analysis
    of Fiber Composites," NADC-78049-60 (September 1979).

18. Ratwani, M.M. and Kan, H.P., "Development of Analytical
    Techniques for Predicting Compression Fatigue Life and
    Residual Strength of Composites," NADC-82104-60 (March
    1982).

19. Porter, P.G., "A Rapid Method to Predict Fatigue Crack
    Initiation," Report No. NADC-81010-60, February 1983.

20. Mullineaux, J.L., Whitehead, R.S. and Ritchie, G.L., "Full
    Scale Testing of Composite Wing Structure," Presented at the
    7th DOD/NASA Conference on Fibrous Composites in Structural
    Design, Denver, Colorado (June 1985).

21. Whitehead, R.S. and Deo, R.B., "A Building Block Approach to
    Design Verification Testing of Primary Composite Struc-
    tures," Paper No. 83-0947, Proceedings of the 24th
    AIAA/ASME/ASCE/AHS Structures, Structural Dynamics and
    Material Conference, Lake Tahoe, Nevada, May 1983, pp. 473-
    477.

APPENDIX

COMPUTER PROGRAMS

The computer programs used in the data analysis and reliability calculations are documented in this Appendix. The programs are written in FORTRAN language and are suitable to be used on the IBM personal computers. Five programs are included below. The theoretical backgrounds are presented in Section 2 of Volume I and Sections 2 and 3 in this volume. The program listings, input and output descriptions and sample problems are given in the following paragraphs.

A.1   Program "WEIBULL"

The program WEIBULL computes the maximum likelihood estimates (MLE) of Weibull shape and scale parameters. The program also computes the mean, standard deviation and coefficient of variation of the data set based on both normal and Weibull distributions. The A- and B-basis allowables are computed based on the MLE Weibull parameters. A $\chi^2$ goodness-of-fit test is also conducted in the program.

The required input to WEIBULL are:

1. A 32-character problem title (INAME)

2. Total number of data points (N)

3. Number of specimens that failed (NF)

4. Data set mode (MODE)   MODE $\leq$ 10 for normal operation mode.

5. Data value and data point ID

   ID = 1    specimen failed
   ID $\neq$ 1    specimen censored

   If N = NF, use ID = 1 for all data points. Maximum number of data points is 60.

6. Number of intervals for goodness-of-fit test (INT).

   The programm listing the input and output of a sample problem are given below.

```
C PROGRAM 'WEIBULL' MLE ESTIMATE OF WEIBULL PARAMETETS FROM DATA
      DOUBLEPRECISION B,Y,CHL,ALPHA,CHIS,GAMMA1,GAM1,GAM2,ARM1,ARM2,P,S
      DIMENSION ID(60),FR(60),RK(60),INAME(8)
      COMMON/GMA/B(101),Y(101)
      COMMON/CHI/CHL(15)
      OPEN(5,FILE='PSI.DAT')
      READ(5,*) (B(I),I=1,101)
      READ(5,*) (Y(I),I=1,101)
      READ(5,*) (CHL(I),I=1,15)
      PAL = -ALOG(0.99)
      PBL = -ALOG(0.90)
      SUM = 0.0
      SUD = 0.0
      KCOT = 0
      WRITE(*,20)
      READ(*,3) INAME
      WRITE(*,10)
   10 FORMAT(2X,'PLEASE ENTER TOTAL NUMBER OF SPECIMENS')
      READ(*,*) N
      WRITE(*,15)
   15 FORMAT(2X,'PLEASE ENTER NUMBER OF SPECIMEN THAT FAILED')
      READ(*,*) NF
   20 FORMAT(2X,'PLEASE ENTER DATASET NAME: 32 CHARACTERS')
    3 FORMAT(8A4)
      WRITE(*,30)
   30 FORMAT(2X,'PLEASE ENTER MODE CODE',
     +/2X,'IF MODE.GE.10 HALF SPECIMENS FAILED, N=NF')
      READ(*,*) MODE
      WRITE(*,41)
   41 FORMAT(2X,'SPECIMEN ID CODE: ID=1 SPECIMEN FAILED'
     +/20X,'ID.NE.1 TEST CENSORED')
      DO 35 J = 1,N
      WRITE(*,40) J
   40 FORMAT(2X,'ENTER STRENGTH VALUE AND ID FOR SPECIMEN',I3)
      READ(*,*) FR(J),ID(J)
   35 CONTINUE
      WRITE(*,50)
   50 FORMAT(2X,'ENTER NUMBER OF INTERVALS FOR GOODNESS TEST')
      READ(*,*) INT
      T = N
      N2 = N*2
      ALPHA = 10.0
      WRITE(*,60) INAME
      WRITE(*,65)
      DO 70 I=1,N
      IF(I.EQ.N) GOTO 70
      I1 = I+1
      DO 71 K = I1,N
      IF(FR(I).LE.FR(K)) GOTO 71
      FRT = FR(I)
      FR(I) = FR(K)
      FR(K) = FRT
      IDT = ID(I)
      ID(I) = ID(K)
      ID(K) = IDT
```

```
   71 CONTINUE
   70 CONTINUE
      FART = ALOG10(FR(1))
      IF(FART.LE.0.5) GOTO 270
      LFA = FART
      RA = 10.0**LFA
      GOTO 271
  270 RA = 1.0
  271 DO 75 I=1,N
      A = N-I+1
      RK(I) = A/(T+1.0)
      SUM = SUM+FR(I)
      SUD = SUD+FR(I)*FR(I)
      WRITE(*,80) FR(I),RK(I),ID(I)
   75 FR(I) = FR(I)/RA
      AVE = SUM/T
      VA = (SUD-T*AVE*AVE)/(T-1.0)
      STD = SQRT(VA)
      CV = STD/AVE
      IF(CV.LT.0.030) ALPHA = 75.0
   65 FORMAT(5X,'ORDERED STRENGTH      ASSIGNED PROBABILITY',4X,'CODE')
   80 FORMAT(5X,F12.5,9X,F12.5,12X,I2)
   60 FORMAT(2X,'WEIBULL ANALYSIS: ',8A4)
      KKT = 0
   85 P = 0.0
      R = 0.0
      S = 0.0
      KCOT = KCOT+1
      IF(KKT.GT.2) GOTO 200
      DO 90 I = 1,N
      ACR = 174.0/ALOG(FR(N))
      IF(ACR.LT.0.0) GOTO 95
      IF(ALPHA.LT.ACR) GOTO 95
      ALPHA = 75.0
      KKT = KKT+1
      GOTO 85
   95 P = P+FR(I)**ALPHA
      S = S+FR(I)**ALPHA*ALOG(FR(I))
      IF(ID(I).NE.1) GOTO 90
      R = R+ALOG(FR(I))
   90 CONTINUE
      IF(MODE.LT.10) GOTO 98
      P = 2.0*P
      S = 2.0*S
   98 ALPH1 = NF*P/(NF*S-P*R)
      DA = ABS(ALPH1-ALPHA)
      IF(DA.LE.0.0001) GOTO 100
      ALPHA = (ALPHA+ALPH1)/2.0
      IF(KCOT.GT.100) GOTO 100
      GOTO 85
  100 ALR = 1.0/ALPHA
      BETA = (P/NF)**ALR
      BETA = BETA*RA
      WRITE(*,105) ALPHA,BETA
      WRITE(*,110) KCOT
```

```
105 FORMAT(2X,'MAX. LIKELIHOOD ESTIMATE OF WEIBULL PARAMETERS:'
    A/2X,'SHAPE PARAMETER ALHPA =',F12.5
    B/2X,'SCALE PARAMETER BETA  =',F12.5)
110 FORMAT(5X,'NUMBER OF ITERATIONS KCOT = ',I5)
    CHSQ = CHIS(N)
115 BETAL = BETA/(CHSQ**ALR)
    ALL = BETAL*PAL**ALR
    BLL = BETAL*PBL**ALR
    WRITE(*,120) BETAL,ALL,BLL
120 FORMAT(2X,'95% CONF. LOWER BETA  =',F12.5,
    A      /2X,'A-BASIS ALLOWABLE     =',F12.5
    B      /2X,'B-BASIS ALLOWABLE     =',F12.5)
    WRITE(*,125) AVE,STD,CV
    ARM1 = 1.0+1.0/ALPHA
    ARM2 = 1.0+2.0/ALPHA
    GAM1 = GAMMA1(ARM1)
    GAM2 = GAMMA1(ARM2)
    XMEAN = BETA*GAM1
    COVS = (GAM2-GAM1*GAM1)/(GAM1*GAM1)
    COV = SQRT(COVS)
    STDW = COV*XMEAN
130 FORMAT(2X,'WEIBULL:')
    WRITE(*,130)
    WRITE(*,125) XMEAN,STDW,COV
135 FORMAT(6X,'STRENGTH VALUE    PROBABILITY OF SURVIVAL')
140 FORMAT(2F13.5)
150 FORMAT(4F12.5)
125 FORMAT(2X,'MEAN STRENGTH    FAVE  =',F12.5
    A      /2X,'STANDARD DEVIATION    =',F12.5
    B      /2X,'COEFF. OF VARIATION   =',F12.5)
    DO 155 M=1,N
155 FR(M) = RA*FR(M)
    DF = FR(N)-FR(1)
    WRITE(*,165)
165 FORMAT(/9X,'RIGHT END',5X,'OBSERVED',10X,'EXPECTED',10X,'CHI-SQ')
    DIN = DF/INT+0.005
    SUCHI = 0.0
    PENL = EXP(-(FR(1)/BETA)**ALPHA)
    ENRT = FR(1)+DIN
    JK = 1
    RIGHT = FR(N)+0.5*DIN
170 INK = 0
    DO 175 I = JK,N
    IF(FR(I).GT.ENRT) GOTO 180
175 INK = INK+1
180 JK = JK+INK
    PENR = (ENRT/BETA)**ALPHA
    PENR = EXP(-PENR)
    EO = (PENL-PENR)*N
    DEO = EO-INK
    CHI = DEO*DEO/EO
    SUCHI = SUCHI+CHI
    WRITE(*,185) ENRT,INK,EO,CHI
185 FORMAT(5X,F12.4,8X,I3,2(9X,F12.5))
    PENL = PENR
```

```
      ENRT = ENRT+DIN
      IF(RIGHT.GE.ENRT) GOTO 170
      NDOF = INT-3
      WRITE(*,190) NDOF,SUCHI
190 FORMAT(/2X,'AT ',I3,'DEGREES OF FREEDOM',
     A/2X,'THE CHI-SQUARE VALUE FOR GOODNESS OF FIT IS',2X,F12.5)
      FI = FR(1)-0.3*DF
      IF(FI.LE.0.0) FI=0.01*DF
      DD = DF/40.0
      WRITE(*,135)
      DO 195 I=1,65
      XX = FI+(I-1)*DD
      AR = (XX/BETA)**ALPHA
      YY = EXP(-AR)
195 WRITE(*,80) XX,YY
200 CONTINUE
      STOP
      END
      FUNCTION CHIS(N)
      DOUBLEPRECISION CHL,CHIS,BE,CL
      COMMON/CHI/CHL(15)
      AN = N
      BN = 2.0*AN
      IF(N.GE.15) GOTO 50
      CHIS = CHL(N)
      GOTO 60
50 BE = 1.0/(9.0*AN)
      CL = 1.0-BE+1.645*SQRT(BE)
      CHIS = CL*CL*CL
60 CONTINUE
      RETURN
      END
      FUNCTION GAMMA1(X)
      DOUBLEPRECISION A,B,X,Y,F,ARG,SLOP,GAMMA1
      COMMON/GMA/B(101),Y(101)
      ARG = X
      A = 1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
      GOTO 30
10 A = A/ARG
      ARG = ARG+1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      GOTO 30
20 ARG = ARG-1.0
      A = A*ARG
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
30 DO 40 I=1,101
      IF(B(I).GT.ARG) GOTO 50
40 CONTINUE
50 SLOP = (Y(I)-Y(I-1))/(B(I)-B(I-1))
```

```
      F = Y(I-1)+(ARG-B(I-1))*SLOP
      GOTO 60
110   F = 1.0
 60   GAMMA1 = F*A
      RETURN
      END
```

PSI.DAT  TAB VALUES OF CHI-SQUARE AND GAMMA FUNCTIONS
1.0,   1.01, 1.02, 1.03, 1.04, 1.05, 1.06, 1.07, 1.08, 1.09, 1.1,   1.11,
1.12, 1.13, 1.14, 1.15, 1.16, 1.17, 1.18, 1.19, 1.2,   1.21, 1.22, 1.23,
1.24, 1.25, 1.26, 1.27, 1.28, 1.29, 1.3,   1.31, 1.32, 1.33, 1.34, 1.35,
1.36, 1.37, 1.38, 1.39, 1.4,   1.41, 1.42, 1.43, 1.44, 1.45, 1.46, 1.47,
1.48, 1.49, 1.5,   1.51, 1.52, 1.53, 1.54, 1.55, 1.56, 1.57, 1.58, 1.59,
1.6,   1.61, 1.62, 1.63, 1.64, 1.65, 1.66, 1.67, 1.68, 1.69, 1.7,   1.71,
1.72, 1.73, 1.74, 1.75, 1.76, 1.77, 1.78, 1.79, 1.8,   1.81, 1.82, 1.83,
1.84, 1.85, 1.86, 1.87, 1.88, 1.89, 1.9,   1.91, 1.92, 1.93, 1.94, 1.95,
1.96, 1.97, 1.98, 1.99, 2.,
1.,      .99433, .98884, .98355, .97844, .9735,   .96874, .96415, .95973,
2.9955, 2.372,  2.09867, 1.93838, 1.8307,  1.75217, 1.69179, 1.6435,
1.60383, 1.5705, 1.542,  1.51729, 1.49558, 1.47632, 1.4591

```
WEIBULL
PLEASE ENTER DATASET NAME: 32 CHARACTERS
WEIBULL ANALYSIS: SAMPLE PROBLEM
PLEASE ENTER TOTAL NUMBER OF SPECIMENS
          10
PLEASE ENTER NUMBER OF SPECIMEN THAT FAILED
          10
PLEASE ENTER MODE CODE
IF MODE.GE.10 HALF SPECIMENS FAILED, N=NF
          0
SPECIMEN ID CODE: ID=1 SPECIMEN FAILED
                  ID.NE.1 TEST CENSORED
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  1
     4574.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  2
     4705.0000000       ·  1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  3
     5305.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  4
     5425.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  5
     6631.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  6
     6692.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  7
     7109.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  8
     7392.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN  9
     7590.0000000          1
ENTER STRENGTH VALUE AND ID FOR SPECIMEN 10
     7757.0000000          1
ENTER NUMBER OF INTERVALS FOR GOODNESS TEST
          5
WEIBULL ANALYSIS: WEIBULL ANALYSIS: SAMPLE PROBLEM
   ORDERED STRENGTH       ASSIGNED PROBABILITY      CODE
      4574.00000               .90909                 1
      4705.00000               .81818                 1
      5305.00000               .72727                 1
      5425.00000               .63636                 1
      6631.00000               .54545                 1
      6692.00000               .45455                 1
      7109.00000               .36364                 1
      7392.00000               .27273                 1
      7590.00000               .18182                 1
      7757.00000               .09091                 1
MAX. LIKELIHOOD ESTIMATE OF WEIBULL PARAMETERS:
SHAPE PARAMETER ALHPA =      6.74585
SCALE PARAMETER BETA  =   6793.37500
   NUMBER OF ITERATIONS KCOT =     10
95% CONF. LOWER BETA  =   6353.67600
A-BASIS ALLOWABLE     =   3212.69700
B-BASIS ALLOWABLE     =   4551.42700
MEAN STRENGTH    FAVE =   6318.00000
STANDARD DEVIATION    =   1209.80300
```

```
COEFF. OF VARIATION      =      .19149
WEIBULL:
MEAN STRENGTH    FAVE  =  6342.53400
STANDARD DEVIATION     =  1102.71200
COEFF. OF VARIATION      =      .17386
```

| RIGHT END | OBSERVED | EXPECTED | CHI-SQ |
|-----------|----------|----------|--------|
| 5210.6050 | 2 | .86846 | 1.47433 |
| 5847.2100 | 2 | 1.50956 | .15934 |
| 6483.8150 | 0 | 2.13309 | 2.13309 |
| 7120.4200 | 3 | 2.28595 | .22305 |
| 7757.0250 | 3 | 1.66725 | 1.06536 |

```
AT    2DEGREES OF FREEDOM
THE CHI-SQUARE VALUE FOR GOODNESS OF FIT IS        5.05516
```

| STRENGTH VALUE | PROBABILITY OF SURVIVAL |
|----------------|-------------------------|
| 3619.10000 | .98581 |
| 3698.67500 | .98359 |
| 3778.25000 | .98107 |
| 3857.82500 | .97825 |
| 3937.40000 | .97508 |
| 4016.97500 | .97153 |
| 4096.55000 | .96756 |
| 4176.12500 | .96315 |
| 4255.70000 | .95826 |
| 4335.27500 | .95283 |
| 4414.85000 | .94684 |
| 4494.42500 | .94024 |
| 4574.00000 | .93299 |
| 4653.57500 | .92504 |
| 4733.15000 | .91634 |
| 4812.72500 | .90686 |
| 4892.30000 | .89655 |
| 4971.87500 | .88536 |
| 5051.45000 | .87326 |
| 5131.02500 | .86020 |
| 5210.60000 | .84614 |
| 5290.17500 | .83106 |
| 5369.75000 | .81492 |
| 5449.32500 | .79771 |
| 5528.90000 | .77940 |
| 5608.47500 | .75999 |
| 5688.05000 | .73947 |
| 5767.62500 | .71787 |
| 5847.20000 | .69519 |
| 5926.77500 | .67147 |
| 6006.35000 | .64676 |
| 6085.92500 | .62112 |
| 6165.50000 | .59461 |
| 6245.07500 | .56732 |
| 6324.65000 | .53936 |
| 6404.22500 | .51084 |
| 6483.80000 | .48188 |
| 6563.37500 | .45263 |
| 6642.95000 | .42325 |

| | |
|---|---|
| 6722.52500 | .39388 |
| 6802.10000 | .36469 |
| 6881.67500 | .33587 |
| 6961.25000 | .30758 |
| 7040.82500 | .28000 |
| 7120.40000 | .25329 |
| 7199.97500 | .22761 |
| 7279.55000 | .20310 |
| 7359.12500 | .17990 |
| 7438.70000 | .15811 |
| 7518.27500 | .13784 |
| 7597.85000 | .11913 |
| 7677.42500 | .10203 |
| 7757.00000 | .08656 |
| 7836.57500 | .07271 |
| 7916.15000 | .06044 |
| 7995.72500 | .04968 |
| 8075.30000 | .04038 |
| 8154.87500 | .03242 |
| 8234.44900 | .02570 |
| 8314.02500 | .02011 |
| 8393.60000 | .01552 |
| 8473.17500 | .01180 |
| 8552.75000 | .00884 |
| 8632.32400 | .00652 |
| 8711.90000 | .00473 |

## A.2  Program "WEIBJNT"

The program WEIBJNT is used for joint Weibull analysis of pooled data. The program computes the joint MLE Weibull shape parameter for the pooled data, the scale parameter, A- and B-basis allowables, mean, standard deviation and coefficient of variation of each individual data set. The program also prints the normalized joint Weibull distribution.

The required input to WEIBJNT are:

1.  A 32-character problem title (INAME)

2.  Number of data sets (M)

3.  Number of data points in the Ith data set (NC(I))

4.  Number of specimens failed in the Ith data set (NF(I))

5.  Data value and ID for the Jth specimen in the Ith data set (FR(I,J)), ID(I,J))

    ID(I,J) = 1    specimen failed
    ID(I,J) ≠ 1    specimen censored

Input item 5 is repeat N(I) times for the N(I) specimens. Input items 3,4 and 5 are repeat M times for M sets of data.

The program listing, the input and output of a sample example are given below:

```
C  PROGRAM 'WEIBJNT' JOINT MLE OF WEIBULL PARAMETERS
       DOUBLEPRECISION B,Y,CHL,ALPHA,CHIS,GAMMA1,GAM1,GAM2,ARM1,ARM2
       DIMENSION ID(10,15),FR(10,15),RA(10),INAME(8),P(10),N(10),NF(10)
       DIMENSION F(150),IDN(150),IMN(150)
       COMMON/GMA/B(101),Y(101)
       COMMON/CHI/CHL(15)
       OPEN(5,FILE='PSI.DAT')
       READ(5,*) (B(I),I=1,101)
       READ(5,*) (Y(I),I=1,101)
       READ(5,*) (CHL(I),I=1,15)
       PAL = -ALOG(0.99)
       PBL = -ALOG(0.90)
       NM = 0
       KCOT = 0
       WRITE(*,20)
   10 FORMAT(2X,'PLEASE ENTER NO. OF SPECIMENS IN DATA SET',I3)
       READ(*,3) INAME
       WRITE(*,501)
       READ(*,*) M
  501 FORMAT(2X,'PLEASE ENTER NUMBER OF DATA SETS')
       DO 502 I =1,M
       WRITE(*,10) I
   15 FORMAT(2X,'PLEASE ENTER NO. OF FAILURES IN DATA SET',I3)
       READ(*,*) N(I)
       NM = NM+N(I)
       WRITE(*,15) I
   20 FORMAT(2X,'PLEASE ENTER A 32-CHARACTER PROBLEM TITLE')
       READ(*,*) NF(I)
    3 FORMAT(8A4)
       WRITE(*,41)
   41 FORMAT(2X,'SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED')
       DO 35 J = 1,N(I)
       WRITE(*,40) J
   40 FORMAT(2X,'ENTER DATA VALUE AND ID FOR SPECIMEN 'I3)
       READ(*,*) FR(I,J),ID(I,J)
   35 CONTINUE
  502 CONTINUE
       WRITE(*,60) INAME
       WRITE(*,503) M
       ALPHA = 10.0
  503 FORMAT(2X,'TOTAL NUMBER OF DATA SETS M = ',I3)
       DO 505 I=1,M
       WRITE(*,504) I
  504 FORMAT(//2X,'DATA SET NO.',I3)
       WRITE(*,506) N(I),NF(I)
  506 FORMAT(2X,'NUMBER OF SPECIMENS ='I3 3X,'NUMBER OF FAILURES',I3)
       SUM = 0.0
       SUD = 0.0
       WRITE(*,65)
   65 FORMAT(/5X,'ORDER DATA',9X,'ASSD. PROB.',13X,'CODE')
       DO 70 J=1,N(I)
       IF(J.EQ.N(I)) GOTO 70
       I1 = J+1
       DO 71 K=I1,N(I)
       IF(FR(I,J).LE.FR(I,K)) GOTO 71
```

```
       FRT = FR(I,J)
       FR(I,J) = FR(I,K)
       FR(I,K) = FRT
       IDT = ID(I,J)
       ID(I,J) = ID(I,K)
       ID(I,K) = IDT
   71 CONTINUE
   70 CONTINUE
       FART = ALOG10(FR(I,1))
       IF(FART.LE.0.5) GOTO 270
       LFA = FART
       RA(I) = 10.0**LFA
       GOTO 271
  270 RA(I) = 1.0
  271 DO 75 J=1,N(I)
       A = N(I)-J+1.0
       RK = A/(N(I)+1.0)
       SUM = SUM+FR(I,J)
       SUD = SUD+FR(I,J)*FR(I,J)
       WRITE(*,80) FR(I,J),RK,ID(I,J)
   75 FR(I,J) = FR(I,J)/RA(I)
       AVE = SUM/N(I)
       VA = (SUD-N(I)*AVE*AVE)/(N(I)-1.0)
       STD = SQRT(VA)
       CV = STD/AVE
   80 FORMAT(5X,F12.3,9X,F12.5,12X,I2)
   60 FORMAT(///2X,'JOINT WEIBULL ANALYSIS: ',8A4)
       WRITE(*,125) AVE,STD,CV
  505 CONTINUE
   85 PM = 0.0
       RM = 0.0
       KCOT = KCOT+1
       DO 510 I=1,M
       P(I) = 0.0
       R = 0.0
       S = 0.0
       DO 90 J=1,N(I)
       T = FR(I,J)**ALPHA
       TL = ALOG(FR(I,J))
       P(I) = P(I)+T
       S = S+T*TL
       IF(ID(I,J).NE.1) GOTO 90
       R = R+TL
   90 CONTINUE
       PM = PM+S/P(I)
       RM = RM+R/NF(I)
  510 CONTINUE
       ALPH1 = M/(PM-RM)
       DA = ABS(ALPH1-ALPHA)
       IF(DA.LE.0.0001) GOTO 100
       ALPHA = (ALPHA+ALPH1)/2.0
       IF(KCOT.GT.100) GOTO 100
       GOTO 85
  100 ALPHA = (ALPHA+ALPH1)/2.0
       ALR = 1.0/ALPHA
```

```
      WRITE(*,110) KCOT
      WRITE(*,508) ALPHA
508 FORMAT(5X,'JOINT MLE WEIBULL ALPHA =',F12.5)
110 FORMAT(5X,'NUMBER OF ITERATIONS KCOT = ',I5)
      CHSQ = CHIS(NM)
120 FORMAT(2X,'SCALE PARAMETER BETA =',F13.5
    A        /2X,'95% CONF. LOWER BETA =',F13.5
    A        /2X,'A-BASIS ALLOWABLE    =',F13.5
    B        /2X,'B-BASIS ALLOWABLE    =',F13.5)
      ARM1 = 1.0+1.0/ALPHA
      ARM2 = 1.0+2.0/ALPHA
      GAM1 = GAMMA1(ARM1)
      GAM2 = GAMMA1(ARM2)
      COVS = (GAM2-GAM1*GAM1)/(GAM1*GAM1)
      COV = SQRT(COVS)
      IN = 0
      DO 511 I=1,M
      WRITE(*,512) I
512 FORMAT(///2X,'WEIBULL STATISTICS FOR DATA SET',I3)
      BETA = (P(I)/NF(I))**ALR
      DO 513 J=1,N(I)
      F(IN+J) = FR(I,J)/BETA
      IDN(IN+J) = ID(I,J)
      IMN(IN+J) = I
513 CONTINUE
      BETA = BETA*RA(I)
      BETAL = BETA/(CHSQ**ALR)
      ALL = BETAL*PAL**ALR
      BLL = BETAL*PBL**ALR
      XMEAN = BETA*GAM1
      STDW = COV*XMEAN
      WRITE(*,120) BETA,BETAL,ALL,BLL
      WRITE(*,125) XMEAN,STDW,COV
125 FORMAT(/2X,'MEAN STRENGTH    FAVE =',F13.5
    A        /2X,'STANDARE DEVIATION    =',F13.5
    B        /2X,'COEFF. OF VARIATION   =',F13.5)
      IN = IN+N(I)
511 CONTINUE
      WRITE(*,533)
533 FORMAT(////2X,'NORMALIZED JOINT WEIBULL DISTRIBUTION',
    A        /2X,'NORM. VALUE',3X,'GROUP',4X,'CODE',3X,'ASSD. PROB.',4X
    B        ,'CALC. PROB.')
      DO 530 I = 1,NM
      IF(I.EQ.NM) GOTO 532
      I1 = I+1
      DO 531 J=I1,NM
      IF(F(I).LE.F(J)) GOTO 531
      FRT = F(I)
      F(I) = F(J)
      F(J) = FRT
      IDT = IDN(I)
      IDN(I) = IDN(J)
      IDN(J) = IDT
      IMT = IMN(I)
      IMN(I) = IMN(J)
```

```
      IMN(J) = IMT
531 CONTINUE
532 RK = (NM-I+1.0)/(NM+1.0)
      PRB = EXP(-F(I)**ALPHA)
      WRITE(*,534) F(I),IMN(I),IDN(I),RK,PRB
530 CONTINUE
534 FORMAT(4X,F9.5,5X,I3,6X,I2,4X,F9.5,4X,F9.5)
      STOP
      END
      FUNCTION CHIS(N)
      DOUBLEPRECISION CHL,CHIS,BE,CL
      COMMON/CHI/CHL(15)
      AN = N
      BN = 2.0*AN
      IF(N.GE.15) GOTO 50
      CHIS = CHL(N)
      GOTO 60
 50 BE = 1.0/(9.0*AN)
      CL = 1.0-BE+1.645*SQRT(BE)
      CHIS = CL*CL*CL
 60 CONTINUE
      RETURN
      END
      FUNCTION GAMMA1(X)
      DOUBLEPRECISION A,B,X,Y,F,ARG,SLOP,GAMMA1
      COMMON/GMA/B(101),Y(101)
      ARG = X
      A = 1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
      GOTO 30
 10 A = A/ARG
      ARG = ARG+1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      GOTO 30
 20 ARG = ARG-1.0
      A = A*ARG
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
 30 DO 40 I=1,101
      IF(B(I).GT.ARG) GOTO 50
 40 CONTINUE
 50 SLOP = (Y(I)-Y(I-1))/(B(I)-B(I-1))
      F = Y(I-1)+(ARG-B(I-1))*SLOP
      GOTO 60
110 F = 1.0
 60 GAMMA1 = F*A
      RETURN
      END
```

PSI.DAT   TAB VALUES OF CHI-SQUARE AND GAMMA FUNCTIONS
1.0,   1.01, 1.02, 1.03, 1.04, 1.05, 1.06, 1.07, 1.08, 1.09, 1.1,  1.11,
1.12, 1.13, 1.14, 1.15, 1.16, 1.17, 1.18, 1.19, 1.2,  1.21, 1.22, 1.23,
1.24, 1.25, 1.26, 1.27, 1.28, 1.29, 1.3,  1.31, 1.32, 1.33, 1.34, 1.35,
1.36, 1.37, 1.38, 1.39, 1.4,  1.41, 1.42, 1.43, 1.44, 1.45, 1.46, 1.47,
1.48, 1.49, 1.5,  1.51, 1.52, 1.53, 1.54, 1.55, 1.56, 1.57, 1.58, 1.59,
1.6,  1.61, 1.62, 1.63, 1.64, 1.65, 1.66, 1.67, 1.68, 1.69, 1.7,  1.71,
1.72, 1.73, 1.74, 1.75, 1.76, 1.77, 1.78, 1.79, 1.8,  1.81, 1.82, 1.83,
1.84, 1.85, 1.86, 1.87, 1.88, 1.89, 1.9,  1.91, 1.92, 1.93, 1.94, 1.95,
1.96, 1.97, 1.98, 1.99, 2.,
1.,      .99433, .98884, .98355, .97844, .9735,  .96874, .96415, .95973,
2.9955, 2.372,  2.09867, 1.93838, 1.8307,  1.75217, 1.69179, 1.6435,
1.60383, 1.5705, 1.542,  1.51729, 1.49558, 1.47632, 1.4591

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```
WEIBJNT
PLEASE ENTER A 32-CHARACTER PROBLEM TITLE
JOINT WEIBULL ANALYSIS SAMPLE
PLEASE ENTER NUMBER OF DATA SETS
          9
PLEASE ENTER NO. OF SPECIMENS IN DATA SET  1
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET  1
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN    1
      673.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    2
      704.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    3
      718.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    4
     1334.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    5
     1511.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET  2
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET  2
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN    1
    181104.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    2
    192966.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    3
    222450.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    4
    248974.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    5
    440585.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET  3
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET  3
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN    1
   1652870.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    2
   1922982.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    3
   2530135.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    4
   2793310.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN    5
   2986175.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET  4
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET  4
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
```

```
ENTER DATA VALUE AND ID FOR SPECIMEN     1
      95914.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     2
     167338.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     3
     399000.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     4
     440228.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     5
     499495.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET   5
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET   5
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN     1
       4000.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     2
       7934.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     3
       8000.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     4
      10703.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     5
      24000.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET   6
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET   6
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN     1
      16747.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     2
      18707.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     3
      21822.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     4
     116786.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     5
     172710.0000000              1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET   7
          5
PLEASE ENTER NO. OF FAILURES IN DATA SET   7
          5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN     1
      20204.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     2
      20440.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     3
      26376.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     4
      44875.0000000              1
ENTER DATA VALUE AND ID FOR SPECIMEN     5
      65151.0000000              1
```

PLEASE ENTER NO. OF SPECIMENS IN DATA SET   8
        2
PLEASE ENTER NO. OF FAILURES IN DATA SET   8
        2
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN    1
   1124317.0000000            1
ENTER DATA VALUE AND ID FOR SPECIMEN    2
   1865990.0000000            1
PLEASE ENTER NO. OF SPECIMENS IN DATA SET   9
        5
PLEASE ENTER NO. OF FAILURES IN DATA SET   9
        5
SPECIMEN ID CODE: ID=1 FAILURE, ID.NE.1 CENSORED
ENTER DATA VALUE AND ID FOR SPECIMEN    1
      1851.0000000            1
ENTER DATA VALUE AND ID FOR SPECIMEN    2
      2581.0000000            1
ENTER DATA VALUE AND ID FOR SPECIMEN    3
      3298.0000000            1
ENTER DATA VALUE AND ID FOR SPECIMEN    4
      5376.0000000            1
ENTER DATA VALUE AND ID FOR SPECIMEN    5
      7530.0000000            1


JOINT WEIBULL ANALYSIS: JOINT WEIBULL ANALYSIS SAMPLE
TOTAL NUMBER OF DATA SETS M =    9


DATA SET NO.   1
NUMBER OF SPECIMENS =  5    NUMBER OF FAILURES   5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 673.000 | .83333 | 1 |
| 704.000 | .66667 | 1 |
| 718.000 | .50000 | 1 |
| 1334.000 | .33333 | 1 |
| 1511.000 | .16667 | 1 |

MEAN STRENGTH     FAVE =    988.00000
STANDARE DEVIATION     =    401.87870
COEFF. OF VARIATION    =       .40676


DATA SET NO.   2
NUMBER OF SPECIMENS =  5    NUMBER OF FAILURES   5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 181104.000 | .83333 | 1 |
| 192966.000 | .66667 | 1 |
| 222450.000 | .50000 | 1 |
| 248974.000 | .33333 | 1 |
| 440585.000 | .16667 | 1 |

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```
MEAN STRENGTH    FAVE = 257215.80000
STANDARE DEVIATION    = 105855.90000
COEFF. OF VARIATION   =       .41155
```

DATA SET NO.  3
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 1652870.000 | .83333 | 1 |
| 1922982.000 | .66667 | 1 |
| 2530135.000 | .50000 | 1 |
| 2793310.000 | .33333 | 1 |
| 2986175.000 | .16667 | 1 |

```
MEAN STRENGTH    F.VE =2377095.00000
STANDARE DEVIATION    = 569726.90000
COEFF. OF VARIATION   =       .23967
```

DATA SET NO.  4
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 95914.000 | .83333 | 1 |
| 167338.000 | .66667 | 1 |
| 399000.000 | .50000 | 1 |
| 440228.000 | .33333 | 1 |
| 499495.000 | .16667 | 1 |

```
MEAN STRENGTH    FAVE = 320395.00000
STANDARE DEVIATION    = 177787.60000
COEFF. OF VARIATION   =       .55490
```

DATA SET NO.  5
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 4000.000 | .83333 | 1 |
| 7934.000 | .66667 | 1 |
| 8000.000 | .50000 | 1 |
| 10703.000 | .33333 | 1 |
| 24000.000 | .15667 | 1 |

```
MEAN STRENGTH    FAVE = 10927.40000
STANDARE DEVIATION    =  7688.66400
COEFF. OF VARIATION   =       .70361
```

DATA SET NO.  6
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|

| 16747.000 | .83333 | 1 |
| 18707.000 | .66667 | 1 |
| 21822.000 | .50000 | 1 |
| 116786.000 | .33333 | 1 |
| 172710.000 | .16667 | 1 |

MEAN STRENGTH    FAVE =  69354.40000
STANDARE DEVIATION     =  71631.28000
COEFF. OF VARIATION    =      1.03283


DATA SET NO.  7
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 20204.000 | .83333 | 1 |
| 20440.000 | .66667 | 1 |
| 26376.000 | .50000 | 1 |
| 44875.000 | .33333 | 1 |
| 65151.000 | .16667 | 1 |

MEAN STRENGTH    FAVE =  35409.20000
STANDARE DEVIATION     =  19436.12000
COEFF. OF VARIATION    =       .54890


DATA SET NO.  8
NUMBER OF SPECIMENS = 2    NUMBER OF FAILURES  2

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 1124317.000 | .66667 | 1 |
| 1865990.000 | .33333 | 1 |

MEAN STRENGTH    FAVE =1495154.00000
STANDARE DEVIATION     = 524442.00000
COEFF. OF VARIATION    =       .35076


DATA SET NO.  9
NUMBER OF SPECIMENS = 5    NUMBER OF FAILURES  5

| ORDER DATA | ASSD. PROB. | CODE |
|---|---|---|
| 1851.000 | .83333 | 1 |
| 2581.000 | .66667 | 1 |
| 3298.000 | .50000 | 1 |
| 5376.000 | .33333 | 1 |
| 7530.000 | .16667 | 1 |

MEAN STRENGTH    FAVE =   4127.20000
STANDARE DEVIATION     =   2312.89600
COEFF. OF VARIATION    =       .56040
    NUMBER OF ITERATIONS KCOT =    11
    JOINT MLE WEIBULL ALPHA =       2.19728

```
WEIBULL STATISTICS FOR DATA SET  1
SCALE PARAMETER BETA =    1063.57000
95% CONF. LOWER BETA =     955.11270
A-BASIS ALLOWABLE    =     117.71500
B-BASIS ALLOWABLE    =     342.97940

MEAN STRENGTH    FAVE =     941.92400
STANDARE DEVIATION   =     452.45610
COEFF. OF VARIATION  =        .48035


WEIBULL STATISTICS FOR DATA SET  2
SCALE PARAMETER BETA = 277660.40000
95% CONF. LOWER BETA = 249346.00000
A-BASIS ALLOWABLE    =  30731.19000
B-BASIS ALLOWABLE    =  89539.72000

MEAN STRENGTH    FAVE = 245902.80000
STANDARE DEVIATION   = 118120.20000
COEFF. OF VARIATION  =        .48035


WEIBULL STATISTICS FOR DATA SET  3
SCALE PARAMETER BETA =2441181.00000
95% CONF. LOWER BETA =2192241.00000
A-BASIS ALLOWABLE    = 270187.50000
B-BASIS ALLOWABLE    = 787230.10000

MEAN STRENGTH    FAVE =2161969.00000
STANDARE DEVIATION   =1038509.00000
COEFF. OF VARIATION  =        .48035


WEIBULL STATISTICS FOR DATA SET  4
SCALE PARAMETER BETA = 363389.80000
95% CONF. LOWER BETA = 326333.10000
A-BASIS ALLOWABLE    =  40219.63000
B-BASIS ALLOWABLE    = 117185.70000

MEAN STRENGTH    FAVE = 321826.80000
STANDARE DEVIATION   = 154590.50000
COEFF. OF VARIATION  =        .48035


WEIBULL STATISTICS FOR DATA SET  5
SCALE PARAMETER BETA =   13299.10000
95% CONF. LOWER BETA =   11942.92000
A-BASIS ALLOWABLE    =    1471.93100
B-BASIS ALLOWABLE    =    4288.68300
```

MEAN STRENGTH     FAVE  =   11778.00000
STANDARE DEVIATION      =    5657.60000
COEFF. OF VARIATION     =        .48035


WEIBULL STATISTICS FOR DATA SET   6
SCALE PARAMETER BETA  =   98244.48000
95% CONF. LOWER BETA  =   88225.99000
A-BASIS ALLOWABLE     =   10873.61000
B-BASIS ALLOWABLE     =   31681.80000

MEAN STRENGTH     FAVE  =   87007.71000
STANDARE DEVIATION      =   41794.42000
COEFF. OF VARIATION     =        .48035


WEIBULL STATISTICS FOR DATA SET   7
SCALE PARAMETER BETA  =   40220.67000
95% CONF. LOWER BETA  =   36119.16000
A-BASIS ALLOWABLE     =    4451.58500
B-BASIS ALLOWABLE     =   12970.33000

MEAN STRENGTH     FAVE  =   35620.41000
STANDARE DEVIATION      =   17110.37000
COEFF. OF VARIATION     =        .48035


WEIBULL STATISTICS FOR DATA SET   8
SCALE PARAMETER BETA  =1549015.00000
95% CONF. LOWER BETA  =1391054.00000
A-BASIS ALLOWABLE     =  171443.50000
B-BASIS ALLOWABLE     =  499525.10000

MEAN STRENGTH     FAVE  =1371845.00000
STANDARE DEVIATION      =  658970.00000
COEFF. OF VARIATION     =        .48035


WEIBULL STATISTICS FOR DATA SET   9
SCALE PARAMETER BETA  =    4707.17900
95% CONF. LOWER BETA  =    4227.16400
A-BASIS ALLOWABLE     =     520.98610
B-BASIS ALLOWABLE     =    1517.96800

MEAN STRENGTH     FAVE  =    4168.79200
STANDARE DEVIATION      =    2002.49200
COEFF. OF VARIATION     =        .48035

## NORMALIZED JOINT WEIBULL DISTRIBUTION

| NORM. VALUE | GROUP | CODE | ASSD. PROB. | CALC. PROB. |
|---|---|---|---|---|
| .17046 | 6 | 1 | .97674 | .97971 |
| .19041 | 6 | 1 | .95349 | .97420 |
| .22212 | 6 | 1 | .93023 | .96400 |
| .26394 | 4 | 1 | .90698 | .94784 |
| .30077 | 5 | 1 | .88372 | .93111 |
| .39323 | 9 | 1 | .86047 | .87930 |
| .46049 | 4 | 1 | .83721 | .83363 |
| .50233 | 7 | 1 | .81395 | .80229 |
| .50820 | 7 | 1 | .79070 | .79773 |
| .54831 | 9 | 1 | .76744 | .76565 |
| .59658 | 5 | 1 | .74419 | .72511 |
| .60154 | 5 | 1 | .72093 | .72084 |
| .63277 | 1 | 1 | .69767 | .69362 |
| .65225 | 2 | 1 | .67442 | .67636 |
| .65578 | 7 | 1 | .65116 | .67321 |
| .66192 | 1 | 1 | .62791 | .66772 |
| .67508 | 1 | 1 | .60465 | .65590 |
| .67708 | 3 | 1 | .58140 | .65410 |
| .69497 | 2 | 1 | .55814 | .63793 |
| .70063 | 9 | 1 | .53488 | .63279 |
| .72583 | 8 | 1 | .51163 | .60984 |
| .78773 | 3 | 1 | .48837 | .55323 |
| .80116 | 2 | 1 | .46512 | .54097 |
| .80479 | 5 | 1 | .44186 | .53766 |
| .89669 | 2 | 1 | .41860 | .45524 |
| 1.03644 | 3 | 1 | .39535 | .33898 |
| 1.09799 | 4 | 1 | .37209 | .29287 |
| 1.11572 | 7 | 1 | .34884 | .28027 |
| 1.14209 | 9 | 1 | .32558 | .26211 |
| 1.14425 | 3 | 1 | .30233 | .26065 |
| 1.18873 | 6 | 1 | .27907 | .23175 |
| 1.20463 | 8 | 1 | .25581 | .22192 |
| 1.21145 | 4 | 1 | .23256 | .21779 |
| 1.22325 | 3 | 1 | .20930 | .21076 |
| 1.25427 | 1 | 1 | .18605 | .19300 |
| 1.37454 | 4 | 1 | .16279 | .13375 |
| 1.42069 | 1 | 1 | .13953 | .11497 |
| 1.58678 | 2 | 1 | .11628 | .06342 |
| 1.59968 | 9 | 1 | .09302 | .06036 |
| 1.61984 | 7 | 1 | .06977 | .05581 |
| 1.75796 | 6 | 1 | .04651 | .03161 |
| 1.80463 | 5 | 1 | .02326 | .02576 |

A.3   Program "LOAD"

This program computes the Load Enhancement Factor by given strength and life scatter parameter, sample size and fatigue test duration.   The required input are

1.   Residual strength Weibull shape parameter (ALPR)

2.   Fatigue life Weibull shape parameter (ALPL)

3.   Sample size (N)

4.   Fatigue test duration (T)

The program listing, input and output of a sample problem are given below:

```
C LOAD.FOR COMPUTES LOAD ENHANCEMENT FACTOR BASED ON WEIBULL
      COMMON/GMA/B(101),Y(101)
      COMMON/CHI/CHL(15)
      OPEN(5,FILE='PSI.DAT')
      READ(5,*) (B(I),I=1,101)
      READ(5,*) (Y(I),I=1,101)
      READ(5,*) (CHL(I),I=1,15)
      PAL = -ALOG(0.99)
      PBL = -ALOG(0.90)
      WRITE(*,1)
1 FORMAT(2X,'THIS PROGRAM COMPUTES THE LOAD ENHANCEMENT FACTOR'
 A/2X,'THE REQUIRED INPUT ARE STRENTH AND LIFE WEIBULL ALPHA,'
 B/2X,'SAMPLE SIZE AND FATIGUE TEST DURATION'
 C/2X,'PLEASE INPUT STRENGTH ALPHA')
      READ(*,*) ALPR
      ALPRI = 1.0/ALPR
      WRITE(*,2)
2 FORMAT(2X,'PLEASE INPUT LIFE ALPHA')
      READ(*,*) ALPL
      ALPLI = 1.0/ALPL
      WRITE(*,3)
3 FORMAT(2X,'PLEASE INPUT SAMPLE SIZE')
      READ(*,*) N
      AN = N
      WRITE(*,4)
4 FORMAT(2X,'PLEASE INPUT TEST DURATION')
      READ(*,*) T
      RALP = ALPL/ALPR
      ARGR = 1.0+ALPRI
      ARGL = 1.0+ALPLI
      GR = GAMMA1(ARGR)
      GL = GAMMA1(ARGL)
      CHSQ = CHIS(N)
      AM = (GL**RALP)/GR
      AA = PAL*(T**ALPL)
      AB = PBL*(T**ALPL)
      BA = (AA/CHSQ)**ALPRI
      BB = (AB/CHSQ)**ALPRI
      FA = AM*GR/BA
      FB = AM*GR/BB
      TA = GL/((PAL/CHSQ)**ALPLI)
      TB = GL/((PBL/CHSQ)**ALPLI)
      WRITE(*,5) ALPR,ALPL,N
      WRITE(*,6) TA,TB
      WRITE(*,7) T,FA,FB
5 FORMAT(2X,'STRENGTH ALPHA = ',F9.5
 A      /2X,'LIFE ALPHA     = ',F9.5
 B      /2X,'SAMPLE SIZE = ',I5)
6 FORMAT(2X,'A-BASIS LIFE FACTOR =',F9.5
 A      /2X,'B-BASIS LIFE FACTOR =',F9.5)
7 FORMAT(2X,'TEST DURATION =',F7.4
 A      /2X,'A-BASIS LOAD ENHANCEMENT FACTOR =',F9.5
 B      /2X,'B-BASIS LOAD ENHANCEMENT FACTOR =',F9.5)
      STOP
      END
```

```
      FUNCTION CHIS(N)
      COMMON/CHI/CHL(15)
      AN = N
      BN = 2.0*AN
      IF(N.GE.15) GOTO 50
      CHIS = CHL(N)
      GOTO 60
   50 B = 1.0/(9.0*AN)
      CL = 1.0-B+1.645*SQRT(B)
      CHIS = CL*CL*CL
   60 CONTINUE
      RETURN
      END
      FUNCTION GAMMA1(X)
      COMMON/GMA/B(101),Y(101)
      ARG = X
      A = 1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
      GOTO 30
   10 A = A/ARG
      ARG = ARG+1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      GOTO 30
   20 ARG = ARG-1.0
      A = A*ARG
      IF(ARG.GT.2.0) GOTO 20
      IF(ARG.EQ.2.0) GOTO 110
   30 DO 40 I=1,101
      IF(B(I).GT.ARG) GOTO 50
   40 CONTINUE
   50 SLOP = (Y(I)-Y(I-1))/(B(I)-B(I-1))
      F = Y(I-1)+(ARG-B(I-1))*SLOP
      GOTO 60
  110 F = 1.0
   60 GAMMA1 = F*A
      RETURN
      END
```

PSI.DAT  TAB VALUES OF CHI-SQUARE AND GAMMA FUNCTIONS
1.0,  1.01, 1.02, 1.03, 1.04, 1.05, 1.06, 1.07, 1.08, 1.09, 1.1,  1.11,
1.12, 1.13, 1.14, 1.15, 1.16, 1.17, 1.18, 1.19, 1.2,  1.21, 1.22, 1.23,
1.24, 1.25, 1.26, 1.27, 1.28, 1.29, 1.3,  1.31, 1.32, 1.33, 1.34, 1.35,
1.36, 1.37, 1.38, 1.39, 1.4,  1.41, 1.42, 1.43, 1.44, 1.45, 1.46, 1.47,
1.48, 1.49, 1.5,  1.51, 1.52, 1.53, 1.54, 1.55, 1.56, 1.57, 1.58, 1.59,
1.6,  1.61, 1.62, 1.63, 1.64, 1.65, 1.66, 1.67, 1.68, 1.69, 1.7,  1.71,
1.72, 1.73, 1.74, 1.75, 1.76, 1.77, 1.78, 1.79, 1.8,  1.81, 1.82, 1.83,
1.84, 1.85, 1.86, 1.87, 1.88, 1.89, 1.9,  1.91, 1.92, 1.93, 1.94, 1.95,
1.96, 1.97, 1.98, 1.99, 2.,
1.,       .99433, .98884, .98355, .97844, .9735,  .96874, .96415, .95973,
2.9955, 2.372,  2.09867, 1.93838, 1.8307,  1.75217, 1.69179, 1.6435,
1.60383, 1.5705, 1.542,  1.51729, 1.49558, 1.47632, 1.4591

```
LOAD
THIS PROGRAM COMPUTES THE LOAD ENHANCEMENT FACTOR
THE REQUIRED INPUT ARE STRENTH AND LIFE WEIBULL ALPHA,
SAMPLE SIZE AND FATIGUE TEST DURATION
PLEASE INPUT STRENGTH ALPHA
        20.0000000
PLEASE INPUT LIFE ALPHA
        1.2500000
PLEASE INPUT SAMPLE SIZE
            5
PLEASE INPUT TEST DURATION
        2.0000000
STRENGTH ALPHA =   20.00000
LIFE ALPHA     =    1.25000
SAMPLE SIZE =         5
A-BASIS LIFE FACTOR = 59.90686
B-BASIS LIFE FACTOR =  9.14278
TEST DURATION = 2.0000
A-BASIS LOAD ENHANCEMENT FACTOR =  1.23674
B-BASIS LOAD ENHANCEMENT FACTOR =  1.09965
```

## A.4 Program "ALLOW"

This program computes the A- and B-basis allowables by either given the Weibull shape parameter ($\alpha$) or the coefficient of variation (CV) and the sample size.

The program listing and the results of two sample problems are given below. The first example uses $\alpha = 20.0$ and n = 5. The second example uses CV = 0.065 and n = 5.

```
C ALLOW.FOR COMPUTES A AND B BASIS ALLOWABLES BASED ON WEIBULL
      COMMON/GMA/B(101),Y(101)
      COMMON/CHI/CHL(15)
      OPEN(5,FILE-'PSI.DAT')
      READ(5,*) (B(I),I-1,101)
      READ(5,*) (Y(I),I-1,101)
      READ(5,*) (CHL(I),I-1,15)
      PAL - -ALOG(0.99)
      PBL - -ALOG(0.90)
      WRITE(*,1)
      WRITE(*,101)
      WRITE(*,102)
      WRITE(*,103)
      WRITE(*,104)
      WRITE(*,105)
    1 FORMAT(2X,'THIS PROGRAM COMPUTES THE A- AND B-BASIS ALLOWABLES')
  101 FORMAT(2X,'THE REQUIRED INPUT ARE COEFFICIENT OF VARIATION OR')
  102 FORMAT(2X,'WEIBULL SHAPE PARAMETER, ALPHA AND SAMPLE SIZE')
  103 FORMAT(2X,'PLEASE ENTER CASE CONTROL ID')
  104 FORMAT(2X,'IF ALPHA IS GIVEN ENTER "0" AND HIT RETURN')
  105 FORMAT(2X,'IF CV IS GIVEN ENTER ANY INTEGER AND HIT RETURN')
      READ(*,*) ID
      WRITE(*,4)
      READ(*,*) N
      IF(ID.EQ.0) GOTO 20
      WRITE(*,2)
    2 FORMAT(2X,'PLEASE INPUT COEFFICIENT OF VARIATION: CV')
      READ(*,*) CV
      DT - 0.0001
      AL1 - (1.0/CV)**1.07
   10 ARG1 - 1.0+1.0/AL1
      ARG2 - 1.0+2.0/AL1
      GM1 - GAMMA1(ARG1)
      GM2 - GAMMA1(ARG2)
      SG - GM2-GM1*GM1
      SG - SQRT(SG)
      CV1 - SG/GM1
      DC - (CV-CV1)/CV
      DC - ABS(DC)
      IF(DC.LT.DT) ALPHA - AL1
      IF(DC.LT.DT) GOTO 30
      AL1 - AL1*CV1/CV
      GOTO 10
   20 WRITE(*,3)
    3 FORMAT(2X,'PLEASE INPUT ALPHA')
      READ(*,*) ALPHA
   30 ARG - 1.0+1.0/ALPHA
      GM - GAMMA1(ARG)
      ALPI - 1.0/ALPHA
    4 FORMAT(2X,'PLEASE INPUT SAMPLE SIZE N')
      AN - N
      BET - 1.0/GM
      CHSQ - CHIS(N)
      ALLA - ((PAL/CHSQ)**ALPI)*BET
      ALLB - ((PBL/CHSQ)**ALPI)*BET
```

```
      WRITE(*,5) ALPHA
      WRITE(*,106) BET
      WRITE(*,107) ALLA
      WRITE(*,108) ALLB
    5 FORMAT(2X,'ALPHA - ',F9.5)
  106 FORMAT(2X,'BETA - ',F9.5)
  107 FORMAT(2X,'A-ALLOWABLE - ',F9.5)
  108 FORMAT(2X,'B-ALLOWABLE - ',F9.5)
      STOP
      END
      FUNCTION CHIS(N)
      COMMON/CHI/CHL(15)
      AN = N
      BN =2.0*AN
      IF(N.GE.15) GOTO 50
      CHIS - CHL(N)
      GOTO 60
   50 B - 1.0/(9.0*AN)
      CL - 1.0-B+1.645*SQRT(B)
      CHIS - CL*CL*CL
   60 CONTINUE
      RETURN
      END
      FUNCTION GAMMA1(X)
      COMMON/GMA/B(101),Y(101)
      ARG = X
      A - 1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
      GOTO 30
   10 A - A/ARG
      ARG = ARG+1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      GOTO 30
   20 ARG - ARG-1.0
      A - A*ARG
      IF(ARG.GT.2.0) GOTO 20
      IF(ARG.EQ.2.0) GOTO 110
   30 DO 40 I-1,101
      IF(B(I).GT.ARG) GOTO 50
   40 CONTINUE
   50 SLOP - (Y(I)-Y(I-1))/(B(I)-B(I-1))
      F - Y(I-1)+(ARG-B(I-1))*SLOP
      GOTO 60
  110 F - 1.0
   60 GAMMA1 - F*A
      RETURN
      END
```

PSI.DAT  TAB VALUES OF CHI-SQUARE AND GAMMA FUNCTIONS
1.0,  1.01, 1.02, 1.03, 1.04, 1.05, 1.06, 1.07, 1.08, 1.09, 1.1,  1.11,
1.12, 1.13, 1.14, 1.15, 1.16, 1.17, 1.18, 1.19, 1.2,  1.21, 1.22, 1.23,
1.24, 1.25, 1.26, 1.27, 1.28, 1.29, 1.3,  1.31, 1.32, 1.33, 1.34, 1.35,
1.36, 1.37, 1.38, 1.39, 1.4,  1.41, 1.42, 1.43, 1.44, 1.45, 1.46, 1.47,
1.48, 1.49, 1.5,  1.51, 1.52, 1.53, 1.54, 1.55, 1.56, 1.57, 1.58, 1.59,
1.6,  1.61, 1.62, 1.63, 1.64, 1.65, 1.66, 1.67, 1.68, 1.69, 1.7,  1.71,
1.72, 1.73, 1.74, 1.75, 1.76, 1.77, 1.78, 1.79, 1.8,  1.81, 1.82, 1.83,
1.84, 1.85, 1.86, 1.87, 1.88, 1.89, 1.9,  1.91, 1.92, 1.93, 1.94, 1.95,
1.96, 1.97, 1.98, 1.99, 2.,
1.,      .99433, .98884, .98355, .97844, .9735,  .96874, .96415, .95973,
2.9955, 2.372,  2.09867, 1.93838, 1.8307,  1.75217, 1.69179, 1.6435,
1.60383, 1.5705, 1.542,  1.51729, 1.49558, 1.47632, 1.4591

```
   ALLOW
THIS PROGRAM COMPUTES THE A- AND B-BASIS ALLOWABLES
THE REQUIRED INPUT ARE COEFFICIENT OF VARIATION OR
WEIBULL SHAPE PARAMETER, ALPHA AND SAMPLE SIZE
PLEASE ENTER CASE CONTROL ID
IF ALPHA IS GIVEN ENTER "0" AND HIT RETURN
IF CV IS GIVEN ENTER ANY INTEGER AND HIT RETURN
             0
PLEASE INPUT SAMPLE SIZE N
             5
PLEASE INPUT ALPHA
        20.0000000
ALPHA =  20.00000
BETA  =   1.02722
A-ALLOWABLE =    .79185
B-ALLOWABLE =    .89057


   ALLOW
THIS PROGRAM COMPUTES THE A- AND B-BASIS ALLOWABLES
THE REQUIRED INPUT ARE COEFFICIENT OF VARIATION OR
WEIBULL SHAPE PARAMETER, ALPHA AND SAMPLE SIZE
PLEASE ENTER CASE CONTROL ID
IF ALPHA IS GIVEN ENTER "0" AND HIT RETURN
IF CV IS GIVEN ENTER ANY INTEGER AND HIT RETURN
             1
PLEASE INPUT SAMPLE SIZE N
             5
PLEASE INPUT COEFFICIENT OF VARIATION: CV
 6.500000E-002
ALPHA =  19.01858
BETA  =   1.02852
A-ALLOWABLE =    .78227
B-ALLOWABLE =    .88515
```

## A.5  Programs "BSRV" and "CSRV"

These two programs compute the structural static and fatigue reliability, taken into consideration of the structural response variability (SRV). The programs are similar, except in their input. BSRV require more input and shorter computation time. Both programs have seven cases of reliability computations. These cases and their corresponding input are discussed below.

Case 1:  Compute static reliability at a specified load level. The required input are:

1. Static strength $\alpha_s$ (ALM)

2. SRV $\alpha_{SRV}$ (ALS)

3. Gamma function value for ALS (GAS) for BSRV
   Number of specimen (N) for CSRV

4. $\overset{\vee}{\beta}$, 95% confidence lower limit of static strength

   $\beta$(BET) for BSRV

   Mean static strength (XB) for CSRV

5. Applied load level (AK)

In using BSRV the parameters GAS and BET must be pre-computed using the following equations

$$\overset{\wedge}{\beta} = \bar{x} / \Gamma(1 + 1/\alpha_s)$$

$$\overset{\vee}{\beta} = \overset{\wedge}{\beta} / [\chi^2 (2n)/2n]^{1/\alpha_s}$$

The values of $\chi^2$ are given in Table A-1 and values of $\Gamma$ are given in Table A-2. The relation between CV and   are given in Table A-3. Example runs for Case 1 static reliability are given below.

TABLE A-1.   VALUES OF $\chi^2(2n)/2n$ AT F = 0.95

| n | $\chi_{.95}(2n)/2n$ |
|---|---|
| 1 | 2.99550 |
| 2 | 2.37200 |
| 3 | 2.09867 |
| 4 | 1.93838 |
| 5 | 1.83070 |
| 6 | 1.75217 |
| 7 | 1.69179 |
| 8 | 1.64350 |
| 9 | 1.60383 |
| 10 | 1.57050 |
| 11 | 1.54200 |
| 12 | 1.51729 |
| 13 | 1.49558 |
| 14 | 1.47632 |
| 15 | 1.45910 |
| 16 | 1.44344 |
| 17 | 1.42935 |
| 18 | 1.41653 |
| 19 | 1.40474 |
| 20 | 1.39388 |
| 22 | 1.37450 |
| 25 | 1.35004 |
| 30 | 1.31800 |
| 40 | 1.27349 |
| 50 | 1.24342 |

TABLE A-2.   VALUES OF $\Gamma(1+\frac{1}{\alpha})$

| $\alpha$ | $\Gamma(1+\frac{1}{\alpha})$ | $\alpha$ | $\Gamma(1+\frac{1}{\alpha})$ |
|---|---|---|---|
| 0.1 | 3,628,800 | 10.0 | 0.95135 |
| 0.125 | 40,320 | 11.0 | 0.95509 |
| 0.15 | 2593.6 | 12.0 | 0.95831 |
| 0.18 | 318.1225 | 13.0 | 0.96109 |
| 0.20 | 120.0 | 14.0 | 0.96352 |
| 0.25 | 24.0 | 15.0 | 0.96568 |
| 0.30 | 9.26067 | 16.0 | 0.96759 |
| 0.40 | 3.32336 | 17.0 | 0.96930 |
| 0.50 | 2.0 | 18.0 | 0.97086 |
| 0.75 | 1.19066 | 19.0 | 0.97225 |
| 1.00 | 1.00000 | 20.0 | 0.97350 |
| | | 22.0 | 0.97575 |
| 1.25 | 0.93138 | 23.2 | 0.97691 |
| 1.50 | 0.90276 | 24.0 | 0.97762 |
| 1.75 | 0.89062 | 25.0 | 0.97844 |
| 2.00 | 0.88623 | 26.0 | 0.97923 |
| 2.17 | 0.88560 | 27.5 | 0.98030 |
| 2.25 | 0.88573 | 30.0 | 0.98185 |
| 2.50 | 0.88726 | 32.5 | 0.98316 |
| 2.75 | 0.88986 | 35.0 | 0.98431 |
| 3.00 | 0.89299 | 37.5 | 0.98531 |
| 3.50 | 0.89976 | 40.0 | 0.98620 |
| 4.00 | 0.90640 | 45.0 | 0.98766 |
| 4.50 | 0.91259 | 50.0 | 0.98884 |
| 5.00 | 0.91817 | 60.0 | 0.99067 |
| 5.50 | 0.92321 | 70.0 | 0.99198 |
| 6.00 | 0.92773 | 80.0 | 0.99296 |
| 6.50 | 0.93179 | 90.0 | 0.99372 |
| 7.00 | 0.93545 | 100.0 | 0.99433 |
| 7.50 | 0.93876 | | |
| 8.00 | 0.94176 | | |
| 8.80 | 0.94601 | | |
| 9.00 | 0.94697 | | |

TABLE A-3   RELATION BETWEEN WEIBULL SHAPE PARAMETER ($\alpha$)
AND COEFFICIENT OF VARIATION (CV)

| $\alpha$ | CV | $\alpha$ | CV |
|---|---|---|---|
| 0.1 | 429.83 | 10.0 | 0.12032 |
| 0.125 | 113.44 | 11.0 | 0.10992 |
| 0.15 | 47.036 | 12.0 | 0.10107 |
| 0.18 | 22.731 | 13.0 | 0.09364 |
| 0.20 | 15.843 | 14.0 | 0.08737 |
| 0.25 | 8.3066 | 15.0 | 0.08168 |
| 0.30 | 5.4076 | 16.0 | 0.07682 |
| 0.40 | 3.1408 | 17.0 | 0.07252 |
| 0.50 | 2.2361 | 18.0 | 0.06838 |
| 0.75 | 1.3528 | 19.0 | 0.06506 |
| 1.00 | 1.0000 | 20.0 | 0.06204 |
|  |  | 21.0 | 0.05906 |
| 1.25 | 0.80501 | 22.0 | 0.05620 |
| 1.50 | 0.67896 | 23.0 | 0.05393 |
| 1.75 | 0.58975 | 23.2 | 0.05351 |
| 2.00 | 0.52271 | 24.0 | 0.05188 |
| 2.17 | 0.48579 | 25.0 | 0.04991 |
| 2.50 | 0.42791 | 26.0 | 0.04797 |
| 3.00 | 0.36342 | 27.0 | 0.04608 |
| 3.50 | 0.31642 | 28.0 | 0.04426 |
| 4.00 | 0.28056 | 29.0 | 0.04269 |
| 4.50 | 0.25208 | 30.0 | 0.04146 |
| 5.00 | 0.22904 | 32.0 | 0.03911 |
| 6.00 | 0.19373 | 34.0 | 0.03689 |
| 7.00 | 0.16796 | 36.0 | 0.03475 |
| 7.50 | 0.15751 | 38.0 | 0.03271 |
| 8.00 | 0.14823 | 40.0 | 0.03075 |
| 8.80 | 0.13561 | 45.0 | 0.02797 |
| 9.00 | 0.13290 | 50.0 | 0.02549 |

Note   $CV = \dfrac{\sqrt{\Gamma(1+2/\alpha) - \Gamma^2(1+1/\alpha)}}{\Gamma(1+1/\alpha)}$

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SKV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          1
CASE 1,STATIC REL. WITH SRV
PLEASE INPUT STATIC STRENGTH ALPHA
        20.0000000
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT SRV GAMMA
9.693000E-001
PLEASE INPUT LOWER STRENGTH BETA
        1.0280000
PLEASE INPUT LOAD LEVEL
        1.1500000
AT LOAD LEVEL OF    1.150
THE STRUCTURAL RELIABILITY IS   .0771640


CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          1
CASE 1,STATIC REL. WITH SRV
PLEASE INPUT STATIC STRENGTH ALPHA
        20.0000000
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT NUMBER OF SPECIMENS
          5
PLEASE INPUT MEAN STATIC STRENGTH
        1.0000000
PLEASE INPUT LOAD LEVEL
        1.1500000
AT LOAD LEVEL OF    1.150
THE STRUCTURAL RELIABILITY IS   .0468499
```

Cases 2 and 3: These two cases compute the fatigue reliability at one lifetime using the ultimate strength approach. Case 2 performs computation with general input parameters and Case 3 uses some fixed parameter values that are frequently encountered. The required input are:

1. Static strength $\alpha_S$ and Gamma value (ALM,GAM)
   GAM is not required in CSRV
   ALM = 20.0 and GAM = 0.9735 in Case 3. These values are not required for Case 3.

2. SRV $\alpha_{SRV}$ and Gamma value
   (ALS,GAS)
   GAS not required in CSRV

3. Fatigue spectrum characteristic parameter
   (CON)
   $$CON = \frac{\sigma_S^M}{\sigma_{TH}^M} \cdot \frac{\sigma_{TH}^M}{\sigma_{TH}^B}$$

   In Case 3 CON is fixed at 1.57369

4. Chi-square value (CHI) for BSRV
   Sample size (N) for CSRV

   Not required in Case 3.
   (Fixed values of CHI = 2.9955 or N = 1)

5. Normalized maximum spectrum load
   (AM = $P_{MSL}/P_{DUL}$) and static failure load F for Case 2.

   For Case 3 these input become AM, FMIN and FMAX where FMIN is the minimum failure load and FMAX is the maximum failure load. The reliability is computed from FMIN to FMAX at an interval of 0.05.

Example runs for cases 2 and 3 are given below.

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
         2
CASE 2, FATIGUE REL. ULT WITH SRV
PLEASE INPUT STATIC ALPHA AND GAMMA
      20.0000000  9.735000E-001
PLEASE INPUT SRV ALPHA AND GAMMA
      17.0000000  9.693000E-001
PLEASE INPUT SPECTRUM CHARACT. CONST.

PLEASE INPUT CHI SQUARE VALUE
      2.9955000
PLEASE INPUT MAX. SPECTRUM LOAD
  AND STATIC FAILURE LOAD
9.600000E-001          1.2200000
AT STATIC FAILURE LOAD      1.220
THE ONE LIFETIME FATIGUE REL. IS   .0035912


CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
         2
CASE 2, FATIGUE REL. ULT WITH SRV
PLEASE INPUT STATIC ALPHA
      20.0000000
PLEASE INPUT SRV ALPHA
      17.0000000
PLEASE INPUT SPECTRUM CHARACT. CONST.
      1.5737000
PLEASE INPUT NUMBER OF SPECIMENS
         1
PLEASE INPUT MAX. SPECTRUM LOAD
  AND STATIC FAILURE LOAD
9.600000E-001          1.2200000
AT STATIC FAILURE LOAD      1.220
THE ONE LIFETIME FATIGUE REL. IS   .0035912
```

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          3
CASE 3, FATIGUE REL. ULT WITH SRV
FIXED CON=1.5736879
STRENGTH ALPHA=20.0, GAMMA=0.9735
SINGLE ARTICLE, CHI=2.9955
REL. COMPUTED AT 0.05 INTERVAL
PLEASE INPUT SRV ALPHA AND GAMMA
        17.0000000   9.693000E-001
PLEASE INPUT MAX. SPECTRUM LOAD
MINIMUM AND MAXIMUM STRENGTH
9.600000E-001           1.2000000           1.5000000
AT STATIC FAILURE LOAD      1.200
THE ONE LIFETIME FATIGUE REL. IS   .0012483
AT STATIC FAILURE LOAD      1.250
THE ONE LIFETIME FATIGUE REL. IS   .0128437
AT STATIC FAILURE LOAD      1.300
THE ONE LIFETIME FATIGUE REL. IS   .0579314
AT STATIC FAILURE LOAD      1.350
THE ONE LIFETIME FATIGUE REL. IS   .1544490
AT STATIC FAILURE LOAD      1.400
THE ONE LIFETIME FATIGUE REL. IS   .2933223
AT STATIC FAILURE LOAD      1.450
THE ONE LIFETIME FATIGUE REL. IS   .4467351
AT STATIC FAILURE LOAD      1.500
THE ONE LIFETIME FATIGUE REL. IS   .5888761
```

CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          3
CASE 3, FATIGUE REL. ULT WITH SRV
FIXED CON=1.5736879
STRENGTH ALPHA=20.0, GAMMA=0.9735
SINGLE ARTICLE, CHI=2.9955
REL. COMPUTED AT 0.05 INTERVAL
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT MAX. SPECTRUM LOAD
MINIMUM AND MAXIMUM STRENGTH
9.600000E-001          1.2000000          1.5000000
AT STATIC FAILURE LOAD      1.200
THE ONE LIFETIME FATIGUE REL. IS   .0012483
AT STATIC FAILURE LOAD      1.250
THE ONE LIFETIME FATIGUE REL. IS   .0128437
AT STATIC FAILURE LOAD      1.300
THE ONE LIFETIME FATIGUE REL. IS   .0579316
AT STATIC FAILURE LOAD      1.350
THE ONE LIFETIME FATIGUE REL. IS   .1544490
AT STATIC FAILURE LOAD      1.400
THE ONE LIFETIME FATIGUE REL. IS   .2933227
AT STATIC FAILURE LOAD      1.450
THE ONE LIFETIME FATIGUE REL. IS   .4467355
AT STATIC FAILURE LOAD      1.500
THE ONE LIFETIME FATIGUE REL. IS   .5888761

Case 4 and 5:  These two cases compute the fatigue reliability at one lifetime using the load enhancement factor approach. Case 4 uses the general input and Case 5 computes the reliability at certain fixed parameters.

1.  Gamma value associated with static strength (GAM) for BSRV
    Static strength $\alpha_S$ (ALM) for CSRV
    Not required for Case 5 ($\alpha_S$ = 20.0)

2.  Fatigue life $\alpha_L$ and Gamma Value. (AL,GAL) for BSRV
    GAL not required for CSRV
    For Case 5 AL = 1.25 and GAL = 0.93139

3.  Chi-square value (CHI) for BSRV
    Sample size (N) for CSRV
    For Case 5 N = 3, CHI = 2.09867

4.  Fatigue test duration (DN)
    For Case 5 DN = 2.0

5.  SRV $\alpha_{SRV}$ and Gamma values (ALS,GAS)
    GAS not required for CSRV

6.  Residual strength $\alpha_R$

7.  Load enhancement factor (PL).

Sample runs for these two cases are given below.

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
        4
CASE 4 FATIGUE REL. LEF WITH SRV
PLEASE INPUT STRENGTH GAMMA
9.735000E-001
PLEASE INPUT LIFE ALPHA AND GAMMA
        1.2500000   9.313800E-001
PLEASE INPUT CHI SQUARE VALUE
        2.0986700
PLEASE INPUT FATIGUE TEST DURATION
        2.0000000
PLEASE INPUT SRV ALPHA AND GAMMA
        17.0000000  9.693000E-001
PLEASE INPUT RES. STRENGTH ALPHA
        20.0000000
PLEASE INPUT LOAD ENHANCEMENT FACTOR
        1.1600000
AT LOAD ENHANCEMENT FACTOR =    1.160
THE ONE LIFETIME REL. IS   .8960821


CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
        4
CASE 4 FATIGUE REL. LEF WITH SRV
PLEASE INPUT STRENGTH ALPHA
        20.0000000
PLEASE INPUT LIFE ALPHA
        1.2500000
PLEASE INPUT NUMBER OF SPECIMENS
        3
PLEASE INPUT FATIGUE TEST DURATION
        2.0000000
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT RES. STRENGTH ALPHA
        20.0000000
PLEASE INPUT LOAD ENHANCEMENT FACTOR
        1.1600000
AT LOAD ENHANCEMENT FACTOR =    1.160
THE ONE LIFETIME REL. IS   .8960822
```

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```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
        5
CASE 5 FATIGUE REL. LEF WITH SRV
FIXED STATIC STRENGTH ALPHA =20.0
FATIGUE LIFE ALPHA = 1.25
TEST DURATION = 2.0 LIFETIME
SAMPLE SIZE = 3, CHI=2.09867
PLEASE INPUT SRV ALPHA AND GAMMA
        17.0000000  9.693000E-001
PLEASE INPUT RES. STRENGTH ALPHA
        20.0000000
PLEASE INPUT LOAD ENHANCEMENT FACTOR
        1.1600000
AT LOAD ENHANCEMENT FACTOR =    1.160
THE ONE LIFETIME REL. IS   .8960821


        CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
        5
CASE 5 FATIGUE REL. LEF WITH SRV
FIXED STATIC STRENGTH ALPHA =20.0
FATIGUE LIFE ALPHA = 1.25
TEST DURATION = 2.0 LIFETIME
SAMPLE SIZE = 3, CHI=2.09867
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT RES. STRENGTH ALPHA
        20.0000000
PLEASE INPUT LOAD ENHANCEMENT FACTOR
        1.1600000
AT LOAD ENHANCEMENT FACTOR =    1.160
THE ONE LIFETIME REL. IS   .8960822
```

Case 6 and 7:   These two cases compute the one lifetime fatigue reliability using the residual strength approach. Case 6 is for general computations and Case 7 is for reliability computations with certain fixed variables. The required input are:

1.   Residual strength $\alpha_R$ (ALPR)
   Fixed at $\alpha_R$ = 20.0 in Case 7.

2.   Fatigue life $\alpha_L$ and associated Gamma value (AL,GAL)
   GAL not required for CSRV
   Fixed at AL = 1.25 and GAL = 0.93138 in Case 7.

3.   Fatigue test duration (AN)
   AN = 2.0 in Case 7.

4.   Static failure strength (SIGU)

5.   Residual strength (SIGR)

6.   Maximum applied stress in fatigue test (SIGA)

7.   Maximum design spectrum load (PM)

8.   SRV $\alpha_{SRV}$ and associated Gamma value (ALS,GAS)
   GAS not required for CSRV.

9.   Chi-square value (CHI) for BSRV
   Sample size (N) for CSRV

Sample runs for Cases 6 and 7 are given below. The program listings follow the sample examples

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          6
CASE 6, FATIGUE REL. RES WITH SRV
PLEASE INPUT RES. STRENGTH ALPHA
       20.0000000
PLEASE INPUT LIFE ALPHA AND GAMMA
       1.2500000   9.313800E-001
PLEASE INPUT FATIGUE TEST DURATION
       2.0000000
PLEASE INPUT STATIC STRENGTH
       1.2500000
PLEASE INPUT RESIDUAL STRENGTH
       1.1800000
PLEASE INPUT TEST APPLIED STRESS
9.600000E-001
PLEASE INPUT MAX. SPECTRUM STRESS
9.200000E-001
PLEASE INPUT SRV ALPHA AND GAMMA
       17.0000000   9.693000E-001
PLEASE INPUT CHI SQUARE VALUE
       2.3720000
AT MAX. SPECTRUM STRESS =     .920
THE ONE LIFETIME REL. =    .7135201
```

```
CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LE  APPROACH WITH SRV
CASE 6--FATIGUE REL., RE . STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
          6
CASE 6, FATIGUE REL. RES WITH SRV
PLEASE INPUT RES. STRENGTH ALPHA
        20.0000000
PLEASE INPUT LIFE ALPHA
        1.2500000
PLEASE INPUT GATIGUE TEST DURATION
        2.0000000
PLEASE INPUT STATIC STRENGTH
        1.2500000
PLEASE INPUT RESIDUAL STRENGTH
        1.1800000
PLEASE INPUT TEST APPLIED STRESS
9.600000E-001
PLEASE INPUT MAX. SPECTRUM STRESS
9.200000E-001
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT NUMBER OF SPECIMENS
            2
AT MAX. SPECTRUM STRESS =    .920
THE ONE LIFETIME REL. =    .7135199
```

```
BSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
        7
CASE 7, FATIGUE REL. RES WITH SRV
FIXED RES. STRENGTH ALPHA =20.0
LIFE ALPHA = 1.25
PLEASE INPUT STATIC STRENGTH
        1.2500000
PLEASE INPUT RESIDUAL STRENGTH
        1.1800000
PLEASE INPUT TEST APPLIED STRESS
9.800000E-001
PLEASE INPUT MAX. SPECTRUM STRESS
9.600000E-001
PLEASE INPUT SRV ALPHA AND GAMMA
        17.0000000   9.693000E-001
PLEASE INPUT CHI SQUARE VALUE
        2.3720000
AT MAX. SPECTRUM STRESS =      .960
THE ONE LIFETIME REL. =    .6190697
```

CSRV
PLEASE INPUT ANALYSIS CASE NUMBER
CASE 1--STATIC RELIABILITY WITY SRV
CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV
CASE 4--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., LEF APPROACH WITH SRV
CASE 5--FATIGUE REL., RES. STRENGTH WITH SRV
CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV
         7
CASE 7, FATIGUE REL. RES WITH SRV
FIXED RES. STRENGTH ALPHA =20.0
LIFE ALPHA = 1.25
PLEASE INPUT STATIC STRENGTH
        1.2500000
PLEASE INPUT RESIDUAL STRENGTH
        1.1800000
PLEASE INPUT TEST APPLIED STRESS
9.800000E-001
PLEASE INPUT MAX. SPECTRUM STRESS
9.600000E-001
PLEASE INPUT SRV ALPHA
        17.0000000
PLEASE INPUT NUMBER OF SPECIMENS
            2
AT MAX. SPECTRUM STRESS =    .960
THE ONE LIFETIME REL. =   .6190694

```
C PROGRAM 'BSRV' STRUCTURAL RELIABILITY COMPUTATIONS--DETAILED INPUT
      DOUBLEPRECISION TEST,PS,SUM,P1,P2,FM,X,DX,DX2,X1,X2,DFI,FME
      SUM = 0.0
      TEST = 0.0000001
      X = 0.0
      DX = 0.001
      DX2 = 0.0005
      AL9 = -ALOG(0.9)
      WRITE(*,101)
  101 FORMAT(3X,'PLEASE INPUT ANALYSIS CASE NUMBER',
     A/3X,'CASE 1--STATIC RELIABILITY WITY SRV',
     B/3X,'CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV',
     C/3X,'CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV',
     D/3X,'CASE 4--FATIGUE REL., LEF APPROACH WITH SRV',
     E/3X,'CASE 5--FATIGUE REL., LEF APPROACH WITH SRV',
     F/3X,'CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV',
     G/3X,'CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV')
      READ(*,*) ICASE
      K = 0
   10 IF(ICASE.EQ.1) GOTO 20
      IF(ICASE.EQ.2.OR.ICASE.EQ.3) GOTO 30
      IF(ICASE.EQ.4.OR.ICASE.EQ.5) GOTO 40
      IF(ICASE.EQ.6.OR.ICASE.EQ.7) GOTO 50
   20 IF(K.NE.0) GOTO 21
      WRITE(*,1)
      WRITE(*,102)
      READ(*,*) ALM
      WRITE(*,103)
      READ(*,*) ALS
      WRITE(*,104)
      READ(*,*) GAS
      WRITE(*,105)
      READ(*,*) BET
      WRITE(*,106)
      READ(*,*) AK
  102 FORMAT(3X,'PLEASE INPUT STATIC STRENGTH ALPHA')
  103 FORMAT(3X,'PLEASE INPUT SRV ALPHA')
  104 FORMAT(3X,'PLEASE INPUT SRV GAMMA')
  105 FORMAT(3X,'PLEASE INPUT LOWER STRENGTH BETA')
  106 FORMAT(3X,'PLEASE INPUT LOAD LEVEL')
      AT = AK/GAS
      DELTX = 1.0
      IF(BET.GE.2.0) DELTX=2.0
      IF(AK.GE.2.0) DELTX=2.0
      GOTO 90
   21 WRITE(*,201) AK,SUM
    1 FORMAT(3X,'CASE 1,STATIC REL. WITH SRV')
  201 FORMAT(3X,'AT LOAD LEVEL OF',F8.3,
     +/3X,'THE STRUCTURAL RELIABILITY IS',F10.7)
      GOTO 99
   30 IF(K.NE.0) GOTO 31
      DELTX = 2.0
      IF(ICASE.EQ.3) GOTO 35
      WRITE(*,3)
      WRITE(*,107)
```

```
      READ(*,*) ALM,GAM
      WRITE(*,108)
      READ(*,*) ALS,GAS
      WRITE(*,109)
      READ(*,*) CON
      WRITE(*,110)
      READ(*,*) CHI
      WRITE(*,111)
      READ(*,*) AM,F
107 FORMAT(3X,'PLEASE INPUT STATIC ALPHA AND GAMMA')
108 FORMAT(3X,'PLEASE INPUT SRV ALPHA AND GAMMA')
109 FORMAT(3X,'PLEASE INPUT SPECTRUM CHARACT. CONST.')
110 FORMAT(3X,'PLEASE INPUT CHI SQUARE VALUE')
111 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM LOAD',
    +/5X,'AND STATIC FAILURE LOAD')
      GOTO 36
 35 WRITE(*,4)
      DFI = 0.05
      CON = 1.5736879
      CHI = 2.9955
      ALM = 20.0
      GAM = 0.9735
      WRITE(*,108)
      READ(*,*) ALS,GAS
      WRITE(*,112)
      READ(*,*) AM,FMIN,FMAX
112 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM LOAD',
    +/3X,'MINIMUM AND MAXIMUM STRENGTH')
      F = FMIN
 36 F1 = AM*CON
      FX = ((AL9/CHI)**(1.0/ALM))*F1/GAM
 32 FG = F/GAS
      AT = FG/(CHI**(1.0/ALS))
      GOTO 90
 31 WRITE(*,202) F,SUM
202 FORMAT(3X,'AT STATIC FAILURE LOAD  ',F8.3,
    +/3X,'THE ONE LIFETIME FATIGUE REL. IS',F10.7)
  3 FORMAT(3X,'CASE 2, FATIGUE REL. ULT WITH SRV')
  4 FORMAT(3X,'CASE 3, FATIGUE REL. ULT WITH SRV',
    +/3X,'FIXED CON=1.5736879',
    +/3X,'STRENGTH ALPHA=20.0, GAMMA=0.9735',
    +/3X,'SINGLE ARTICLE, CHI=2.9955',
    +/3X,'REL. COMPUTED AT 0.05 INTERVAL')
      IF(ICASE.EQ.2) GOTO 99
      F = F+DFI
      X =0.0
      SUM = 0.0
      IF(F.LT.FMAX) GOTO 32
      GOTO 99
 40 IF(K.NE.0) GOTO 41
      DELTX = 2.0
      IF(ICASE.EQ.5) GOTO 45
      WRITE(*,5)
      WRITE(*,113)
      READ(*,*) GAM
```

```
      WRITE(*,114)
      READ(*,*) AL,GAL
      WRITE(*,115)
      READ(*,*) CHI
      WRITE(*,116)
      READ(*,*) DN
  113 FORMAT(3X,'PLEASE INPUT STRENGTH GAMMA')
  114 FORMAT(3X,'PLEASE INPUT LIFE ALPHA AND GAMMA')
  115 FORMAT(3X,'PLEASE INPUT CHI SQUARE VALUE')
  116 FORMAT(3X,'PLEASE INPUT FATIGUE TEST DURATION')
      GOTO 46
   45 WRITE(*,6)
      GAM = 0.9735
      GAL = 0.93138
      CHI = 2.09867
      DN = 2.0
      AL = 1.25
   46 WRITE(*,117)
      READ(*,*) ALS,GAS
      WRITE(*,118)
      READ(*,*) ALPR
      WRITE(*,119)
      READ(*,*) PL
  117 FORMAT(3X,'PLEASE INPUT SRV ALPHA AND GAMMA')
  118 FORMAT(3X,'PLEASE INPUT RES. STRENGTH ALPHA')
  119 FORMAT(3X,'PLEASE INPUT LOAD ENHANCEMENT FACTOR')
      ARG = CHI*((GAL/DN)**AL)
      AT = PL/GAS
      GOTO 90
   41 WRITE(*,203) PL,SUM
  203 FORMAT(3X,'AT LOAD ENHANCEMENT FACTOR = ',F8.3,
     +/3X,'THE ONE LIFETIME REL. IS ',F10.7)
    5 FORMAT(3X,'CASE 4 FATIGUE REL. LEF WITH SRV')
    6 FORMAT(3X,'CASE 5 FATIGUE REL. LEF WITH SRV',
     +/3X,'FIXED STATIC STRENGTH ALPHA =20.0',
     +/3X,'FATIGUE LIFE ALPHA = 1.25',
     +/3X,'TEST DURATION = 2.0 LIFETIME',
     +/3X,'SAMPLE SIZE = 3, CHI=2.09867')
      GOTO 99
   50 IF(K.NE.0) GOTO 51
      DELTX = 1.0
      X = 0.05-DX2
      IF(ICASE.EQ.7) GOTO 55
      WRITE(*,7)
      WRITE(*,120)
      READ(*,*) ALPR
      WRITE(*,121)
      READ(*,*) AL,GAL
      WRITE(*,122)
      READ(*,*) AN
  120 FORMAT(3X,'PLEASE INPUT RES. STRENGTH ALPHA')
  121 FORMAT(3X,'PLEASE INPUT LIFE ALPHA AND GAMMA')
  122 FORMAT(3X,'PLEASE INPUT FATIGUE TEST DURATION')
      GOTO 56
   55 WRITE(*,8)
```

```
      ALPR = 20.0
      AL = 1.25
      GAL = 0.93138
      AN = 2.0
   56 WRITE(*,123)
      READ(*,*) SIGU
      WRITE(*,124)
      READ(*,*) SIGR
      WRITE(*,125)
      READ(*,*) SIGA
      WRITE(*,126)
      READ(*,*) PM
      WRITE(*,127)
      READ(*,*) ALS,GAS
      WRITE(*,128)
      READ(*,*) CHI
  123 FORMAT(3X,'PLEASE INPUT STATIC STRENGTH')
  124 FORMAT(3X,'PLEASE INPUT RESIDUAL STRENGTH')
  125 FORMAT(3X,'PLEASE INPUT TEST APPLIED STRESS')
  126 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM STRESS')
  127 FORMAT(3X,'PLEASE INPUT SRV ALPHA AND GAMMA')
  128 FORMAT(3X,'PLEASE INPUT CHI SQUARE VALUE')
      CH = CHI**(1.0/AL)
      S1 = ALPR/AL
      C = ((SIGU/SIGA)**S1-(SIGR/SIGA)**S1)/AN
      AT = PM/GAS
      IF(SIGU.LT.DELTX) DELTX = 0.9*SIGU
      GOTO 90
   51 WRITE(*,204) PM,SUM
  204 FORMAT(3X,'AT MAX. SPECTRUM STRESS =',F8.3,
     +/3X,'THE ONE LIFETIME REL. = ',F10.7)
    7 FORMAT(3X,'CASE 6, FATIGUE REL. RES WITH SRV')
    8 FORMAT(3X,'CASE 7, FATIGUE REL. RES WITH SRV',
     +/3X,'FIXED RES. STRENGTH ALPHA =20.0',
     +/3X,'LIFE ALPHA = 1.25')
      GOTO 99
   90 X = X+DX
      X1 = X-DX2
      X2 = X+DX2
      P1 = EXP(-(X1/AT)**ALS)
      P2 = EXP(-(X2/AT)**ALS)
      IF(ICASE.EQ.2.OR.ICASE.EQ.3) GOTO 91
      IF(ICASE.EQ.4.OR.ICASE.EQ.5) GOTO 92
      IF(ICASE.EQ.6.OR.ICASE.EQ.7) GOTO 93
      FME = (X/BET)**ALM
      IF(FME.GT.40.) GOTO 190
      FM = EXP(-FME)
      GOTO 95
  190 FM = 0.0
      GOTO 95
   91 FME = ((FX*GAM/X)**ALM)*CHI
      IF(FME.GT.40.) GOTO 191
      FM = EXP(-FME)
      GOTO 95
  191 FM = 0.0
```

```
        GOTO 95
   92 FME = ARG/(X**ALPR)
        IF(FME.GT.40.) GOTO 192
        FM = EXP(-FME)
        GOTO 95
  192 FM = 0.0
        GOTO 95
   93 F = ((SIGU/X)**S1-1.0)/C
        BN = F/GAL
        BU = BN/CH
        FME = (1.0/BU)**AL
        IF(FME.GT.40.) GOTO 193
        FM = EXP(-FME)
        GOTO 95
  193 FM = 0.0
   95 PS = FM*(P1-P2)
        SUM = SUM+PS
        IF(X.LE.DELTX) GOTO 90
        IF(PS.GT.TEST) GOTO 90
        K = K+1
        GOTO 10
   99 STOP
        END
```

```
C PROGRAM 'CSRV' STRUCTURAL RELIABILITY COMPUTATIONS--SHORT INPUT
      DOUBLEPRECISION TEST,PS,SUM,P1,P2,FM,X,DX,DX2,X1,X2,DFI,FME
      COMMON/GMA/B(101),Y(101)
      COMMON/CHI/CHL(15)
      OPEN(5,FILE='PSI.DAT')
      READ(5,*) (B(I),I=1,101)
      READ(5,*) (Y(I),I=1,101)
      READ(5,*) (CHL(I),I=1,15)
      SUM = 0.0
      TEST = 0.0000001
      X = 0.0
      DX = 0.001
      DX2 = 0.0005
      AL9 = -ALOG(0.9)
      WRITE(*,101)
  101 FORMAT(3X,'PLEASE INPUT ANALYSIS CASE NUMBER',
     A/3X,'CASE 1--STATIC RELIABILITY WITY SRV',
     B/3X,'CASE 2--FATIGUE REL., ULT. STRENGTH WITH SRV',
     C/3X,'CASE 3--FATIGUE REL., ULT. STRENGTH WITH SRV',
     D/3X,'CASE 4--FATIGUE REL., LEF APPROACH WITH SRV',
     E/3X,'CASE 5--FATIGUE REL., LEF APPROACH WITH SRV',
     F/3X,'CASE 6--FATIGUE REL., RES. STRENGTH WITH SRV',
     G/3X,'CASE 7--FATIGUE REL., RES. STRENGTH WITH SRV')
      READ(*,*) ICASE
      K = 0
   10 IF(ICASE.EQ.1) GOTO 20
      IF(ICASE.EQ.2.OR.ICASE.EQ.3) GOTO 30
      IF(ICASE.EQ.4.OR.ICASE.EQ.5) GOTO 40
      IF(ICASE.EQ.6.OR.ICASE.EQ.7) GOTO 50
   20 IF(K.NE.0) GOTO 21
      WRITE(*,1)
      WRITE(*,102)
      READ(*,*) ALM
      WRITE(*,103)
      READ(*,*) ALS
      WRITE(*,104)
      READ(*,*) N
      WRITE(*,105)
      READ(*,*) XB
      WRITE(*,106)
      READ(*,*) AK
  102 FORMAT(3X,'PLEASE INPUT STATIC STRENGTH ALPHA')
  103 FORMAT(3X,'PLEASE INPUT SRV ALPHA')
  104 FORMAT(3X,'PLEASE INPUT NUMBER OF SPECIMENS')
  105 FORMAT(3X,'PLEASE INPUT MEAN STATIC STRENGTH ')
  106 FORMAT(3X,'PLEASE INPUT LOAD LEVEL')
      ARG = 1.0+1.0/ALM
      GAM = GAMMA(ARG)
      CHI = CHIS(N)
      BET = XB/GAM
      BET = BET/(CHI**(1.0/ALM))
      ARG = 1.0+1.0/ALS
      GAS = GAMMA(ARG)
      AT = AK/GAS
      DELTX = 1.0
```

```
      IF(BET.GE.2.0) DELTX = 2.0
      IF(AK.GE.2.0) DELTX = 2.0
      GOTO 90
   21 WRITE(*,201) AK,SUM
    1 FORMAT(3X,'CASE 1,STATIC REL. WITH SRV')
  201 FORMAT(3X,'AT LOAD LEVEL OF',F8.3,
     +/3X,'THE STRUCTURAL RELIABILITY IS',F10.7)
      GOTO 99
   30 IF(K.NE.0) GOTO 31
      DELTX = 2.0
      IF(ICASE.EQ.3) GOTO 35
      WRITE(*,3)
      WRITE(*,107)
      READ(*,*) ALM
      WRITE(*,108)
      READ(*,*) ALS
      WRITE(*,109)
      READ(*,*) CON
      WRITE(*,110)
      READ(*,*) N
      WRITE(*,111)
      READ(*,*) AM,F
  107 FORMAT(3X,'PLEASE INPUT STATIC ALPHA')
  108 FORMAT(3X,'PLEASE INPUT SRV ALPHA')
  109 FORMAT(3X,'PLEASE INPUT SPECTRUM CHARACT. CONST.')
  110 FORMAT(3X,'PLEASE INPUT NUMBER OF SPECIMENS')
  111 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM LOAD',
     +/5X,'AND STATIC FAILURE LOAD')
      CHI = CHIS(N)
      ARG = 1.0+1.0/ALM
      GAM = GAMMA(ARG)
      ARG = 1.0+1.0/ALS
      GAS = GAMMA(ARG)
      GOTO 36
   35 WRITE(*,4)
      DFI = 0.05
      CON = 1.5736879
      CHI = 2.9955
      ALM = 20.0
      GAM = 0.9735
      WRITE(*,108)
      READ(*,*) ALS
      WRITE(*,112)
      READ(*,*) AM,FMIN,FMAX
  112 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM LOAD',
     +/3X,'MINIMUM AND MAXIMUM STRENGTH')
      ARG = 1.0+1.0/ALS
      GAS = GAMMA(ARG)
      F = FMIN
   36 F1 = AM*CON
      FX = ((AL9/CHI)**(1.0/ALM))*F1/GAM
   32 FG = F/GAS
      AT = FG/(CHI**(1.0/ALS))
      GOTO 90
   31 WRITE(*,202) F,SUM
```

```
202 FORMAT(3X,'AT STATIC FAILURE LOAD  ',F8.3,
   +/3X,'THE ONE LIFETIME FATIGUE REL. IS',F10.7)
  3 FORMAT(3X,'CASE 2, FATIGUE REL. ULT WITH SRV')
  4 FORMAT(3X,'CASE 3, FATIGUE REL. ULT WITH SRV',
   +/3X,'FIXED CON=1.5736879',
   +/3X,'STRENGTH ALPHA=20.0, GAMMA=0.9735',
   +/3X,'SINGLE ARTICLE, CHI=2.9955',
   +/3X,'REL. COMPUTED AT 0.05 INTERVAL')
    IF(ICASE.EQ.2) GOTO 99
    F = F+DFI
    X =0.0
    SUM = 0.0
    IF(F.LT.FMAX) GOTO 32
    GOTO 99
 40 IF(K.NE.0) GOTO 41
    DELTX = 2.0
    IF(ICASE.EQ.5) GOTO 45
    WRITE(*,5)
    WRITE(*,113)
    READ(*,*) ALM
    WRITE(*,114)
    READ(*,*) AL
    WRITE(*,115)
    READ(*,*) N
    WRITE(*,116)
    READ(*,*) DN
113 FORMAT(3X,'PLEASE INPUT STRENGTH ALPHA')
114 FORMAT(3X,'PLEASE INPUT LIFE ALPHA')
115 FORMAT(3X,'PLEASE INPUT NUMBER OF SPECIMENS')
116 FORMAT(3X,'PLEASE INPUT FATIGUE TEST DURATION')
    ARG = 1.0+1.0/ALM
    GAM = GAMMA(ARG)
    ARG = 1.0+1.0/AL
    GAL = GAMMA(ARG)
    CHI = CHIS(N)
    GOTO 46
 45 WRITE(*,6)
    GAM = 0.9735
    GAL = 0.93138
    CHI = 2.09867
    DN = 2.0
    AL = 1.25
 46 WRITE(*,117)
    READ(*,*) ALS
    WRITE(*,118)
    READ(*,*) ALPR
    WRITE(*,119)
    READ(*,*) PL
117 FORMAT(3X,'PLEASE INPUT SRV ALPHA')
118 FORMAT(3X,'PLEASE INPUT RES. STRENGTH ALPHA')
119 FORMAT(3X,'PLEASE INPUT LOAD ENHANCEMENT FACTOR')
    ARG = 1.0+1.0/ALS
    GAS = GAMMA(ARG)
    ARG = CHI*((GAL/DN)**AL)
    AT = PL/GAS
```

```
      GOTO 90
   41 WRITE(*,203) PL,SUM
  203 FORMAT(3X,'AT LOAD ENHANCEMENT FACTOR = ',F8.3,
     +/3X,'THE ONE LIFETIME REL. IS ',F10.7)
    5 FORMAT(3X,'CASE 4 FATIGUE REL. LEF WITH SRV')
    6 FORMAT(3X,'CASE 5 FATIGUE REL. LEF WITH SRV',
     +/3X,'FIXED STATIC STRENGTH ALPHA =20.0',
     +/3X,'FATIGUE LIFE ALPHA = 1.25',
     +/3X,'TEST DURATION = 2.0 LIFETIME',
     +/3X,'SAMPLE SIZE = 3, CHI=2.09867')
      GOTO 99
   50 IF(K.NE.0) GOTO 51
      DELTX = 1.0
      X = 0.05-DX2
      IF(ICASE.EQ.7) GOTO 55
      WRITE(*,7)
      WRITE(*,120)
      READ(*,*) ALPR
      WRITE(*,121)
      READ(*,*) AL
      WRITE(*,122)
      READ(*,*) AN
  120 FORMAT(3X,'PLEASE INPUT RES. STRENGTH ALPHA')
  121 FORMAT(3X,'PLEASE INPUT LIFE ALPHA')
  122 FORMAT(3X,'PLEASE INPUT GATIGUE TEST DURATION')
      ARG = 1.0+1.0/AL
      GAL = GAMMA(ARG)
      GOTO 56
   55 WRITE(*,8)
      ALPR = 20.0
      AL = 1.25
      GAL = 0.93138
      AN = 2.0
   56 WRITE(*,123)
      READ(*,*) SIGU
      WRITE(*,124)
      READ(*,*) SIGR
      WRITE(*,125)
      READ(*,*) SIGA
      WRITE(*,126)
      READ(*,*) PM
      WRITE(*,127)
      READ(*,*) ALS
      WRITE(*,128)
      READ(*,*) N
  123 FORMAT(3X,'PLEASE INPUT STATIC STRENGTH')
  124 FORMAT(3X,'PLEASE INPUT RESIDUAL STRENGTH')
  125 FORMAT(3X,'PLEASE INPUT TEST APPLIED STRESS')
  126 FORMAT(3X,'PLEASE INPUT MAX. SPECTRUM STRESS')
  127 FORMAT(3X,'PLEASE INPUT SRV ALPHA')
  128 FORMAT(3X,'PLEASE INPUT NUMBER OF SPECIMENS')
      ARG = 1.0+1.0/ALS
      GAS = GAMMA(ARG)
      CHI = CHIS(N)
      CH = CHI**(1.0/AL)
```

```
      S1 = ALPR/AL
      C = ((SIGU/SIGA)**S1-(SIGR/SIGA)**S1)/AN
      AT = PM/GAS
      IF(SIGU.LT.DELTX) DELTX = 0.9*SIGU
      GOTO 90
   51 WRITE(*,204) PM,SUM
  204 FORMAT(3X,'AT MAX. SPECTRUM STRESS =',F8.3,
     +/3X,'THE ONE LIFETIME REL. = ',F10.7)
    7 FORMAT(3X,'CASE 6, FATIGUE REL. RES WITH SRV')
    8 FORMAT(3X,'CASE 7, FATIGUE REL. RES WITH SRV',
     +/3X,'FIXED RES. STRENGTH ALPHA =20.0',
     +/3X,'LIFE ALPHA = 1.25')
      GOTO 99
   90 X = X+DX
      X1 = X-DX2
      X2 = X+DX2
      P1 = EXP(-(X1/AT)**ALS)
      P2 = EXP(-(X2/AT)**ALS)
      IF(ICASE.EQ.2.OR.ICASE.EQ.3) GOTO 91
      IF(ICASE.EQ.4.OR.ICASE.EQ.5) GOTO 92
      IF(ICASE.EQ.6.OR.ICASE.EQ.7) GOTO 93
      FME = (X/BET)**ALM
      IF(FME.GT.40.) GOTO190
      FM = EXP(-FME)
      GOTO 95
  190 FM = 0.0
      GOTO 95
   91 FME = ((FX*GAM/X)**ALM)*CHI
      IF(FME.GT.40.) GOTO 191
      FM = EXP(-FME)
      GOTO 95
  191 FM = 0.0
      GOTO 95
   92 FME = ARG/(X**ALPR)
      IF(FME.GT.40.) GOTO 192
      FM = EXP(-FME)
      GOTO 95
  192 FM = 0.0
      GOTO 95
   93 F = ((SIGU/X)**S1-1.0)/C
      BN = F/GAL
      BU = BN/CH
      FME = (1.0/BU)**AL
      IF(FME.GT.40.) GOTO 193
      FM = EXP(-FME)
      GOTO 95
  193 FM = 0.0
   95 PS = FM*(P1-P2)
      SUM = SUM+PS
      IF(X.LE.DELTX) GOTO 90
      IF(PS.GT.TEST) GOTO 90
      K = K+1
      GOTO 10
   99 STOP
      END
```

```
      FUNCTION CHIS(N)
      COMMON/CHI/CHL(15)
      AN = N
      BN = 2.0*AN
      IF(N.GE.15) GOTO 50
      CHIS = CHL(N)
      GOTO 60
   50 B = 1.0/(9.0*AN)
      CL = 1.0-B+1.645*SQRT(B)
      CHIS = CL*CL*CL
   60 CONTINUE
      RETURN
      END
      FUNCTION GAMMA(X)
      COMMON/GMA/B(101),Y(101)
      ARG = X
      A = 1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      IF(ARG.EQ.2.0) GOTO 110
      IF(ARG.GT.2.0) GOTO 20
      GOTO 30
   10 A = A/ARG
      ARG = ARG+1.0
      IF(ARG.LT.1.0) GOTO 10
      IF(ARG.EQ.1.0) GOTO 110
      GOTO 30
   20 ARG = ARG-1.0
      A = A*ARG
      IF(ARG.GT.2.0) GOTO 20
      IF(ARG.EQ.2.0) GOTO 110
   30 DO 40 I=1,101
      IF(B(I).GT.ARG) GOTO 50
   40 CONTINUE
   50 SLOP = (Y(I)-Y(I-1))/(B(I)-B(I-1))
      F = Y(I-1)+(ARG-B(I-1))*SLOP
      GOTO 60
  110 F = 1.0
   60 GAMMA = F*A
      RETURN
      END
```

PSI.DAT  TAB VALUES OF CHI-SQUARE AND GAMMA FUNCTIONS
1.0,  1.01, 1.02, 1.03, 1.04, 1.05, 1.06, 1.07, 1.08, 1.09, 1.1,  1.11,
1.12, 1.13, 1.14, 1.15, 1.16, 1.17, 1.18, 1.19, 1.2,  1.21, 1.22, 1.23,
1.24, 1.25, 1.26, 1.27, 1.28, 1.29, 1.3,  1.31, 1.32, 1.33, 1.34, 1.35,
1.36, 1.37, 1.38, 1.39, 1.4,  1.41, 1.42, 1.43, 1.44, 1.45, 1.46, 1.47,
1.48, 1.49, 1.5,  1.51, 1.52, 1.53, 1.54, 1.55, 1.56, 1.57, 1.58, 1.59,
1.6,  1.61, 1.62, 1.63, 1.64, 1.65, 1.66, 1.67, 1.68, 1.69, 1.7,  1.71,
1.72, 1.73, 1.74, 1.75, 1.76, 1.77, 1.78, 1.79, 1.8,  1.81, 1.82, 1.83,
1.84, 1.85, 1.86, 1.87, 1.88, 1.89, 1.9,  1.91, 1.92, 1.93, 1.94, 1.95,
1.96, 1.97, 1.98, 1.99, 2.,
1.,       .99433, .98884, .98355, .97844, .9735,  .96874, .96415, .95973,
2.9955, 2.372,  2.09867, 1.93838, 1.8307,  1.75217, 1.69179, 1.6435,
1.60383, 1.5705, 1.542,  1.51729, 1.49558, 1.47632, 1.4591

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